National Institute for Health and Care Excellence

Guideline version (Draft)

Lung Cancer Update

Evidence reviews for the clinical and cost effectiveness of treatment regimen for the treatment of Stage IIIA-N2 NSCLC

NICE guideline <number>

Evidence reviews

October 2018

Draft for Consultation

These evidence reviews were developed by the NICE Guideline Updates Team

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ISBN:

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Evidence reviews for the clinical and cost effectiveness of treatment regimens for the treatment of Stage IIIA-N2 NSCLC

Review questions

RQ3.1: What is the clinical and cost effectiveness of chemoradiotherapy or surgery with adjuvant treatment for the treatment for N2 stage NSCLC?

Introduction

The aim of the review is to provide clearer guidance regarding the treatment of stage IIIA-N2 NSCLC. This is because the roles of surgery and chemoradiotherapy in this setting are extensively debated.

Table 1: PICO table

Population	People with stage N2 M0 NSCLC
Interventions	Surgery (S) with or without chemotherapy (C)
Comparators	• Chemoradiotherapy (radiotherapy and chemotherapy (CR)) • Tri-modality treatment (radiotherapy, chemotherapy and surgery (CRS))
Outcomes	• Mortality • Quality of life • Length of stay • Exercise tolerance • Adverse events • Treatment-related dropout rates • Pain

Methods and process

This evidence review was developed using the methods and process described in Developing NICE guidelines: the manual (2014). Methods specific to this review question are described in the review protocol in appendix A, and the methods section in appendix B. In particular, the minimally important differences (MIDs) used in this review are summarised in appendix B.

Declarations of interest were recorded according to NICE's 2018 conflicts of interest policy.

One thousand abstracts were screened manually.

This review includes several network meta-analysis performed by the NICE Guidelines Technical Support Unit (TSU), which is based at the University of Bristol and the University of Leicester.

Clinical evidence

Included studies

This review was conducted as part of a larger update of the NICE Lung cancer: diagnosis and management guideline (CG121). A systematic literature search for randomised controlled trials (RCTs) with a no date limit yielded 4,241 references.

Papers returned by the literature search were screened on title and abstract, with 21 full-text papers ordered as potentially relevant systematic reviews or RCTs.

Eleven papers representing 10 unique RCTs were included after full text screening. The RCTs were: Albain 2009 (n=396, follow-up period was a minimum of 2.5 years), Eberhardt 2015 (n=161, follow-up period was a minimum of 1 year), Girard 2010 (n=46, the median follow-up period was 31.4 months), Johnstone 2002 (n=61, follow up period was a minimum of 4 years), Katakami 2012 (n=56, follow-up period was a minimum of 5 years), Pless 2015 (n=231, the median follow-up period was 52 months), Shepherd 1998 (n=31, follow-up was 24 months in one arm and 31 months in the other), Stephens 2005 (n=48, the median follow-up period was 14 months), Thomas 2008 (n=524, the median follow-up period was 70 months) and van Meerbeeck 2007 (n=208, the median follow-up period was 6 years).

For the search strategy, please see appendix C. For the clinical evidence study selection flowchart, see appendix D. For the full evidence tables and full GRADE profiles for included studies, please see appendices E and F.

Excluded studies

Details of the studies excluded at full-text review are given in appendix G along with a reason for their exclusion.

Summary of clinical studies included in the evidence review

Study locations

One randomised controlled study was from the UK (Stephens 2005), 1 was from France (Girard 2010), 2 were from Germany (Eberhardt 2015, Thomas 2008), 1 was from Switzerland, Germany and Serbia (Pless 2015), 1 was from the Netherlands (van Meerbeeck 2007), 1 was from the USA (Johnstone 2002), 1 was from Canada (Shepherd 1998), 1 was from the USA and Canada (Albain 2009) and 1 was from Japan (Katakami 2012).

Outcomes and sample sizes

The reported outcomes with extractable data were mortality and adverse events. The sample sizes ranged from 31 participants to 524 across studies.

See full evidence tables and Grade profiles in appendices E and F.

Quality assessment of clinical studies included in the evidence review

See appendix E for full GRADE tables.

Economic evidence

Standard health economic filters were applied to the clinical search for this question, and a total of 956 citations were returned. Following review of titles and abstracts,

two full text studies were retrieved for detailed consideration, but these were subsequently excluded as not relevant. Therefore, no relevant cost–utility analyses were identified for this question.

This review question was prioritised for economic modelling, and an original economic model was developed.

Summary of original economic model

The de novo cost-utility analysis developed for this guideline included three strategies; chemoradiotherapy (CR), chemotherapy and surgery (CS) and chemoradiotherapy and surgery (CRS). It was based on a hybrid structure where the amount of time that patients spent in the progression free and progressed states, the probability of survival and the adverse events during the first five years were drawn from network meta-analyses conducted for this guideline. Survival in patients still alive after five years was modelled using patient registry data. The model included costs for the initial interventions and for treatment on progression, deaths, adverse events and routine costs associated with the progression free and progressed states. The model included utility estimates for both states as well as longer term survival and a disutility adjustment in the surgical arm. In accordance with data from the underpinning trials, not all patients in surgical strategies went on to receive surgery following chemoradiotherapy. Patients entered the model at age 60, which reflected the average age in the underpinning trials. The cycle length was one month and costs and health benefits were discounted at 3.5% per year.

The model found that CS was extendedly dominated by CR and CRS and had an ICER of £53,500/QALY versus CR. CRS was cost-effective compared to CR with an ICER of £17,800/QALY. These results were robust to a wide range of sensitivity and scenario analyses. The probabilistic sensitivity analysis showed that CRS produced more QALYs than CR in 97% and 87% of iterations respectively. There were, however, key uncertainties in the underpinning clinical data with no individual pairwise studies having reported significant differences in overall or progression free survival. No subgroup analyses were performed. The full modelling report is available in Appendix K.

Evidence statements

The outcomes reported in network meta-analyses were not directly reported in the underpinning trials and therefore, although the trials are the same, there are no corresponding evidence statements for pairwise comparisons. Progression free survival time, post-progression survival time and the probability of survival were calculated using data extracted from survival graphs and ‘number at risk’ tables available in the underpinning studies.

C = chemotherapy, R = radiotherapy, S = surgery.

CRS vs CR vs CS (network meta-analysis)

Moderate quality evidence from 1 network meta-analysis that included more than 1,000 patients across 6 RCTs could not distinguish the odds of survival at 4 years between the interventions.

Moderate quality evidence from 1 network meta-analysis that included more than 1,000 patients across 5 RCTs could not distinguish the odds of survival at 5 years between the interventions.

High quality evidence from 1 network meta-analysis that included more than 1,000 patients across 6 RCTs found that CRS was associated with a longer progression-free survival time than both CS and CR at 4 years. The data could not differentiate CS from CR.

High quality evidence from 1 network meta-analysis that included more than 1,000 patients across 5 RCTs found that CRS was associated with a longer progression-free survival time than both CS and CR at 5 years. The data could not differentiate CS from CR.

High quality evidence from 1 network meta-analysis that included more than 1,000 patients across 6 RCTs found that CS and CRS were both associated with a shorter post-progression survival time than CR at 4 years.

High quality evidence from 1 network meta-analysis that included more than 1,000 patients across 5 RCTs found that CRS was associated with a shorter post-progression survival time than CR at 5 years. The data could not differentiate CS and CR.

Moderate quality evidence from 1 network meta-analysis that included more than 1,000 patients across 6 RCTs could not distinguish total life years at 4 years between the interventions.

Moderate quality evidence from 1 network meta-analysis that included more than 1,000 patients across 5 RCTs could not distinguish total life years at 5 years between the interventions.

High quality evidence from 1 network meta-analysis that included more than 1,000 patients across 4 RCTs found that CCRS was associated with a lower hazard ratio of adverse events at grade 3+ than both CS and CR.

CRS vs CR

Moderate-quality evidence from 1 RCT reporting data on 396 people with N2 NSCLC found that the data could not differentiate for mortality (all-cause hazard ratio). However, high to moderate-quality evidence found there were a greater number of participants who experienced anaemia, nausea and/or emesis, oesophagitis and pulmonary (adverse events grade 3 or above) in the CR group compared to the CRS group. The data could not differentiate for eukopenia, neutropenia, thrombocytopenia, worst haematologic toxicity per patient, neuropathy, stomatitis and/or mucositis, other gastrointestinal or renal, cardiac, miscellaneous infection, haemorrhage, fatigue, anorexia or allergy (adverse events grade 3 or above).

CRS vs CS

Very low to moderate-quality evidence from 3 RCTs reporting data on 333 people with NSCLC found that the data could not differentiate for mortality (all-cause hazard ratio and risk ratio for survival at 1, 2 and 3 years), stomatitis, dyspnoea and pneumonitis (adverse events grade 3 or above).

C, CRS vs C, CR boost

Moderate to high-quality evidence from 1 RCT reporting data from 161 people with potentially resectable stage IIIA (N2) or selected stage IIIB NSCLC found that the data could not differentiate for mortality at 1 year, 2 years, 3 years, 4 years, 5 years and 6 years. However, there were a greater number of participants who experienced oesophagitis in the C, CR boost group compared to the C, CRS group. The data could not differentiate for leukopenia, anaemia, thrombocytopenia, nausea/vomiting, neuropathy, mucositis/stomatitis, pulmonary, other GI or renal, cardiac, miscellaneous infection, fatigue, pain (adverse events grade 3 or above) or dropout during treatment.

CS vs CR

Very low to moderate-quality evidence from 2 RCTs reporting data from 369 people with N2 NSCLC found that the data could not differentiate for mortality at 1 year, 2 years, 3 years and 4 years. Neither could the data differentiate for treatment-related mortality nor dropout during treatment.

CS vs CRS (cisplatin + docetaxel)

Moderate to high-quality evidence from 1 RCT reporting data from 231 people who had stage IIIA (T1-3) N2 NSCLC found the CS group had a greater number of people who experienced infection compared to the CRS (cisplatin + docetaxel) group. The data could not differentiate for mortality (all-cause hazard ratio), alopecia, nausea/vomiting, fatigue, diarrhoea, neurotoxic effects, stomatitis, skin toxic effects, dyspnoea, fluid retention, constipation, febrile neutropenia, fever, allergic reaction, neutropenia, leukopenia, thrombocytopenia, anaemia (adverse events grade 3 or above), or dropout during treatment.

CS vs R

Very low to low-quality evidence from 2 RCTs reporting data from 79 people who had NSCLC T3, N1, M0 or T1-3, N2, M0 found that the data could not differentiate for mortality, lethargy (this adverse event was grade 2 or above) or dropout during treatment.

C, CRS, R vs CRS

Very low-quality evidence from 1 RCT reporting data from 524 people with NSCLC stage IIIA (T1-3, N2, M0 or central T3, N0-1, M0) or stage IIIB (T4, N1-3, M0 or T1-4, N3, M0) found that the data could not differentiate for mortality (all-cause hazard ratio or treatment related). However, there were a greater number of people who experienced haemotoxicity in the C, CRS, R group compared to the CRS group. There were a greater number of people who experienced pneumonitis in the CRS compared to the C, CRS, R group. The data could not differentiate for oesophagitis and peri-operative complications (adverse events were grade 3 or above).

Health economics evidence statements

Evidence from one directly applicable original health economic model with minor limitations built for this guideline showed that chemoradiotherapy with surgery is very likely to be more cost-effective than chemoradiotherapy (pairwise ICER = £17,800/QALY) and chemotherapy with surgery (pairwise ICER = £6,800) per QALY. The model's conclusions were largely insensitive to changes in model parameters and assumptions.

Recommendations

1.4.40 For people with stage IIIA–N2 NSCLC who are well enough for multimodality therapy and who can have surgery, consider chemoradiotherapy with surgery. **[2019]**

1.4.41 For people with stage IIIA–N2 NSCLC who are having chemoradiotherapy and surgery, ensure that their surgery is scheduled for 3–5 weeks after the chemoradiotherapy. **[2019]**

1.4.42 Centres performing lung resections for lung cancer should validate their data for the Lung Cancer Clinical Outcomes publication. **[2019]**

Research recommendation

What is the effectiveness and cost effectiveness of immunotherapy in people with stage IIIA-N2 NSCLC following multimodality treatment including surgery?

Rationale and impact

Why the committee made the recommendations

The available evidence showed that chemoradiotherapy and surgery are more effective than chemoradiotherapy alone in people who are well enough for surgery. For chemotherapy and surgery, there was no evidence that outcomes were better than for chemoradiotherapy, so the additional costs outweighed the benefits. The key benefit associated with chemoradiotherapy and surgery is the longer progression free survival time. However, there are some uncertainties in the evidence:

- it was not possible to tell whether chemoradiotherapy alone or chemotherapy and surgery provide better survival outcomes
- the evidence in favour of chemoradiotherapy and surgery involved indirect comparisons, and no head-to-head trials showed meaningful differences in outcomes for any of the interventions.

The 3–5 week wait for surgery is recommended to give people time to recover from the chemoradiotherapy.

Immunotherapy has been shown to be effective in a variety of NSCLC indications but there is currently no evidence on whether it is clinically or cost effective for people with stage IIIA-N2 non-small-cell lung cancer following surgery. The committee made a research recommendation to address this.

Impact of the recommendations on practice

The committee felt that chemoradiotherapy and surgery is offered far less often than chemoradiotherapy alone or chemotherapy and surgery for people with NSCLC stage IIIA-N2. Therefore, these recommendations could lead to a change in current practice.

Interpreting the evidence

The outcomes that matter most

The committee agreed that the outcome that matters the most is mortality. This is because the purpose of chemotherapy, radiotherapy and surgery is to reduce mortality as much as possible. Secondary outcomes were severe adverse events and quality of life.

The quality of the evidence

The committee agreed that the aim of the review question was to try to establish a standard approach to managing NSCLC stage IIIA-N2. Ten of the 11 RCTs included in this review question could not differentiate mortality.

The committee agreed that the six trials most relevant to current practice were Pless 2015, Katakami 2012, Albain 2009, Eberhardt 2015, Girard 2010 and van Meerbeeck 2007. For the first four of these trials, outcomes were largely graded as moderate quality evidence. For the final two, outcomes were largely graded as low quality evidence. Overall survival time, progression-free survival time, probability of survival at study endpoint and adverse event data were then combined in network meta-analyses (NMA). The fixed effects network meta-analyses found that patients receiving chemoradiotherapy and surgery spent significantly longer progression free than those receiving chemotherapy and surgery or chemoradiotherapy alone, that patients receiving chemoradiotherapy alone spent significantly longer in the post-progression state than those receiving the surgical options and that there was a strong but statistically insignificant trend favouring chemoradiotherapy and surgery over the other two interventions for overall survival time and probability of survival at study endpoint. While model fit statistics did not suggest that it fit the data any better, the random effects network meta-analyses used in sensitivity analysis found no statistically significant difference for any outcome between any of the interventions. See Appendix J for more details on the NMAs conducted for this question.

Benefits and harms

Based on the NMA, the committee agreed that it is likely that (particularly) progression-free survival and overall survival are better for chemoradiotherapy and surgery (CRS) than the other two options if patients are well enough for it. The NMA found that CRS was associated with a 4.5 month (0.38 year) improvement in progression-free survival versus chemoradiotherapy (CR). The adverse event profile of the different interventions is uncertain but pairwise and network meta-analyses estimates conducted for the health economic model favoured CRS. The committee were unsure about the clinical plausibility of this, given that CRS is the most intensive intervention but agreed that there was no evidence that it was more harmful than the other two interventions. The committee agreed it was likely that there would be some quality of life loss in the months following the interventions as patients recovered. This was expected to be particularly true of the interventions including surgery.

Cost effectiveness and resource use

An original health economic model was developed to answer this question (the full modelling report is available in Appendix K). Outcomes in the first five years of this model were calculated via the network meta-analyses conducted for this guideline (Appendix I), which showed that chemoradiotherapy and surgery (CRS) was associated with a statistically significantly longer progression free survival time than chemoradiotherapy alone (CR) and that CRS showed a high probability of being associated with the greatest overall survival. After the first five years, it was assumed

that those patients who were still alive would continue progression free until the end of the model. Their overall survival was estimated using data from an epidemiological dataset on NSCLC stage IIIA-N2 patients who had survived five years after diagnosis.

The model found that while CRS was the most expensive intervention, it was also the most cost-effective, with a base case ICER of less than £20,000/QALY gained versus CR. Chemotherapy and surgery (CS) was extendedly dominated by the combination of CRS and CR and was itself not cost-effective compared to CR with highly uncertain ICERs that were consistently above £30,000/QALY gained in sensitivity analyses.

The committee discussed the limitations of the model and the assumptions that had been needed through lack of high quality directly available data and decided that the analysis was robust for decision making purposes because its results were quite insensitive to realistic variations in uncertain data and assumptions. They noted, however, that none of the RCTs included in the NMAs found any difference in overall survival, which was the most important outcome. Taking all the above considerations together, they decided that a 'consider' recommendation in favour of CRS was justified by the evidence. This is because while they thought that CRS is likely to be the most cost-effective intervention and that CS was unlikely to be cost-effective compared to the other two interventions, there were a number of key uncertainties in the clinical data.

Surgery and radical radiotherapy are expensive interventions, costing approximately £7,500 and £2,500 respectively. The committee thought that only a small number of stage IIIA-N2 patients are currently treated with CRS and that these recommendations therefore represent an increase in resource use, which will depend on the extent of take-up.

Other factors the committee took into account

The committee noted that none of the trials underpinning the network meta-analysis and health economic model were conducted in a UK setting and many recruited before the widespread adoption of newer and more effective treatments for advanced NSCLC such as targeted and immunotherapies. There have also been significant innovations in surgery and radiotherapy techniques in recent years. The survival data might therefore not reflect outcomes that would be seen in UK practice today although none of these things in themselves provide reasons to reject the differential effectiveness observed in the network meta-analyses.

The committee noted that patient fitness and patient choice were important factors in deciding between interventions and tried to reflect this in their recommendations. The recommendations for a 3-5 week wait between CR and surgery reflect current clinical practice. This is similar to the waiting period between CR and surgery in the most relevant studies: Pless 2015, 21-28 days; Katakami 2012, 3-5 weeks; Albain 2009, 3-5 weeks; Eberhardt 2015, median of 37 days (20-61 day range); Girard 2010, 4-6 weeks.

Appendix A – Review protocols

Review protocol for the clinical and cost effectiveness of chemoradiotherapy or surgery with adjuvant treatment for the treatment for N2 stage NSCLC

Field (based on PRISMA-P	Content
Review question	What is the clinical and cost effectiveness of chemoradiotherapy or surgery with adjuvant treatment for the treatment for N2 stage NSCLC?
Type of review question	Intervention
Objective of the review	To provide clearer guidance regarding the treatment of N2 stage NSCLC. This question was identified during scoping meeting 2. Variation in practice has also been identified.
Eligibility criteria – population/ disease/ condition/ issue/ domain	People with stage N2 M0 NSCLC.
Eligibility criteria –	Surgery with/ without chemotherapy

intervention(s)/ exposure(s)/ prognostic factor(s)	
Eligibility criteria – comparator(s)/ control or reference (gold) standard	1. Chemoradiotherapy (radiotherapy and chemotherapy) versus 2. Tri-modality treatment

Outcomes and prioritisation	• Mortality
	  ○ Cancer-related
	  ○ Treatment-related
	  ○ All-cause
	• Quality of life (as measured by QoL instrument, for example)
	  ○ ECOG score
	  ○ EORTC score
	  ○ EQ-5D
	• Length of stay
	  ○ hospital
	  ○ ICU
	• Exercise tolerance
	• Adverse events
	  ○ Oesophagitis, pneumonitis, sepsis (grading)
	  ○ Dyspnoea

	○ Hypoxia and need for home oxygen ○ Stroke ○ Cardiovascular disease • Treatment-related dropout rates • Pain (continuous pain scales and/ or proportions of people in pain)
Eligibility criteria – study design	• RCT data. • Systematic reviews of RCTs
Other inclusion exclusion criteria	• Non English-language papers • Unpublished evidence/ conference proceedings
Proposed sensitivity/sub-group analysis, or meta-regression	No subgroup analysis identified
Selection process – duplicate screening/selection/analysis	10% of the abstracts were reviewed by two reviewers, with any disagreements resolved by discussion or, if necessary, a third independent reviewer. If meaningful disagreements were found between the different reviewers, a further 10% of the abstracts were reviewed by two reviewers, with this process continued until agreement is achieved between the two reviewers. From this point, the remaining abstracts will be screened by a single reviewer.

	This review made use of the priority screening functionality with the EPPI-reviewer systematic reviewing software. See Appendix B for more details.
Data management (software)	See appendix B.
Information sources – databases and dates	No date limit. See appendix C. Main Searches: • Cochrane Database of Systematic Reviews – CDSR • Cochrane Central Register of Controlled Trials – CENTRAL • Database of Abstracts of Reviews of Effects – DARE • Health Technology Assessment Database – HTA • EMBASE (Ovid) • MEDLINE (Ovid) • MEDLINE In-Process (Ovid) Citation searching will be carried out in addition on analyst/committee selected papers. The search will not be date limited because this is a new review question.

Identify if an update	Update. Original Question (linked): What is the most effective treatment for patients with resectable non-small cell lung cancer? **Recommendations that may be affected:** 1.4.27 Patients with stage I or II NSCLC who are medically inoperable but suitable for radical radiotherapy should be offered the CHART regimen. **[2005]**
Author contacts	Guideline update
Highlight if amendment to previous protocol	For details please see section 4.5 of Developing NICE guidelines: the manual
Search strategy – for one database	For details please see appendix C
Data collection process – forms/ duplicate	A standardised evidence table format will be used, and published as appendix G (clinical evidence tables) or H (economic evidence tables) of the full guideline.

Data items – define all variables to be collected	For details please see evidence tables in appendix G (clinical evidence tables) or H (economic evidence tables) of the full guideline.
Methods for assessing bias at outcome/study level	Standard study checklists were used to critically appraise individual studies. For details please see section 6.2 of Developing NICE guidelines: the manual The risk of bias across all available evidence was evaluated for each outcome using an adaptation of the 'Grading of Recommendations Assessment, Development and Evaluation (GRADE) toolbox' developed by the international GRADE working group http://www.gradeworkinggroup.org/ For further detail see Appendix B.
Criteria for quantitative synthesis (where suitable)	For details please see section 6.4 of Developing NICE guidelines: the manual
Methods for analysis – combining studies and exploring (in)consistency	For details please see the methods chapter of the full guideline. See appendix B.

Meta-bias assessment – publication bias, selective reporting bias	For details please see section 6.2 of Developing NICE guidelines: the manual. See appendix B.
Assessment of confidence in cumulative evidence	For details please see sections 6.4 and 9.1 of Developing NICE guidelines: the manual See appendix B.
Rationale/ context – Current management	For details please see the introduction to the evidence review in the full guideline.
Describe contributions of authors and guarantor	A multidisciplinary committee developed the guideline. The committee was convened by NICE Guideline Updates Team and chaired by Gary McVeigh in line with section 3 of Developing NICE guidelines: the manual. Staff from NICE Guideline Updates Team undertook systematic literature searches, appraised the evidence, conducted meta-analysis and cost-effectiveness analysis where appropriate, and drafted the guideline in collaboration with the committee. For details please see the methods chapter of the full guideline.

Sources of funding/support	The NICE Guideline Updates Team is an internal team within NICE.
Name of sponsor	The NICE Guideline Updates Team is an internal team within NICE.
Roles of sponsor	The NICE Guideline Updates Team is an internal team within NICE.
PROSPERO registration number	N/A

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Appendix B – Methods

Priority screening

The reviews undertaken for this guideline all made use of the priority screening functionality with the EPPI-reviewer systematic reviewing software. This uses a machine learning algorithm (specifically, an SGD classifier) to take information on features (1, 2 and 3 word blocks) in the titles and abstract of papers marked as being 'includes' or 'excludes' during the title and abstract screening process, and re-orders the remaining records from most likely to least likely to be an include, based on that algorithm. This re-ordering of the remaining records occurs every time 25 additional records have been screened.

Research is currently ongoing as to what are the appropriate thresholds where reviewing of abstract can be stopped, assuming a defined threshold for the proportion of relevant papers it is acceptable to miss on primary screening. As a conservative approach until that research has been completed, the following rules were adopted during the production of this guideline:

- In every review, at least 50% of the identified abstract (or 1,000 records, if that is a greater number) were always screened.
- After this point, screening was only terminated when the threshold was reached for a number of abstracts being screened without a single new include being identified. This threshold was set according to the expected proportion of includes in the review (with reviews with a lower proportion of includes needing a higher number of papers without an identified study to justify termination), and was always a minimum of 250.
- A random 10% sample of the studies remaining in the database when the threshold were additionally screened, to check if a substantial number of relevant studies were not being correctly classified by the algorithm, with the full database being screened if concerns were identified.

As an additional check to ensure this approach did not miss relevant studies, the included studies lists of included systematic reviews were searched to identify any papers not identified through the primary search.

Incorporating published systematic reviews

For all review questions where a literature search was undertaken looking for a particular study design, systematic reviews containing studies of that design were also included. All included studies from those systematic reviews were screened to identify any additional relevant primary studies not found as part of the initial search.

Quality assessment

Individual systematic reviews were quality assessed using the ROBIS tool, with each classified into one of the following three groups:

- High quality – It is unlikely that additional relevant and important data would be identified from primary studies compared to that reported in the review, and unlikely that any relevant and important studies have been missed by the review.
- Moderate quality – It is possible that additional relevant and important data would be identified from primary studies compared to that reported in the review, but unlikely that any relevant and important studies have been missed by the review.

- Low quality – It is possible that relevant and important studies have been missed by the review.

Each individual systematic review was also classified into one of three groups for its applicability as a source of data, based on how closely the review matches the specified review protocol in the guideline. Studies were rated as follows:

- Fully applicable – The identified review fully covers the review protocol in the guideline.
- Partially applicable – The identified review fully covers a discrete subsection of the review protocol in the guideline (for example, some of the factors in the protocol only).
- Not applicable – The identified review, despite including studies relevant to the review question, does not fully cover any discrete subsection of the review protocol in the guideline.

Using systematic reviews as a source of data

If systematic reviews were identified as being sufficiently applicable and high quality, and were identified sufficiently early in the review process (for example, from the surveillance review or early in the database search), they were used as the primary source of data, rather than extracting information from primary studies. The extent to which this was done depended on the quality and applicability of the review, as defined in Table 2. When systematic reviews were used as a source of primary data, and unpublished or additional data included in the review which is not in the primary studies was also included. Data from these systematic reviews was then quality assessed and presented in GRADE/CERQual tables as described below, in the same way as if data had been extracted from primary studies. In questions where data was extracted from both systematic reviews and primary studies, these were cross-referenced to ensure none of the data had been double counted through this process.

Table 2: Criteria for using systematic reviews as a source of data

Quality	Applicability	Use of systematic review
High	Fully applicable	Data from the published systematic review were used instead of undertaking a new literature search or data analysis. Searches were only done to cover the period of time since the search date of the review.
High	Partially applicable	Data from the published systematic review were used instead of undertaking a new literature search and data analysis for the relevant subsection of the protocol. For this section, searches were only done to cover the period of time since the search date of the review. For other sections not covered by the systematic review, searches were undertaken as normal.
Moderate	Fully applicable	Details of included studies were used instead of undertaking a new literature search. Full-text papers of included studies were still retrieved for the purposes of data analysis. Searches were only done to cover the period of time since the search date of the review.
Moderate	Partially applicable	Details of included studies were used instead of undertaking a new literature search for the relevant subsection of the protocol. For this section, searches were only done to cover the period of time since the search date of the review. For other sections not covered by the systematic review, searches were undertaken as normal.

1.3 Evidence synthesis and meta-analyses

Where possible, meta-analyses were conducted to combine the results of quantitative studies for each outcome. For continuous outcomes analysed as mean differences, where change from baseline data were reported in the trials and were accompanied by a measure of spread (for example standard deviation), these were extracted and used in the meta-analysis. Where measures of spread for change from baseline values were not reported, the corresponding values at study end were used and were combined with change from baseline values to produce summary estimates of effect. These studies were assessed to ensure that baseline values were balanced across the treatment groups; if there were significant differences at baseline these studies were not included in any meta-analysis and were reported separately. For continuous outcomes analysed as standardised mean differences, where only baseline and final time point values were available, change from baseline standard deviations were estimated, assuming a correlation coefficient of 0.5.

1.4 Evidence of effectiveness of interventions

1.4.1 Quality assessment

Individual RCTs and quasi-randomised controlled trials were quality assessed using the Cochrane Risk of Bias Tool. Other study were quality assessed using the ROBINS-I tool. Each individual study was classified into one of the following three groups:

- Low risk of bias – The true effect size for the study is likely to be close to the estimated effect size.
- Moderate risk of bias – There is a possibility the true effect size for the study is substantially different to the estimated effect size.
- High risk of bias – It is likely the true effect size for the study is substantially different to the estimated effect size.

Each individual study was also classified into one of three groups for directness, based on if there were concerns about the population, intervention, comparator and/or outcomes in the study and how directly these variables could address the specified review question. Studies were rated as follows:

- Direct – No important deviations from the protocol in population, intervention, comparator and/or outcomes.
- Partially indirect – Important deviations from the protocol in one of the population, intervention, comparator and/or outcomes.
- Indirect – Important deviations from the protocol in at least two of the following areas: population, intervention, comparator and/or outcomes.

1.4.2 Methods for combining intervention evidence

Meta-analyses of interventional data were conducted with reference to the Cochrane Handbook for Systematic Reviews of Interventions (Higgins et al. 2011).

Where different studies presented continuous data measuring the same outcome but using different numerical scales (e.g. a 0-10 and a 0-100 visual analogue scale), these outcomes were all converted to the same scale before meta-analysis was conducted on the mean differences. Where outcomes measured the same underlying construct but used different instruments/metrics, data were analysed using standardised mean differences (Hedges' g).

A pooled relative risk was calculated for dichotomous outcomes (using the Mantel–Haenszel method) reporting numbers of people having an event, and a pooled incidence rate ratio was calculated for dichotomous outcomes reporting total numbers of events. Both relative and absolute risks were presented, with absolute risks calculated by applying the relative risk to the pooled risk in the comparator arm of the meta-analysis (all pooled trials).

Fixed- and random-effects models (der Simonian and Laird) were fitted for all syntheses, with the presented analysis dependent on the degree of heterogeneity in the assembled evidence. Fixed-effects models were the preferred choice to report, but in situations where the assumption of a shared mean for fixed-effects model were clearly not met, even after appropriate pre-specified subgroup analyses were conducted, random-effects results are presented. Fixed-effects models were deemed to be inappropriate if one or both of the following conditions was met:

- Significant between study heterogeneity in methodology, population, intervention or comparator was identified by the reviewer in advance of data analysis. This decision was made and recorded before any data analysis was undertaken.
- The presence of significant statistical heterogeneity in the meta-analysis, defined as $I^2 \geq 50\%$.

In any meta-analyses where some (but not all) of the data came from studies at high risk of bias, a sensitivity analysis was conducted, excluding those studies from the analysis. Results from both the full and restricted meta-analyses are reported. Similarly, in any meta-analyses where some (but not all) of the data came from indirect studies, a sensitivity analysis was conducted, excluding those studies from the analysis.

Meta-analyses were performed in Cochrane Review Manager V5.3, with the exception of incidence rate ratio analyses which were carried out in R version 3.3.4.

Minimal clinically important differences (MIDs)

The Core Outcome Measures in Effectiveness Trials (COMET) database was searched to identify published minimal clinically important difference thresholds relevant to this guideline. However, no relevant MIDs were found. In addition, the Guideline Committee were asked to specify any outcomes where they felt a consensus MID could be defined from their experience. In particular, any questions looking to evaluate non-inferiority (that one intervention is not meaningfully worse than another) required an MID to be defined to act as a non-inferiority margin. However, the committee agreed that in their experience, they could not define any MIDs. This is because the committee were not aware of evidence supporting the use of MIDs for the protocol's outcomes. Therefore, the line of no effect was used as the MID for risk ratios, hazard ratios and mean differences.

GRADE for pairwise meta-analyses of interventional evidence

GRADE was used to assess the quality of evidence for the selected outcomes as specified in 'Developing NICE guidelines: the manual (2014)'. Data from all study designs was initially rated as high quality and the quality of the evidence for each outcome was downgraded or not from this initial point, based on the criteria given in **Error! Reference source not found.**

198 **Table 3: Rationale for downgrading quality of evidence for intervention studies**

GRADE criteria	Reasons for downgrading quality
Risk of bias	Not serious: If less than 33.3% of the weight in a meta-analysis came from studies at moderate or high risk of bias, the overall outcome was not downgraded. Serious: If greater than 33.3% of the weight in a meta-analysis came from studies at moderate or high risk of bias, the outcome was downgraded one level. Very serious: If greater than 33.3% of the weight in a meta-analysis came from studies at high risk of bias, the outcome was downgraded two levels. Outcomes meeting the criteria for downgrading above were not downgraded if there was evidence the effect size was not meaningfully different between studies at high and low risk of bias.
Indirectness	Not serious: If less than 33.3% of the weight in a meta-analysis came from partially indirect or indirect studies, the overall outcome was not downgraded. Serious: If greater than 33.3% of the weight in a meta-analysis came from partially indirect or indirect studies, the outcome was downgraded one level. Very serious: If greater than 33.3% of the weight in a meta-analysis came from indirect studies, the outcome was downgraded two levels. Outcomes meeting the criteria for downgrading above were not downgraded if there was evidence the effect size was not meaningfully different between direct and indirect studies.
Inconsistency	Concerns about inconsistency of effects across studies, occurring when there is unexplained variability in the treatment effect demonstrated across studies (heterogeneity), after appropriate pre-specified subgroup analyses have been conducted. This was assessed using the I^2 statistic. N/A: Inconsistency was marked as not applicable if data on the outcome was only available from one study. Not serious: If the I^2 was less than 33.3%, the outcome was not downgraded. Serious: If the I^2 was between 33.3% and 66.7%, the outcome was downgraded one level. Very serious: If the I^2 was greater than 66.7%, the outcome was downgraded two levels. Outcomes meeting the criteria for downgrading above were not downgraded if there was evidence the effect size was not meaningfully different between studies with the smallest and largest effect sizes.
Imprecision	If an MID other than the line of no effect was defined for the outcome, the outcome was downgraded once if the 95% confidence interval for the effect size crossed one line of the MID, and twice if it crosses both lines of the MID. If the line of no effect was defined as an MID for the outcome, it was downgraded once if the 95% confidence interval for the effect size crossed the line of no effect (i.e. the outcome was not statistically significant), and twice if the sample size of the study was sufficiently small that it is not plausible any realistic effect size could have been detected. Outcomes meeting the criteria for downgrading above were not downgraded if the confidence interval was sufficiently narrow that the upper and lower bounds would correspond to clinically equivalent scenarios.

199 The quality of evidence for each outcome was upgraded if any of the following three
200 conditions were met:

201 • Data from non-randomised studies showing an effect size sufficiently large that it cannot
202 be explained by confounding alone.

- Data showing a dose-response gradient.
- Data where all plausible residual confounding is likely to increase our confidence in the effect estimate.

Publication bias

Publication bias was assessed in two ways. First, if evidence of conducted but unpublished studies was identified during the review (e.g. conference abstracts, trial protocols or trial records without accompanying published data), available information on these unpublished studies was reported as part of the review. Secondly, where 10 or more studies were included as part of a single meta-analysis, a funnel plot was produced to graphically assess the potential for publication bias.

Evidence statements

Evidence statements for pairwise intervention data are classified in to one of four categories:

- Situations where the data are only consistent, at a 95% confidence level, with an effect in one direction (i.e. one that is 'statistically significant'), and the magnitude of that effect is most likely to meet or exceed the MID (i.e. the point estimate is not in the zone of equivalence). In such cases, we state that the evidence showed that there is an effect.
- Situations where the data are only consistent, at a 95% confidence level, with an effect in one direction (i.e. one that is 'statistically significant'), but the magnitude of that effect is most likely to be less than the MID (i.e. the point estimate is in the zone of equivalence). In such cases, we state that the evidence could not demonstrate a meaningful difference.
- Situations where the confidence limits are smaller than the MIDs in both directions. In such cases, we state that the evidence demonstrates that there is no meaningful difference.
- In all other cases, we state that the evidence could not differentiate between the comparators.

For outcomes without a defined MID or where the MID is set as the line of no effect (for example, in the case of mortality), evidence statements are divided into 2 groups as follows:

- We state that the evidence showed that there is an effect if the 95% CI does not cross the line of no effect.
- The evidence could not differentiate between comparators if the 95% CI crosses the line of no effect.

Methods for combining direct and indirect evidence (network meta-analysis) for interventions

Conventional ‘pairwise’ meta-analysis involves the statistical combination of direct evidence about pairs of interventions that originate from two or more separate studies (for example, where there are two or more studies comparing A vs B).

In situations where there are more than two interventions, pairwise meta-analysis of the direct evidence alone is of limited use. This is because multiple pairwise comparisons need to be performed to analyse each pair of interventions in the evidence, and these results can be difficult to interpret. Furthermore, direct evidence about interventions of interest may not be available. For example studies may compare A vs B and B vs C, but there may be no

direct evidence comparing A vs C. Network meta-analysis overcomes these problems by combining all evidence into a single, internally consistent model, synthesising data from direct and indirect comparisons, and providing estimates of relative effectiveness for all comparators and the ranking of different interventions. Network meta-analyses were undertaken in all situations where the following three criteria were met:

- At least three treatment alternatives.
- A sufficiently connected network to enable valid estimates to be made.
- The aim of the review was to produce recommendations on the most effective option, rather than simply an unordered list of treatment alternatives.

Synthesis

Two separate frameworks and software packages were used for undertaking network-meta analyses in this guideline, with the chosen method dependent on the specifics of the question (for certain datasets, it may be possible to run the preferred analysis in one program but not the other, or it may be particularly more efficient to use one package over another):

- Hierarchical Bayesian Network Meta-Analysis (NMA) was performed using WinBUGS version 1.4.3. The models used reflected the recommendations of the NICE Decision Support Unit's Technical Support Documents (TSDs) on evidence synthesis, particularly TSD 2 ('A generalised linear modelling framework for pairwise and network meta-analysis of randomised controlled trials'; see http://www.nicedsu.org.uk). The WinBUGS code provided in the appendices of TSD 2 was used without substantive alteration to specify synthesis models.

  Results were reported summarising 10,000 samples from the posterior distribution of each model, having first run and discarded 50,000 ‘burn-in’ iterations. Three separate chains with different initial values were used.

  Non-informative prior distributions were used in all models. Unless otherwise specified, trial-specific baselines and treatment effects were assigned N(0,1000) priors, and the between-trial standard deviations used in random-effects models were given U(0,5) priors. These are consistent with the recommendations in TSD 2 for dichotomous outcomes.

  Fixed- and random-effects models were explored for each outcome, with the final choice of model based on deviance information criterion (DIC): if DIC was at least 3 points lower for the random-effects model, it was preferred; otherwise, the fixed effects model was considered to provide an equivalent fit to the data in a more parsimonious analysis, and was preferred.

  In studies where there was residual unexplained heterogeneity (defined as when a random-effects model has been preferred), consideration was given to running a bias-adjusted meta-analysis, in line with recommendations from the NICE Technical Support Unit. Such an analysis was undertaken only when sufficient data were available, meaning that there needed to be a sufficiently high ratio of studies to nodes in the network, and a sufficient number of studies at both low and high risk of bias. When conducting a bias-adjusted NMA it is necessary to dichotomise studies into high and low risk of bias, and this was done by individual studies rated as being either moderate or high risk of bias being classed under high risk of bias.

- Frequentist NMAs were undertaken using the `netmeta` package in R v3.4.0. This uses a graph-theoretical method which is mathematically equivalent to frequentist network meta-analysis (Rücker 2012). Inconsistency was assessed using the overall I^2 value for the whole network, which is a weighted average of the I^2 value for all comparisons where there are multiple trials (both direct and indirect), and random-effects models were used if

the I^2 value was above 50% (as for pairwise meta-analyses, this was interpreted as showing the assumption of consistent, shared underlying means was not met, and therefore a fixed-effects model was inappropriate).

Because different approaches and software had been applied, sensitivity analysis have previously been undertaken to establish whether this might have led to any substantive differences in output. Specimen dichotomous and continuous NMAs from the Bayesian analysis were rerun in the frequentist framework and generated results that were materially indistinguishable from the Bayesian version.

In any meta-analyses where some (but not all) of the data came from studies at high risk of bias, a sensitivity analysis was conducted, excluding those studies from the analysis. Results from both the full and restricted meta-analyses are reported. Similarly, in any meta-analyses where some (but not all) of the data came from indirect studies, a sensitivity analysis was conducted, excluding those studies from the analysis. Where sufficient studies were available, meta-regression was undertaken to explore the effect of study level covariates.

Modified GRADE for network meta-analyses

A modified version of the standard GRADE approach for pairwise interventions was used to assess the quality of evidence across the network meta-analyses undertaken. While most criteria for pairwise meta-analyses still apply, it is important to adapt some of the criteria to take into consideration additional factors, such as how each 'link' or pairwise comparison within the network applies to the others. As a result, the following was used when modifying the GRADE framework to a network meta-analysis. It is designed to provide a single overall quality rating for an NMA, which can then be combined with pairwise quality ratings for individual comparisons (if appropriate), to judge the overall strength of evidence for each comparison.

Table 4: Rationale for downgrading quality of evidence for intervention studies

GRADE criteria	Reasons for downgrading quality
Risk of bias	Not serious: If fewer than 33.3% of the studies in the network meta-analysis were at moderate or high risk of bias, the overall network was not downgraded. Serious: If greater than 33.3% of the studies in the network meta-analysis were at moderate or high risk of bias, the network was downgraded one level. Very serious: If greater than 33.3% of the studies in the network meta-analysis were at high risk of bias, the network was downgraded two levels.
Indirectness	Not serious: If fewer than 33.3% of the studies in the network meta-analysis were partially indirect or indirect, the overall network was not downgraded. Serious: If greater than 33.3% of the studies in the network meta-analysis were partially indirect or indirect, the network was downgraded one level. Very serious: If greater than 33.3% of the studies in the network meta-analysis were indirect, the network was downgraded two levels.
Inconsistency	N/A: Inconsistency was marked as not applicable if there were no links in the network where data from multiple studies (either direct or indirect) were synthesised. For network meta-analyses conducted under a Bayesian framework, the network was downgraded one level if the DIC for a random-effects model was lower than the DIC for a fixed-effects model. For network meta-analyses conducted under a frequentist framework, the network was downgraded one level if the I^2 was greater than 50%.

GRADE criteria	Reasons for downgrading quality
	In addition, under both frameworks, the direct and indirect treatment estimates were compared as a check on the consistency of the network.
Imprecision	The overall network was downgraded for imprecision if it was not possible to differentiate between any meaningfully distinct treatments options in the network (based on 95% confidence/credible intervals). Whether two options were meaningfully distinct was judged using the MIDs defined above for pairwise meta-analysis of the outcomes, if available; or statistical significance if MIDs were not available.

Quality assessment

Individual cohort and case-control studies were quality assessed using the CASP cohort study and case-control checklists, respectively. Each individual study was classified into one of the following three groups:

- Low risk of bias – The true effect size for the study is likely to be close to the estimated effect size.
- Moderate risk of bias – There is a possibility the true effect size for the study is substantially different to the estimated effect size.
- High risk of bias – It is likely the true effect size for the study is substantially different to the estimated effect size.

Individual cross-sectional studies were quality assessed using the Joanna Briggs Institute critical appraisal checklist for analytical cross sectional studies (2016), which contains 8 questions covering: inclusion criteria, description of the sample, measures of exposure, measures of outcomes, confounding factors, and statistical analysis. Each individual study was classified into one of the following groups:

- Low risk of bias – Evidence of non-serious bias in zero or one domain.
- Moderate risk of bias – Evidence of non-serious bias in two domains only, or serious bias in one domain only.
- High risk of bias – Evidence of bias in at least three domains, or of serious bias in at least two domains.

Each individual study was also classified into one of three groups for directness, based on if there were concerns about the population, predictors and/or outcomes in the study and how directly these variables could address the specified review question. Studies were rated as follows:

- Direct – No important deviations from the protocol in population, predictors and/or outcomes.
- Partially indirect – Important deviations from the protocol in one of the population, predictors and/or outcomes.
- Indirect – Important deviations from the protocol in at least two of the population, predictors and/or outcomes.

Methods for combining association studies

Where appropriate, hazard ratios were pooled using the inverse-variance method, and odds ratios were pooled using the Mantel-Haenszel method. Adjusted odds ratios from multivariate models were only pooled if the same set of predictor variables were used across multiple studies and if the same thresholds to measure predictors were used across studies.

Fixed- and random-effects models (der Simonian and Laird) were fitted for all syntheses, with the presented analysis dependent on the degree of heterogeneity in the assembled evidence. Fixed-effects models were the preferred choice to report, but in situations where the assumption of a shared mean for fixed-effects model were clearly not met, even after appropriate pre-specified subgroup analyses were conducted, random-effects results are presented. Fixed-effects models were deemed to be inappropriate if one or both of the following conditions was met:

- Significant between study heterogeneity in methodology, population, intervention or comparator was identified by the reviewer in advance of data analysis. This decision would need to be made and recorded before any data analysis is undertaken.
- The presence of significant statistical heterogeneity, defined as $I^2 \geq 50\%$.

In any meta-analyses where some (but not all) of the data came from studies at high risk of bias, a sensitivity analysis was conducted, excluding those studies from the analysis. Results from both the full and restricted meta-analyses are reported. Similarly, in any meta-analyses where some (but not all) of the data came from indirect studies, a sensitivity analysis was conducted, excluding those studies from the analysis.

Meta-analyses were performed in Cochrane Review Manager v 5.3.

Minimal clinically important differences (MIDs)

The Core Outcome Measures in Effectiveness Trials (COMET) database was searched to identify published minimal clinically important difference thresholds relevant to this guideline. Identified MIDs were assessed to ensure they had been developed and validated in a methodologically rigorous way, and were applicable to the populations, interventions and outcomes specified in this guideline. In addition, the Guideline Committee were asked to prospectively specify any outcomes where they felt a consensus MID could be defined from their experience. In particular, any questions looking to evaluate non-inferiority (that one treatment is not meaningfully worse than another) required an MID to be defined to act as a non-inferiority margin.

MIDs found through this process and used to assess imprecision in the guideline are given in Table 5.

Table 5: Identified MIDs

Outcome	MID	Source

When decisions were made in situations where MIDs were not available, the 'Evidence to Recommendations' section of that review should make explicit the committee's view of the expected clinical importance and relevance of the findings.

Modified GRADE for association studies

GRADE has not been developed for use with predictive studies; therefore a modified approach was applied using the GRADE framework. Data from cohort studies was initially

387 rated as high quality, and data from case-control studies as low quality, with the quality of the
388 evidence for each outcome then downgraded or not from this initial point.

389 **Table 6: Rationale for downgrading quality of evidence for association studies**

GRADE criteria	Reasons for downgrading quality
Risk of bias	Not serious: If less than 33.3% of the weight in a meta-analysis came from studies at moderate or high risk of bias, the overall outcome was not downgraded. Serious: If greater than 33.3% of the weight in a meta-analysis came from studies at moderate or high risk of bias, the outcome was downgraded one level. Very serious: If greater than 33.3% of the weight in a meta-analysis came from studies at high risk of bias, the outcome was downgraded two levels. Outcomes meeting the criteria for downgrading above were not downgraded if there was evidence the effect size was not meaningfully different between studies at high and low risk of bias. In addition, unadjusted odds ratio outcomes from univariate analyses were downgraded one level, in addition to any downgrading for risk of bias in individual studies. Adjusted odds ratios from multivariate analyses were not similarly downgraded.
Indirectness	Not serious: If less than 33.3% of the weight in a meta-analysis came from partially indirect or indirect studies, the overall outcome was not downgraded. Serious: If greater than 33.3% of the weight in a meta-analysis came from partially indirect or indirect studies, the outcome was downgraded one level. Very serious: If greater than 33.3% of the weight in a meta-analysis came from indirect studies, the outcome was downgraded two levels. Outcomes meeting the criteria for downgrading above were not downgraded if there was evidence the effect size was not meaningfully different between direct and indirect studies.
Inconsistency	Concerns about inconsistency of effects across studies, occurring when there is unexplained variability in the treatment effect demonstrated across studies (heterogeneity). This was assessed using the I^2 statistic. N/A: Inconsistency was marked as not applicable if data on the outcome was only available from one study. Not serious: If the I^2 was less than 33.3%, the outcome was not downgraded. Serious: If the I^2 was between 33.3% and 66.7%, the outcome was downgraded one level. Very serious: If the I^2 was greater than 66.7%, the outcome was downgraded two levels. Outcomes meeting the criteria for downgrading above were not downgraded if there was evidence the effect size was not meaningfully different between studies with the smallest and largest effect sizes.
Imprecision	If an MID other than the line of no effect was defined for the outcome, the outcome was downgraded once if the 95% confidence interval for the effect size crossed one line of the MID, and twice if it crosses both lines of the MID. If the line of no effect was defined as an MID for the outcome, it was downgraded once if the 95% confidence interval for the effect size crossed the line of no effect (i.e. the outcome was not statistically significant), and twice if the sample size of the study was sufficiently small that it is not plausible any realistic effect size could have been detected.

GRADE criteria	Reasons for downgrading quality
	Outcomes meeting the criteria for downgrading above were not downgraded if the confidence interval was sufficiently narrow that the upper and lower bounds would correspond to clinically equivalent scenarios.

The quality of evidence for each outcome was upgraded if either of the following conditions were met:

- Data showing an effect size sufficiently large that it cannot be explained by confounding alone.
- Data where all plausible residual confounding is likely to increase our confidence in the effect estimate.

Publication bias

Publication bias was assessed in two ways. First, if evidence of conducted but unpublished studies was identified during the review (e.g. conference abstracts or protocols without accompanying published data), available information on these unpublished studies was reported as part of the review. Secondly, where 10 or more studies were included as part of a single meta-analysis, a funnel plot was produced to graphically assess the potential for publication bias.

1.6 Health economics

Literature reviews seeking to identify published cost–utility analyses of relevance to the issues under consideration were conducted for all questions. In each case, the search undertaken for the clinical review was modified, retaining population and intervention descriptors, but removing any study-design filter and adding a filter designed to identify relevant health economic analyses. In assessing studies for inclusion, population, intervention and comparator, criteria were always identical to those used in the parallel clinical search; only cost–utility analyses were included. Economic evidence profiles, including critical appraisal according to the Guidelines manual, were completed for included studies.

Economic studies identified through a systematic search of the literature are appraised using a methodology checklist designed for economic evaluations (NICE guidelines manual; 2014). This checklist is not intended to judge the quality of a study per se, but to determine whether an existing economic evaluation is useful to inform the decision-making of the committee for a specific topic within the guideline.

There are 2 parts of the appraisal process. The first step is to assess applicability (that is, the relevance of the study to the specific guideline topic and the NICE reference case); evaluations are categorised according to the criteria in Table 7.

Table 7 Applicability criteria

Level	Explanation
Directly applicable	The study meets all applicability criteria, or fails to meet one or more applicability criteria but this is unlikely to change the conclusions about cost effectiveness
Partially applicable	The study fails to meet one or more applicability criteria, and this could change the conclusions about cost effectiveness

Level	Explanation
Not applicable	The study fails to meet one or more applicability criteria, and this is likely to change the conclusions about cost effectiveness. These studies are excluded from further consideration

In the second step, only those studies deemed directly or partially applicable are further assessed for limitations (that is, methodological quality); see categorisation criteria in Table 8.

Table 8 Methodological criteria

Level	Explanation
Minor limitations	Meets all quality criteria, or fails to meet one or more quality criteria but this is unlikely to change the conclusions about cost effectiveness
Potentially serious limitations	Fails to meet one or more quality criteria and this could change the conclusions about cost effectiveness
Very serious limitations	Fails to meet one or more quality criteria and this is highly likely to change the conclusions about cost effectiveness. Such studies should usually be excluded from further consideration

Where relevant, a summary of the main findings from the systematic search, review and appraisal of economic evidence is presented in an economic evidence profile alongside the clinical evidence.

Appendix C – Literature search strategies

Scoping search strategies

Scoping searches Scoping searches were undertaken on the following websites and databases (listed in alphabetical order) in April 2017 to provide information for scope development and project planning. Browsing or simple search strategies were employed.

Guidelines/website
American Cancer Society
American College of Chest Physicians
American Society for Radiation Oncology
American Thoracic Society
Association for Molecular Pathology
British Lung Foundation
British Thoracic Society
Canadian Medical Association Infobase
Canadian Task Force on Preventive Health Care
Cancer Australia
Cancer Care Ontario
Cancer Control Alberta
Cancer Research UK
Care Quality Commission
College of American Pathologists
Core Outcome Measures in Effectiveness Trials (COMET)
Department of Health & Social Care
European Respiratory Society
European Society for Medical Oncology
European Society of Gastrointestinal Endoscopy
European Society of Thoracic Surgery
General Medical Council
Guidelines & Audit Implementation Network (GAIN)
Guidelines International Network (GIN)
Healthtalk Online
International Association for the Study of Lung Cancer
MacMillan Cancer Support
Medicines and Products Regulatory Agency (MHRA)
National Audit Office
National Cancer Intelligence Network
National Clinical Audit and Patient Outcomes Programme
National Health and Medical Research Council - Australia
National Institute for Health and Care Excellence (NICE) - published & in development guidelines
National Institute for Health and Care Excellence (NICE) - Topic Selection
NHS Choices
NHS Digital
NHS England
NICE Clinical Knowledge Summaries (CKS)
NICE Evidence Search

Guidelines/website
Office for National Statistics
Patient UK
PatientVoices
Public Health England
Quality Health
Royal College of Anaesthetists
Royal College of General Practitioners
Royal College of Midwives
Royal College of Nursing
Royal College of Pathologists
Royal College of Physicians
Royal College of Radiologists
Royal College of Surgeons
Scottish Government
Scottish Intercollegiate Guidelines Network (SIGN)
UK Data Service
US National Guideline Clearinghouse
Walsall community Health NHS Trust
Welsh Government

Clinical search literature search strategy

Main searches

Bibliographic databases searched for the guideline

- Cochrane Database of Systematic Reviews – CDSR (Wiley)
- Cochrane Central Register of Controlled Trials – CENTRAL (Wiley)
- Database of Abstracts of Reviews of Effects – DARE (Wiley)
- Health Technology Assessment Database – HTA (Wiley)
- EMBASE (Ovid)
- MEDLINE (Ovid)
- MEDLINE Epub Ahead of Print (Ovid)
- MEDLINE In-Process (Ovid)

Identification of evidence for review questions

The searches were conducted between October 2017 and April 2018 for 9 review questions (RQ).

Searches were re-run in May 2018.

Where appropriate, in-house study design filters were used to limit the retrieval to, for example, randomised controlled trials. Details of the study design filters used can be found in section 3.

Search strategy

Medline Strategy, searched 26th February 2018 Database: Ovid MEDLINE(R) 1946 to Present with Daily Update Search Strategy:
1 exp Lung Neoplasms/
2 ((lung* or pulmonary or bronch*) adj3 (cancer* or neoplasm* or carcinoma* or tumo?r* or lymphoma* or metast* or malignan* or blastoma* or carcinogen* or adenocarcinoma* or angiosarcoma* or chrondosarcoma* or sarcoma* or teratoma* or microcytic*)).tw.
3 ((pancoast* or superior sulcus or pulmonary sulcus) adj4 (tumo?r* or syndrome*)).tw.
4 ((lung* or pulmonary or bronch*) adj4 (oat or small or non-small) adj4 cell*).tw.
5 (SCLC or NSCLC).tw.
6 or/1-5
7 (N2* or cN2* or pN2* or ypN2* or T*N2* or N0-2* or IIIA* or cIIIA* or IIIB*).tw.
8 (stag* adj3 (three or III or four or IV or late* or advance*)).tw.
9 (stag* adj3 ("3" or "4")).tw.
10 (local* advanc* adj3 (non-small or NSCLC)).tw.
11 LA-NSCLC.tw.
12 Mediastinum/
13 Mediastinal Neoplasms/
14 (mediastin* or subcarinal).tw.
15 or/7-14
16 Thoracic Surgery/
17 Thoracic Surgical Procedures/
18 Pulmonary Surgical Procedures/
19 Pneumonectomy/
20 Thoracotomy/
21 exp Thoracoscopy/
22 ((lung* or pulmonary or bronch* or thorax or thorac*) adj4 (surg* or operation* or reoperation* or resection* or excision*)).tw.
23 (surg* adj1 resection*).tw.
24 (pneumonectom* or pneumoresect* or pulmonectom* or thoracotom* or pleuracotom* or pleurotom* or pleuroscop* or rethoracotom* or pneumolobectom* or segmentectom* or thoracoscop* or videothoracoscop* or bilobectom*).tw.
25 (EPP or PNE or VATS).tw.
26 (pleura* adj4 (endoscop* or incision*)).tw.
27 ((lung* or pulmonary or bronch*) adj4 lobect*).tw.
28 ((wedge or triangl*) adj4 (resect* or excision*)).tw.
29 or/16-28
30 exp Chemoradiotherapy/
31 (chemoradiotherap* or radiochemotherap* or chemoradiation*).tw.
32 (CRT or CRTx or CCRT or NCRT or RCTx or RT-CT or chemoRT).tw.
33 Combined Modality Therapy/
34 (combine* adj4 modal* adj4 (treat* or therap* or regimen* or manag* or intervention*)).tw.
35 ((tri-modal* or trimodal* or multi-modal* or multimodal*) adj4 (treat* or therap* or regimen* or manag* or intervention*)).tw.
36 TMT.tw.
37 or/30-36
38 29 or 37

Medline Strategy, searched 26th February 2018 Database: Ovid MEDLINE(R) 1946 to Present with Daily Update Search Strategy:
39 6 and 15 and 38
40 Animals/ not Humans/
41 39 not 40
42 limit 41 to english language

Note: In-house RCT and systematic review filters were appended. No date limit was used due to additional terminology to that in the searches carried out in the 2011 guideline update.

Study Design Filters

The MEDLINE SR, RCT, and observational studies filters are presented below.
Systematic Review
1. Meta-Analysis.pt.
2. Meta-Analysis as Topic/
3. Review.pt.
4. exp Review Literature as Topic/
5. (metaanaly$ or metanaly$ or (meta adj3 analy$)).tw.
6. (review$ or overview$).ti.
7. (systematic$ adj5 (review$ or overview$)).tw.
8. ((quantitative$ or qualitative$) adj5 (review$ or overview$)).tw.
9. ((studies or trial$) adj2 (review$ or overview$)).tw.
10. (integrat$ adj3 (research or review$ or literature)).tw.
11. (pool$ adj2 (analy$ or data)).tw.
12. (handsearch$ or (hand adj3 search$)).tw.
13. (manual$ adj3 search$).tw.
14. or/1-13
15. animals/ not humans/
16. 14 not 15
RCT
1 Randomized Controlled Trial.pt.
2 Controlled Clinical Trial.pt.
3 Clinical Trial.pt.
4 exp Clinical Trials as Topic/
5 Placebos/
6 Random Allocation/
7 Double-Blind Method/
8 Single-Blind Method/
9 Cross-Over Studies/
10 ((random$ or control$ or clinical$) adj3 (trial$ or stud$)).tw.
11 (random$ adj3 allocat$).tw.
12 placebo$.tw.
13 ((singl$ or doubl$ or trebl$ or tripl$) adj (blind$ or mask$)).tw.
14 (crossover$ or (cross adj over$)).tw.
15 or/1-14
16 animals/ not humans/
17 15 not 16
Observational

The MEDLINE SR, RCT, and observational studies filters are presented below.
1 Observational Studies as Topic/
2 Observational Study/
3 Epidemiologic Studies/
4 exp Case-Control Studies/
5 exp Cohort Studies/
6 Cross-Sectional Studies/
7 Controlled Before-After Studies/
8 Historically Controlled Study/
9 Interrupted Time Series Analysis/
10 Comparative Study.pt.
11 case control$.tw.
12 case series.tw.
13 (cohort adj (study or studies)).tw.
14 cohort analy$.tw.
15 (follow up adj (study or studies)).tw.
16 (observational adj (study or studies)).tw.
17 longitudinal.tw.
18 prospective.tw.
19 retrospective.tw.
20 cross sectional.tw.
21 or/1-20

Health Economics literature search strategy

Sources searched to identify economic evaluations

- NHS Economic Evaluation Database – NHS EED (Wiley) last updated Apr 2015
- Health Technology Assessment Database – HTA (Wiley) last updated Oct 2016
- Embase (Ovid)
- MEDLINE (Ovid)
- MEDLINE In-Process (Ovid)

Search filters to retrieve economic evaluations and quality of life papers were appended to the review question search strategies. For some health economics strategies additional terms were added to the original review question search strategies (see sections 4.2, 4.3 and 4.4) The searches were conducted between October 2017 and April 2018 for 9 review questions (RQ).

Searches were re-run in May 2018.

Searches were limited to those in the English language. Animal studies were removed from results.

Economic evaluation and quality of life filters

Medline Strategy
Economic evaluations
1 Economics/
2 exp "Costs and Cost Analysis"/
3 Economics, Dental/
4 exp Economics, Hospital/

Medline Strategy
5 exp Economics, Medical/
6 Economics, Nursing/
7 Economics, Pharmaceutical/
8 Budgets/
9 exp Models, Economic/
10 Markov Chains/
11 Monte Carlo Method/
12 Decision Trees/
13 econom$.tw.
14 cba.tw.
15 cea.tw.
16 cua.tw.
17 markov$.tw.
18 (monte adj carlo).tw.
19 (decision adj3 (tree$ or analys$)).tw.
20 (cost or costs or costing$ or costly or costed).tw.
21 (price$ or pricing$).tw.
22 budget$.tw.
23 expenditure$.tw.
24 (value adj3 (money or monetary)).tw.
25 (pharmacoeconomic$ or (pharmaco adj economic$)).tw.
26 or/1-25
Quality of life
1 "Quality of Life"/
2 quality of life.tw.
3 "Value of Life"/
4 Quality-Adjusted Life Years/
5 quality adjusted life.tw.
6 (qaly$ or qald$ or qale$ or qtime$).tw.
7 disability adjusted life.tw.
8 daly$.tw.
9 Health Status Indicators/
10 (sf36 or sf 36 or short form 36 or shortform 36 or sf thirtysix or sf thirty six or shortform thirtysix or shortform thirty six or short form thirtysix or short form thirty six).tw.
11 (sf6 or sf 6 or short form 6 or shortform 6 or sf six or sfsix or shortform six or short form six).tw.
12 (sf12 or sf 12 or short form 12 or shortform 12 or sf twelve or sftwelve or shortform twelve or short form twelve).tw.
13 (sf16 or sf 16 or short form 16 or shortform 16 or sf sixteen or sfsixteen or shortform sixteen or short form sixteen).tw.
14 (sf20 or sf 20 or short form 20 or shortform 20 or sf twenty or sftwenty or shortform twenty or short form twenty).tw.
15 (euroqol or euro qol or eq5d or eq 5d).tw.
16 (qol or hql or hqol or hrqol).tw.
17 (hye or hyes).tw.
18 health$ year$ equivalent$.tw.
19 utilit$.tw.

Medline Strategy
20 (hui or hui1 or hui2 or hui3).tw.
21 disutili$.tw.
22 rosser.tw.
23 quality of wellbeing.tw.
24 quality of well-being.tw.
25 qwb.tw.
26 willingness to pay.tw.
27 standard gamble$.tw.
28 time trade off.tw.
29 time tradeoff.tw.
30 tto.tw.
31 or/1-30

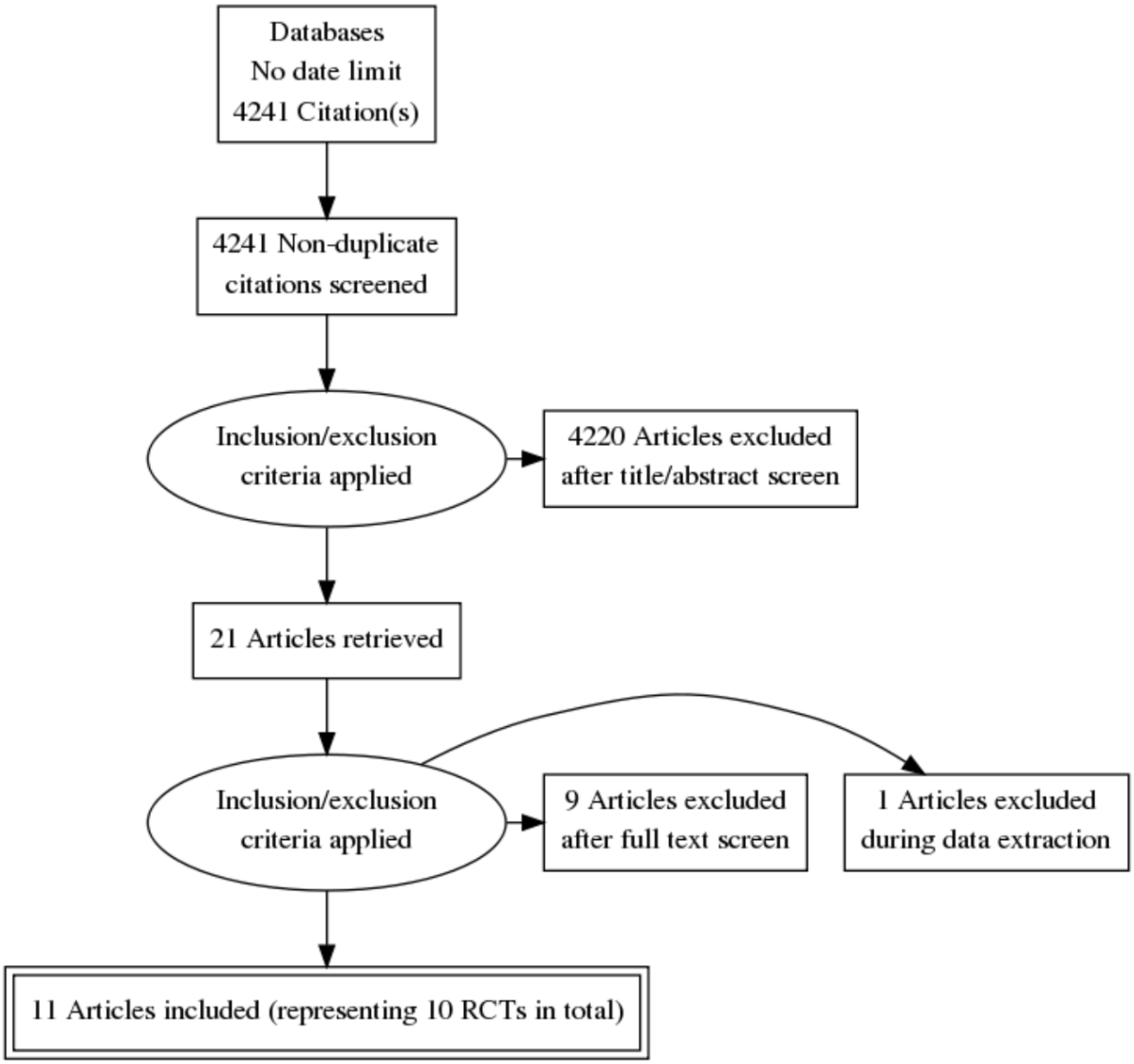
474 **Health economics search strategy**

Medline Strategy, searched 13th February 2018 Database: Ovid MEDLINE(R) 1946 to Present with Daily Update Search Strategy:
1 Small Cell Lung Carcinoma/
2 Carcinoma, Small Cell/
3 SCLC.tw.
4 ((pancoast* or superior sulcus or pulmonary sulcus) adj4 (tumo?r* or syndrome*)).tw.
5 or/1-4
6 ((small or oat or reserve or round) adj1 cell adj1 (lung* or pulmonary or bronch*) adj3 (cancer* or neoplasm* or carcinoma* or tumo?r* or lymphoma* or metast* or malignan* or blastoma* or carcinogen* or adenocarcinoma* or angiosarcoma* or chrondosarcoma* or sarcoma* or teratoma* or microcytic*)).tw.
7 (non adj1 small adj1 cell adj1 (lung* or pulmonary or bronch*) adj3 (cancer* or neoplasm* or carcinoma* or tumo?r* or lymphoma* or metast* or malignan* or blastoma* or carcinogen* or adenocarcinoma* or angiosarcoma* or chrondosarcoma* or sarcoma* or teratoma* or microcytic*)).tw.
8 6 not 7
9 5 or 8
10 exp Radiotherapy/
11 Radiation Oncology/
12 exp Radiography, Thoracic/
13 radiotherapy.fs.
14 (radiotherap* or radiotreat* or roentgentherap* or radiosurg*).tw.
15 ((radiat* or radio* or irradiat* or roentgen or x-ray or xray) adj4 (therap* or treat* or repair* or oncolog* or surg*)).tw.
16 (RT or RTx or XRT or TRT or TCRT).tw.
17 or/10-16
18 9 and 17
19 limit 18 to english language
20 Animals/ not Humans/
21 19 not 20

475
476
477

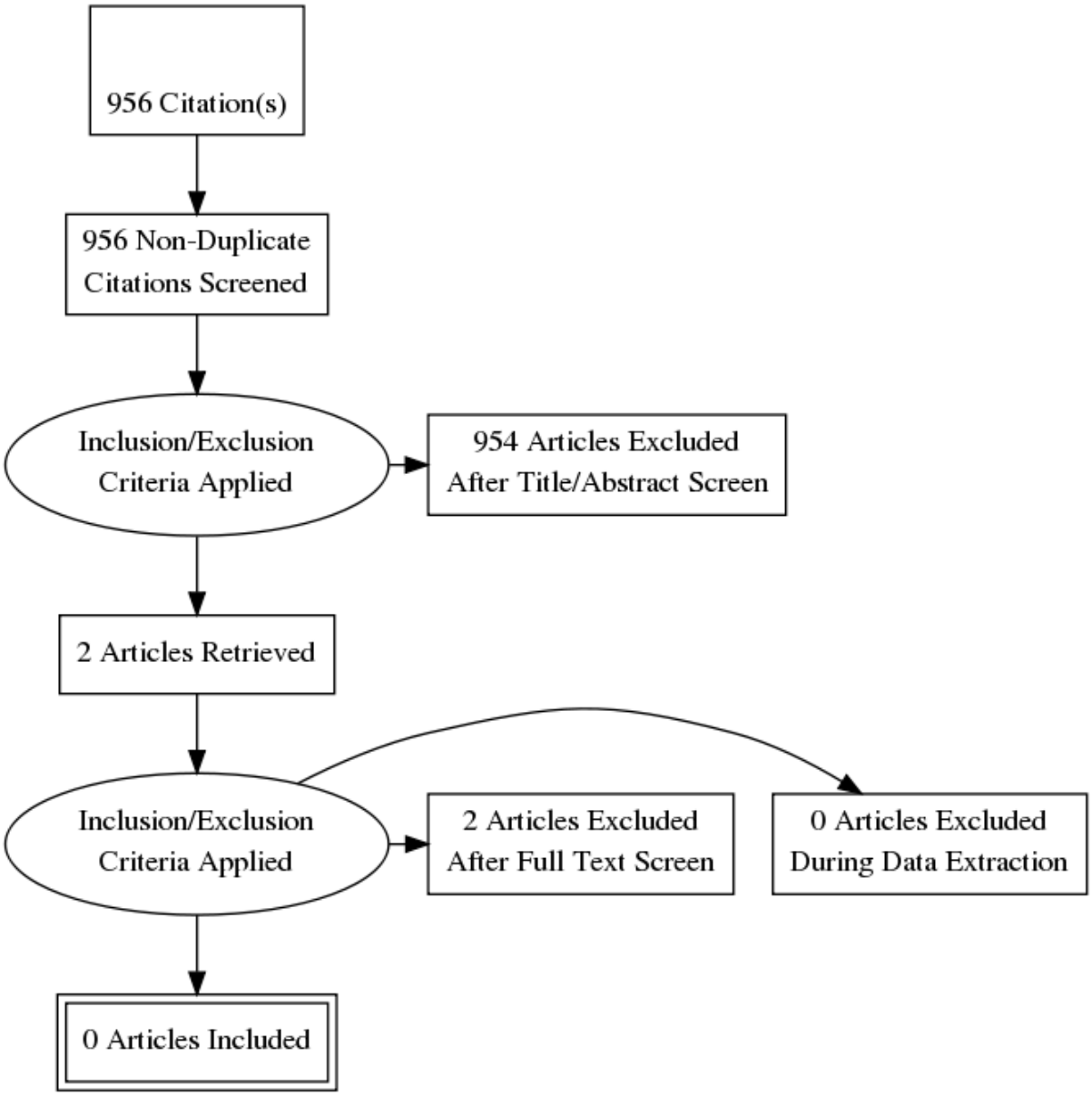
Appendix D – Evidence study selection

Clinical Evidence study selection



483 **Economic Evidence study selection**

484



485

Appendix E – Clinical evidence tables

Short Title	Title	Study Characteristics	Risk of Bias
Albain 2009	Radiotherapy plus chemotherapy with or without surgical resection for stage III non-small-cell lung cancer: a phase III randomised controlled trial	**Study type** • Randomised controlled trial **Study details** • Study location *USA and Canada* • Study setting *Hospitals* • Study dates *Recruitment was between 1994 to 2001* • Duration of follow-up *A minimum of 2.5 years. Participants were followed every 2 months for 1 year, every 3 months for 2 years, then every 6 months indefinitely. The median follow-up was 22.5 months.* • Sources of funding *National Cancer Institute and the Canadian Cancer Society.* **Inclusion criteria** • Pathologic proof of N2 involvement *All patients had stage IIIA (pN2) disease: T1, T2 or T3 primary NSCLC. If contralateral mediastinal nodes larger than 1 cm were visible on the CT scan, biopsy was required to exclude N3 (stage IIIB) disease.* • Staging CT of chest, abdomen, head *CT brain or MRI brain* • Potentially resectable **Exclusion criteria**	**Quality assessment (RCT)** Random sequence generation • Low risk of bias Allocation concealment • Unclear risk of bias No blinding. However, this is probably not possible. Blinding of participants and personnel • Unclear risk of bias No blinding. However, this is probably not possible. Blinding of outcome assessment • Unclear risk of bias No blinding. However, this is probably not possible. Incomplete outcome data • Low risk of bias Selective reporting • Low risk of bias Other sources of bias • Low risk of bias Overall risk of bias • Low

Short Title	Title	Study Characteristics	Risk of Bias
		• If overall FEV1 was less than 2000 cc, a predicted post-resection FEV1 of <800 cc • Karnofsky performance status <90 • If Karnofsky performance status 70 or 80, albumin <0.85 x normal or weight loss >10% within previous 3 months **Sample characteristics** • Sample size *396 people* • Split between study groups *Induction chemotherapy + radiotherapy, followed by surgery = 202; Induction chemotherapy + radiotherapy = 194* • Loss to follow-up *None were lost to follow-up. However, of the 202 people in the surgery arm, 9 did not have surgery. There was no explanation given.* • % female *Induction chemotherapy + radiotherapy, followed by surgery = 35.1%; Induction chemotherapy + radiotherapy = 37.6%* • Average age *Median (range): Induction chemotherapy + radiotherapy, followed by surgery = 59 (31-77); Induction chemotherapy + radiotherapy = 61 (32-78)* **Interventions** • Chemoradiotherapy, surgery *The induction chemoRT was cisplatin (50 mg/m2 days 1, 8, 29, 36), and etoposide (50 mg/m2 days 1-5 and 29-33), plus 45 Gy thoracic RT beginning day 1, in 1.8 Gy daily fractions. Disease re-evaluation by CT scan plus repeat pulmonary function tests was done 2-4 weeks after completion of RT. If there was no disease progression and the patient remained medically fit, a complete surgical resection (with protocol-specified mediastinal lymph node sampling/dissection) was performed*	Directness • Directly applicable

Short Title	Title	Study Characteristics	Risk of Bias
		3-5 weeks after completion of RT. Patients received 2 cycles of consolidation chemotherapy (same doses and schedule as during induction). Dose reduction guidelines were specified for chemoRT, with central quality control. A chest CT scan was scheduled 4-6 weeks after completion of the last chemotherapy cycle. Patients were followed every 2 months for 1 year, every 3 months for 2 years, then every 6 months indefinitely. CT scans of the thorax and upper abdomen and MRI or CT of the brain were done at 12, 18, and 24 months and annually thereafter. • Chemoradiotherapy *The induction chemoRT was cisplatin (50 mg/m2 days 1, 8, 29, 36), and etoposide (50 mg/m2 days 1-5 and 29-33), plus 45 Gy thoracic RT beginning day 1, in 1.8 Gy daily fractions. Disease re-evaluation by CT scan plus repeat pulmonary function tests was done 7 days before completion of induction chemoRT. If there was no disease progression and the patient remained medically fit, the RT was continued to 61 Gy. Patients received 2 cycles of consolidation chemotherapy (same doses and schedule as during induction). Dose reduction guidelines were specified for chemoRT, with central quality control. A chest CT scan was scheduled 4-6 weeks after completion of the last chemotherapy cycle. Patients were followed every 2 months for 1 year, every 3 months for 2 years, then every 6 months indefinitely. CT scans of the thorax and upper abdomen and MRI or CT of the brain were done at 12, 18, and 24 months and annually thereafter.* **Outcome measures** • Mortality, all-cause • Adverse events grade 3 or above	
Eberhardt 2015	Phase III Study of Surgery Versus Definitive Concurrent Chemoradiotherapy Boost in Patients	**Study type** • Randomised controlled trial **Study details** • Study location	**Quality assessment (RCT)** Random sequence generation • Low risk of bias Allocation concealment

Short Title	Title	Study Characteristics	Risk of Bias
	With Resectable Stage IIIA(N2) and Selected IIIB Non-Small-Cell Lung Cancer After Induction Chemotherapy and Concurrent Chemoradiotherapy (ESPATUE)	*Germany* • Study setting *Hospitals* • Study dates *Recruitment was from 2004 to 2013* • Duration of follow-up *Follow-up visits were scheduled every 3 months after random assignment. Follow-up was a minimum of 1 year.* • Sources of funding *German Cancer Aid* **Inclusion criteria** • Pathologically proven NSCLC • Potentially resectable stage IIIA(N2) or selected stage IIIB *N2 disease had to be pathologically proven during mediastinoscopy (recommended), endobronchial ultrasonography, or parasternal mediastinotomy. Selected resectable IIIB disease was defined as N3 disease with contralateral mediastinal nodes and proven T4 disease with involvement of the pulmonary artery, carina, left atrium, vena cava, or mediastinum. Positron emission tomographic (PET) or PET–computed tomographic staging, which was performed in 97%, and brain imaging investigations were routinely recommended.* **Exclusion criteria** • ECOG performance status 2 or above • >10% weight loss in the 6 months before diagnosis • Inadequate renal, hepatic or haematologic functions **Sample characteristics** • Sample size *161 people*	• Unclear risk of bias No blinding. However, this is probably not possible in this instance. Blinding of participants and personnel • Unclear risk of bias No blinding. However, this is probably not possible in this instance. Blinding of outcome assessment • Unclear risk of bias No blinding. However, this is probably not possible in this instance. Incomplete outcome data • Low risk of bias Selective reporting • Low risk of bias Other sources of bias • Low risk of bias Overall risk of bias • Low Directness • Partially directly applicable 30% in the surgery arm and 35% in the non-surgery arm were T4, N0 or N1. (They were not N2)

Short Title	Title	Study Characteristics	Risk of Bias
		• Split between study groups *Induction chemotherapy, chemoradiotherapy + surgery = 81; induction chemotherapy, chemoradiotherapy = 80* • Loss to follow-up *None* • %female *Induction chemotherapy, chemoradiotherapy + surgery = 31%; induction chemotherapy, chemoradiotherapy = 34%* • Average age *Median (range): Induction chemotherapy, chemoradiotherapy + surgery = 58 years (33-72); induction chemotherapy, chemoradiotherapy = 59 years (42-74)* **Interventions** • Chemotherapy, chemoradiotherapy + surgery *Induction chemotherapy consisted of three cycles of dose-dense cisplatin and paclitaxel in a 21-day cycle. Neoadjuvant radiotherapy was delivered to a total cumulative dose of 45 Gy, as two 1.5-Gy fractions per day, given 5 days a week. The minimum interval between daily fractions was 6 hours. Three dimensional treatment planning was mandatory. Intensity-modulated radiotherapy was not allowed. Concurrent chemotherapy consisted of one cycle of cisplatin and vinorelbine: cisplatin 50 mg/m2 on days 2 and 9 and vinorelbine 20 mg/m2 on days 2 and 9 of neoadjuvant radiotherapy.* • Chemotherapy, chemoradiotherapy boost *Induction chemotherapy consisted of three cycles of dose-dense cisplatin and paclitaxel in a 21-day cycle. Neoadjuvant radiotherapy was delivered to a total cumulative dose of 45 Gy, as two 1.5-Gy fractions per day, given 5 days a week. The minimum interval between daily fractions was 6 hours. Three dimensional treatment planning was mandatory. Intensity-modulated radiotherapy was not allowed. Concurrent chemotherapy consisted of one cycle of cisplatin and vinorelbine: cisplatin 50 mg/m2 on days 2 and 9 and vinorelbine 20*	

Short Title	Title	Study Characteristics	Risk of Bias
		mg/m2 on days 2 and 9 of neoadjuvant radiotherapy. The chemoradiotherapy boost was risk adapted to between 65 and 71 Gy. This was done in the following way: Definitive boost radiotherapy was given at 2 Gy per fraction, five fractions per week, to a cumulative dose of 20 to 26 Gy without a treatment break from neoadjuvant radiotherapy. A 26-Gy boost dose was recommended if deliverable within the normal tissue constraints. Specific radiation parameters, techniques, concurrent chemotherapy application given to the boost (cisplatin 40 mg/m2 on day 2 and vinorelbine 15mg/m2 on days 2 and 9 of the boost radiotherapy). The maximum allowed mean dose to the lung was 18 Gy, and the maximum dose at the spinal cord had to be less than 42 Gy. To avoid increased toxicities during the concurrent chemoradiotherapy boost, and given the previous experience in the pilot phase II study, concurrent chemotherapy to the boost was reduced in doses of cisplatin and vinorelbine. **Outcome measures** • Mortality, all-cause • Adverse events grade 3 or above • Dropout during treatment	
Girard 2010	Is neoadjuvant chemoradiotherapy a feasible strategy for stage IIIA-N2 non-small cell lung cancer? Mature results of the randomized IFCT-0101 phase II trial	**Study type** • Randomised controlled trial **Study details** • Study location *France* • Study setting *Hospitals* • Study dates *Recruitment was from 2003 to 2007* • Duration of follow-up *Median follow-up of 31.4 months.*	**Quality assessment (RCT)** Random sequence generation • High risk of bias Randomization was stratified by clinical centre and histological type (squamous cell carcinoma vs. others). However, the 3 groups were not balanced in terms of gender or pN2/cN2. This might be because of the relatively low numbers of participants. Nevertheless, they were not balanced. Allocation concealment • Unclear risk of bias Blinding is probably not possible in this sort of study.

Short Title	Title	Study Characteristics	Risk of Bias
		• Sources of funding *Programme Hospitalier de Recherche Clinique, Ligue National contre le Cancer and the Lilly Laboratories.* **Inclusion criteria** • Staging CT of chest, abdomen, head *CT brain or MRI brain. Fiberoptic bronchoscopy, mediastinoscopy.* • Pathologically proven NSCLC • Stage IIIA (T1-3)-N2 • Potentially resectable **Exclusion criteria** • ECOG performance status 2 or above • Inadequate renal, hepatic or haematologic functions • Age <18 years • Age >70 years • Unsatisfactory medical condition for chemotherapy, thoracic radiotherapy and surgery • Predicted post-operative FEV1 <35% of predicted value • High probability of stage IIIB NSCLC *In other words, if the tumour was suspected to invade the carina, the superior vena cava, the phrenic nerves, the aorta, the oesophagus, the vertebrae, the heart, the chest wall, or the contra-lateral mediastinal or supra-clavicular lymph nodes.* • Previous chemotherapy or thoracic radiotherapy • History of respiratory, cardiac failure, or invasive cancer **Sample characteristics** • Sample size *46 people* • Split between study groups	Blinding of participants and personnel • Unclear risk of bias Blinding is probably not possible in this sort of study. Blinding of outcome assessment • Unclear risk of bias Blinding is probably not possible in this sort of study. Incomplete outcome data • Low risk of bias Selective reporting • Low risk of bias Other sources of bias • Low risk of bias Overall risk of bias • Moderate Directness • Directly applicable

Short Title	Title	Study Characteristics	Risk of Bias
		Induction chemotherapy, surgery = 14; induction chemoradiotherapy (cisplatin + vinorelbine), surgery = 17; induction chemoradiotherapy (carboplatin + paclitaxel), surgery = 15 • Loss to follow-up *None* • %female *Induction chemotherapy, surgery = 35.7%; induction chemoradiotherapy (cisplatin + vinorelbine), surgery = 11.8%; induction chemoradiotherapy (carboplatin + paclitaxel), surgery = 13.3%* • Average age *Not provided* • Numbers of participants with pN2 and cN2 *Induction chemotherapy, surgery = 6 & 8; induction chemoradiotherapy (cisplatin and vinorelbine), surgery = 15 & 2; induction chemoradiotherapy (carboplatin and paclitaxel), surgery = 12 & 3* **Interventions** • Chemotherapy, surgery *This arm consisted of chemotherapy with cisplatin (80mg/m2 on days 1, 22, 43) and gemcitabine (1250mg/m2 on days 1, 8, 22, 29, 43, 50). Surgery was scheduled between week 11 and week 14 after randomisation. Lobectomy or pneumonectomy was performed. After surgery, post-operative treatment depended on the completion of the resection. In case of complete resection (R0), no adjuvant treatment was administered; in case of microscopically incomplete resection (R1), adjuvant radiotherapy was done to a total dose of 60 Gy for patients assigned this arm. After macroscopically incomplete resection (R2), radiotherapy was administered to a total dose of 60 Gy after a pneumonectomy, and of 66Gy after a lobectomy for patients in this arm.* • Chemoradiotherapy (cisplatin + vinorelbine), surgery *Participants received induction chemotherapy followed by chemoradiotherapy. This arm consisted of the combination of cisplatin*	

Short Title	Title	Study Characteristics	Risk of Bias
		(80mg/m2 on days 1, 22, 43) and vinorelbine (25mg/m2 on days 1, 8, 15, and 15mg/m2 on days 22, 29, 43, 50), with radiotherapy to a total dose of 46 grays delivered from week 4 to week 8. Conformal radiotherapy was delivered using a standard fractionation scheme (2 Gy/day, 5 days/week), after a three-dimensional treatment planning. Patients were immobilized using a cervico-thoracic immobilization device. The gross tumor volume (GTV) was defined as the primary tumor mass including any hilar or mediastinal lymph node ≥1 cmin short axis dimension. A 6–8mmmargin was added to the GTV to account for microscopic extension. Additional margins for tumor motion, ranging from 10 to 20mm were added based on radioscopy to define the Planned Tumor Volume (PTV). Dose–volume histograms for normal lung were calculated using total lung volume excluding the PTV. The lung V20 had to be lower than 30%. Total dose to the spinal cord was limited to 46 Gy. The maximal dose delivered to more than 15cm of the oesophagus was 40 Gy. Treatment plans included corrections for lung tissue inhomogeneity. The 100%-isodose line was defined at the isocenter of the treatment plan, and total dose was prescribed to this point. Beam-eye-view display was used to ensure optimal target volume coverage and normal tissue sparing. After surgery, post-operative treatment depended on the completion of the resection. In case of complete resection (R0), no adjuvant treatment was administered; in case of microscopically incomplete resection (R1), a dose of 14 Gy was delivered post-operatively. After macroscopically incomplete resection (R2), radiotherapy was administered to a total dose of 60 Gy after a pneumonectomy. For patients initially assigned to this arm, the decision about adjuvant treatment was left to the discretion of the local investigator. *• Chemoradiotherapy (carboplatin + paclitaxel), surgery* *Participants received induction chemotherapy followed by chemoradiotherapy. This arm consisted of the association of carboplatin (Calvert AUC 6 on day 1, and AUC 2 on days 22, 29, 36, 43, 50) and paclitaxel (200mg/m2 on day 1, and 40mg/m2 on days 22, 29, 36, 43, 50), with radiotherapy to a total dose of 46 grays delivered from week 4 to week 8. Conformal radiotherapy was delivered using a*	

Short Title	Title	Study Characteristics	Risk of Bias
		standard fractionation scheme (2 Gy/day, 5 days/week), after a three-dimensional treatment planning. Patients were immobilized using a cervico-thoracic immobilization device. The gross tumour volume (GTV) was defined as the primary tumour mass including any hilar or mediastinal lymph node ≥1 cmin short axis dimension. A 6–8mmmargin was added to the GTV to account for microscopic extension. Additional margins for tumour motion, ranging from 10 to 20mm, were added based on radioscopy to define the Planned Tumour Volume (PTV). Dose–volume histograms for normal lung were calculated using total lung volume excluding the PTV. The lung V20 had to be lower than 30%. Total dose to the spinal cord was limited to 46 Gy. The maximal dose delivered to more than 15cm of the oesophagus was 40 Gy. Treatment plans included corrections for lung tissue inhomogeneity. The 100%-isodose line was defined at the isocenter of the treatment plan, and total dose was prescribed to this point. Beam-eye-view display was used to ensure optimal target volume coverage and normal tissue sparing. After surgery, post-operative treatment depended on the completion of the resection. In case of complete resection (R0), no adjuvant treatment was administered; in case of microscopically incomplete resection (R1), a dose of 14 Gy was delivered post-operatively. After macroscopically incomplete resection (R2), radiotherapy was administered to a total dose of 60 Gy after a pneumonectomy. For patients initially assigned to this arm, the decision about adjuvant treatment was left to the discretion of the local investigator. **Outcome measures** • Mortality, all-cause • Adverse events grade 3 or above	
Johnstone 2002	Phase III study comparing chemotherapy and radiotherapy with preoperative	**Study type** • Randomised controlled trial **Study details**	**Quality assessment (RCT)** Random sequence generation • High risk of bias

Short Title	Title	Study Characteristics	Risk of Bias
	chemotherapy and surgical resection in patients with non-small-cell lung cancer with spread to mediastinal lymph nodes (N2); final report of RTOG 89-01. Radiation Therapy Oncology Group	• Study location *USA* • Study setting *Hospitals* • Study dates *1990 to 1994* • Duration of follow-up *Follow-up was for at least 48 months.* • Sources of funding *Not stated* **Inclusion criteria** • Pathologic proof of N2 involvement • Stage IIIA (T1-3)-N2 *And M0* **Exclusion criteria** • None **Sample characteristics** • Sample size *61 people* • Split between study groups *Induction chemotherapy, surgery = 29; induction chemotherapy, radiotherapy = 32* • Loss to follow-up *2 people. It is not specified which arms they were in.* • %female *Induction chemotherapy, surgery = 38%; induction chemotherapy, radiotherapy = 22%*	Some participants were not randomised but were included in the mortality results: 7/29 in the surgery arm and 9/32 in the radiotherapy arm. Allocation concealment • Unclear risk of bias No blinding. However, this may not be possible for these participants. Blinding of participants and personnel • Unclear risk of bias No blinding. However, this may not be possible for these participants. Blinding of outcome assessment • Unclear risk of bias No blinding. However, this may not be possible for these participants. Incomplete outcome data • High risk of bias There was a narrative description of the adverse events. However, there should have been a table because the investigators' definition of what is "equivalent" might not be the same as other people's definition of equivalence. Selective reporting • High risk of bias The mortality data included non-randomised participants. The mortality data might have been

Short Title	Title	Study Characteristics	Risk of Bias
		• Average age *Percentage <60 years, percentage 60+ years: Induction chemotherapy, surgery = 59%, 41%; induction chemotherapy, radiotherapy = 50%, 50%* **Interventions** • Chemotherapy, surgery *Induction chemotherapy consisted of cisplatin 120 mg/m2 on Days 1 and 29, vinblastine 4.5 mg/m2 on Days 1, 15, 29, and 43, and mitomycin-C 8 mg/m2 on Days 1 and 29. Patients were randomised to surgery on Day 71 followed by cisplatin on Days 99 and 127, vinblastine on Days 99, 113, 127, and 141. 7/29 participants were not randomised and had mitomycin-C in addition to the induction chemotherapy described above.* • Chemotherapy, radiotherapy *Induction chemotherapy consisted of cisplatin 120 mg/m2 on Days 1 and 29, vinblastine 4.5 mg/m2 on Days 1, 15, 29, and 43, and mitomycin-C 8 mg/m2 on Days 1 and 29. Participants were randomised to radiotherapy starting on Day 71, given to 64 Gy in 2.0 Gy fractions, followed by cisplatin on Days 141 and 169 and vinblastine on Days 141, 155, 169, and 183. 9/32 participants were not randomised and had mitomycin-C in addition to the induction chemotherapy described above. Radiotherapy (50 Gy at 2.0-Gy fractions/d, 5 fractions/wk) to the primary and regional nodes began 2–4 weeks after the completion of induction chemotherapy. A boost dose of 14 Gy was delivered at 2.0-Gy fractions/d, 5 fractions/wk, to gross disease as seen on the original CT scan, for a total dose of 64 Gy to all involved sites. All doses were calculated at the center of the target volume; the maximal dose could not exceed the target dose by >15%. The primary site and hilar/mediastinal nodes were treated with a 2-cm margin to a minimal dose of 50 Gy; the boost volume included only gross disease in these sites, with the fields defined by custom lead blocking. Beam energies >1 MeV were required, and posterior spinal*	different if only randomised participants had been included. Other sources of bias • High risk of bias The non-randomised participants that were included in the mortality data had different chemotherapy regimens compared to the randomised participants. Overall risk of bias • High Directness • Directly applicable

Short Title	Title	Study Characteristics	Risk of Bias
		cord blocks were not allowed. All simulation and portal films were centrally reviewed for protocol compliance. **Outcome measures** • Mortality, all-cause	
Katakami 2012	A phase 3 study of induction treatment with concurrent chemoradiotherapy versus chemotherapy before surgery in patients with pathologically confirmed N2 stage IIIA nonsmall cell lung cancer (WJTOG9903)	**Study type** • Randomised controlled trial **Study details** • Study location *Japan* • Study setting *Multiple academic and community hospitals.* • Study dates *2000 to 2005* • Duration of follow-up *Patients were scheduled for a chest CT scan 4 to 6 weeks after completion of the last chemotherapy cycle and were followed up every 2 months for at least 5 years. During this time, the patients received CT scans of the chest and upper abdomen, CT or MRI scans of the brain, and bone scans every 6 months.* • Sources of funding *No specific funding was disclosed.* **Inclusion criteria** • Pathologic proof of N2 involvement *From biopsy samples of the ipsilateral mediastinal nodes that were visible on a CT scan.* • Staging CT of chest, abdomen, head *Also included a bone scan. CT brain or MRI brain.* • Pathologically proven NSCLC	**Quality assessment (RCT)** Random sequence generation • Unclear risk of bias The randomisation method was not provided. However, the baseline characteristics of both arms were roughly equal. Allocation concealment • Unclear risk of bias There was no blinding in this study. However, blinding might not be realistically possible for these participants. Blinding of participants and personnel • Unclear risk of bias There was no blinding in this study. However, blinding might not be realistically possible for these participants. Blinding of outcome assessment • Unclear risk of bias There was no blinding in this study. However, blinding might not be realistically possible for these participants. Incomplete outcome data

Short Title	Title	Study Characteristics	Risk of Bias
		• Stage IIIA (T1-3)-N2 • Potentially resectable **Exclusion criteria** • ECOG performance status 2 or above • Inadequate renal, hepatic or haematologic functions *And unsatisfactory cardiac function.* • Age >70 years • Partial pressure of arterial oxygen <70 Torr • FEV1 <1.5 L • Prior malignancy other than non-melanoma skin cancer or adequately treated stage I in situ cervical cancer • Uncontrolled angina pectoris or a history of congestive heart failure or myocardial infarction within 3 months • Pulmonary fibrosis detectable by CT scan • COPD (FEV1 <65%) • >10% weight loss within the previous 6 months • Age <20 years **Sample characteristics** • Sample size *56 people* • Split between study groups *Induction chemotherapy, surgery = 29; induction chemoradiotherapy, surgery = 31* • Loss to follow-up *None* • %female *Induction chemotherapy, surgery = 32%; induction chemoradiotherapy, surgery = 34%* • Average age	• Low risk of bias Selective reporting • Low risk of bias Other sources of bias • Low risk of bias Overall risk of bias • Low Directness • Directly applicable

Short Title	Title	Study Characteristics	Risk of Bias
		Median age (range): Induction chemotherapy, surgery = 58.0 years (34-69); induction chemoradiotherapy, surgery = 57.0 years (36-70) **Interventions** • Chemotherapy, surgery *Induction chemotherapy involved 2 cycles of carboplatin (area under the receiver operating curve [AUC] = 5 on days 1, 22, intravenous infusions) and docetaxel (60 mg/m2 on days 1, 22, intravenous infusions). The patients were reassessed using CT scan plus repeat pulmonary function tests 2 to 4 weeks after completion of the induction therapy. The response to induction was assessed by WHO criteria without the need for a second confirmation of response. If the disease had not progressed and the patient remained medically healthy, a complete surgical resection with a mediastinal lymph node dissection was performed 3 or 4 weeks after the induction therapy was completed. No consolidation chemotherapy was administered after surgery. Dose reduction guidelines were specified in the protocol.* • *Chemoradiotherapy (carboplatin + docetaxel), surgery* *Induction chemotherapy involved 2 cycles of carboplatin (area under the receiver operating curve [AUC] = 5 on days 1, 22, intravenous infusions) and docetaxel (60 mg/m2 on days 1, 22, intravenous infusions). Thoracic radiotherapy (40 Gy in 20 fractions of 2 Gy over 4 weeks) was also administered from day 1. All patients were treated with a linear accelerator photon beam of 6MV or more. At the commencement of this multi-institutional study, a 3-dimensional (3D) treatment planning system using CT was not available at some of the participating institutions. Hence, 2-dimensional (2D) treatment planning techniques were allowed. Radiation doses were specified at the centre of the target volume, and doses were calculated assuming tissue homogeneity without correction for lung tissues. The primary tumour and involved nodal disease received 40 Gy in 2 Gy fractions over 4 weeks via the anterior and posterior opposing portals. Radiation fields included the primary tumour with a margin of at least 1.0 cm, and the ipsilateral hilum and mediastinal nodal areas with a margin of 0.5 to 1.0*	

Short Title	Title	Study Characteristics	Risk of Bias
		cm from the paratracheal lymph nodes (#2) to 4.5 cm below the tracheal bifurcation including subcarinal lymph nodes (#7). The contralateral hilum was not included. The supraclavicular areas were not treated routinely, but the ipsilateral supraclavicular area was treated when the primary tumour was located in the upper lobe. The patients were reassessed using CT scan plus repeat pulmonary function tests 2 to 4 weeks after completion of the induction therapy. The response to induction was assessed by WHO criteria without the need for a second confirmation of response. If the disease had not progressed and the patient remained medically healthy, a complete surgical resection with a mediastinal lymph node dissection was performed 3 or 4 weeks after the induction therapy was completed. No consolidation chemotherapy was administered after surgery. Dose reduction guidelines were specified in the protocol. Patients in the CRS arm who could not be treated surgically within 6 weeks after induction therapy received further radiotherapy of up to 66 Gy in 33 fractions in total. In this boost radiotherapy procedure, the spinal cord was excluded from the radiation fields. **Outcome measures** • Mortality, all-cause • Adverse events grade 3 or above	
Pless 2015	Induction chemoradiation in stage IIIA/N2 non-small-cell lung cancer: a phase 3 randomised trial	**Study type** • Randomised controlled trial **Study details** • Study location *Switzerland, Germany and Serbia* • Study setting *Cancer centres* • Study dates *Enrolment was from 2001 to 2012*	**Quality assessment (RCT)** Random sequence generation • Low risk of bias Allocation concealment • Unclear risk of bias There was no blinding. However, blinding may not be realistically possible with these participants. Blinding of participants and personnel • Unclear risk of bias

Short Title	Title	Study Characteristics	Risk of Bias
		• Duration of follow-up *Patients attended follow-up visits 1 month after surgery, then every 3 months for 2 years, every 6 months for 2 years, and then every 12 months. During visits patients were assessed for toxic effects. They also underwent chest radiography or chest CT at alternate visits for 5 years. The trial was stopped after the third interim analysis and 134 events, on the advice of the independent data monitoring board, because the futility boundary had been crossed. At the time of data cut-off, the median follow-up time was 52·4 months (IQR 32·0–85·2).* • Sources of funding *This study was funded by the Swiss State Secretariat for Education, Research and Innovation, the Swiss Cancer League and Sanofi.* **Inclusion criteria** • Pathologic proof of N2 involvement *Participants with histological or cytological proof of non-small-cell lung cancer but N2 lymph nodes not accessible to biopsy (eg, aortic node regions 5 and 6) were eligible, provided that the N2 node had a diameter greater than 1 cm and was PET positive, and the N3 nodes had diameters less than 1 cm and were PET negative.* • Pathologically proven NSCLC • Stage IIIA (T1-3)-N2 *And M0* • Staging PET-CT and brain MRI **Exclusion criteria** • ECOG performance status 2 or above • Age <18 years • Age >75 years • Unacceptable lung and cardiac function according to local standards • Inadequate liver, bone marrow and kidney functions *Creatinine clearance less than 1·00 mL/s [60 mL/min]*	There was no blinding. However, blinding may not be realistically possible with these participants. Blinding of outcome assessment • Unclear risk of bias There was no blinding. However, blinding may not be realistically possible with these participants. Incomplete outcome data • Low risk of bias Selective reporting • Low risk of bias Other sources of bias • Low risk of bias Overall risk of bias • Low Directness • Directly applicable

Short Title	Title	Study Characteristics	Risk of Bias
		Sample characteristics • Sample size *231 people* • Split between study groups *Induction chemotherapy, surgery = 115; induction chemoradiotherapy, surgery = 117* • Loss to follow-up *Induction chemotherapy, surgery = 8; induction chemoradiotherapy, surgery = 2* • %female *Induction chemotherapy, surgery = 33%; induction chemoradiotherapy, surgery = 33%* • Average age *Median age (range): Induction chemotherapy, surgery = 59.0 years (30.0-74.0); induction chemoradiotherapy, surgery = 60.0 years (37.0-76.0)* **Interventions** • Chemotherapy, surgery *Chemotherapy consisted of three cycles of 100 mg/m² intravenous cisplatin and 85 mg/m² docetaxel given every 3 weeks. The administration of prophylactic granulocyte-colony stimulating factor was compulsory. Dose reductions were not allowed for cisplatin. Switch to carboplatin (target area under the curve 6) was possible if patients developed renal insufficiency (creatinine clearance lower than 0·83 mL/s [50 mL/ min]), hearing loss worse than grade 1, or peripheral neuropathy worse than grade 2. Dose reductions for docetaxel to 55 mg/m² were possible if patients developed impaired liver function (worse than grade 1), grade 3 diarrhoea, or peripheral neuropathy (worse than grade 1). If toxic effects did not recover to grade 1 severity or resolve within 2 weeks, chemotherapy was stopped. Surgery was scheduled 21 days after the last chemotherapy cycle for patients in the*	

Short Title	Title	Study Characteristics	Risk of Bias
		chemotherapy group. Surgery included tumour resection and systematic lymph node dissection. Patients in the chemotherapy group in whom resection was incomplete (R1 or R2) were allowed to receive postoperative radiotherapy. • Chemoradiotherapy (cisplatin + docetaxel), surgery *Chemotherapy consisted of three cycles of 100 mg/m² intravenous cisplatin and 85 mg/m² docetaxel given every 3 weeks. The administration of prophylactic granulocyte-colony stimulating factor was compulsory. Dose reductions were not allowed for cisplatin. Switch to carboplatin (target area under the curve 6) was possible if patients developed renal insufficiency (creatinine clearance lower than 0·83 mL/s [50 mL/ min]), hearing loss worse than grade 1, or peripheral neuropathy worse than grade 2. Dose reductions for docetaxel to 55 mg/m² were possible if patients developed impaired liver function (worse than grade 1), grade 3 diarrhoea, or peripheral neuropathy (worse than grade 1). If toxic effects did not recover to grade 1 severity or resolve within 2 weeks, chemotherapy was stopped. Three weeks after day 1 of the last planned date of chemotherapy, radiotherapy was started in patients in the chemoradiotherapy group. Patients received 44 Gy in 22 fractions over a 3 week period, delivered with a concomitant boost technique. Planning target volumes were defined according to the results of CT scans done after induction chemotherapy. Planning target volume 1, representing the original volume, included the primary tumour, lymph nodes, ipsilateral hilus, and ipsilateral and contralateral mediastinum at risk of subclinical disease, with a 1·5–2·0 cm margin. Planning target volume 2 included the primary tumour (gross disease) with a 1·5–2·0 cm margin and lymph node metastases in the mediastinum and represented the boost volume. Arrangement of fields was at the discretion of the investigators as long as the target volumes were clearly outlined. The dose to the spinal cord had to remain lower than 36 Gy. The prescribed dose was specified at the International Commission on Radiation Units and Measurements reference point. Computer assisted three-dimensional treatment planning was used in all cases, and the selection of a collapsed cone or Monte Carlo algorithm was recommended for photon*	

Short Title	Title	Study Characteristics	Risk of Bias
		energies greater than 6 MV. The reference isodose had to be within 10% of that prescribed, and hot spots were delineated and recorded. Central review of three random patients from each centre was done to ensure radiotherapy quality control. Surgery was scheduled 21–28 days after completion of radiotherapy for patients in the chemoradiotherapy group. Surgery included tumour resection and systematic lymph node dissection. **Outcome measures** • Mortality, all-cause • Adverse events grade 3 or above	
Shepherd 1998	Randomized study of chemotherapy and surgery versus radiotherapy for stage IIIA non-small-cell lung cancer: a National Cancer Institute of Canada Clinical Trials Group Study	**Study type** • Randomised controlled trial **Study details** • Study location *Canada* • Study setting *Hospital* • Study dates *Not provided. This study was received by the publishers in 1997.* • Duration of follow-up *Looking at the survival chart, participants were followed up for 24 months in the radiotherapy arm and 31 months in the surgery arm.* • Sources of funding *Not stated* **Inclusion criteria** • Stage IIIA N2 NSCLC with biopsy-proven mediastinal node involvement	**Quality assessment (RCT)** Random sequence generation • High risk of bias Method of randomisation was not given. In addition, the median age of participants was 9 years older in the chemotherapy, surgery group compared to the radiotherapy group. Allocation concealment • Unclear risk of bias There was no blinding in this study. However, blinding may not have been realistically possible due to the nature of the condition. Blinding of participants and personnel • Unclear risk of bias There was no blinding in this study. However, blinding may not have been realistically possible due to the nature of the condition. Blinding of outcome assessment

Short Title	Title	Study Characteristics	Risk of Bias
		Exclusion criteria • Stage IIIB • Not able to tolerate planned surgery • Post-operative predicted FEV1 <0.8 L • ECOG performance status >2 • Haemoglobin <100 g/L • Granulocytes <2.0 x 109 /L • Platelets <100 x 109 /L • Serum creatinine >150 micro mol / L • Liver enzymes >1.25 x upper limit of normal **Sample characteristics** • Sample size *31 people* • Split between study groups *Chemotherapy, surgery = 16; radiotherapy = 15* • Loss to follow-up *None* • %female *Chemotherapy, surgery = 25%; radiotherapy = 33%* • Average age *Median (range): chemotherapy, surgery = 61 years (49-70); radiotherapy = 52 years (44-72)* **Interventions** • Chemotherapy, surgery *Patients received cisplatin 120 mg m2 on days 1 and 29 and vinblastine 6 mg m2 on days 1. 15. 22. 29 and 43. Cisplatin was administered in hospital with vigorous hydration and mannitol diuresis and dexamethasone. Ondansetron and lorazepam were given to prevent vomiting. Patients proceeded to surgery between days 51 and*	• Unclear risk of bias There was no blinding in this study. However, blinding may not have been realistically possible due to the nature of the condition. Incomplete outcome data • High risk of bias A narrative description of adverse events was given in such a way that it is not possible to compare groups. For example, there was either no grading or no participant numbers provided and it is not clear which adverse events occurred in which arm. A table of adverse events was not provided. Median survival in both arms was provided. However, follow-up lasted for 32 months and about 1/3 of participants were still alive at this time. Selective reporting • High risk of bias A narrative description of adverse events was given in such a way that it is not possible to compare groups. For example, there was either no grading or no participant numbers provided and it is not clear which adverse events occurred in which arm. A table of adverse events was not provided. Median survival in both arms was provided. However, follow-up lasted for 32 months and about 1/3 of participants were still alive at this time. Other sources of bias • Low risk of bias Overall risk of bias

Short Title	Title	Study Characteristics	Risk of Bias
		64 if they achieved partial or complete response or stable disease after chemotherapy. An attempt was made to excise all tissue felt to have been involved before chemotherapy and radical lymph node dissection was required. Patients who had complete resection received the same chemotherapy starting 6 weeks post-operatively. • Radiotherapy *A total dose of 60 Gy was planned to be riven as 2 Gy daily 5 days a week with the dose prescribed to the centre of the target volume (ICRU 29). The initial target volume (50 Gy) included the primary tumour and ipsilateral hilar, subcarinal, tracheobronchial and paratracheal nodes. The reduced target volume (10 Gy) included the tumour and involved nodes as determined by computerized tomography or mediastinoscopy. The spinal cord dose was limited to 48 Gy and real time review was performed.* **Outcome measures** • Mortality, all-cause • Dropout during treatment	• High Directness • Directly applicable
Stephens 2005	A randomised controlled trial of pre-operative chemotherapy followed, if feasible, by resection versus radiotherapy in patients with inoperable stage T3, N1, M0 or T1-3, N2, M0 non-small cell lung cancer	**Study type** • Randomised controlled trial **Study details** • Study location *UK* • Study setting *Christie Hospital NHS Trust, Manchester* • Study dates *Randomisation occurred between 1995 to 1999* • Duration of follow-up *The SF-36 quality of life questionnaire was used at baseline, 12 weeks and at 6 months. Adverse events were measured for the first 6 months.*	**Quality assessment (RCT)** Random sequence generation • Low risk of bias Allocation concealment • Unclear risk of bias No blinding. However, blinding these participants and the staff involved with them may not be realistically possible. Blinding of participants and personnel • Unclear risk of bias

Short Title	Title	Study Characteristics	Risk of Bias
		Of the 48 patients, 39 died. The median follow-up for the 9 survivors was 14 months (range 5—68 months). • Sources of funding *Not provided. However, the MRC Clinical Trials Unit co-ordinated and analysed the results of the trial.* **Inclusion criteria** • NSCLC (T3, N1, M0 or T1-3, N2, M0) • Currently unresectable but have the potential to become resectable following chemotherapy • Thoracotomy or CT thorax & abdomen + mediastinoscopy or mediastinotomy **Exclusion criteria** • Not able to tolerate planned surgery • WHO performance status >2 • Creatinine clearance <50 ml/min • Full blood count outside the normal range • Previous or current other malignancy • Other disease or condition likely to interfere with the protocol treatments or comparisons • Contraindications to either of the treatment regimens **Sample characteristics** • Sample size *48 people* • Split between study groups *Chemotherapy, surgery = 24; radiotherapy = 24* • Loss to follow-up *None* • %female	No blinding. However, blinding these participants and the staff involved with them may not be realistically possible. Blinding of outcome assessment • Unclear risk of bias No blinding. However, blinding these participants and the staff involved with them may not be realistically possible. Incomplete outcome data • High risk of bias With the exception of lethargy, it was not possible to compare the other adverse events. This is because numbers and grades were not provided for each arm. In addition, quality of life data for each arm was not provided (it was only narratively described in the vaguest terms, e.g. – no statistically significant differences). Selective reporting • High risk of bias With the exception of lethargy, it was not possible to compare the other adverse events. This is because numbers and grades were not provided for each arm. In addition, quality of life data for each arm was not provided (it was only narratively described in the vaguest terms, e.g. – no statistically significant differences). Other sources of bias • Low risk of bias

Short Title	Title	Study Characteristics	Risk of Bias
		Chemotherapy, surgery = 29%; radiotherapy = 38% • Average age *Median (range): chemotherapy, surgery = 58 years (44-76); radiotherapy = 61 years (42-71)* **Interventions** • Chemotherapy, surgery *Chemotherapy, surgery patients received 4 cycles of chemotherapy at 3-week intervals with either MVP (mitomycin 6mg/m2 by IV injection, vinblastine 6mg/m2 by IV injection (maximum dose 10 mg), and cisplatin 50mg/m2 by IV infusion over 4 hours) or MIC (mitomycin 6mg/m2 by IV injection, ifosfamide 3 g/m2 by IV injection, with mesna, and cisplatin 50mg/m2 by IV infusion over 1 hour), with standard hydration and anti-emetics. Surgical resection, if considered feasible, was carried out between 4 and 6 weeks after the final cycle of chemotherapy. The surgical technique was decided by the local surgeon according to the site and extent of the tumour and local practice. Patients considered to have unresectable disease following chemotherapy received thoracic radiotherapy, the details of which were decided by the local radiation oncologist. One patient was withdrawn from the trial, and so the data below relate to 23 patients. Twenty-one patients were treated with MIC and two with MVP; 21 received all four cycles and two three cycles. Only four patients were treated surgically (two pneumonectomies), one lobectomy, one sleeve resection), although three further patients had a thoracotomy but did not proceed to resection. The 16 remaining patients were all reported to have progressive disease post-chemotherapy, although it may be that most of these patients simply did not respond sufficiently to be considered for resection. Of the 19 patients whose tumour was not resected, 13 received radiotherapy.* • Radiotherapy *Radiotherapy participants received thoracic radiotherapy, the details of which were to be decided by the local radiation oncologist according to the site and extent of the tumour and local practice, starting as soon as*	Overall risk of bias • High Directness • Partially directly applicable In the chemotherapy, surgery group, 4/24 were T3, N1, M0. In the radiotherapy group, 3/24 were T3, N1, M0 (not N2).

Short Title	Title	Study Characteristics	Risk of Bias
		possible after randomisation. It was recommended that the radiotherapy regimen be chosen in accordance with the recommendations of the 1994 Department of Health Standing Medical Advisory Committee, which stated that patients should receive 50—60 Gy to their tumour over a period of 3—6 weeks. Twenty of the 24 patients received radiotherapy, the commonest schedules used being 50 Gy/20f, 50 Gy/15f, 40 Gy/20f, 37 Gy/26f and 28 Gy/8f. The reasons for not receiving radiotherapy were: one patient refused treatment, one was considered unsuitable for radiotherapy, the diagnosis for one patient was changed to SCLC, and for the remaining patient the reason is not known. **Outcome measures** • Mortality, all-cause • Adverse events grade 2 or above *However, only enough data for a direct comparison was provided for lethargy.* • Dropout during treatment	
Thomas 2008	Effect of preoperative chemoradiation in addition to preoperative chemotherapy: a randomised trial in stage III non-small-cell lung cancer	**Study type** • Randomised controlled trial **Study details** • Study location *Germany* • Study setting *Hospitals* • Study dates *Randomisation occurred between 1995 to 2003* • Duration of follow-up *After the end of treatment, follow-up assessments (physical assessment, chest radiography, abdominal ultrasonography, and blood chemistry) were done every 3 months for the first 2 years, then every 6*	**Quality assessment (RCT)** Random sequence generation • Unclear risk of bias Randomisation was done by a coordinating member in the Department of Medical Informatics. However, the method used was not described. Nevertheless, the baseline characteristics of both arms appear balanced. Allocation concealment • Unclear risk of bias There was no blinding. However, given the nature of the participants, blinding them and/or the staff may not be realistically possible.

Short Title	Title	Study Characteristics	Risk of Bias
		months. Additionally, for 5 years at every 6-month follow-up visit, a CT scan of the thorax was done. The median follow-up was 70 months. • Sources of funding *German Cancer Aid* **Inclusion criteria** • Pathologically proven NSCLC *Assessment of mediastinal lymph nodes by mediastinoscopy (occasionally by thoracoscopy, thoracotomy, or needle biopsy) was mandatory.* • Stage IIIA (T1-3, N2, M0) NSCLC • Stage IIIA (central T3, N0-1, M0) NSCLC • Stage IIIB (T4, N1-3, M0) NSCLC *T4 tumours were deemed potentially resectable if they involved the superior vena cava, left atrium, carina, distant trachea, or the great vessels.* • Stage IIIB (T1-4, N3, M0) NSCLC **Exclusion criteria** • ECOG performance status 2 or above • Age >70 years • Participants with T4 tumours with a malignant effusion, supraclavicular lymph node involvement, or invasion of the heart, oesophagus or vertebra. **Sample characteristics** • Sample size *524 people* • Split between study groups *Chemotherapy, chemoradiotherapy, surgery, radiotherapy = 264; chemotherapy, surgery, radiotherapy = 260*	Blinding of participants and personnel • Unclear risk of bias There was no blinding. However, given the nature of the participants, blinding them and/or the staff may not be realistically possible. Blinding of outcome assessment • Unclear risk of bias There was no blinding. However, given the nature of the participants, blinding them and/or the staff may not be realistically possible. Incomplete outcome data • High risk of bias The adverse events of leukocytopenia, thrombocytopenia and anaemia are not reported separately for each arm. In addition, many participants were missing adverse events data: chemotherapy, chemoradiotherapy, surgery, radiotherapy = 58/264; chemotherapy, surgery, radiotherapy = 73/260. Some adverse events may not have been reported altogether. For example, it's hard to believe that no participants experienced nausea or vomiting. Selective reporting • High risk of bias The adverse events of leukocytopenia, thrombocytopenia and anaemia are not reported separately for each arm. Some adverse events may not have been reported altogether. For example, it's

Short Title	Title	Study Characteristics	Risk of Bias
		• Loss to follow-up *Many participants were missing adverse events data: chemotherapy, chemoradiotherapy, surgery, radiotherapy = 58/264; chemotherapy, surgery, radiotherapy = 73/260.* • %female *Chemotherapy, chemoradiotherapy, surgery, radiotherapy = 18%; chemotherapy, surgery, radiotherapy = 17%* • Average age *Median (range): chemotherapy, chemoradiotherapy, surgery, radiotherapy = 59 years (33-69); chemotherapy, surgery, radiotherapy = 59 years (35-69)* **Interventions** • Chemotherapy, chemoradiotherapy, surgery, radiotherapy *In this arm, after three cycles of chemotherapy with cisplatin (55 mg/m²) and etoposide (100 mg/m²), patients without progressive disease (assessed with the same imaging techniques as used at baseline) were scheduled to continue with twice-daily radiotherapy and concurrent chemotherapy 3–5 weeks after the start of the third cycle of chemotherapy. All patients received CT-based three-dimensional planning. Two 1.5 Gy fractions per day, with an inter-treatment interval of at least 6 hours, were administered 5 days per week to a total dose of 45 Gy. The target volume included the primary lesion with margins of 1.5 cm, and the ipsilateral hilum and ipsilateral mediastinum extending inferiorly 5 cm below the tracheal bifurcation with a margin of 0.5–1 cm. For patients with N3 disease, the contralateral mediastinal lymph nodes, but not the contralateral hilum, were included with margins of 0.5 cm. Carboplatin (100 mg/m²) and vindesine (3 mg absolute) were administered once-weekly during treatment with twice-daily radiotherapy on days 1, 8, and 15 from the start of this phase. Surgery was scheduled 4–6 weeks after the completion of radiotherapy and concurrent chemotherapy in this arm. Extensive removal of the mediastinal lymph nodes was done, preferably by mediastinal lymph-node dissection (en-block removal of the mediastinal fatty tissue*	hard to believe that no participants experienced nausea or vomiting. Other sources of bias • High risk of bias Over 20% of participants were 'lost to follow-up' with regards to adverse events data: chemotherapy, chemoradiotherapy, surgery, radiotherapy = 58/264 (22%); chemotherapy, surgery, radiotherapy = 73/260 (28%). Overall risk of bias • High Directness • Indirectly applicable Participants who were N2 were in the minority: chemo, chemoradiotherapy, surgery = 17%; chemo, surgery = 12%.

Short Title	Title	Study Characteristics	Risk of Bias
		containing the lymphatics). Lymph-node levels to be removed were decided in accordance with the guidelines of the American Thoracic Society. If mediastinal lymph-node dissection was not done, at least mediastinal lymph-node sampling (removal or sampling of at least one lymph node) of the respective levels would have been done. Complete resection was defined as resection with negative margins and no metastatic involvement of the removed uppermost mediastinal lymph node. Histological diagnosis of the biopsies of the primary lesion and further histopathological assessment was done by the local pathologist and reviewed centrally by an experienced pneumopathologist. Also, mediastinal down-staging (initially documented N2 or N3 disease changing to N0 or N1 disease assessed by surgery) and tumour regression of more 90% was assessed centrally. Histopathological response was defined as fewer than 10% residual tumour cells in the sections of the primary lesion and no or only focal involvement with microscopic disease in the sections of mediastinal lymph nodes (tumour regression >90%). Patients deemed to have unresectable tumours or who were receiving an exploratory thoracotomy were scheduled to start twice-daily radiotherapy (total dose 24 Gy) as soon as possible after surgery. The target volume included the primary tumour with margins of 1.5 cm, the ipsilateral hilum, and ipsilateral mediastinum extending inferiorly 5 cm below the tracheal bifurcation with a margin of 0.5 to 1 cm. For patients with N3 disease, the contralateral mediastinal lymph nodes, but not the contralateral hilum, were included with margins of 0.5 cm. Additionally, patients with positive resection margins were given further radiotherapy (total dose 24 Gy). The target volume included the bronchial stump and the ipsilateral hilum. • Chemotherapy, surgery, radiotherapy *Participants had 3 cycles of chemotherapy with cisplatin (55 mg/m²) and etoposide (100 mg/m²). Surgery was scheduled after the third cycle of chemotherapy in this arm of the trial. Extensive removal of the mediastinal lymph nodes was done, preferably by mediastinal lymph-node dissection (en-block removal of the mediastinal fatty tissue containing the lymphatics). Lymph-node levels to be removed were*	

Short Title	Title	Study Characteristics	Risk of Bias
		decided in accordance with the guidelines of the American Thoracic Society. If mediastinal lymph-node dissection was not done, at least mediastinal lymph-node sampling (removal or sampling of at least one lymph node) of the respective levels would have been done. Complete resection was defined as resection with negative margins and no metastatic involvement of the removed uppermost mediastinal lymph node. Histological diagnosis of the biopsies of the primary lesion and further histopathological assessment was done by the local pathologist and reviewed centrally by an experienced pneumopathologist. Also, mediastinal down-staging (initially documented N2 or N3 disease changing to N0 or N1 disease assessed by surgery) and tumour regression of more 90% was assessed centrally. Histopathological response was defined as fewer than 10% residual tumour cells in the sections of the primary lesion and no or only focal involvement with microscopic disease in the sections of mediastinal lymph nodes (tumour regression >90%). Patients who were resected received conventionally fractionated radiotherapy (1.8 Gy per day) 4–6 weeks after surgery. All patients received CT-based three-dimensional planning. The target volume included the bronchial stump, the ipsilateral hilum, and ipsilateral mediastinum extending inferiorly 5 cm below the tracheal bifurcation with a margin of 0.5–1 cm. For patients with N3 disease, the contralateral mediastinal lymph nodes, but not the contralateral hilum, were included with margins of 0.5 cm. Patients with negative resection margins received a target volume dose of 54 Gy; those with positive margins received 68.4 Gy. Patients deemed unresectable or those with an exploratory thoracotomy were scheduled to start radiotherapy as soon as possible up to a total dose of 68.4 Gy. The target volume included the primary tumour with margins of 1.5 cm, the ipsilateral hilum, and ipsilateral mediastinum extending inferiorly 5 cm below the tracheal bifurcation with a margin of 0.5–1 cm. For patients with N3 disease, the contralateral mediastinal lymph nodes, but not the contralateral hilum, were included with margins of 0.5 cm. **Outcome measures** • Mortality, all-cause	

Short Title	Title	Study Characteristics	Risk of Bias
		• Adverse events grade 3 or above	
Van Meerbeeck 2007	Randomized controlled trial of resection versus radiotherapy after induction chemotherapy in stage IIIA-N2 non-small-cell lung cancer	**Study type** • Randomised controlled trial **Study details** • Study location *The Netherlands* • Study setting *Hospitals* • Study dates *Recruitment was from 1994 to 2002* • Duration of follow-up *Patients underwent follow-up visits every 3 months for 2 years and every 6 months thereafter, which included clinical evaluation, a chest-x-ray, and additional investigations when clinically indicated. The median follow-up was approximately 6 years.* • Sources of funding *National Cancer Institute. The study was supported by unrestricted educational grants of Eli Lilly, Bristol-Myers Squibb and Aventis.* **Inclusion criteria** • Pathologic proof of N2 involvement *Eligible patients had to have cytologic or histologic proof of unresectable stage IIIA-N2 NSCLC.* • Staging CT of chest, abdomen, head *Guidelines for unresectability were as follows: 1) any N2 involvement by a non-squamous carcinoma; 2) in case of squamous cell carcinoma, any N2 nodal involvement exceeding level 4R for a right-sided tumour and level 5 and 6 for a left-sided tumour. N2 found only at thoracotomy after a negative staging mediastinoscopy was not necessarily considered to be unresectable. Tumors and/or any involved	**Quality assessment (RCT)** Random sequence generation • Low risk of bias Allocation concealment • Unclear risk of bias No blinding. However, it may not be realistically possible to blind participants and staff given the nature of the disease. Blinding of participants and personnel • Unclear risk of bias No blinding. However, it may not be realistically possible to blind participants and staff given the nature of the disease. Blinding of outcome assessment • Unclear risk of bias No blinding. However, it may not be realistically possible to blind participants and staff given the nature of the disease. Incomplete outcome data • High risk of bias The adverse events are reported narratively in such a way that it is not possible to compare the arms of the trial. It is hard to believe that no participant experienced nausea or vomiting. Selective reporting

Short Title	Title	Study Characteristics	Risk of Bias
		mediastinal lymph node(s) had to be unidimensionally measurable on CT scan. • Pathologically proven NSCLC **Exclusion criteria** • Age <18 years • Unsatisfactory medical condition for chemotherapy, thoracic radiotherapy and surgery • WHO performance status >2 • Previous or current other malignancy • Evidence of pulmonary fibrosis • Pre-existing neurotoxicity • Pre-existing infection • Previous therapy for NSCLC **Sample characteristics** • Sample size *308 people* • Split between study groups *Chemotherapy, surgery = 154; chemotherapy, radiotherapy = 154* • Loss to follow-up *None* • %female *Chemotherapy, surgery = 29%; chemotherapy, radiotherapy = 23%* • Average age *Median (range): chemotherapy, surgery = 61 years (29-78); chemotherapy, radiotherapy = 62 years (33-76)* **Interventions** • Chemotherapy, surgery	• High risk of bias The adverse events are reported narratively in such a way that it is not possible to compare the arms of the trial. It is hard to believe that no participant experienced nausea or vomiting. Other sources of bias • Low risk of bias Overall risk of bias • High Directness • Directly applicable

Short Title	Title	Study Characteristics	Risk of Bias
		Induction chemotherapy consisted of three cycles of cisplatin, at a dose of at least 80 mg/m 2 per cycle, or carboplatin, at a target area under the curve of at least 5 per cycle, combined with at least one other chemotherapy drug. Response was evaluated with CT scan after at least two cycles of induction chemotherapy and scored according to WHO criteria, but confirmation was not required. Eligibility was reassessed before random assignment. Only patients showing a response (complete, partial, or minor) to induction chemotherapy were eligible for random assignment. Surgery had to start within 6 weeks of random assignment. Postoperative radiotherapy consisting of 56 Gy in once-daily fractions of 2 Gy was recommended in cases of incomplete resection and had to start between the 4th and 10th postoperative week. • Chemotherapy, radiotherapy *Induction chemotherapy consisted of three cycles of cisplatin, at a dose of at least 80 mg/m 2 per cycle, or carboplatin, at a target area under the curve of at least 5 per cycle, combined with at least one other chemotherapy drug. Response was evaluated with CT scan after at least two cycles of induction chemotherapy and scored according to WHO criteria, but confirmation was not required. Eligibility was reassessed before random assignment. Only patients showing a response (complete, partial, or minor) to induction chemotherapy were eligible for random assignment. Radiotherapy had to start within 6 weeks of random assignment. The dosage administered to the primary tumour and involved mediastinum was 60–62.5 Gy and to the uninvolved mediastinum it was 40–46 Gy. The fractionation size was 1.95 – 2.05 Gy. A number of fractions were 30-32. The total treatment duration was 40-46 days.* **Outcome measures** • Mortality, all-cause • Dropout during treatment	

Appendix F – GRADE tables

Network meta-analyses[1]: chemoradiotherapy, surgery vs chemoradiotherapy vs chemotherapy, surgery

Quality assessment						Effect estimate	Quality
No of studies	**Design**	**Risk of bias**	**Indirectness**	**Inconsistency**	**Imprecision**	**Summary of results (95% CI)**	
Progression free life years at 4 years							
6 RCTs (Albain 2009, Eberhard 2015, Pless 2015, Girard 2009, Katakami 2012, van Meerbeeck 2007)	RCTs	Not Serious	Not Serious	Not Serious	Not Serious	CS vs CR: 0.00 (-0.21, 0.22) CRS vs CR: 0.25 (0.06,0.44)	High
Post progression life years at 4 years							
6 RCTs (as above)	RCTs	Not Serious	Not Serious	Not Serious	Not Serious	CS vs CR: -0.11 (-0.32,0.11) CRS vs CR: -0.18 (-0.28,-0.08)	High
Total life years at 4 years							
6 RCTs (as above)	RCTs	Not Serious	Not Serious	Not Serious	Serious[2]	CS vs CR: -0.11 (-0.19,-0.03) CRS vs CR: 0.07 (-0.13,0.27)	Moderate
Odds ratio of being alive at 4 years							
6 RCTs (as above)	RCTs	Not Serious	Not Serious	Not Serious	Serious[2]	CS vs CR: 1.18 (0.76,1.86) CRS vs CR: 1.28 (0.86,1.90)	Moderate
Progression free life years at 5 years							
5 RCTs (Albain 2009, Eberhard 2015, Pless 2015, Katakami 2012, van Meerbeeck 2007)	RCTs	Not Serious	Not Serious	Not Serious	Not Serious	CS vs CR: 0.01 (-0.27, 0.3) CRS vs CR: 0.38 (0.12,0.63)	High
Post progression life years at 5 years							
5 RCTs (as above)	RCTs	Not Serious	Not Serious	Not Serious	Not Serious	CS vs CR: -0.09 (-0.18, 0.01) CRS vs CR: -0.2 (-0.33,0.07)	High
Total life years at 5 years							
5 RCTs (as above)	RCTs	Not Serious	Not Serious	Not Serious	Serious[2]	CS vs CR: -0.07 (-0.36, 0.22) CRS vs CR: 0.17 (-0.11,0.45)	Moderate
Odds ratio of being alive at 5 years							

Quality assessment						Effect estimate	Quality
No of studies	**Design**	**Risk of bias**	**Indirectness**	**Inconsistency**	**Imprecision**	**Summary of results (95% CI)**	
5 RCTs (as above)	RCTs	Not Serious	Not Serious	Not Serious	Serious[2]	CS vs CR: 1.32 (0.77, 2.14) CRS vs CR: 1.28 (0.83,1.92)	Moderate
Total adverse events of grade 3+ hazard ratio							
4 RCTs (Albain 2009, Eberhard 2015, Pless 2015, van Meerbeeck 2007)	RCTs	Not Serious	Not Serious	Not Serious	Not Serious	CR vs CRS: 1.24 (1.13,1.38) CS vs CRS: 1.39 (1.18,1.67)	High

1. Effect sizes for CS vs CRS are not shown for outcomes other than total adverse event hazard ratio. This was the only outcome for which there was a statistically significant difference between CS and CRS.
2. Not possible to distinguish any meaningfully distinct treatment options in the network

Chemoradiotherapy, surgery vs chemoradiotherapy

Quality assessment						No of patients		Effect estimate	Quality
No of studies	**Design**	**Risk of bias**	**Indirectness**	**Inconsistency**	**Imprecision**	**Chemoradio, surgery**	**Chemoradio**	**Summary of results (95% CI)**	
Mortality: all-cause hazard ratio (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Serious[1]	202	194	HR 0.87 (0.69, 1.09)	Moderate
Adverse events grade 3 or above: leukopenia (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Serious[1]	202	194	RR 0.87 (0.72, 1.05)	Moderate
Adverse events grade 3 or above: neutropenia (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Serious[1]	202	194	RR 0.92 (0.72, 1.18)	Moderate
Adverse events grade 3 or above: anaemia (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Not serious	202	194	RR 0.53 (0.34, 0.82)	High
Adverse events grade 3 or above: thrombocytopenia (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Serious[1]	202	194	RR 0.58 (0.31, 1.10)	Moderate
Adverse events grade 3 or above: worst haematologic toxicity per patient (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Serious[1]	202	194	RR 0.90 (0.77, 1.05)	Moderate

Quality assessment						No of patients		Effect estimate	Quality
No of studies	Design	Risk of bias	Indirectness	Inconsistency	Imprecision	Chemoradio, surgery	Chemoradio	Summary of results (95% CI)	
Adverse events grade 3 or above: nausea and/or emesis (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Not serious	202	194	RR 0.44 (0.27, 0.71)	High
Adverse events grade 3 or above: neuropathy (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Serious[1]	202	194	RR 1.37 (0.53, 3.53)	Moderate
Adverse events grade 3 or above: oesophagitis (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Not serious	202	194	RR 0.44 (0.27, 0.71)	High
Adverse events grade 3 or above: stomatitis and/or mucositis (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Serious[1]	202	194	RR 1.15 (0.36, 3.71)	Moderate
Adverse events grade 3 or above: pulmonary (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Not serious	202	194	RR 0.58 (0.39, 0.87)	High
Adverse events grade 3 or above: other gastrointestinal or renal (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Serious[1]	202	194	RR 1.37 (0.53, 3.53)	Moderate
Adverse events grade 3 or above: cardiac (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Serious[1]	202	194	RR 1.07 (0.44, 2.57)	Moderate
Adverse events grade 3 or above: miscellaneous infection (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Serious[1]	202	194	RR 0.72 (0.25, 2.04)	Moderate
Adverse events grade 3 or above: haemorrhage (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Serious[1]	202	194	RR 0.96 (0.06, 15.25)	Moderate
Adverse events grade 3 or above: fatigue (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Serious[1]	202	194	RR 1.17 (0.50, 2.77)	Moderate
Adverse events grade 3 or above: anorexia (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Serious[1]	202	194	RR 0.41 (0.11, 1.57)	Moderate
Adverse events grade 3 or above: allergy (values greater than 1 favour chemoradio)									
1 (Albain 2009)	RCT	Not serious	Not serious	N/A	Serious[1]	202	194	RR 0.32 (0.03, 3.05)	Moderate

3. 95% CI of the effect size crosses the line of no effect

Chemoradiotherapy, surgery vs chemotherapy, surgery

Quality assessment						No of people		Effect estimate	Quality
No of studies	Design	Risk of bias	Indirectness	Inconsistency	Imprecision	Chemo, surgery	Chemoradiotherapy, surgery	Summary of results	
Mortality: all-cause hazard ratio (values below 1 favour chemoradiotherapy, surgery)									
2 (Katakami 2012, Pless 2015)	RCT	Not serious	Not serious	Not serious	Serious[1]	149	138	HR 0.94 (0.69, 1.27)	Moderate
Mortality: risk ratio for survival at 1 year (values below 1 favour chemoradiotherapy, surgery)									
1 (Girard 2010)	RCT	Serious[2]	Not serious	Not serious	Serious[1]	14	32	RR 1.10 (0.89, 1.36)	Low
Mortality: risk ratio for survival at 2 years (values below 1 favour chemoradiotherapy, surgery)									
1 (Girard 2010)	RCT	Serious[2]	Not serious	Not serious	Serious[1]	14	32	RR 0.87 (0.52, 1.46)	Low
Mortality: risk ratio for survival at 3 years (values below 1 favour chemoradiotherapy, surgery)									
2 (Girard 2010, Katakami 2012)	RCT	Serious[2]	Not serious	Serious[4]	Serious[1]	42	60	RR 0.76 (0.49, 1.18)	Very low
Adverse events grade 3 or above: stomatitis (values above 1 favour chemoradiotherapy, surgery)									
1 (Pless 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	121	110	RR 4.55 (0.54, 38.30)	Moderate
Adverse events grade 3 or above: dyspnoea (values above 1 favour chemoradiotherapy, surgery)									
2 (Katakami 2012, Pless 2015)	RCT	Not serious	Not serious	Not serious	Serious[1]	149	138	RR 8.19 (0.45, 150.38)	Moderate
Adverse events grade 3 or above: pneumonitis (values above 1 favour chemoradiotherapy, surgery)									
1 (Girard 2010)	RCT	Serious[2]	Not serious	Not serious	Serious[1]	14	32	RR 0.73 (0.03, 16.97)	Low

1. 95% CI of the effect size crosses the line of no effect
2. Girard 2010: Randomisation was stratified by clinical centre and histological type (squamous cell carcinoma vs. others). However, the groups were not balanced in terms of gender or pN2/cN2. This might be because of the relatively low numbers of participants. Nevertheless, they were not balanced.

Chemotherapy, chemoradiotherapy + surgery vs chemotherapy, chemoradiotherapy boost

Quality assessment						No of patients		Effect estimate	Quality
No of studies	Design	Risk of bias	Indirectness	Inconsistency	Imprecision	Chemo, chemorad + surgery	Chemo, chemorad boost	Summary of results (95% CI)	
Mortality: risk ratio for survival at 1 year (values over 1 favour chemo, chemorad + surgery)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 0.94 (0.81, 1.10)	Moderate
Mortality: risk ratio for survival at 2 years (values over 1 favour chemo, chemorad + surgery)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.07 (0.84, 1.37)	Moderate
Mortality: risk ratio for survival at 3 years (values over 1 favour chemo, chemorad + surgery)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.08 (0.75, 1.56)	Moderate
Mortality: risk ratio for survival at 4 years (values over 1 favour chemo, chemorad + surgery)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.23 (0.75, 2.04)	Moderate
Mortality: risk ratio for survival at 5 years (values over 1 favour chemo, chemorad + surgery)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.23 (0.69, 2.21)	Moderate
Mortality: risk ratio for survival at 6 years (values over 1 favour chemo, chemorad + surgery)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.12 (0.60, 2.08)	Moderate
Adverse events grade 3 or above: leukopenia (values over 1 favour chemo, chemorad boost)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.01 (0.78, 1.30)	Moderate
Adverse events grade 3 or above: anaemia (values over 1 favour chemo, chemorad boost)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.10 (0.47, 2.56)	Moderate
Adverse events grade 3 or above: thrombocytopenia (values over 1 favour chemo, chemorad boost)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.11 (0.45, 2.74)	Moderate
Adverse events grade 3 or above: nausea/vomiting (values over 1 favour chemo, chemorad boost)									

Quality assessment						No of patients		Effect estimate	Quality
No of studies	Design	Risk of bias	Indirectness	Inconsistency	Imprecision	Chemo, chemorad + surgery	Chemo, chemorad boost	Summary of results (95% CI)	
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.55 (0.63, 3.80)	Moderate
Adverse events grade 3 or above: neuropathy (values over 1 favour chemo, chemorad boost)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 0.99 (0.30, 3.28)	Moderate
Adverse events grade 3 or above: oesophagitis (values over 1 favour chemo, chemorad boost)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Not serious	81	80	RR 0.52 (0.27, 1.00)	High
Adverse events grade 3 or above: mucositis/stomatitis (values over 1 favour chemo, chemorad boost)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.48 (0.25, 8.63)	Moderate
Adverse events grade 3 or above: pulmonary (values over 1 favour chemo, chemorad boost)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.78 (0.62, 5.07)	Moderate
Adverse events grade 3 or above: other GI or renal (values over 1 favour chemo, chemorad boost)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.58 (0.54, 4.62)	Moderate
Adverse events grade 3 or above: cardiac (values over 1 favour chemo, chemorad boost)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.98 (0.37, 10.48)	Moderate
Adverse events grade 3 or above: miscellaneous infection (values over 1 favour chemo, chemorad boost)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 2.30 (0.62, 8.60)	Moderate
Adverse events grade 3 or above: fatigue (values over 1 favour chemo, chemorad boost)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 0.62 (0.21, 1.81)	Moderate
Adverse events grade 3 or above: pain (values over 1 favour chemo, chemorad boost)									
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.17 (0.65, 2.11)	Moderate
Dropout during treatment (values over 1 favour chemo, chemorad boost)									

Quality assessment						No of patients		Effect estimate	Quality
No of studies	Design	Risk of bias	Indirectness	Inconsistency	Imprecision	Chemo, chemorad + surgery	Chemo, chemorad boost	Summary of results (95% CI)	
1 (Eberhardt 2015)	RCT	Not serious	Not serious	N/A	Serious[1]	81	80	RR 1.65 (0.41, 6.66)	Moderate

1. 95% CI of the effect size crosses the line of no effect

Chemotherapy, surgery vs chemotherapy, radiotherapy

Quality assessment						No of patients		Effect estimate	Quality
No of studies	Design	Risk of bias	Indirectness	Inconsistency	Imprecision	Chemo, surgery	Chemo, radio	Summary of results (95% CI)	
Mortality: all-cause hazard ratio (values greater than 1 favour chemo, radio)									
1 (van Meerbeeck 2007)	RCT	Not serious	Not serious	N/A	Serious[2]	154	154	HR 1.06 (0.85, 1.33)	Moderate
Mortality: risk ratio of being alive at 1 year (values greater than 1 favour chemo, surgery)									
1 (Johnstone 2002)	RCT	Very serious[1,3]	Not serious	N/A	Serious[2]	29	32	RR 1.00 (0.69, 1.44)	Very low
Mortality: risk ratio of being alive at 2 years (values greater than 1 favour chemo, surgery)									
1 (Johnstone 2002)	RCT	Very serious[1,3]	Not serious	N/A	Serious[2]	29	32	RR 1.30 (0.70, 2.44)	Very low
Mortality: risk ratio of being alive at 3 years (values greater than 1 favour chemo, surgery)									
1 (Johnstone 2002)	RCT	Very serious[1,3]	Not serious	N/A	Serious[2]	29	32	RR 1.42 (0.61, 3.32)	Very low
Mortality: risk ratio of being alive at 4 years (values greater than 1 favour chemo, surgery)									
1 (Johnstone 2002)	RCT	Very serious[1,3]	Not serious	N/A	Serious[2]	29	32	RR 0.95 (0.36, 2.49)	Very low
Mortality: risk ratio of treatment-related mortality									
1 (Johnstone 2002)	RCT	Very serious[1,3]	Not serious	N/A	Serious[2]	29	32	RR 3.30 (0.14, 77.95)	Very low
Dropout during treatment									

Quality assessment						No of patients		Effect estimate	Quality
No of studies	Design	Risk of bias	Indirectness	Inconsistency	Imprecision	Chemo, surgery	Chemo, radio	Summary of results (95% CI)	
1 (van Meerbeeck 2007)	RCT	Serious[1]	Not serious	N/A	Serious[2]	165	167	HR 0.85 (0.37, 1.95)	Low

1. Incomplete and selective reporting of data
2. 95% CI of the effect size crosses the line of no effect
3. Some participants were not randomised and had different chemotherapy regimens

Chemotherapy, surgery vs radiotherapy

Quality assessment						No of patients		Effect estimate	Quality
No of studies	Design	Risk of bias	Indirectness	Inconsistency	Imprecision	Chemo, surgery	Radio	Summary of results (95% CI)	
Mortality: all-cause									
1 (Shepherd 1998)	RCT	Very serious[1,2]	Not serious	N/A	Very serious[3,4]	16	15	Median survival 18.7 months in chemo, surgery arm (12.9 – 32) Median survival 16.2 months in radio arm (10.7 – 32.3)[5]	Very low
Mortality: all-cause hazard ratio									
1 (Stephens 20015)	RCT	Very serious[6]	Not serious	N/A	Serious[7]	24	24	HR 0.91 (0.49, 1.70)	Very low
Mortality: treatment-related deaths									
1 (Stephens 20015)	RCT	Serious[1]	Not serious	N/A	Serious[7]	24	24	RR 5.00 (0.25, 98.96)	Low
Adverse events grade 2 or above: lethargy									
1 (Stephens 20015)	RCT	Serious[1]	Not serious	N/A	Serious[7]	24	24	RR 1.44 (0.77, 2.72)	Low
Dropout during treatment (values greater than 1 favour radiotherapy)									

Quality assessment						No of patients		Effect estimate	Quality
No of studies	Design	Risk of bias	Indirectness	Inconsistency	Imprecision	Chemo, surgery	Radio	Summary of results (95% CI)	
1 (Shepherd 1998)	RCT	Very serious[1,2]	Not serious	N/A	Very serious[4]	16	15	RR 3.75 (0.47, 29.87)	Very low
Dropout during treatment (values greater than 1 favour radiotherapy)									
1 (Stephens 20015)	RCT	Serious[1]	Not serious	N/A	Serious[7]	24	24	RR 0.11 (0.01, 1.96)	Low

1. Incomplete and selective reporting of data
2. Method of randomisation not given and arms were not balanced at baseline
3. The 95% CIs for the median values overlap
4. Sample size is 25 to 40. Therefore, downgraded once for imprecision
5. However, according to the survival chart, follow-up was only 21 months for radiotherapy (~34% were still alive) and 32 months for chemotherapy, surgery (30% were still alive)
6. High risk of bias
7. 95% CI of the effect size crosses the line of no effect

Chemotherapy, chemoradiotherapy, surgery, radiotherapy vs chemotherapy, surgery, radiotherapy

Quality assessment						No of patients		Effect estimate	Quality
No of studies	Design	Risk of bias	Indirectness	Inconsistency	Imprecision	Chemo, chemorad, surgery, radio	Chemo, surgery, radio	Summary of results (95% CI)	
Mortality: all-cause hazard ratio (values greater than 1 favour chemo, chemorad, surgery, radio)									
1 (Thomas 2008)	RCT	Very serious[1]	Very serious[2]	N/A	Serious[3]	264	260	HR 0.91 (0.49, 1.70)	Very low
Mortality: treatment related: all (values greater than 1 favour chemo, surgery, radio)									
1 (Thomas 2008)	RCT	Very serious[1]	Very serious[2]	N/A	Serious[3]	264	260	RR 1.12 (0.57, 2.19)	Very low
Mortality: treatment related: fatal events after neutropenia caused by chemotherapy (values greater than 1 favour chemo, surgery, radio)									
1 (Thomas 2008)	RCT	Very serious[1]	Very serious[2]	N/A	Serious[3]	264	260	RR 0.66 (0.11, 3.90)	Very low
Mortality: treatment related: oesophagitis (values greater than 1 favour chemo, surgery, radio)									
1 (Thomas 2008)	RCT	Very serious[1]	Very serious[2]	N/A	Serious[3]	206	187	RR 2.72 (0.11, 66.48)	Very low

Quality assessment						No of patients		Effect estimate	Quality
No of studies	Design	Risk of bias	Indirectness	Inconsistency	Imprecision	Chemo, chemorad, surgery, radio	Chemo, surgery, radio	Summary of results (95% CI)	
Mortality: treatment related: pneumonitis (values greater than 1 favour chemo, surgery, radio)									
1 (Thomas 2008)	RCT	Very serious[1]	Very serious[2]	N/A	Serious[3]	206	187	RR 0.08 (0.00, 1.48)	Very low
Mortality: treatment related: surgical mortality (values greater than 1 favour chemo, surgery, radio)									
1 (Thomas 2008)	RCT	Very serious[1]	Very serious[2]	N/A	Serious[3]	142	154	RR 2.01 (0.83, 4.91)	Very low
Adverse events grade 3 or above: haemotoxicity (values greater than 1 favour chemo, surgery, radio)									
1 (Thomas 2008)	RCT	Very serious[1]	Very serious[2]	N/A	Not serious	206	187	RR 18.16 (2.46, 133.96)	Very low
Adverse events grade 3 or above: oesophagitis (values greater than 1 favour chemo, surgery, radio)									
1 (Thomas 2008)	RCT	Very serious[1]	Very serious[2]	N/A	Serious[3]	206	187	RR 5.06 (2.32, 11.03)	Very low
Adverse events grade 3 or above: pneumonitis (values greater than 1 favour chemo, surgery, radio)									
1 (Thomas 2008)	RCT	Very serious[1]	Very serious[2]	N/A	Not serious	206	187	RR 0.21 (0.06, 0.72)	Very low
Adverse events: peri-operative complications (values greater than 1 favour chemo, surgery, radio)									
1 (Thomas 2008)	RCT	Very serious[1]	Very serious[2]	N/A	Serious[3]	142	154	RR 1.51 (0.86, 2.64)	Very low

1. Incomplete and selective reporting of data. Over 20% of participants were lost to follow-up with regards to adverse events data
2. Participants who were N2 were in the minority: chemo, chemoradio, surgery = 17%; chemo, surgery = 12%. 349 of 524 patients (67%) had stage IIIB disease and comprised a substantial proportion of 113 of 524 patients (22%) with pathologically confirmed N3 disease
3. 95% CI of the effect size crosses the line of no effect

Appendix G – Meta-analyses

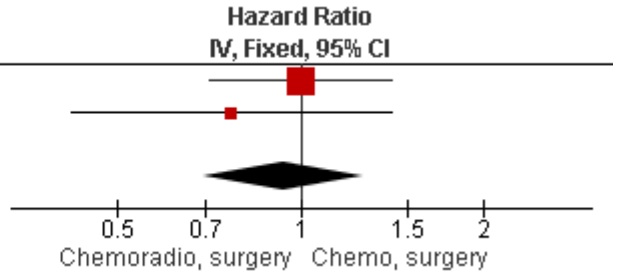
Randomised controlled trials

Chemoradiotherapy, surgery vs chemotherapy, surgery

Mortality: all-cause hazard ratio

Study or Subgroup	log[Hazard Ratio]	SE	Weight	Hazard Ratio IV, Fixed, 95% CI
Pless 2015	0	0.176823	75.3%	1.00 [0.71, 1.41]
Katakami 2012	-0.26136	0.308952	24.7%	0.77 [0.42, 1.41]
Total (95% CI)			**100.0%**	**0.94 [0.69, 1.27]**

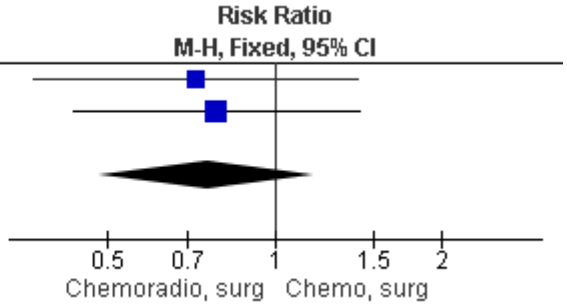
Heterogeneity: Chi² = 0.54, df = 1 (P = 0.46); I² = 0%
Test for overall effect: Z = 0.42 (P = 0.67)



Mortality: risk ratio for survival at 3 years

Study or Subgroup	Chemo, surgery Events	Total	Chemoradio, surg Events	Total	Weight	Risk Ratio M-H, Fixed, 95% CI
Girard 2010	6	14	19	32	45.2%	0.72 [0.37, 1.41]
Katakami 2012	11	28	14	28	54.8%	0.79 [0.44, 1.42]
Total (95% CI)		**42**		**60**	**100.0%**	**0.76 [0.49, 1.18]**
Total events	17		33			

Heterogeneity: Chi² = 0.03, df = 1 (P = 0.85); I² = 0%
Test for overall effect: Z = 1.23 (P = 0.22)



1 Appendix H – Excluded Studies

2 Excluded clinical studies

Study	Title	Reason for exclusion
Billiet 2016	Postoperative radiotherapy for lung cancer: Is it worth the controversy?	Paper on postoperative radiotherapy, not tri-modality treatment.
Chen 2018	Comparing the benefits of chemoradiotherapy and chemotherapy for resectable stage III A/N2 non-small cell lung cancer: a meta-analysis	The studies used in this systematic review were checked to ensure that we included all relevant ones.
Cheng 2005	Predicting efficacy of neoadjuvant cheomotherapy on resectable stage IIIA non-small cell lung cancer by multi-gene expressions	This study is not written in English. In addition, it is on the prognostic value of gene expressions
Guberina 2017	Heart dose exposure as prognostic marker after radiotherapy for resectable stage IIIA/B non-small-cell lung cancer: secondary analysis of a randomized trial	This is a secondary analysis of Eberhardt 2015. However, the data was not analysed as an RCT. Both arms were placed into the same group
Pass 1992	Randomized trial of neoadjuvant therapy for lung cancer: interim analysis	The comparison of 'surgery, radiotherapy vs chemotherapy, surgery, chemotherapy' is not in the protocol
Pezzetta 2005	Comparison of neoadjuvant cisplatin-based chemotherapy versus radiochemotherapy followed by resection for stage III (N2) NSCLC	Retrospective study
Pottgen 2017	Definitive radiochemotherapy versus surgery within multimodality treatment in stage III non-small cell lung cancer (NSCLC) - a cumulative meta-analysis of the randomized evidence	Not a systematic review. This is a meta-analysis of selected studies. This meta-analysis also includes a study that is conference proceedings. The studies used in this meta-analysis were checked to ensure that we included all relevant ones.
Shah 2011	Induction chemoradiotherapy is not superior to induction chemotherapy alone in patients with stage IIIA(N2) non-small cell lung cancer: a systematic review and meta-analysis	Conference proceedings. This abstract has a lot of information. However, this systematic review used 2 studies that were abstracts (conference proceedings). It also includes 2 retrospective studies. The studies used in this systematic review were checked to ensure that we included all relevant ones.
Shah 2012	Induction chemoradiation is not superior to induction chemotherapy alone in stage IIIA lung cancer	Systematic review contains mostly retrospective studies and conference proceedings. This systematic review used 2 studies that were abstracts (conference proceedings). It also includes 3 retrospective studies. The studies used in this systematic review were checked to ensure that we included all relevant ones.
Sorensen 2013	Surgery for NSCLC stages T1-3N2M0 having preoperative pathologically verified N2 involvement: a prospective	Conference proceedings

Study	Title	Reason for exclusion
	randomized multinational phase III trial by the Nordic Thoracic Oncology Group	

Excluded economic studies

Paper	Primary reason for exclusion
Bongers, M.L., de Ruysscher, D., Oberije, C., Lambin, P., Uyl-de Groot, C.A., Belderbos, J. and Coupe, V.M., 2017. Model-based cost-effectiveness of conventional and innovative chemo-radiation in lung cancer. *International journal of technology assessment in health care*, *33*(6), pp.681-690.	Not a cost-utility paper that met the PICOS criteria.
Louie, A.V., Rodrigues, G.B., Palma, D.A. and Senan, S., 2014. Measuring the population impact of introducing stereotactic ablative radiotherapy for stage I non-small cell lung cancer in Canada. *The oncologist*, *19*(8), pp.880-885.	Not a cost-utility paper that met the PICOS criteria.

Appendix I – References

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Health Economic studies – Included

None

Health Economic studies – Excluded

Bongers, M.L., de Ruysscher, D., Oberije, C., Lambin, P., Uyl-de Groot, C.A., Belderbos, J. and Coupe, V.M., 2017. Model-based cost-effectiveness of conventional and innovative chemo-radiation in lung cancer. *International journal of technology assessment in health care*, *33*(6), pp.681-690.

Louie, A.V., Rodrigues, G.B., Palma, D.A. and Senan, S., 2014. Measuring the population impact of introducing stereotactic ablative radiotherapy for stage I non-small cell lung cancer in Canada. *The oncologist*, *19*(8), pp.880-885.

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Appendix J – Network Meta-analysis

Background

Evidence synthesis was performed for survival outcomes and for adverse events associated with the three interventions of interest; chemoradiotherapy (CR), chemotherapy and surgery (CS) and chemoradiotherapy and surgery (CRS). In this review, all studies provided Kaplan Meier curves for progression free survival (PFS) and overall survival (OS). Visual inspection of the Kaplan Meier curves revealed that the proportional hazards assumption did not appear to hold, and so traditional pooling of hazards ratios was not considered appropriate. Furthermore, the shapes of the survival curves were different across studies, suggesting that it was not appropriate to synthesise the evidence under an assumption of a single parametric model. A non-parametric approach to evidence synthesis was therefore required.

An alternative measure of treatment effect for time-to-event outcomes is the difference in the restricted mean survival time (RMST) [1], where RMST is the mean survival time accrued from randomisation up to *T* years. RMST can be estimated by the area under the survival curve (AUC) up to time T, and the treatment effect estimated as the difference in AUCs between treatments. This measure does not assume proportional hazards and can be calculated regardless of the curve fitted to the data, including directly from the Kaplan-Meier curve, and so can allow for different survival distributions across studies.

In addition, the PFS and OS outcomes are related, because OS is a sum of progression free survival (PFS) and post-progression survival (PPS). Joint modelling of OS and PFS, where the synthesis model is given to PFS and PPS, ensures that predictions from the model conform to the natural constraint that OS is always greater than PFS.

We begin by describing the Network Meta-Analysis (NMA) methods used to estimate the treatment effects on the area under the Kaplan Meier curves for OS and PFS jointly. We then describe how these estimates can be combined with external evidence on longer-term survival to estimate mean time in PFS and PPS on each treatment. Because the non-parametric approach taken means that it is not straightforward to apply discounting in the economic model, we describe how the NMA is adapted to obtain discounted mean survival times required for the economic model. We also describe the NMA model used to synthesis evidence on adverse events. We then describe how we selected models on the basis of model fit and checked for inconsistency in the NMAs. We then present the results from the NMAs and the estimates to be inputted into the economic model.

Synthesising the Clinical Evidence: Methods

Data extraction

Data was extracted from the Kaplan Meier curves using a validated algorithm that makes use of the digitized curves as well as data on the numbers at risk and total number of events [2]. For each treatment group within each study, this produces a set of individual patient data (survival times and censor times) that produce Kaplan-Meier curves similar to those published. This was done for both the PFS and OS curves.

Calculating the Area Under the Kaplan Meier Curves

Kaplan Meier curves were fitted to the extracted data using the survfit function from the survival package in R (v. 3.4.2)[3, 4]. The area under the Kaplan Meier curves from randomisation $t_0 = 0$ to a truncated follow up time t_T was calculated as a Reimann sum

$$AUC_{KM} = \sum_{i=1}^{N} \left(t_i - t_{i-1}\right) \hat{S}_{KM}\left(t_{i-1}\right)$$

where

$$N = \begin{cases} \text{number of distinct event times between } \mathrm{t}_0 \text{ and } t_T & \text{if an event occurs at } \mathrm{t}_T \\ \left(\text{number of distinct event times between } \mathrm{t}_0 \text{ and } t_T\right) + 1 & \textit{otherwise} \end{cases},$$

t_i are the ordered event times, and $\hat{S}_{KM}\left(t_{i-1}\right)$ is the probability of survival at time t_{i-1}. The variance of the AUC was estimated as [5]

$$\hat{V}\left(AUC_{KM}\right) = \sum_{i=1}^{N-1} \frac{d_{(i)}}{n_{(i)}\left(n_{(i)} - d_{(i)}\right)} \left(\sum_{j=i}^{N-1} \left(t_{j+1} - t_j\right) \hat{S}_{KM}\left(t_{j+1}\right) \right)^2$$

where $d_{(i)}$ is the number of patients who experienced an event at time t_i and $n_{(i)}$ is the number of people at risk at time t_i.

All studies report Kaplan Meier curves up until T=5 years, with the exception of Girard (2009) which reports up to *T*=4 years. We use *T*=5 years to estimate differences in the restricted mean survival time in the base-case (which excludes Girard 2009) and use *T*=4 years in a sensitivity analysis (which includes all studies).

The areas under the Kaplan Meier curves for each RCT are provided in Model Critique

Assessing model fit

The posterior mean of the residual deviance, which measures the magnitude of the differences between the observed data and the model predictions of the data, was used to assess the goodness of fit of each model [12]. Smaller values are preferred, and in a well-fitting model the posterior mean residual deviance should be close to the number of data points in the network (each study arm contributes 1 data point) [12].

In addition to comparing how well the models fit the data using the posterior mean of the residual deviance, models were compared using the deviance information criterion (DIC). This is equal to the sum of the posterior mean deviance and the effective number of parameters, and thus penalizes model fit with model complexity [12]. Lower values are preferred and differences of at least 5 points were considered meaningful [12].

Assessing heterogeneity and inconsistency

Heterogeneity concerns the differences in treatment effects between trials within each treatment contrast, while consistency concerns the differences between the direct and indirect evidence informing the treatment contrasts [9, 10].

Heterogeneity is assessed by comparing the fit of fixed and random effects NMA models. The fixed effect model assumes that all trials are estimating the same treatment effect, regardless of any differences in the conduct of the trials, populations, or treatments. The random effects NMA model on the other hand accounts for any differences in treatment effects between trials, that are beyond sampling error, by assuming a distribution of study-specific treatment effects with a pooled mean and between-study standard deviation. The estimated between study standard deviation in treatment effects is also inspected to assess heterogeneity.

Inconsistency was assessed by comparing the fit of the chosen consistency model (fixed or random effects) to an “inconsistency”, or unrelated mean effects, model [9, 10]. The latter is equivalent to having separate, unrelated, meta-analyses for every pairwise contrast, with a common variance parameter assumed in the case of random effects models. Note that inconsistency can only be assessed when there are closed loops of direct evidence on 3 treatments that are informed by at least 3 distinct trials [11].

Network meta-analysis: Results of Clinical Evidence Synthesis

5-year Follow-up

Five studies presented survival data up to 5-years, and a network diagram summarizing the evidence is given in Figure 2

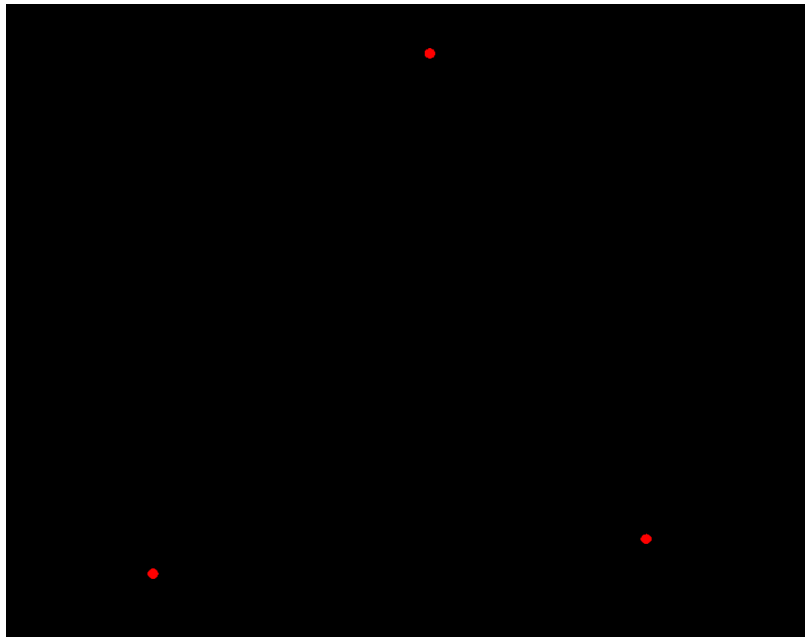


Figure 2: Network diagram of comparisons for which direct evidence on differences in restricted mean survival time up to 5-years is available. Lines are proportional to the number of studies that compare the two connected treatments.

Model fit statistics for the area under the Kaplan Meier curves up to 5-years, as well as the probability of survival are given in Table 12. Convergence was satisfactory for the fixed effect model after a burn-in of 20,000 iterations and results are based on a further 40,000 samples on two chains. For the random effects model, convergence was satisfactory after a burn-in of 30,000 iterations and results are based on a further 60,000 samples on two chains.

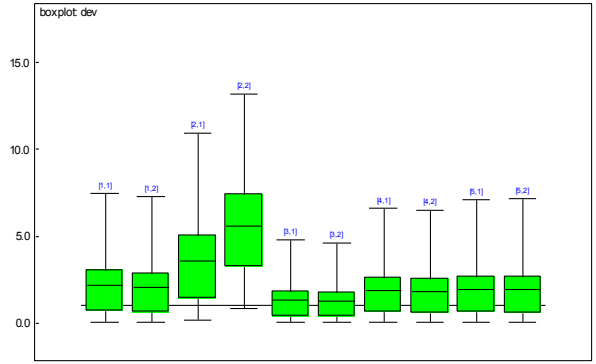
Table 12: Model fit statistics based on 5-year follow-up data

Model		**Median Between-Study SD (95% CrI)**	**Posterior mean residual deviance**	**DIC**
Fixed effect	P(Survival)	---	9.267	-24.852
	AUC		23.47	-11.075

Random effects	P(Survival)	0.35 (0.02, 2.41)	9.618	-22.809
	AUC	PFS: 0.18 (0.01, 1.32) PPS: 0.25 (0.03, 1.46)	18.95	-11.781

Total number of data points for P(survival) is 10 and for AUC is 20.

There were no meaningful differences between the fixed and random effects models in terms of the posterior mean residual deviance and DIC for both NMAs (Table 12). The box plots of the posterior deviance values for each study arm in Figure 3 show that the area under the Kaplan Meier curves up to 5 years in Eberhardt 2015 is not predicted well and this study is a possible outlier. Although the prediction of this study improves in the random effects model (Figure 4), this comes at a cost of slight overfit of the model (posterior mean residual deviance = 18.95, compared to 20 datapoints) and additional parameters in the model. In addition, progression events and deaths were rare in the chemoradiotherapy group of this study after 3-years and 4-years, respectively. The simpler fixed effect model was therefore selected in the base-case.



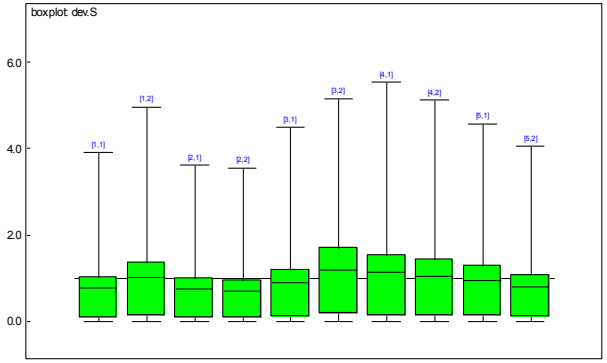


Figure 3: Posterior deviance values for each study arm for the area under the Kaplan Meier curves (left) and probability of survival (right) – fixed effect model.

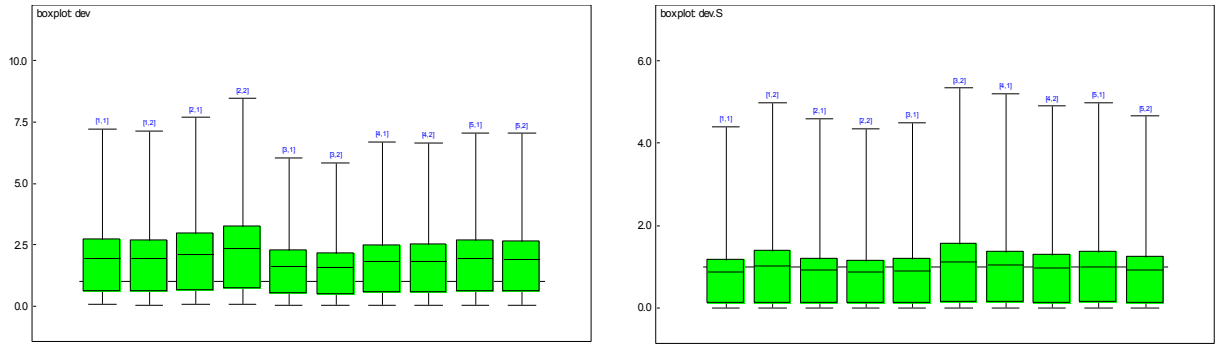


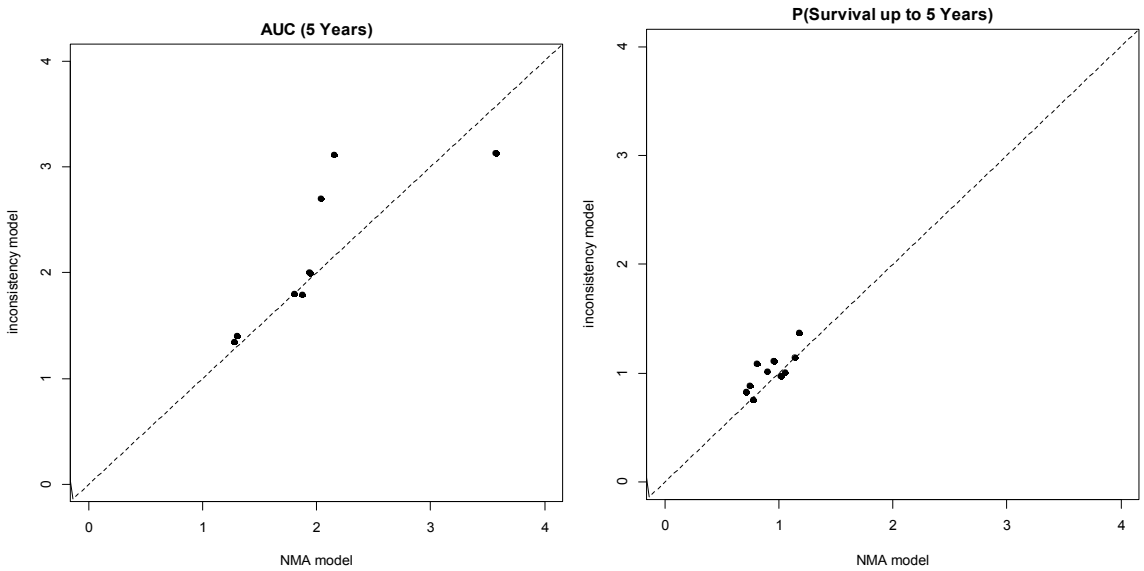
Figure 4: Posterior deviance values for each study arm for the area under the Kaplan Meier curves (left) and probability of survival (right) – random effects model.

No evidence of inconsistency was found, with model fit (posterior mean residual deviance) similar for the consistency and inconsistency (unrelated means) fixed effect models, and a lower DIC for the consistency model (Table 13). The area below the line of equality in Figure 5 highlights where the inconsistency model better predicted data points, and any improvement is minimal.

Table 13: Model fit statistics for consistency and inconsistency fixed effect models based on 5-year follow-up data

Model		Posterior mean residual deviance	DIC
Fixed effect - consistency	P(Survival)	9.267	-24.852
	AUC	23.47	-11.075
Fixed effect - inconsistency	P(Survival)	10.17	-22.867
	AUC	23.65	-8.882

Total number of data points for P(survival) is 10 and for AUC is 20.



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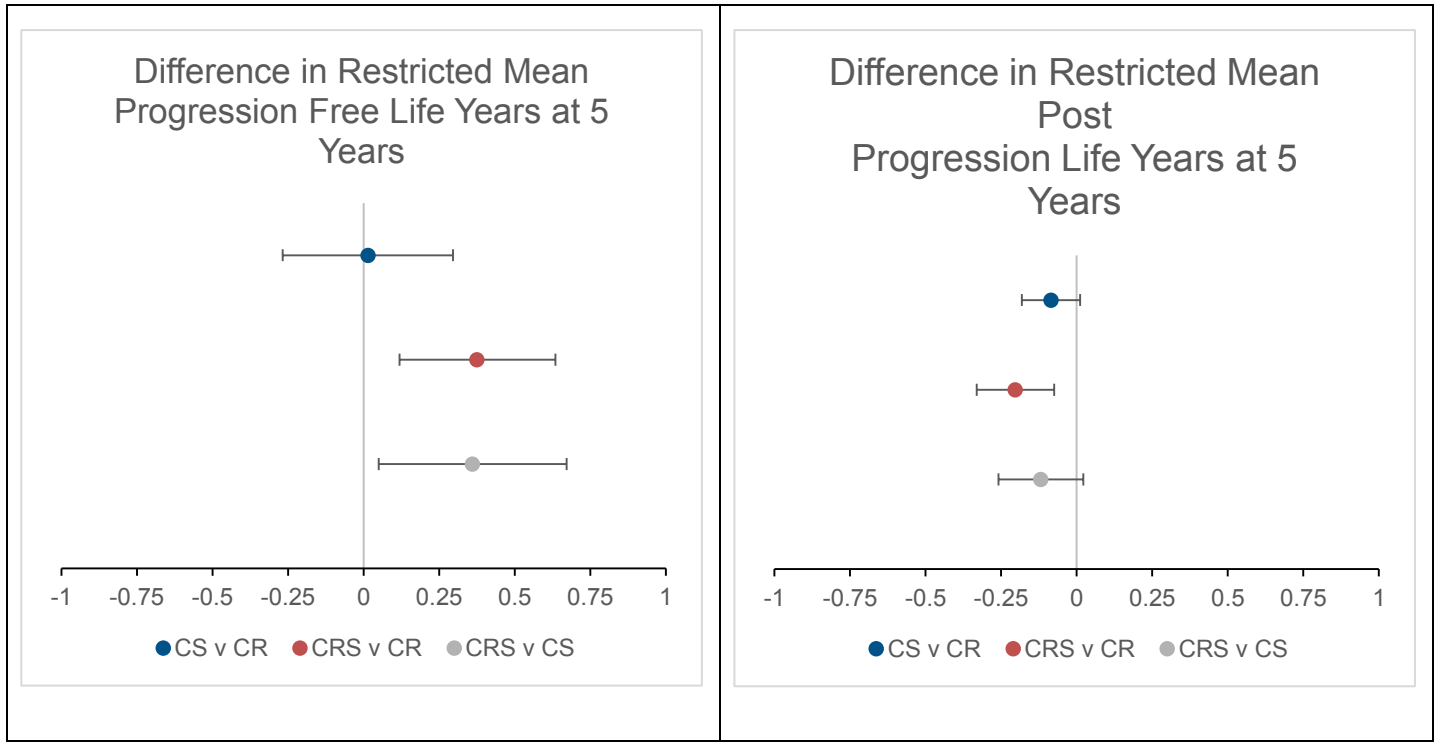
199 **Figure 5: Deviance contributions from the fixed effect consistency and inconsistency models for area under the Kaplan Meier curves**
200 **(left) and probability of survival (right).**

201 There is evidence to suggest that chemoradiotherapy + surgery is more effective in increasing progression free life years at 5-year follow-up
202 compared to chemoradiotherapy alone, while there is no evidence to suggest the effect of chemotherapy + surgery is any different from
203 chemoradiotherapy (


Difference in Restricted Mean Progression Free Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS

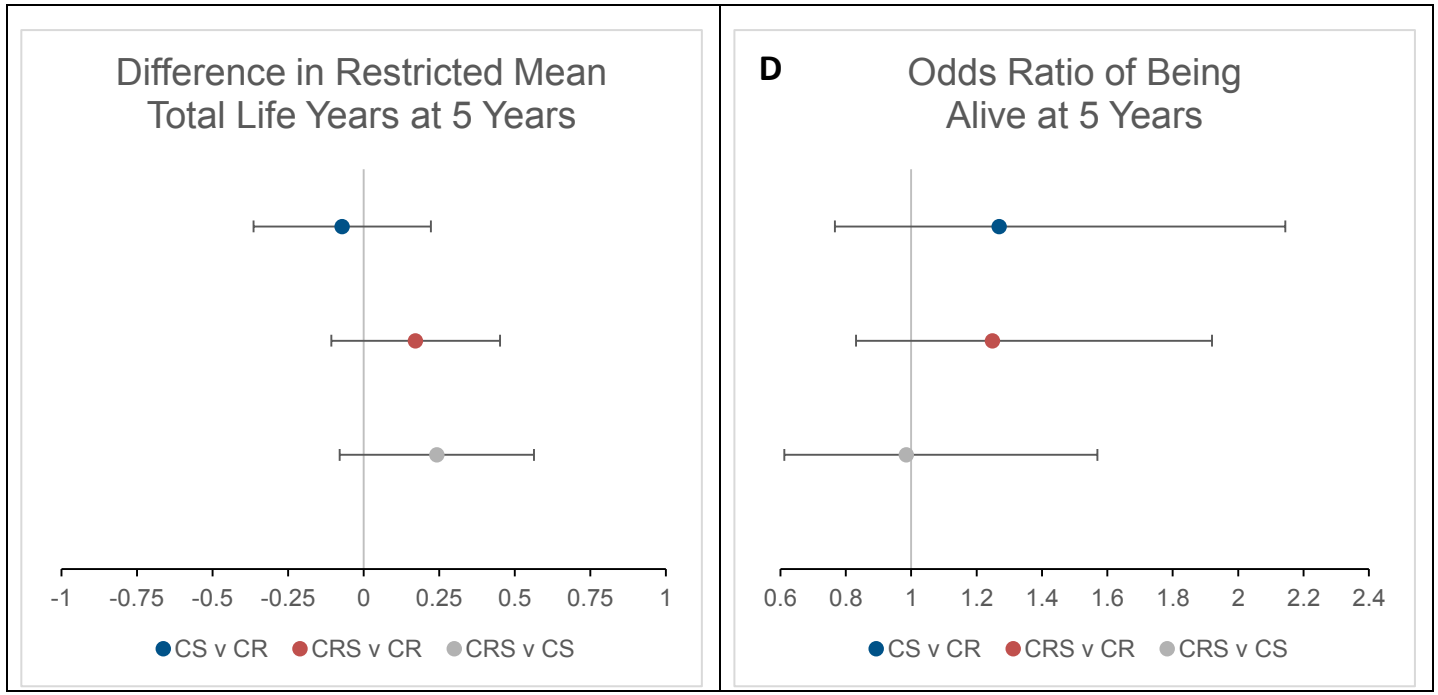
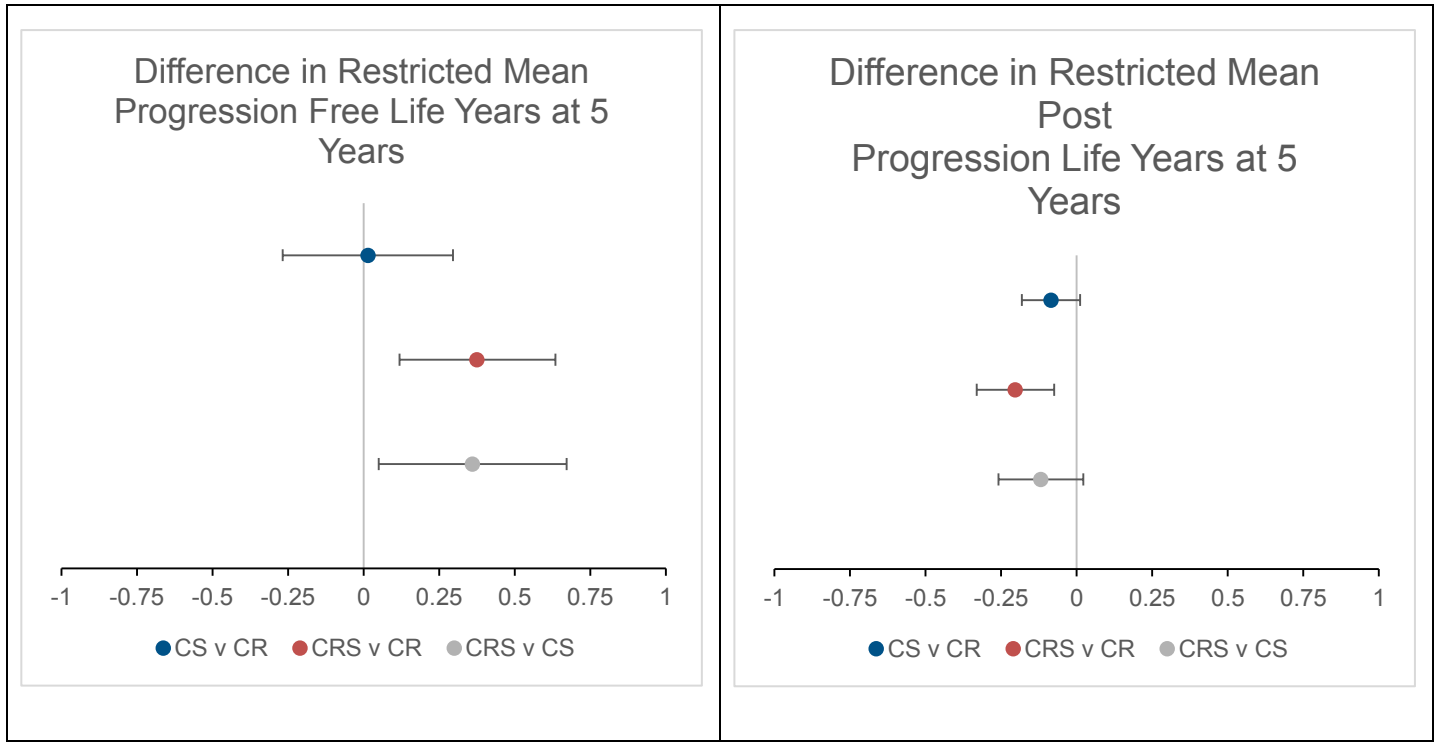


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

A, Table 14). There is also evidence to suggest that chemoradiotherapy + surgery improves progression free life years compared to chemotherapy + surgery (posterior median difference in RMST: 0.36 (95% CrI: 0.05, 0.67)) and it ranked the most effective intervention in increasing progression free life years (Table 14).

213 In terms of post progression life years at 5-year follow-up, there is evidence suggesting that chemoradiotherapy is more effective than
214 chemoradiotherapy + surgery (


Difference in Restricted Mean Progression Free Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS

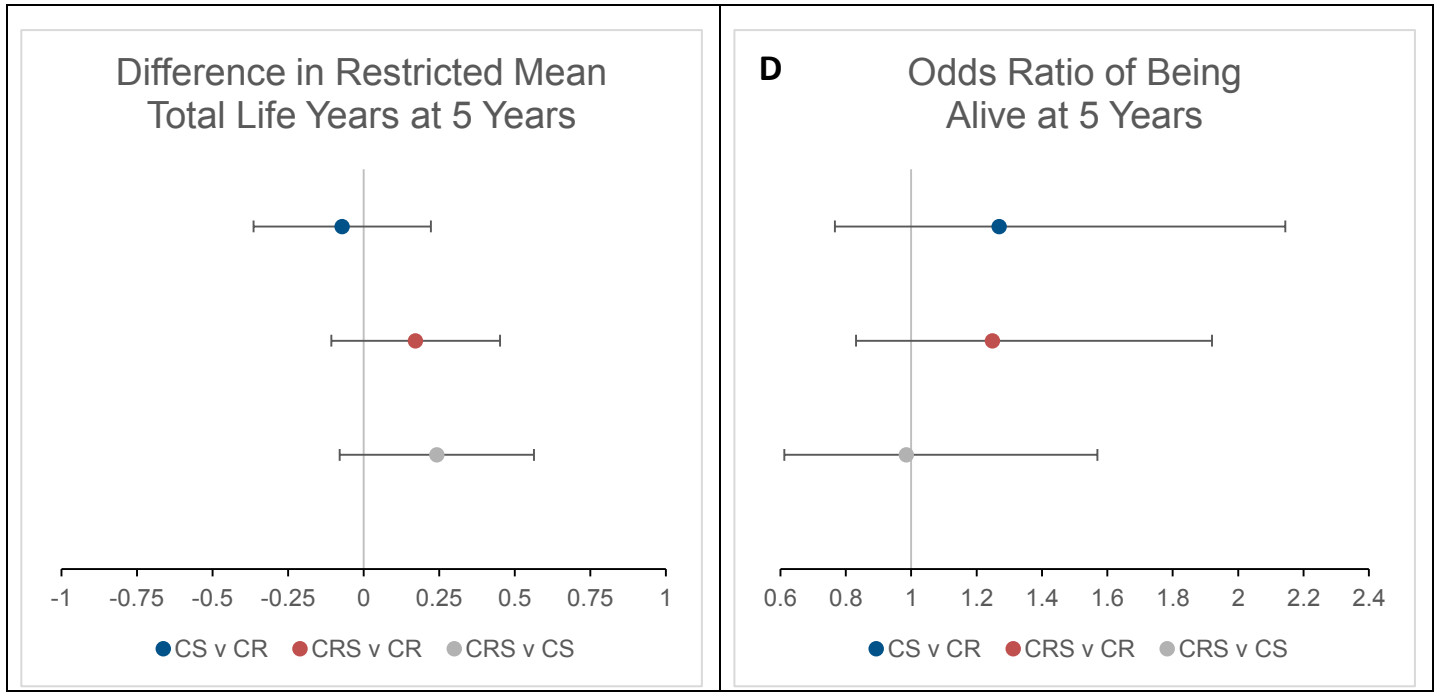
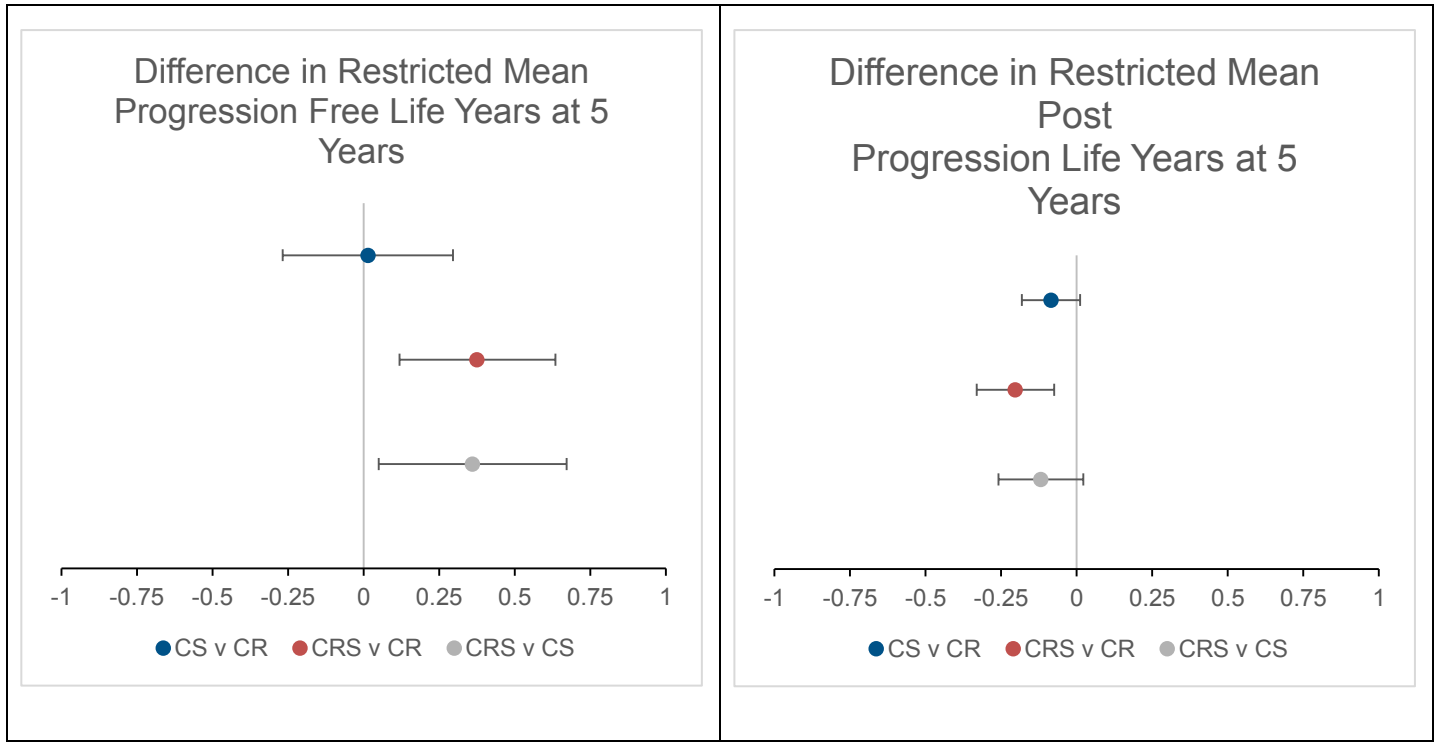


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

B, Table 14). Chemoradiotherapy appears to be more effective than chemotherapy + surgery as well, but this cannot be concluded with high certainty (


Difference in Restricted Mean Progression Free Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS

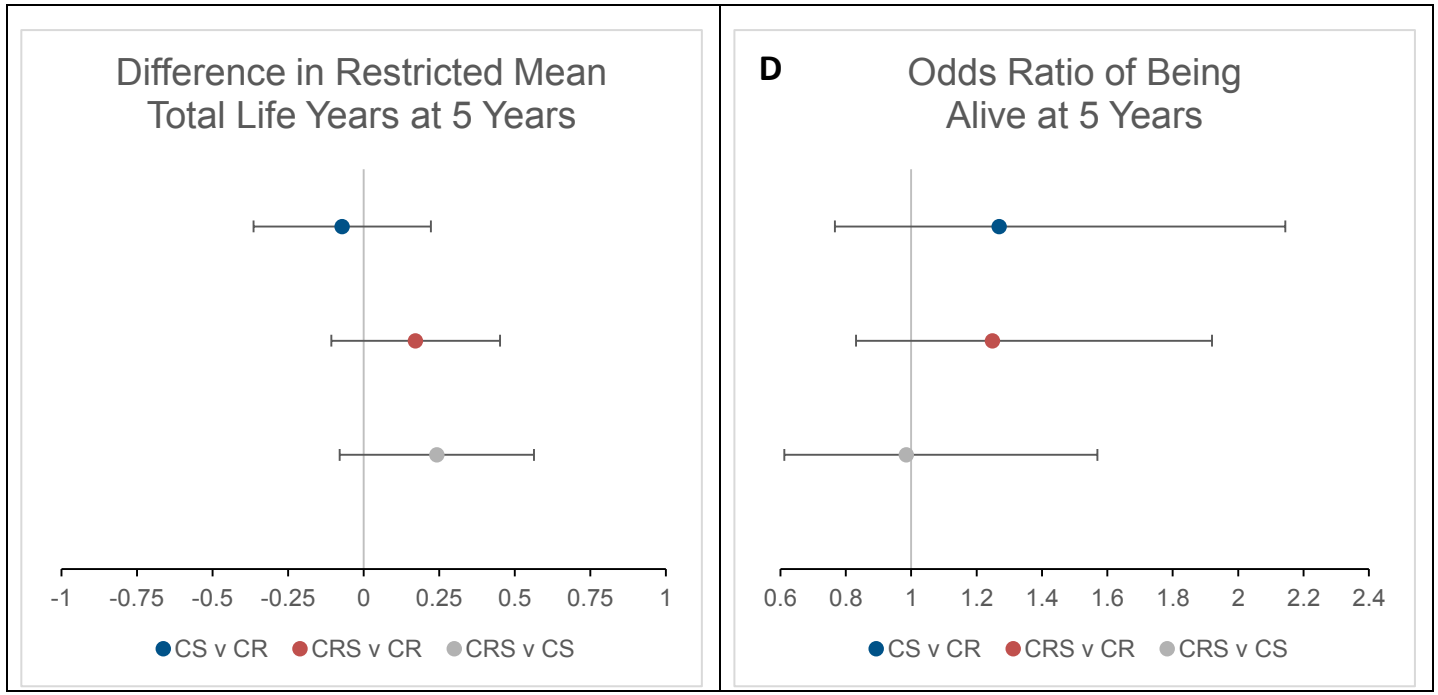
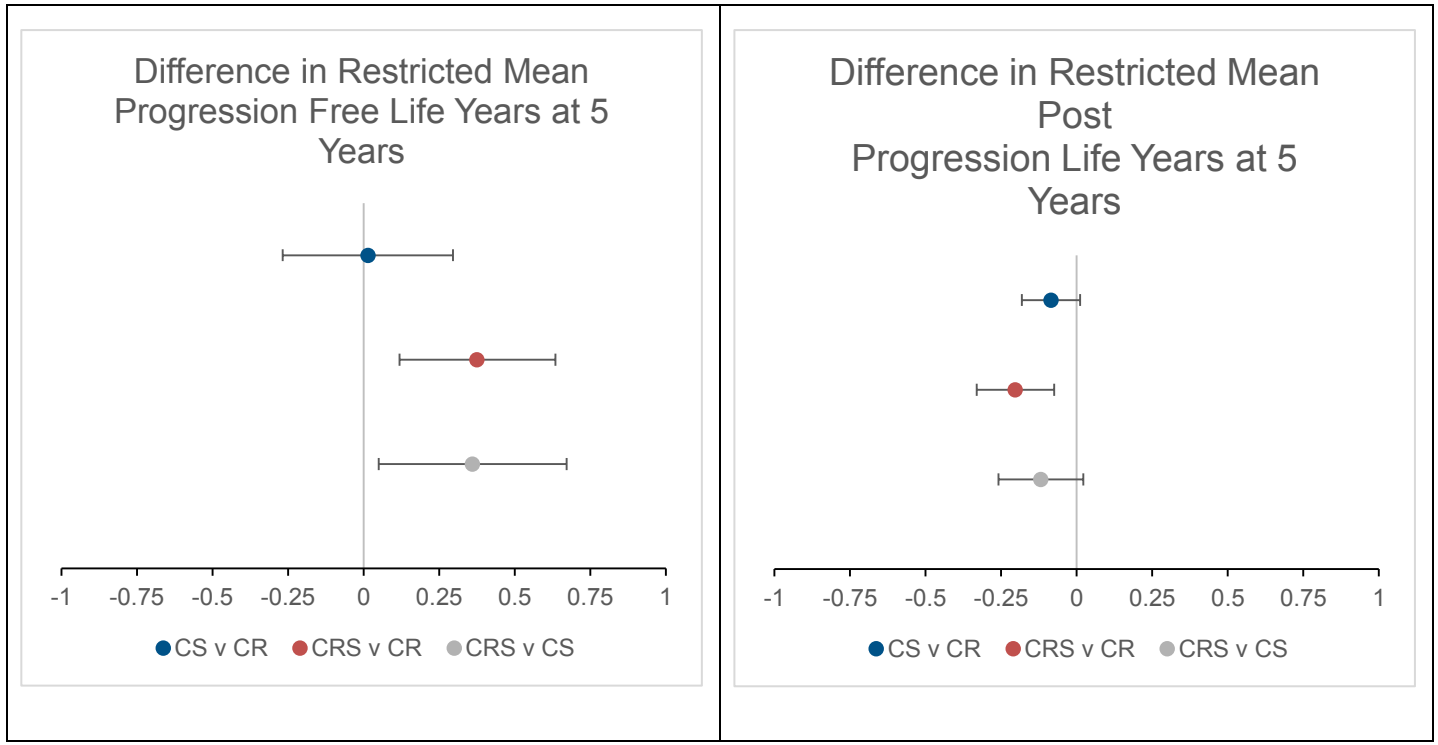


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

B, Table 14). There was not enough evidence to suggest any of the three treatments were different from each other in terms of improving total life years at 5-year follow- up, which is the sum of the progression free and post progression life years (


Difference in Restricted Mean Progression Free Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS

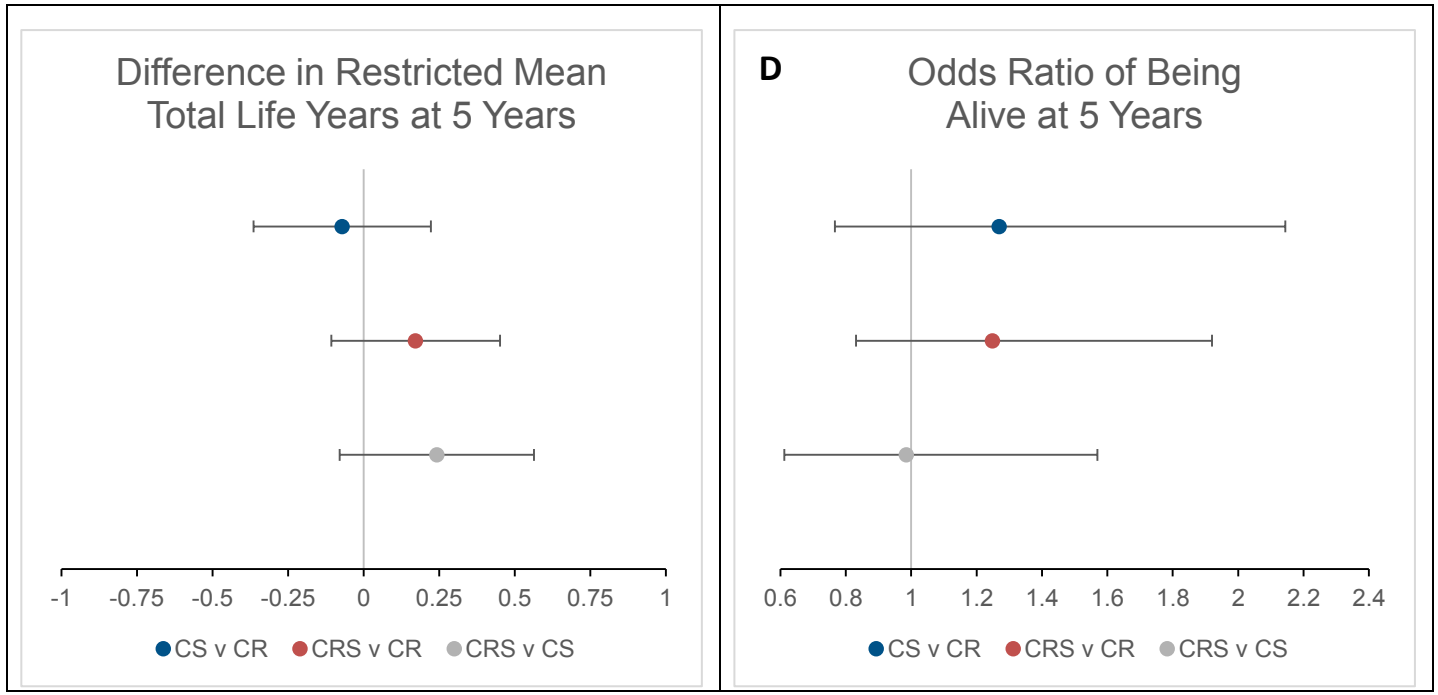
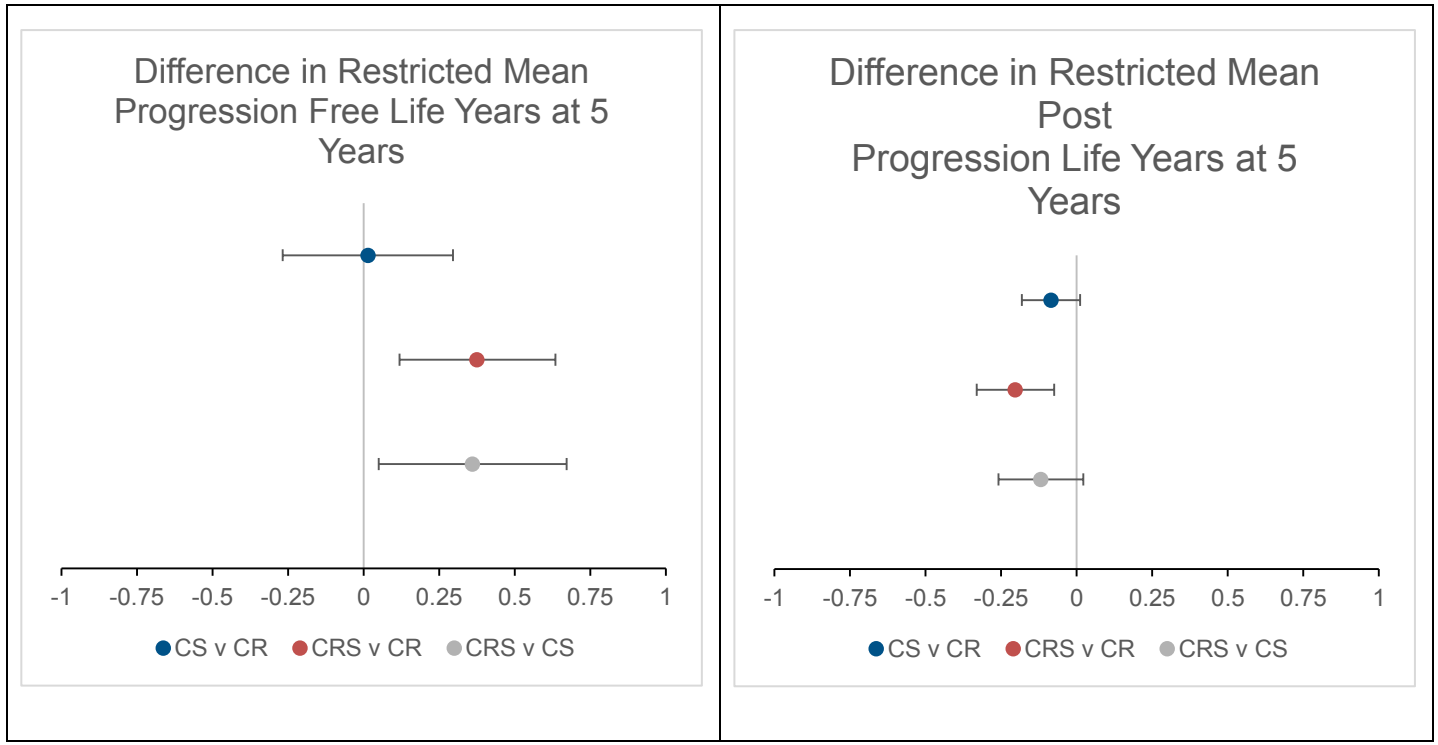
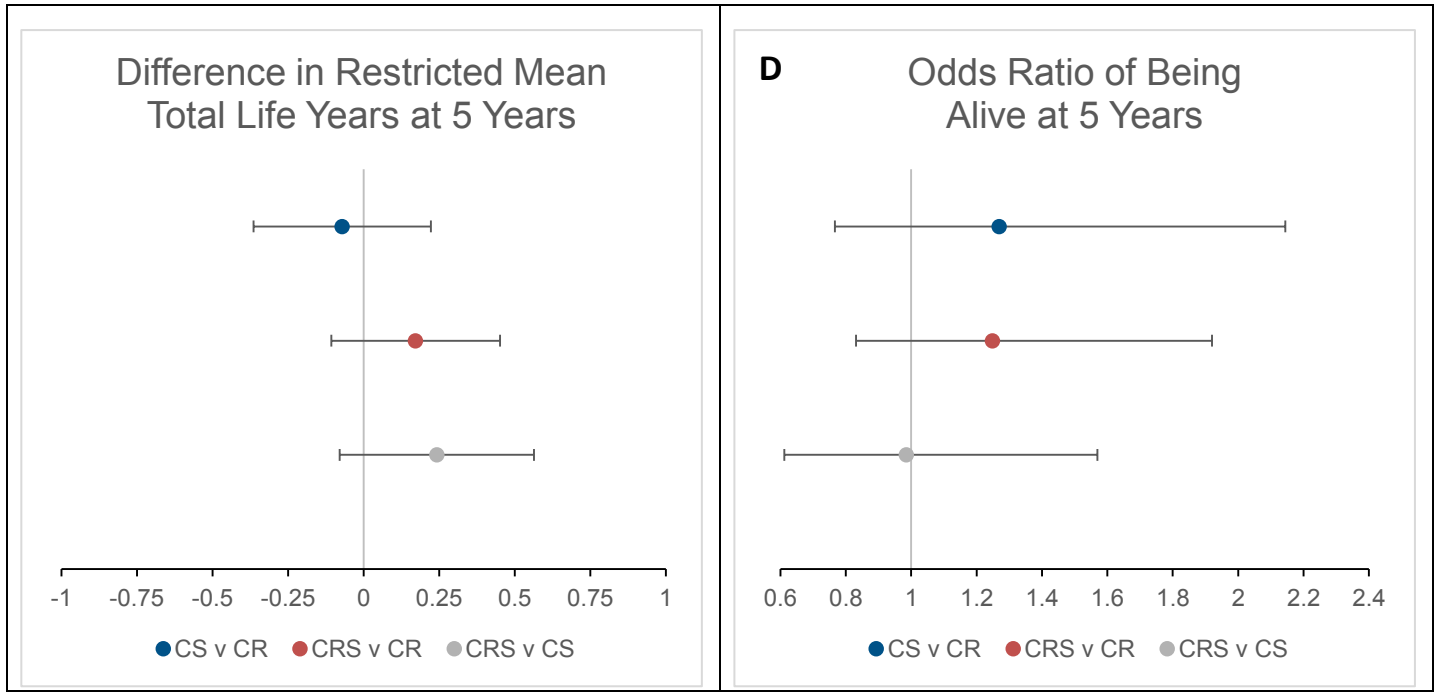


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

C, Table 14).

Chemotherapy + surgery and chemoradiotherapy + surgery appear to be more likely to improve the odds of being alive at 5-years compared to chemoradiotherapy alone, but there is not enough evidence to infer the direction of effects with certainty (


Difference in Restricted Mean Progression Free Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS



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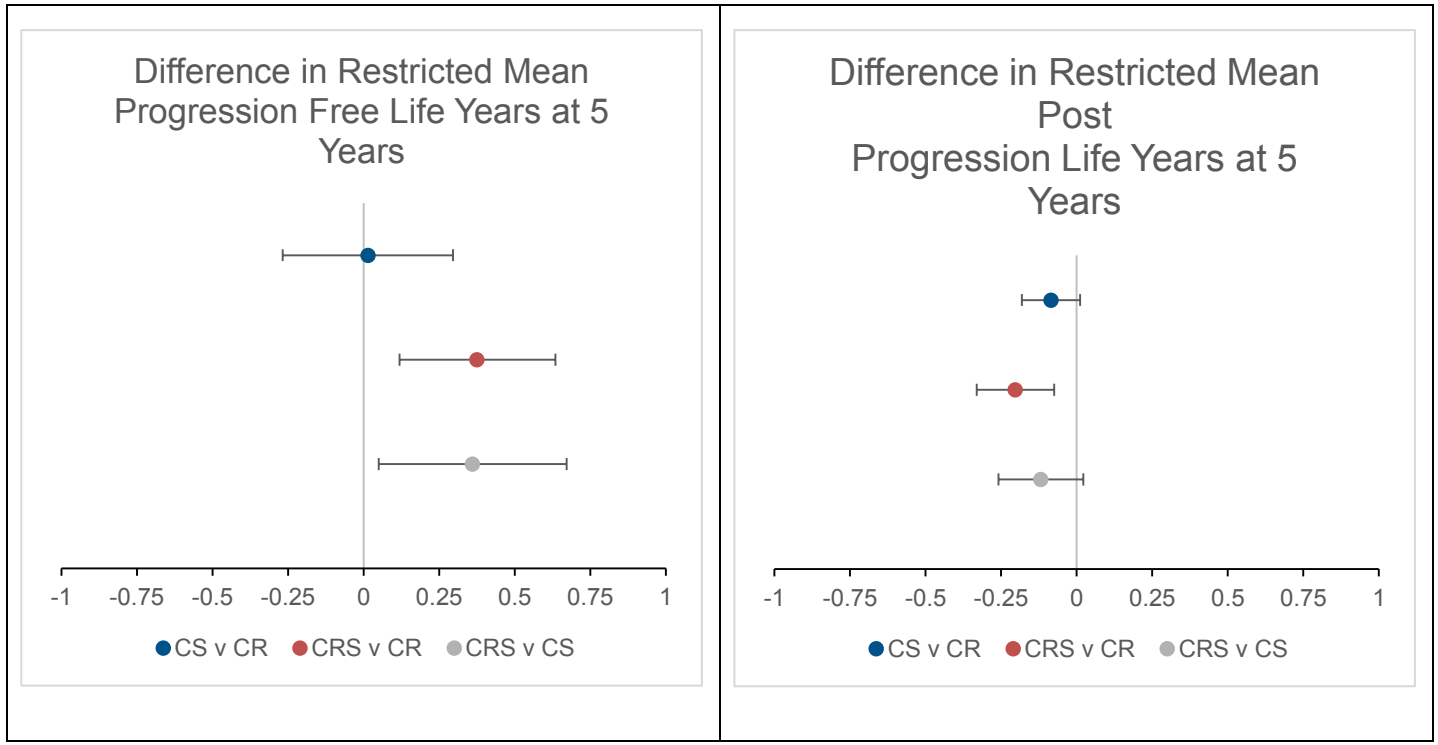
D, Table 14).


Difference in Restricted Mean Progression Free Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS

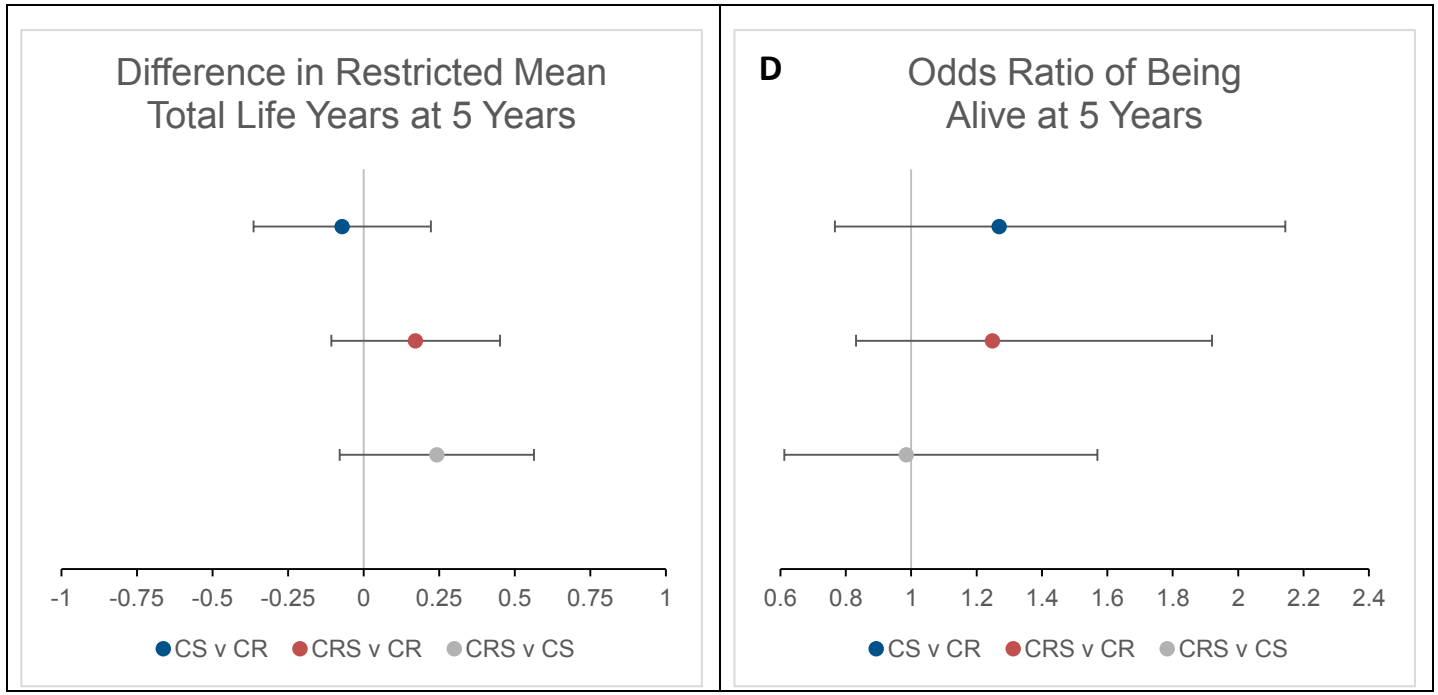


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

Table 14: Treatment differences in restricted mean survival times (RMST) up to 5 years, odds ratios of being alive at 5-years, probabilities of ranking best, ranks, and predicted RMST and probability of being alive at 5-years in the UK population for the three interventions.

		Intervention		
		Chemoradiotherapy[a]	Chemotherapy + Surgery	Chemoradiotherapy + Surgery
Difference in RMST (95% CrI[b])	Progression Free Life Years at 5 Years	Reference Treatment	0.01 (-0.27, 0.3)	0.38 (0.12, 0.63)
	Post Progression Life Years at 5 Years		-0.09 (-0.18, 0.01)	-0.2 (-0.33, -0.07)
	Total Life Years at 5 Years		-0.07 (-0.36, 0.22)	0.17 (-0.11, 0.45)
Odds Ratio (95% CrI)	Being Alive at 5 Years		1.27 (0.77, 2.14)	1.25 (0.83, 1.92)
Probability of Ranking Best	Progression Free Life Years at 5 Years	0.2%	1.1%	98.7%
	Post Progression Life Years at 5 Years	95.8%	4.1%	0.1%
	Total Life Years at 5 Years	9.9%	5.4%	84.7%
	Being Alive at 5 Years	6.3%	50.2%	43.6%
Median Rank (95% CrI)	Progression Free Life Years at 5 Years	3 (2, 3)	2 (2, 3)	1 (1, 1)

	Post Progression Life Years at 5 Years	1 (1, 2)	2 (1, 3)	3 (2, 3)
	Total Life Years at 5 Years	3 (1, 3)	2 (1, 3)	1 (1, 3)
	Being Alive at 5 Years	3 (1, 3)	1 (1, 3)	2 (1, 3)
Predicted RMST and Probability of Being Alive in UK at 5 Years[c]	Mean Progression Free Life Years	1.5 (1.28, 1.71)	1.51 (1.29, 1.73)	1.87 (1.57, 2.17)
	Mean Post Progression Life Years	0.58 (0.51, 0.65)	0.49 (0.42, 0.56)	0.37 (0.24, 0.51)
	Mean Total Life Years	2.07 (1.85, 2.29)	2 (1.77, 2.23)	2.24 (1.93, 2.56)
	Probability of Being Alive at 5 Years	0.13 (0.08, 0.18)	0.16 (0.11, 0.21)	0.16 (0.1, 0.23)

[a] Relative treatment effects presented for comparisons versus chemoradiotherapy. Point estimates are based on posterior medians.
[b] CrI = Credible Interval
[c] Baseline based on posterior distributions of outcomes for van Meerbeeck 2007.

Sensitivity analyses

As part of an assessment of the sensitivity of the results to the selected follow-up time, we also synthesised data based on a shorter follow-up period of 4-years, which allowed the inclusion of all 6 studies, including Girard 2009. Model fit statistics for the fixed and random effects models based on the 4-year follow-up data are given in Table 15Table 15. Convergence was satisfactory for the both models after a burn-in of 20,000 iterations and results are based on a further 40,000 samples on two chains.

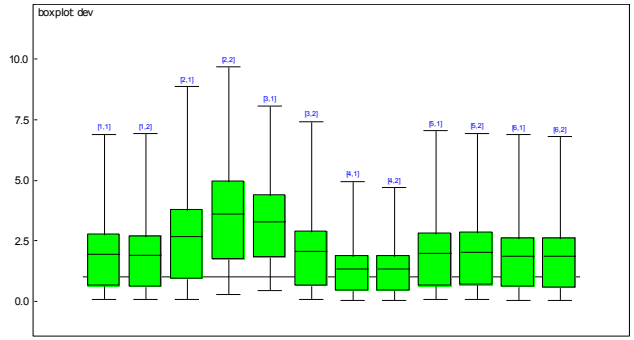
Table 15: Model fit statistics based on 4-year follow-up data

Model			DIC

		Posterior Median Between-Study SD (95% CrI)	Posterior mean residual deviance	
Fixed effect	P(Survival)	---	13.22	-27.429
	AUC		25.84	-20.356
Random effects	P(Survival)	0.24 (0.02, 1.63)	14.29	-25.090
	AUC	PFS: 0.12 (0.01, 0.76) PPS: 0.14 (0.01, 0.59)	23.61	-18.623

Total number of data points for P(survival) is 12 and for AUC is 24.

There were no meaningful differences between the fixed and random effects models in terms of the posterior mean residual deviance and DIC (Table 15). The plots of the posterior deviance values for each study arm in Figure 7 show that the probability of survival up to 4 years in Girard 2009 is not predicted well and this study is a possible outlier. Fitting a random effects model did not help in the prediction of data points in this study (Figure 8). The simpler fixed effect model is therefore preferred.



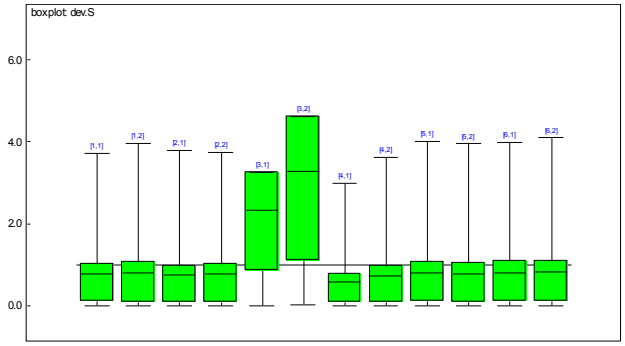


Figure 7: Posterior deviance values for each study arm for the area under the Kaplan Meier curves (left) and probability of survival (right) – fixed effect model.

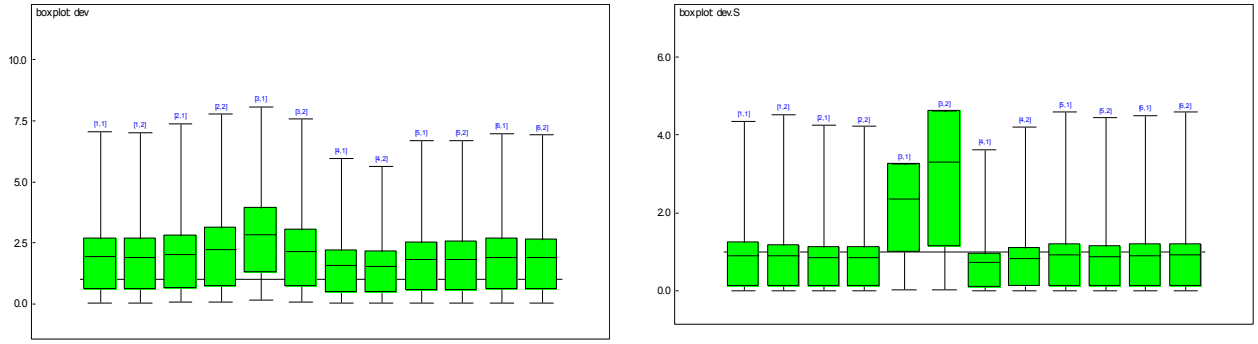


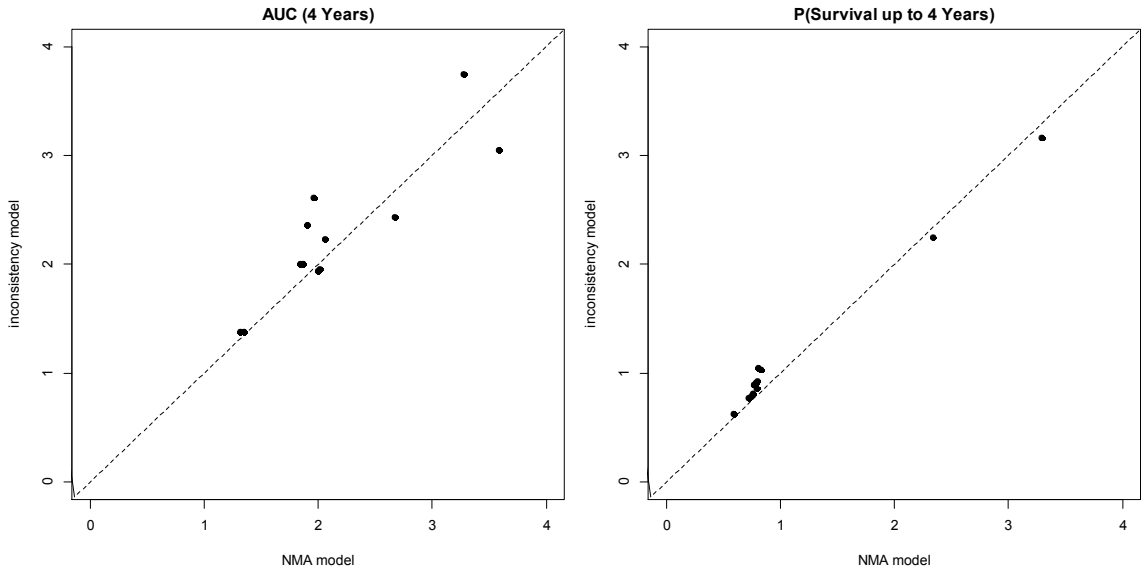
Figure 8: Posterior deviance values for each study arm for the area under the Kaplan Meier curves (left) and probability of survival (right) – random effects model.

No evidence of inconsistency was found through comparison of the consistency and inconsistency random effects models, as little difference was observed between the fit of the models (Table 16). The area below the line of equality in Figure 9 highlights where the inconsistency model better predicted data points, but any improvements were minimal.

Table 16: Model fit statistics for consistency and inconsistency fixed effect models based on 4-year follow-up data

Model		Posterior mean residual deviance	DIC
Fixed effect - consistency	P(Survival)	13.22	-27.429
	AUC	25.84	-20.356
Fixed effect - inconsistency	P(Survival)	14.07	-25.773
	AUC	27.07	-17.115

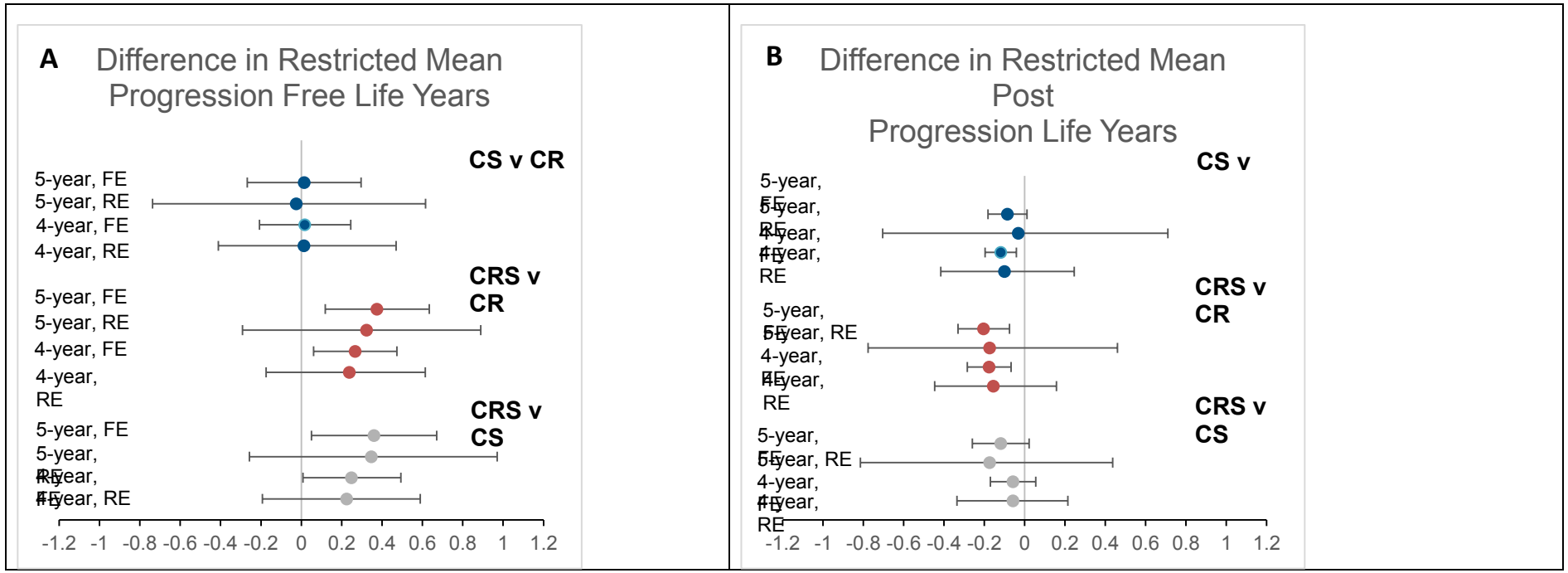
Total number of data points for P(survival) is 12 and for AUC is 24.



282

283 **Figure 9: Deviance contributions from the fixed effect consistency and inconsistency models for area under the Kaplan Meier curves**
284 **(left) and probability of survival (right).**

285 Treatment effects estimated by the fixed and random effects models based on the 4- and 5-year follow up data are presented in


A
Difference in Restricted Mean Progression Free Life Years
CS v CR
5-year, FE
5-year, RE
4-year, FE
4-year, RE
CRS v CR
5-year, FE
5-year, RE
4-year, FE
4-year, RE
CRS v CS
5-year, FE
5-year, RE
4-year, FE
4-year, RE
-1.2 -1 -0.8 -0.6 -0.4 -0.2 0 0.2 0.4 0.6 0.8 1 1.2
B
Difference in Restricted Mean Post Progression Life Years
CS v
CRS v CR
CRS v CS
-1.2 -1 -0.8 -0.6 -0.4 -0.2 0 0.2 0.4 0.6 0.8 1 1.2

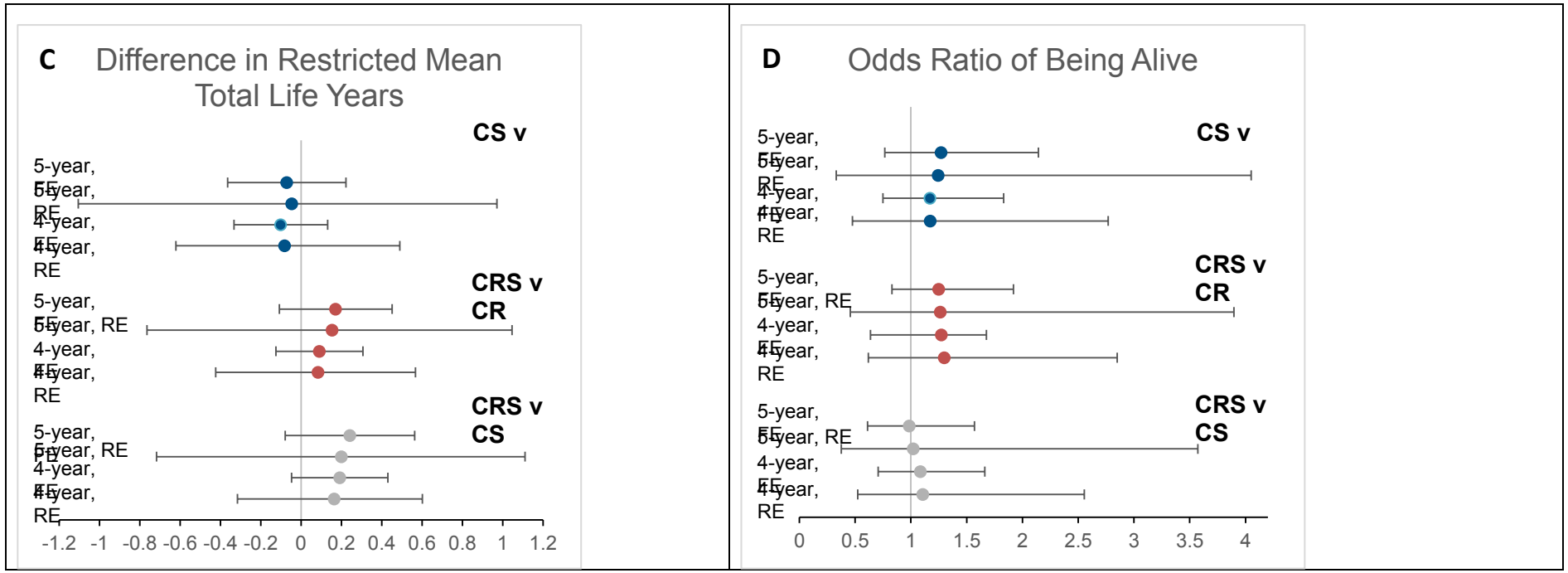
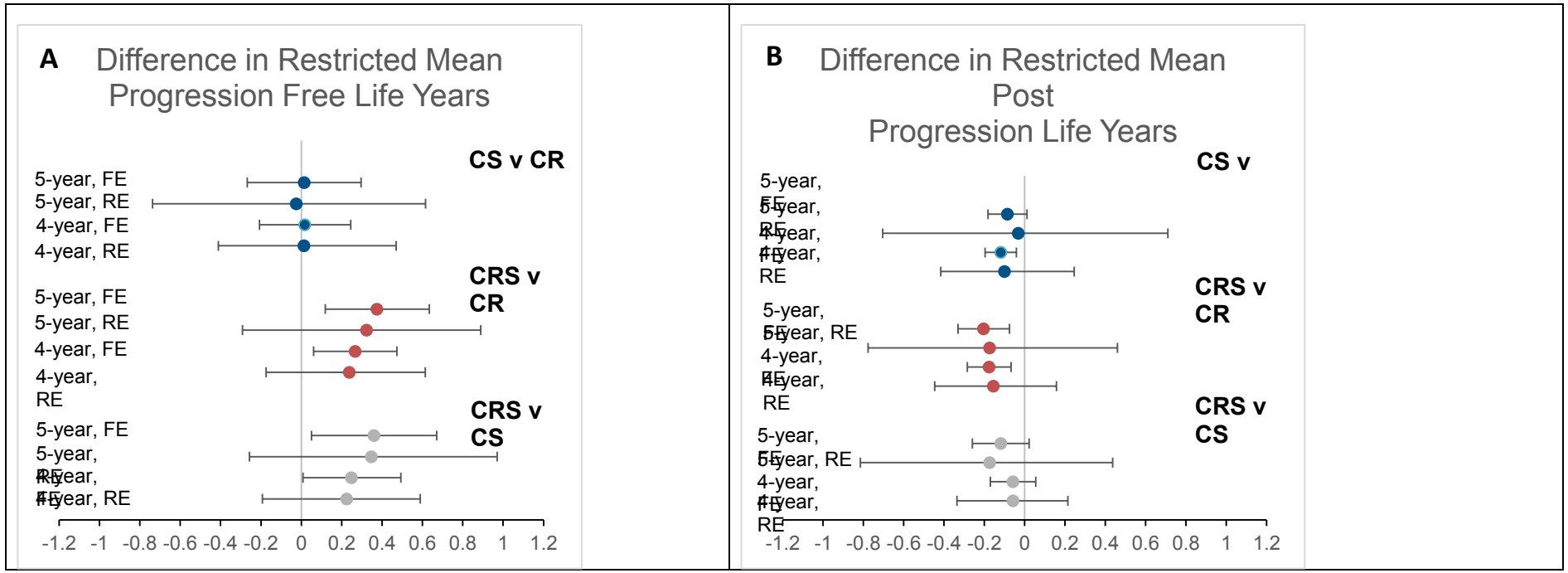


Figure 10. The point estimates of the treatment effects are similar, and the width of the credible intervals reflect that random effects models estimate the treatment effects with more uncertainty, and that there is additional data included in the 4-dataset compared with the 5-year dataset.

Noting that

1. the model fit assessment supports the use of the fixed effect model in both datasets,
2. the assumption that non-progressors by *T*-years do not progress (are "cured") is more reasonable at 5-years than at 4-years,
3. the 5-year dataset excludes the Girard (2009) study, which seems to be an outlier and is based on small numbers

supports the use of the fixed effect model based on the 5-year dataset for the base-case. Results from the random effects model based on the 5-year dataset are presented as a sensitivity analysis.


A
Difference in Restricted Mean Progression Free Life Years
CS v CR
5-year, FE
5-year, RE
4-year, FE
4-year, RE
CRS v CR
5-year, FE
5-year, RE
4-year, FE
4-year, RE
CRS v CS
5-year, FE
5-year, RE
4-year, FE
4-year, RE
-1.2 -1 -0.8 -0.6 -0.4 -0.2 0 0.2 0.4 0.6 0.8 1 1.2
B
Difference in Restricted Mean Post Progression Life Years
CS v
CRS v CR
CRS v CS
-1.2 -1 -0.8 -0.6 -0.4 -0.2 0 0.2 0.4 0.6 0.8 1 1.2

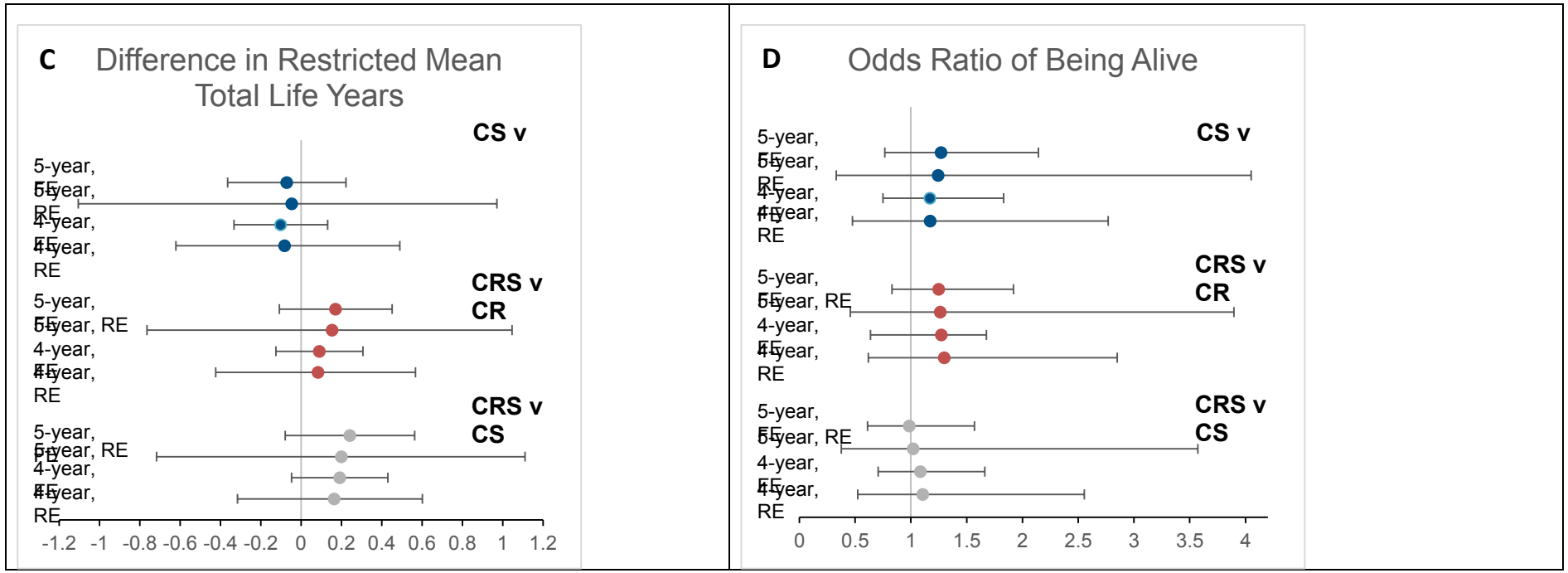


Figure 10: Forest plots of fixed and random effects estimates at 5- and 4-year follow up for (A) differences in restricted mean progression free life years at T-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at T-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at T-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at T-years follow-up relative to

chemoradiotherapy. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

Results: Inputs for Economic Model

Discounted Area Under the Kaplan Meier Curves and Probability of Survival

The fit of the NMA models based on the discounted AUC was also assessed and were in line with the results presented in Section 0 For both the 4-year and 5-year follow-up data, there were no meaningful differences between the fit of the fixed and random effects models (Table 17), and thus the fixed effect model was preferred.

Table 17: Model fit statistics based on 5-year follow-up data, discounted at 3.5% annual rate

Follow-Up Period	Model		Posterior Median Between-Study SD (95% CrI)	Posterior mean residual deviance	DIC
5 years [a]	Fixed effect[c]	P(Survival)	---	9.27	-24.85
		AUC		23.18	-14.69
	Random effects[d]	P(Survival)	0.33 (0.01, 2.34)	9.57	-22.94
		AUC	PFS: 0.17 (0.01, 1.25) PPS: 0.23 (0.03, 1.29)	18.86	-15.24
4 years [b]	Fixed effect[c]	P(Survival)	---	13.35	-27.18
		AUC		24.86	-23.87
	Random effects[e]	P(Survival)	0.22 (0.01, 1.56)	14.31	-25.08
		AUC	PFS: 0.11 (0.00, 0.68) PPS: 0.12 (0.01, 0.54)	23.34	-21.59

[a] Total posterior mean residual deviance compared to total number of data points for P(survival): 10 and AUC: 20
[b] Total posterior mean residual deviance compared to total number of data points for P(survival): 12 and AUC: 24
[c] Burn-in: 20,000 iterations, results based on: 40,000 samples, 2 chains
[d] Burn-in: 50,000 iterations, results based on: 100,000 samples, 2 chains
[e] Burn-in: 30,000 iterations, results based on: 60,000 samples, 2 chains

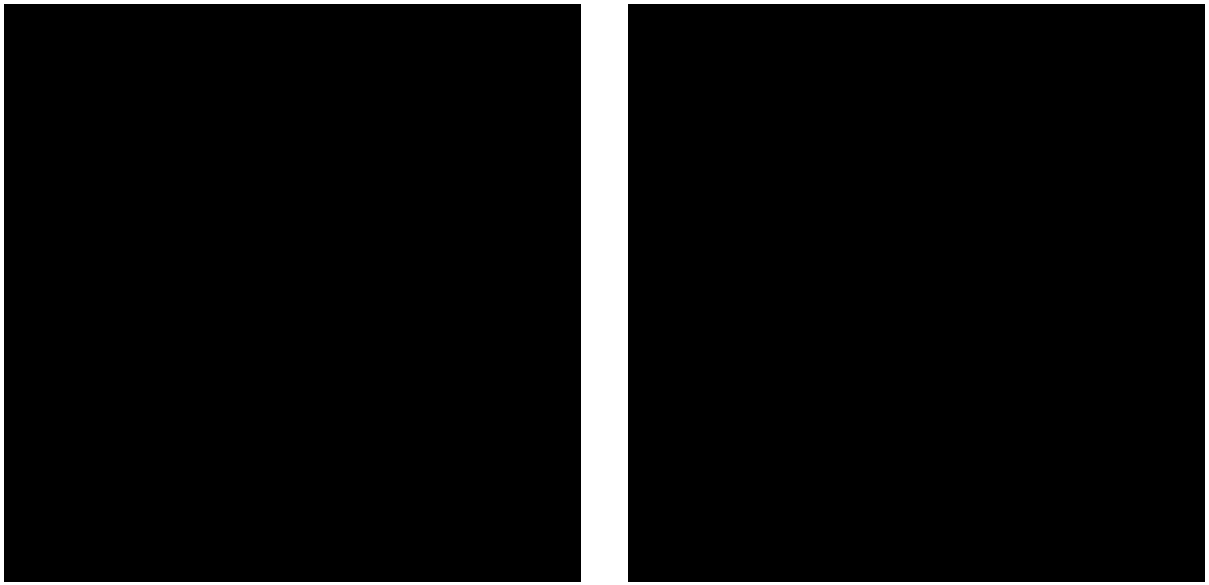
Similarly, the fit of the consistency and inconsistency models for both 4- and 5-year follow-up data were compared (Table 18). There is no evidence of inconsistency as no meaningful differences were found in the fit of the models for both datasets. The area below the line of equality in Figure 11 and Figure 12 highlights where the inconsistency model better predicted data points, but any improvements were minimal.

Table 18: Model fit statistics for consistency and inconsistency fixed effect models based on 4-year follow-up data, discounted at 3.5% annual rate

Follow-Up Period	Model[c]		Posterior mean residual deviance	DIC
5 years [a]	Fixed effect - consistency	P(Survival)	9.27	-24.85
		AUC	23.18	-14.69
	Fixed effect – inconsistency	P(Survival)	10.17	-22.87
		AUC	23.43	-12.42
4 years [b]	Fixed effect – consistency	P(Survival)	13.35	-27.18
		AUC	24.86	-23.87
	Random effects - inconsistency	P(Survival)	14.15	-25.62
		AUC	26.12	-20.59

[a] Total posterior mean residual deviance compared to total number of data points for P(survival): 10 and AUC: 20
[b] Total posterior mean residual deviance compared to total number of data points for P(survival): 12 and AUC: 24
[c] Burn-in: 20,000 iterations, results based on: 40,000 samples, 2 chains

322 **Figure 11: Deviance contributions from the fixed effect consistency and inconsistency models for area under the Kaplan Meier curves**
323 **discounted at 3.5% annual rate (left) and probability of survival (right).**



324 **Figure 12: Deviance contributions from the fixed effect consistency and inconsistency models for area under the Kaplan Meier curves**
325 **discounted at 3.5% annual rate (left) and probability of survival (right).**

326 ***Proportion of Events Occurring each Year***

327 The proportion of events occurring each year pooled across studies is given in Table 19. The estimated proportions are similar across the 5-year
328 and 4-year follow-up datasets.

329

330

Table 19: Pooled proportion of events occurring each year

Follow-Up Period	Event Type	Year	Median Proportion of Events (95% CrI)
5-year	PFS[a]	1	0.63 (0.59, 0.67)
		2	0.23 (0.19, 0.28)
		3	0.08 (0.03, 0.13)
		4	0.04 (0.00, 0.09)
		5	0.01 (0.00, 0.07)
	OS[b]	1	0.38 (0.34, 0.42)
		2	0.32 (0.27, 0.38)
		3	0.16 (0.10, 0.22)
		4	0.11 (0.04, 0.17)
		5	0.03 (0.00, 0.10)
4-year	PFS[c]	1	0.65 (0.61, 0.69)
		2	0.24 (0.19, 0.30)
		3	0.09 (0.00, 0.14)
		4	0.01 (0.00, 0.08)
	OS[c]	1	0.39 (0.35, 0.43)
		2	0.35 (0.29, 0.41)
		3	0.17 (0.11, 0.23)
		4	0.10 (0.00, 0.15)

[a] Burn-in: 500,000 iterations, results based on: 1,000,000 samples, 2 chains
[b] Burn-in: 2,000,000 iterations, results based on: 4,000,000 samples, 2 chains
[c] Burn-in: 100,000 iterations, results based on: 100,000 samples, 2 chains

NMA for Adverse Events

The base case approach used in the economic model for adverse events used pairwise meta-analyses but data then became available that allowed us to fit an NMA for use in sensitivity analyses.

The studies had reported adverse events heterogeneously; in some studies the reporting was comprehensive and in others scant or no details were available. Furthermore, events were classified heterogeneously across studies, being grouped under narrow or broad classes that made event-specific pooling difficult. The committee decided that adverse events should be included in the economic model if possible and we agreed an aggregate approach with them. This involved grouping all adverse events of grade 3+ as homogenously requiring one hospital admission, but having no long term clinical effects or detriment to quality of life. The committee thought it possible that grade 4 adverse events would affect quality of life but these occurred to sparsely to be meaningfully included in the model. Because of the wide disparity between the frequency of adverse events reported among the studies, we selected Pless 2015, Eberhardt 2015, Albain 2009 and van Meerbeeck 2007 for the analysis. These studies were the largest and best conducted studies in the network and had reported event rates that the committee found credible. The data from van Meerbeeck was not reported in the published paper but provided to us upon request by the EORTC, who hold the trial data. We obtained the person years at risk by multiplying the total number of patients in each arm by the mean AUC for total life years at 5 years. The data are in Table 20.

Table 20: Adverse Event NMA Input Data

Treatment Arm 1	Events Arm 1	TatRisk Arm 1	Treatment Arm 2	Events Arm 2	TatRisk Arm 2	Study		Treatments
2	182	285.2	3	141	299.52	Pless 2015		1=CR
3	482	434.3	1	608	409.34	Albain 2009		2=CS
1	137	214.4	3	150	230.04	Eberhardt 2015		3=CRS
1	98	321.75	2	108	298.93	van Meerbeerck 2007		

We assumed that adverse events were treatment related and therefore that it was appropriate to assume a homogenous follow-up time. Since this meant that we did not have to account for variable study endpoints in our pooling of the data, we selected a poisson likelihood, log link NMA model and copied the code directly from NICE TSD2 (citation). The results of the fixed and random effects models are in Table 21. Models were run using 50,000 burn-in iterations and 50,000 iterations to generate the posterior distributions.

Table 21: Adverse Event NMA Results

All Adverse Events	estimate	LCL	UCL	DIC
Fixed effects				74.44
HR of CS vs CR	1.132	0.9382	1.354	
HR of CR vs CRS	1.2425447	1.125112511	1.377221	
HR of CS vs CRS	1.3970383	1.174950065	1.67336	

Random effects				72.627
HR of CR vs CS	1.166	0.3146	4.654	
HR of CR vs CRS	1.176886	0.374531835	3.354579	
HR of CS vs CRS	1.3696754	0.361663653	5.186722	

The DIC for the random effects model was not more than 3-5 points lower than the fixed effects model so we preferred it in the base case. The results show that both CR and CS are associated with more adverse events than CRS.

As discussed in the economic modelling report (Appendix J), the NMA data agreed well with the pairwise estimates of adverse events.

References and Code

References

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7. National Institute for Health and Clinical Excellence, *The guidelines Manual (November 2012). Available from http://publications.nice.org.uk/the-guidelines-manual-pmg6*. 2012, National Institute of Health and Clinical Excellence: London.
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Code

SEER dataset

Selection criteria:

{Age at Diagnosis.Age recode with <1 year olds} = '30-34 years','35-39 years','40-44 years','45-49 years','50-54 years','55-59 years','60-64 years','65-69 years','70-74 years','75-79 years'

AND ({Site and Morphology.CS Schema v0204+} = 'Lung'

OR {Site and Morphology.CS Schema - AJCC 6th Edition} = 'Lung')

AND ({Stage - AJCC.Derived AJCC Stage Group, 7th ed (2010+)} = 'IIIA'

OR {Stage - AJCC.Derived AJCC Stage Group, 6th ed (2004+)} = 'IIIA'

OR {Stage - AJCC.AJCC stage 3rd edition (1988-2003)} = ' 31'

OR {Stage - AJCC.SEER modified AJCC stage 3rd (1988-2003)} = ' 31')

AND ({Stage - TNM.Derived AJCC N, 7th ed (2010+)} = 'N2','N2a','N2b','N2c'

OR {Stage - TNM.Derived AJCC N, 6th ed (2004+)} = 'N2','N2a','N2b','N2c'

OR {Stage - TNM.N value - based on AJCC 3rd (1988-2003)} = 'N2')

NMA Model for Adverse Events – Fixed Effects

```
# Poisson likelihood, log link
# Fixed effects model for multi-arm trials
model{ # *** PROGRAM STARTS
```

```
401  for(i in 1:ns){ # LOOP THROUGH STUDIES
402   mu[i] ~ dnorm(0,.0001) # vague priors for all trial baselines
403   for (k in 1:na[i]) { # LOOP THROUGH ARMS
404   r[i,k] ~ dpois(theta[i,k]) # Poisson likelihood
405   theta[i,k] <- lambda[i,k]*E[i,k] # event rate * exposure
406   log(lambda[i,k]) <- mu[i] + d[t[i,k]] - d[t[i,1]] # model for linear predictor
407   dev[i,k] <- 2*((theta[i,k]-r[i,k]) + r[i,k]*log(r[i,k]/theta[i,k])) #Deviance contribution
408   }
409   resdev[i] <- sum(dev[i,1:na[i]]) # summed residual deviance contribution for this trial
410   }
411  totresdev <- sum(resdev[]) #Total Residual Deviance
412  d[1]<-0 # treatment effect is zero for reference treatment
413  for (k in 2:nt){ d[k] ~ dnorm(0,.0001) } # vague priors for treatment effects
414
415
416
417
418  sd ~ dunif(0,5) # vague prior for between-trial SD
419  tau <- pow(sd,-2) # between-trial precision = (1/between-trial variance)
420
```

```
421  # pairwise HRs and LHRs for all possible pair-wise comparisons, if nt>2
422  for (c in 1:(nt-1)) {
423   for (k in (c+1):nt) {
424   lhr[c,k] <- (d[k]-d[c])
425   log(hr[c,k]) <- lhr[c,k]
426   }
427   }
428
429  } # *** PROGRAM ENDS
430
431  list(ns=4, nt=3)
432
433  t[,1]   r[,1]   E[,1]   t[,2]   r[,2]   E[,2]   na[]
434  2       182     285.2   3       141     299.52  2
435  3       482     434.3   1       608     409.34  2
436  1       137     214.4   3       150     230.04  2
437  1       98      321.75  2       108     298.93  2
438
439  END
440
```

```
#chain 1
list(d=c( NA, 0, 0), mu=c(0, 0, 0, 0))
#chain 2
list(d=c( NA, -1, 1), mu=c(-3, -3, -3, -3))
#chain 3
list(d=c( NA, 2, 2),  mu=c(-3, 5, -1, -3))
```

NMA Model for Adverse Events - Random Effects

```
# Poisson likelihood, log link
# Random effects model for multi-arm trials
model{ # *** PROGRAM STARTS
for(i in 1:ns){ # LOOP THROUGH STUDIES
 w[i,1] <- 0 # adjustment for multi-arm trials is zero for control arm
 delta[i,1] <- 0 # treatment effect is zero for control arm
 mu[i] ~ dnorm(0,.0001) # vague priors for all trial baselines
 for (k in 1:na[i]) { # LOOP THROUGH ARMS
 r[i,k] ~ dpois(theta[i,k]) # Poisson likelihood
 theta[i,k] <- lambda[i,k]*E[i,k] # failure rate * exposure
 log(lambda[i,k]) <- mu[i] + delta[i,k] # model for linear predictor
```

```
461     dev[i,k] <- 2*((theta[i,k]-r[i,k]) + r[i,k]*log(r[i,k]/theta[i,k])) #Deviance contribution
462     }
463     resdev[i] <- sum(dev[i,1:na[i]]) # summed residual deviance contribution for this trial
464     for (k in 2:na[i]) { # LOOP THROUGH ARMS
465     delta[i,k] ~ dnorm(md[i,k],taud[i,k]) # trial-specific LOR distributions
466     md[i,k] <- d[t[i,k]] - d[t[i,1]] + sw[i,k] # mean of LOR distributions (with multi-arm trial correction)
467     taud[i,k] <- tau *2*(k-1)/k # precision of LOR distributions (with multi-arm trial correction)
468     w[i,k] <- (delta[i,k] - d[t[i,k]] + d[t[i,1]]) # adjustment for multi-arm RCTs
469     sw[i,k] <- sum(w[i,1:k-1])/(k-1) # cumulative adjustment for multi-arm trials
470     }
471     }
472
473
474   totresdev <- sum(resdev[]) #Total Residual Deviance
475   d[1]<-0 # treatment effect is zero for reference treatment
476   for (k in 2:nt){ d[k] ~ dnorm(0,.0001) } # vague priors for treatment effects
477   sd ~ dunif(0,5) # vague prior for between-trial SD
478   tau <- pow(sd,-2) # between-trial precision = (1/between-trial variance)
479   # pairwise HRs and LHRs for all possible pair-wise comparisons, if nt>2
480   for (c in 1:(nt-1)) {
```

```
481     for (k in (c+1):nt) {
482     lhr[c,k] <- (d[k]-d[c])
483     log(hr[c,k]) <- lhr[c,k]
484     }
485     }
486
487   } # *** PROGRAM ENDS
488
489   list(ns=4, nt=3)
490
491   t[,1]   r[,1]   E[,1]   t[,2]   r[,2]   E[,2]   na[]
492   2       182     285.2   3       141     299.52 2
493   3       482     434.3   1       608     409.34 2
494   1       137     214.4   3       150     230.04 2
495   1       98      321.75 2        108     298.93 2
496
497   END
498
499   #chain 1
500   list(d=c( NA, 0, 0), sd=1, mu=c(0, 0, 0, 0))
```

501 #chain 2

502 list(d=c(NA, -1, 1), sd=4, mu=c(-3, -3, -3, -3))

503 #chain 3

504 list(d=c(NA, 2, 2), sd=2, mu=c(-3, 5, -1, -3))

505

506 .

507 **Table 9: Trial data for evidence synthesis (Treatment 1=CR, 2=CS and 3=CRS)**

			PFS		OS			Survival	
	Study	Treatment	AUC	SE	AUC	SE	AUC Correlation	Probability[a]	SE
4-year data	Albain	1	1.42	0.09	2.11	0.12	0.82	0.25	0.04
		3	1.72	0.11	2.15	0.12	0.87	0.28	0.04
	Eberhardt	1	2.05	0.18	2.68	0.16	0.92	0.41	0.06
		3	2.16	0.17	2.84	0.17	0.85	0.50	0.06
	Girard	2	2.21	0.42	2.47	0.32	0.95	0.26	0.15
		3	1.65	0.34	2.14	0.32	0.97	0.24	0.11
	Katakami	2	1.47	0.24	2.60	0.23	0.79	0.31	0.09
		3	1.89	0.28	2.82	0.23	0.82	0.38	0.09
	Pless	2	1.63	0.14	2.48	0.14	0.87	0.43	0.05
		3	1.89	0.15	2.56	0.14	0.85	0.43	0.05
	van Meerbeeck	1	1.39	0.09	1.95	0.10	0.94	0.18	0.03
		2	1.36	0.10	1.79	0.11	0.97	0.20	0.03
5-year data	Albain	1	1.55	0.11	2.33	0.15	0.87	0.19	0.04
		3	1.95	0.13	2.42	0.15	0.91	0.26	0.04
	Eberhardt	1	2.41	0.23	3.09	0.21	0.95	0.41	0.06

		3	2.49	0.22	3.30	0.21	0.88	0.44	0.06
	Katakami	2	1.60	0.28	2.88	0.30	0.85	0.26	0.09
		3	2.15	0.35	3.19	0.30	0.88	0.38	0.09
	Pless	2	1.86	0.18	2.90	0.19	0.87	0.41	0.05
		3	2.13	0.19	2.94	0.18	0.87	0.35	0.05
	van Meerbeeck	1	1.52	0.12	2.11	0.12	0.95	0.14	0.03
		2	1.48	0.12	1.96	0.13	0.96	0.16	0.03

Abbreviations: AUC – area under the curve, OS – overall survival, PFS – progression free survival, SE – standard error.

[a] Probability of surviving up to 4- or 5-years.

Correlation between AUCs for PFS and OS

The AUCs for progression free and overall survival are correlated because the AUC for OS must be greater than for PFS. We estimated this correlation using non-parametric bootstrapping, constrained to samples where the AUC for OS was greater than that for PFS [6]. These correlations are provided in Model Critique

Assessing model fit

The posterior mean of the residual deviance, which measures the magnitude of the differences between the observed data and the model predictions of the data, was used to assess the goodness of fit of each model [12]. Smaller values are preferred, and in a well-fitting model the posterior mean residual deviance should be close to the number of data points in the network (each study arm contributes 1 data point) [12].

In addition to comparing how well the models fit the data using the posterior mean of the residual deviance, models were compared using the deviance information criterion (DIC). This is equal to the sum of the posterior mean deviance and the effective number of parameters, and thus penalizes model fit with model complexity [12]. Lower values are preferred and differences of at least 5 points were considered meaningful [12].

Assessing heterogeneity and inconsistency

Heterogeneity concerns the differences in treatment effects between trials within each treatment contrast, while consistency concerns the differences between the direct and indirect evidence informing the treatment contrasts [9, 10].

Heterogeneity is assessed by comparing the fit of fixed and random effects NMA models. The fixed effect model assumes that all trials are estimating the same treatment effect, regardless of any differences in the conduct of the trials, populations, or treatments. The random effects

NMA model on the other hand accounts for any differences in treatment effects between trials, that are beyond sampling error, by assuming a distribution of study-specific treatment effects with a pooled mean and between-study standard deviation. The estimated between study standard deviation in treatment effects is also inspected to assess heterogeneity.

Inconsistency was assessed by comparing the fit of the chosen consistency model (fixed or random effects) to an “inconsistency”, or unrelated mean effects, model [9, 10]. The latter is equivalent to having separate, unrelated, meta-analyses for every pairwise contrast, with a common variance parameter assumed in the case of random effects models. Note that inconsistency can only be assessed when there are closed loops of direct evidence on 3 treatments that are informed by at least 3 distinct trials [11].

Network meta-analysis: Results of Clinical Evidence Synthesis

5-year Follow-up

Five studies presented survival data up to 5-years, and a network diagram summarizing the evidence is given in Figure 2

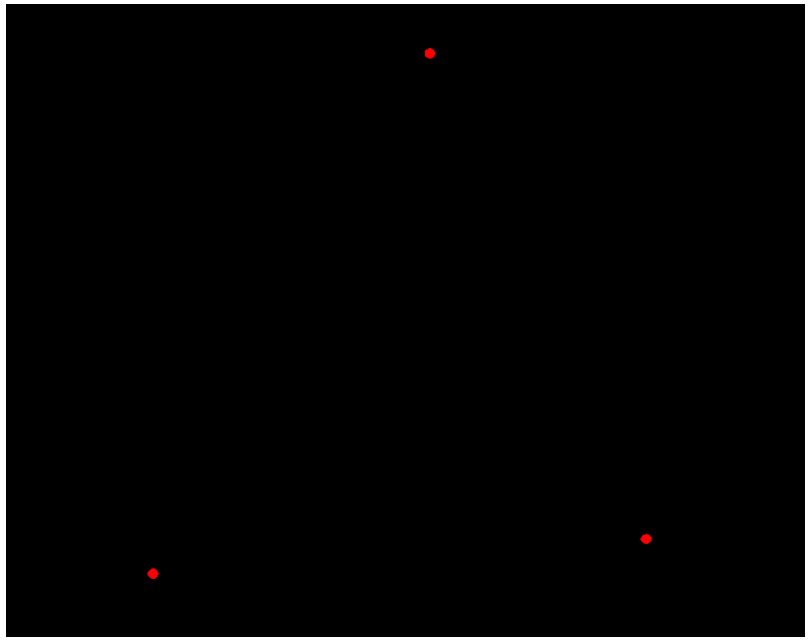


Figure 2: Network diagram of comparisons for which direct evidence on differences in restricted mean survival time up to 5-years is available. Lines are proportional to the number of studies that compare the two connected treatments.

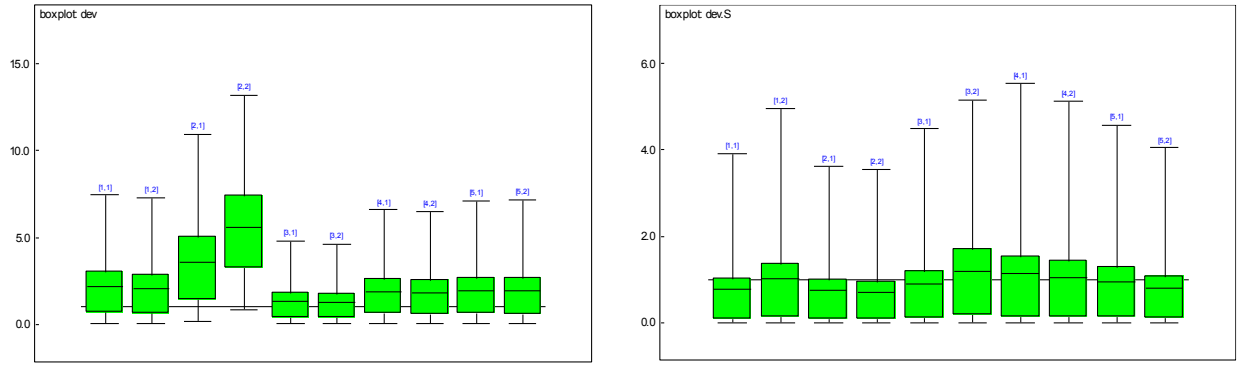
Model fit statistics for the area under the Kaplan Meier curves up to 5-years, as well as the probability of survival are given in Table 12. Convergence was satisfactory for the fixed effect model after a burn-in of 20,000 iterations and results are based on a further 40,000 samples on two chains. For the random effects model, convergence was satisfactory after a burn-in of 30,000 iterations and results are based on a further 60,000 samples on two chains.

Table 12: Model fit statistics based on 5-year follow-up data

Model		**Median Between-Study SD (95% CrI)**	**Posterior mean residual deviance**	**DIC**
Fixed effect	P(Survival)	---	9.267	-24.852
	AUC		23.47	-11.075
Random effects	P(Survival)	0.35 (0.02, 2.41)	9.618	-22.809
	AUC	PFS: 0.18 (0.01, 1.32) PPS: 0.25 (0.03, 1.46)	18.95	-11.781

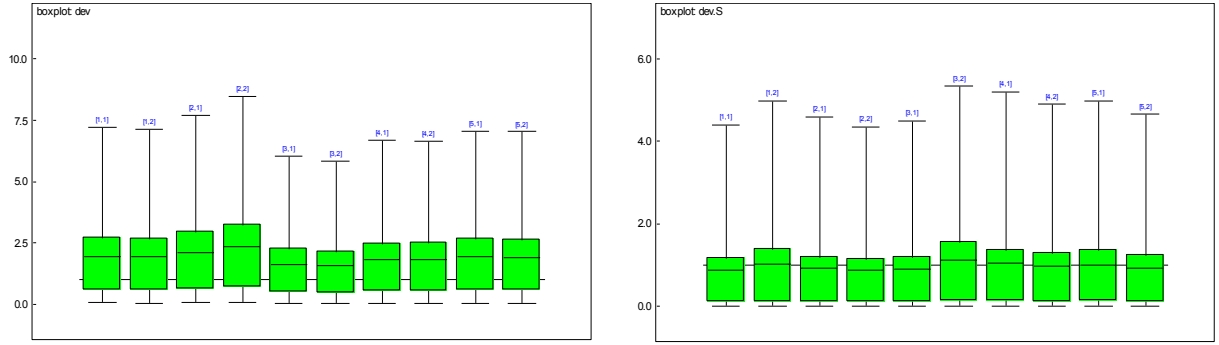
Total number of data points for P(survival) is 10 and for AUC is 20.

There were no meaningful differences between the fixed and random effects models in terms of the posterior mean residual deviance and DIC for both NMAs (Table 12). The box plots of the posterior deviance values for each study arm in Figure 3 show that the area under the Kaplan Meier curves up to 5 years in Eberhardt 2015 is not predicted well and this study is a possible outlier. Although the prediction of this study improves in the random effects model (Figure 4), this comes at a cost of slight overfit of the model (posterior mean residual deviance = 18.95, compared to 20 datapoints) and additional parameters in the model. In addition, progression events and deaths were rare in the chemoradiotherapy group of this study after 3-years and 4-years, respectively. The simpler fixed effect model was therefore selected in the base-case.



551 **Figure 3: Posterior deviance values for each study arm for the area under the Kaplan Meier curves (left) and probability of survival**
552 **(right) – fixed effect model.**

553



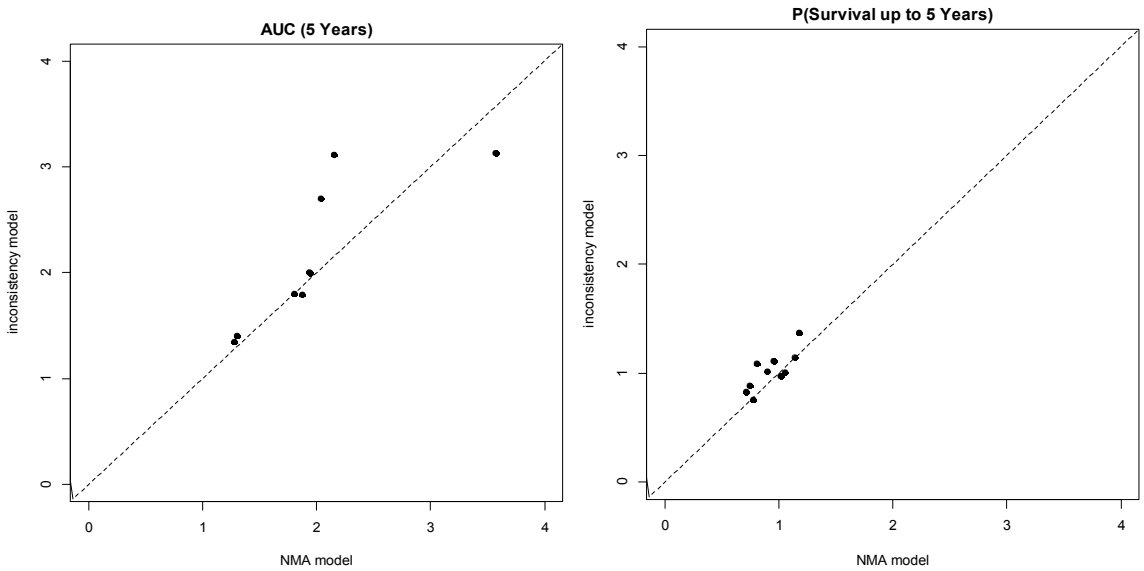
554 **Figure 4: Posterior deviance values for each study arm for the area under the Kaplan Meier curves (left) and probability of survival**
555 **(right) – random effects model.**

556 No evidence of inconsistency was found, with model fit (posterior mean residual deviance) similar for the consistency and inconsistency (unrelated
557 means) fixed effect models, and a lower DIC for the consistency model (Table 13). The area below the line of equality in Figure 5 highlights where
558 the inconsistency model better predicted data points, and any improvement is minimal.

559 **Table 13: Model fit statistics for consistency and inconsistency fixed effect models based on 5-year follow-up data**

Model		Posterior mean residual deviance	DIC
Fixed effect - consistency	P(Survival)	9.267	-24.852
	AUC	23.47	-11.075
Fixed effect - inconsistency	P(Survival)	10.17	-22.867
	AUC	23.65	-8.882

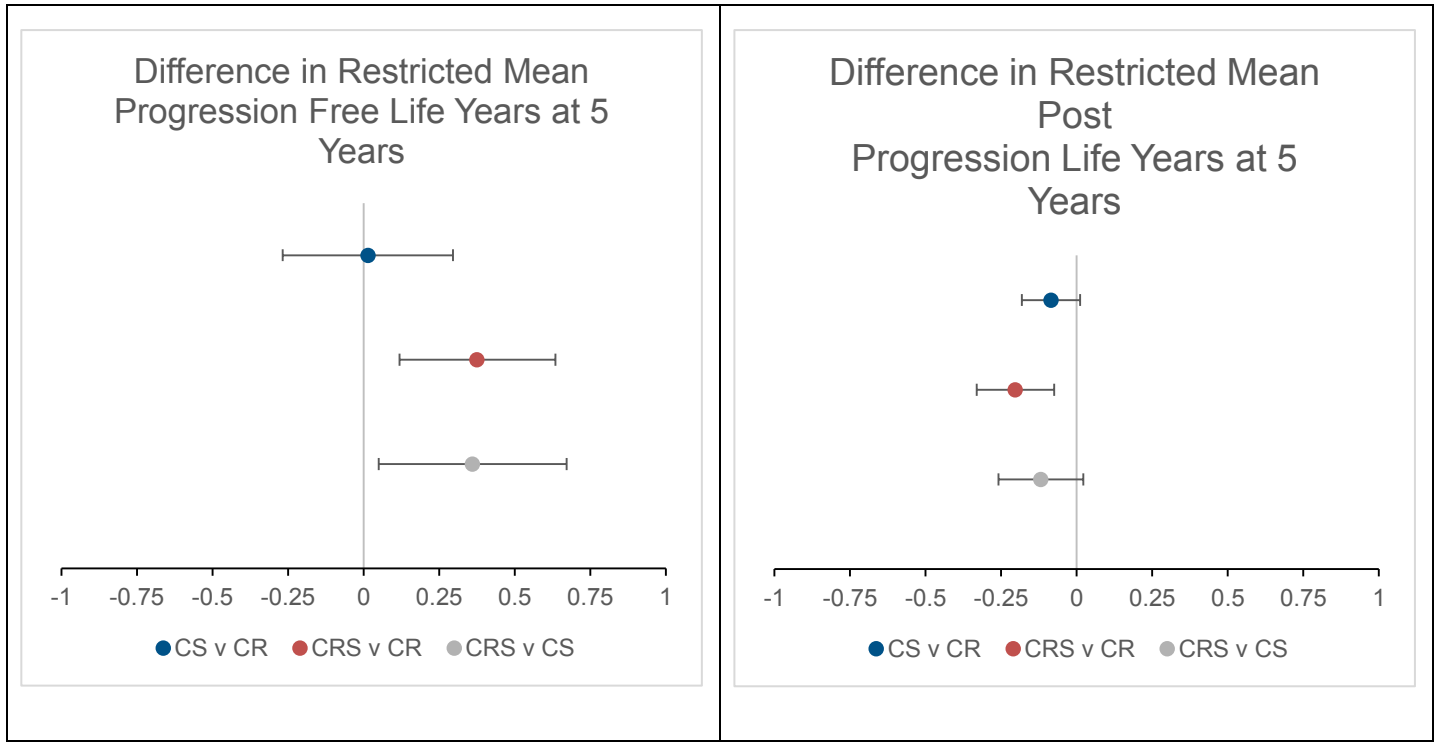
560 Total number of data points for P(survival) is 10 and for AUC is 20.



561

562 **Figure 5: Deviance contributions from the fixed effect consistency and inconsistency models for area under the Kaplan Meier curves**
563 **(left) and probability of survival (right).**

564 There is evidence to suggest that chemoradiotherapy + surgery is more effective in increasing progression free life years at 5-year follow-up
565 compared to chemoradiotherapy alone, while there is no evidence to suggest the effect of chemotherapy + surgery is any different from
566 chemoradiotherapy (


Difference in Restricted Mean Progression Free Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS

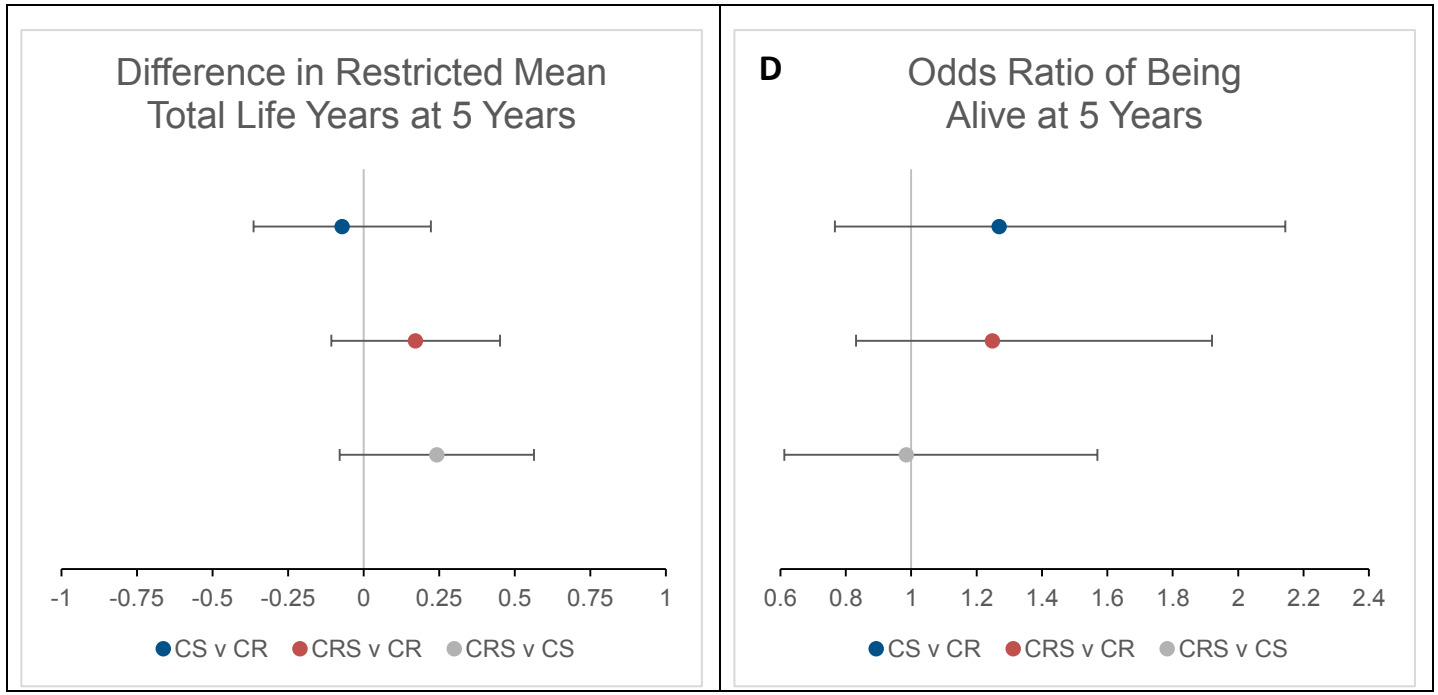
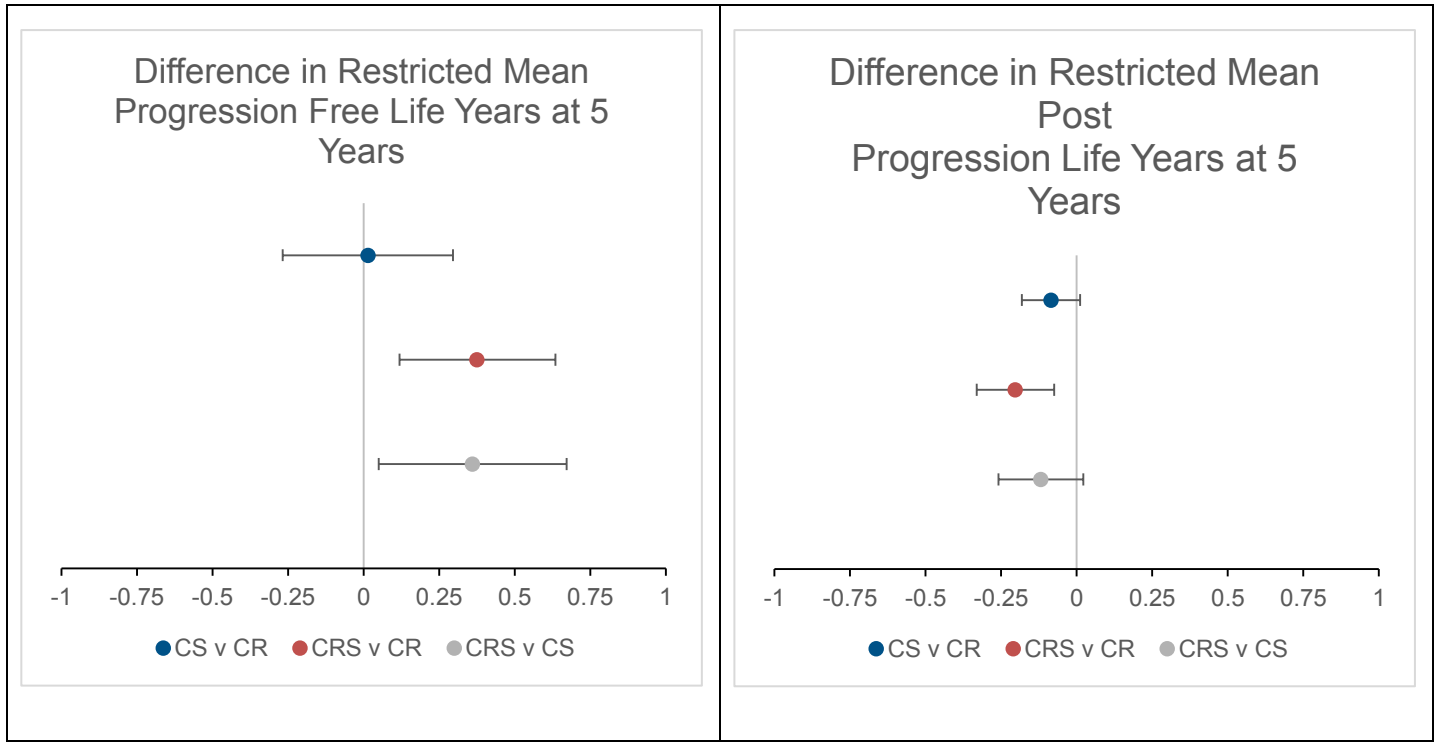


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

A, Table 14). There is also evidence to suggest that chemoradiotherapy + surgery improves progression free life years compared to chemotherapy + surgery (posterior median difference in RMST: 0.36 (95% CrI: 0.05, 0.67)) and it ranked the most effective intervention in increasing progression free life years (Table 14).

In terms of post progression life years at 5-year follow-up, there is evidence suggesting that chemoradiotherapy is more effective than chemoradiotherapy + surgery (


Difference in Restricted Mean Progression Free Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS

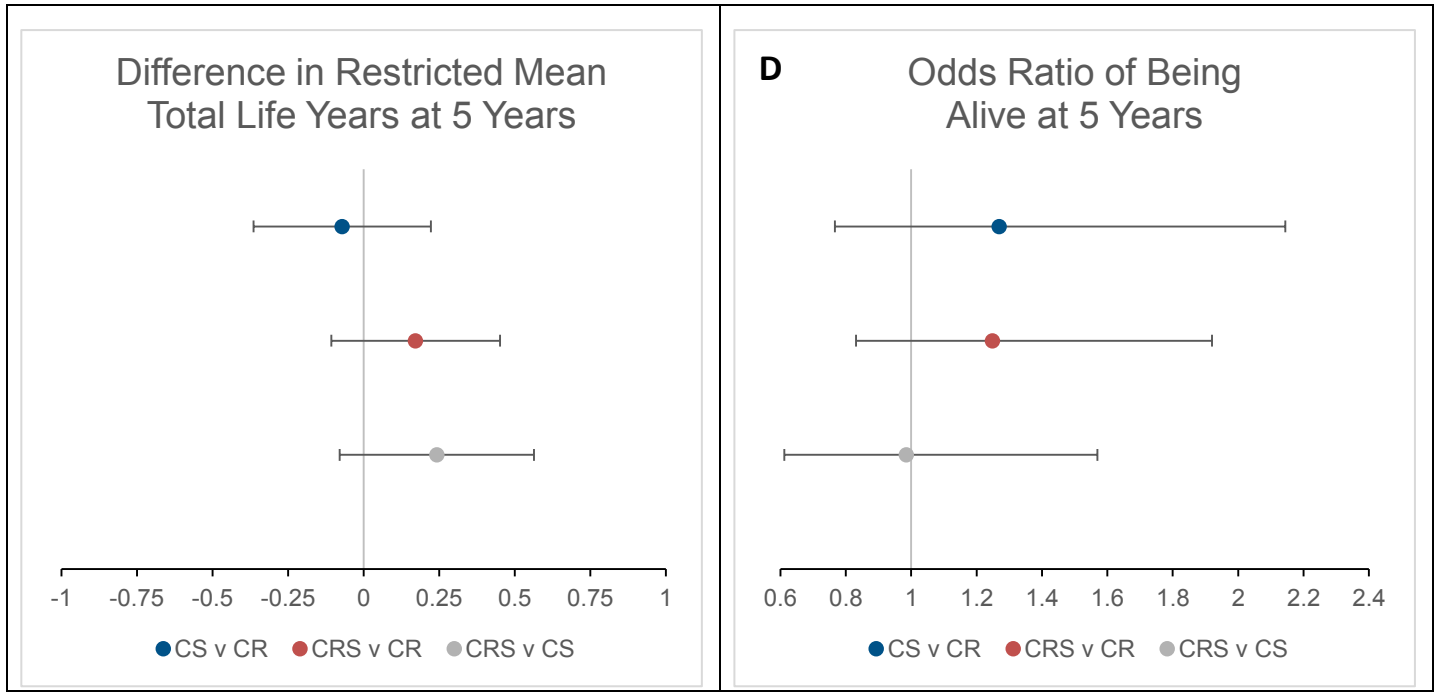
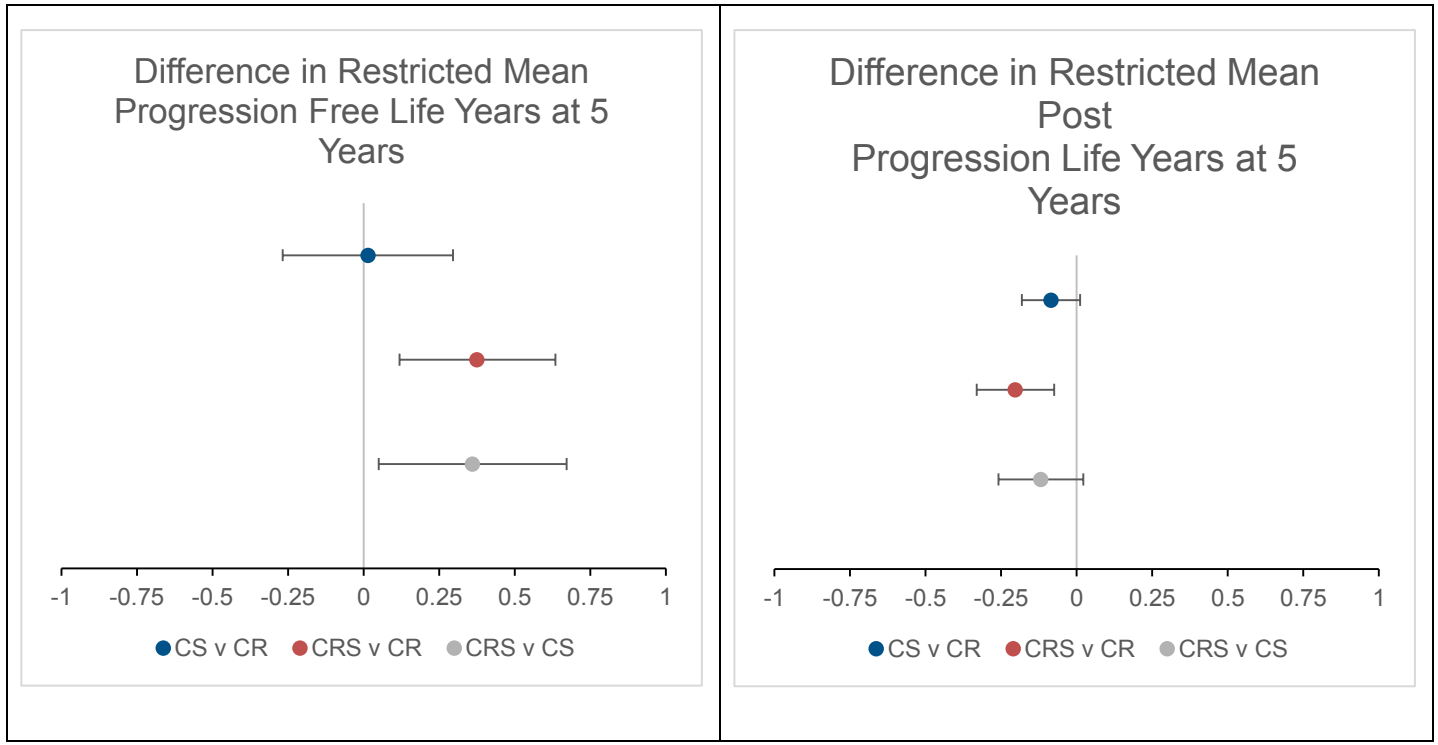


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

B, Table 14). Chemoradiotherapy appears to be more effective than chemotherapy + surgery as well, but this cannot be concluded with high certainty (


Difference in Restricted Mean Progression Free Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS

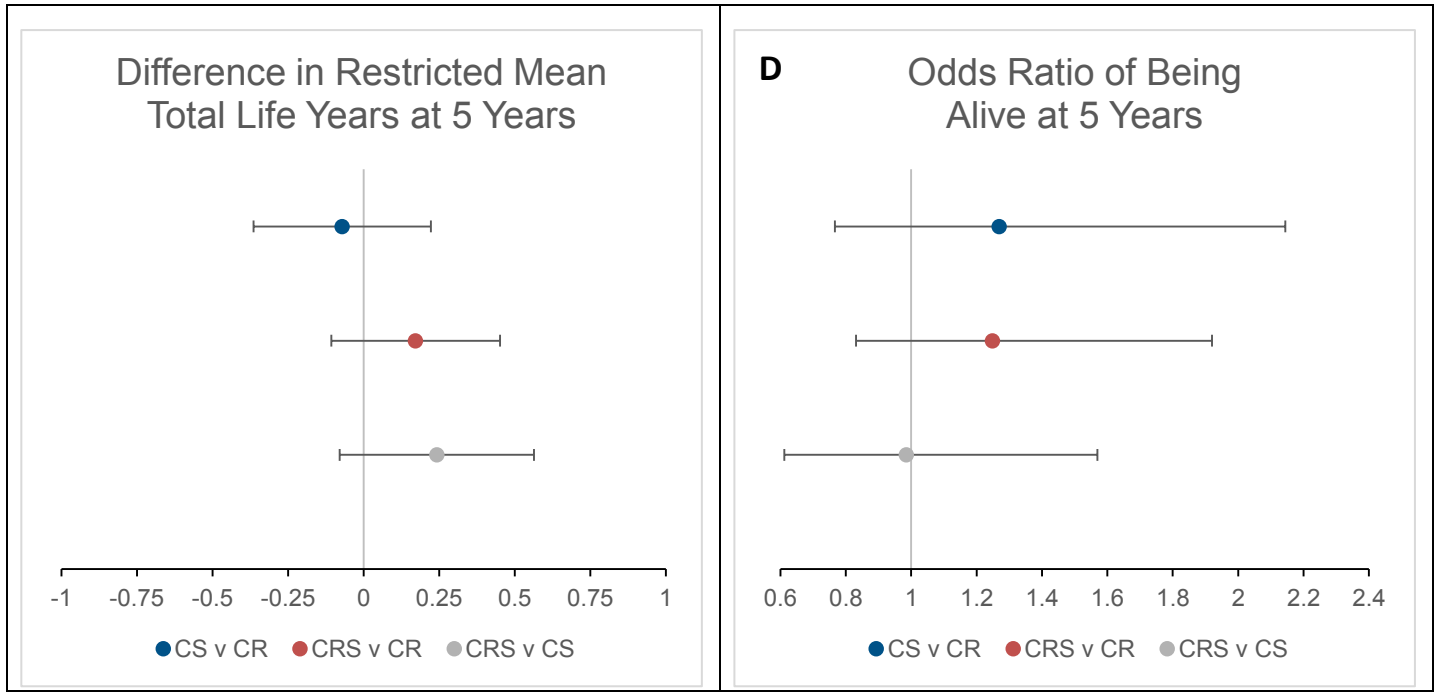
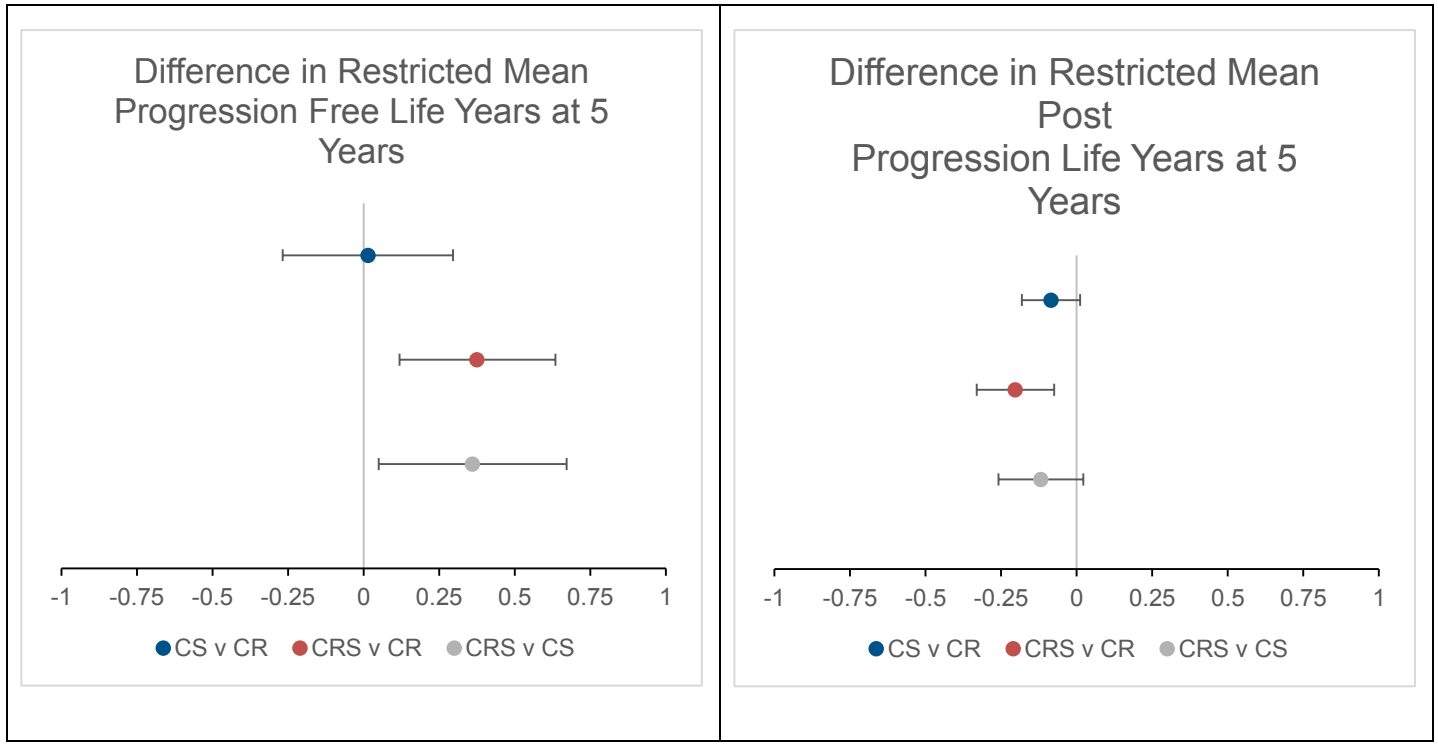


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

B, Table 14). There was not enough evidence to suggest any of the three treatments were different from each other in terms of improving total life years at 5-year follow- up, which is the sum of the progression free and post progression life years (


Difference in Restricted Mean Progression Free Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS

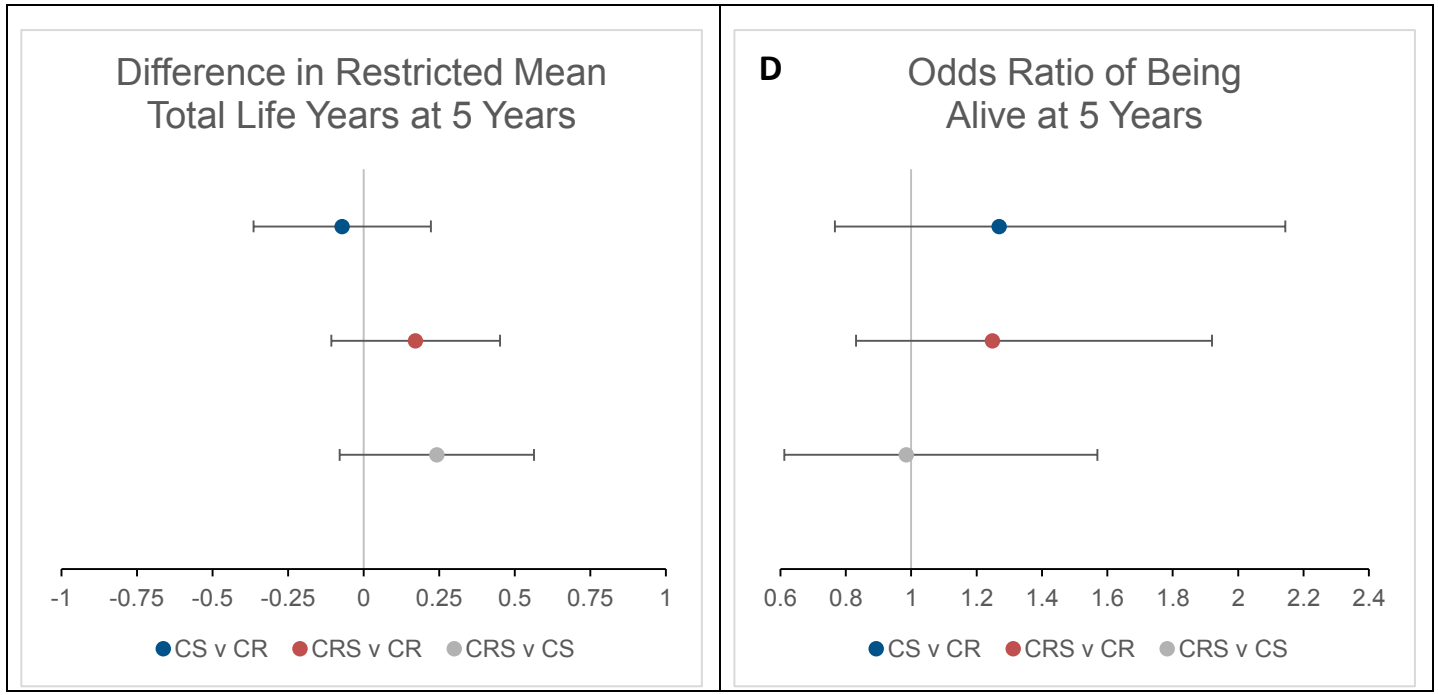
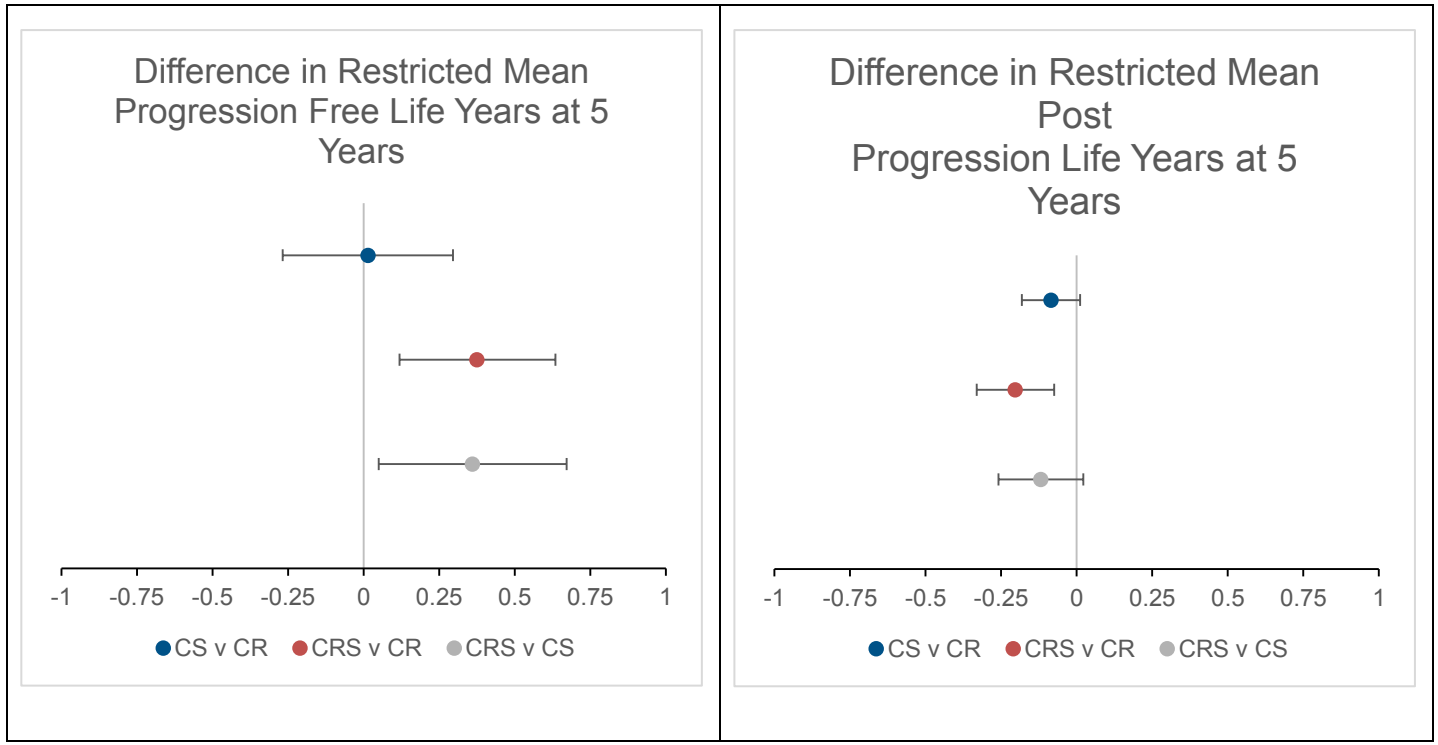


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

C, Table 14).

Chemotherapy + surgery and chemoradiotherapy + surgery appear to be more likely to improve the odds of being alive at 5-years compared to chemoradiotherapy alone, but there is not enough evidence to infer the direction of effects with certainty (


Difference in Restricted Mean Progression Free Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS

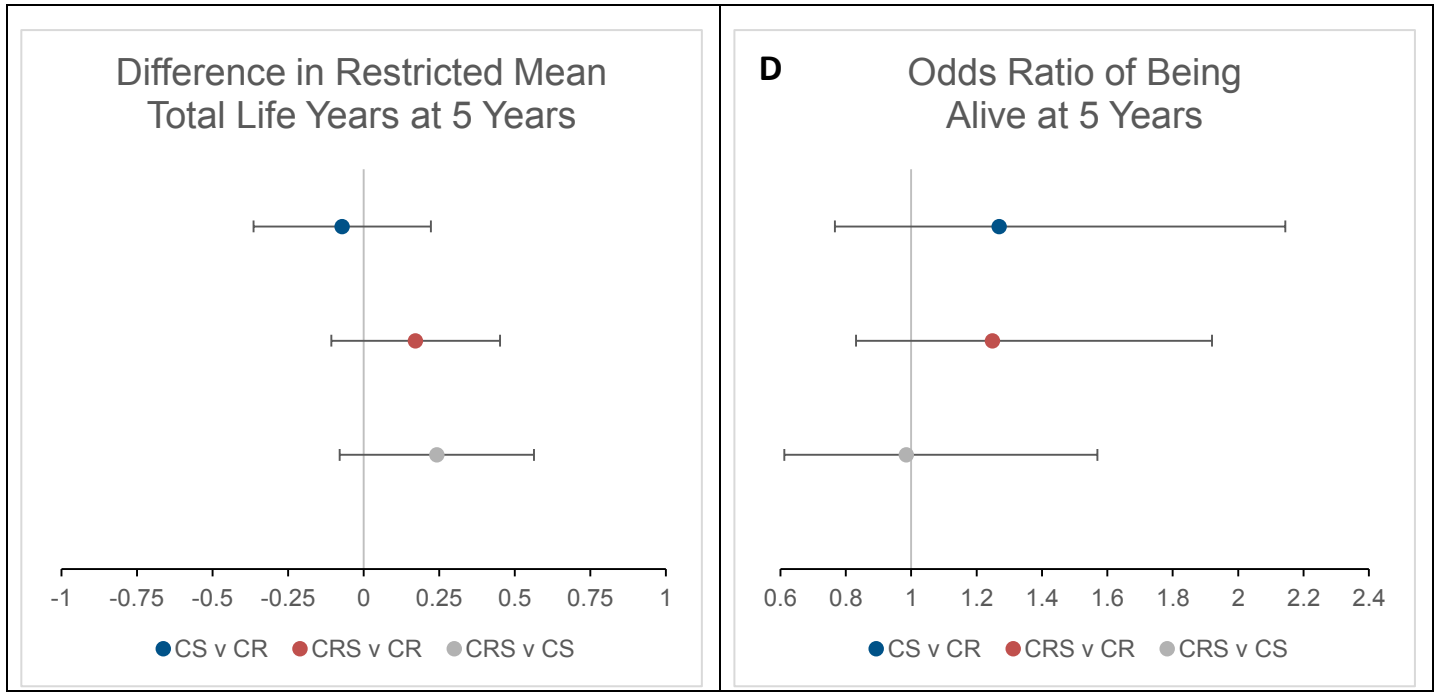
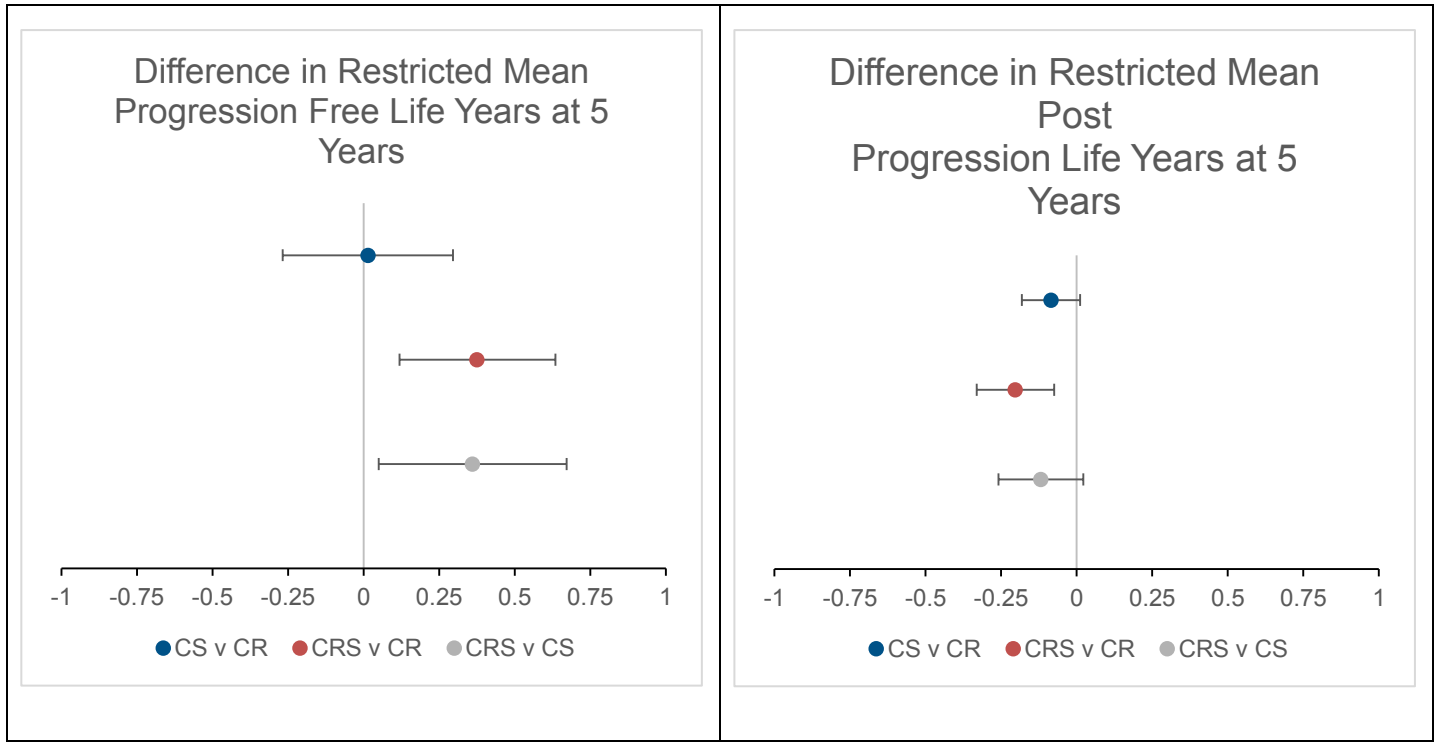


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

D, Table 14).


Difference in Restricted Mean Progression Free Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS

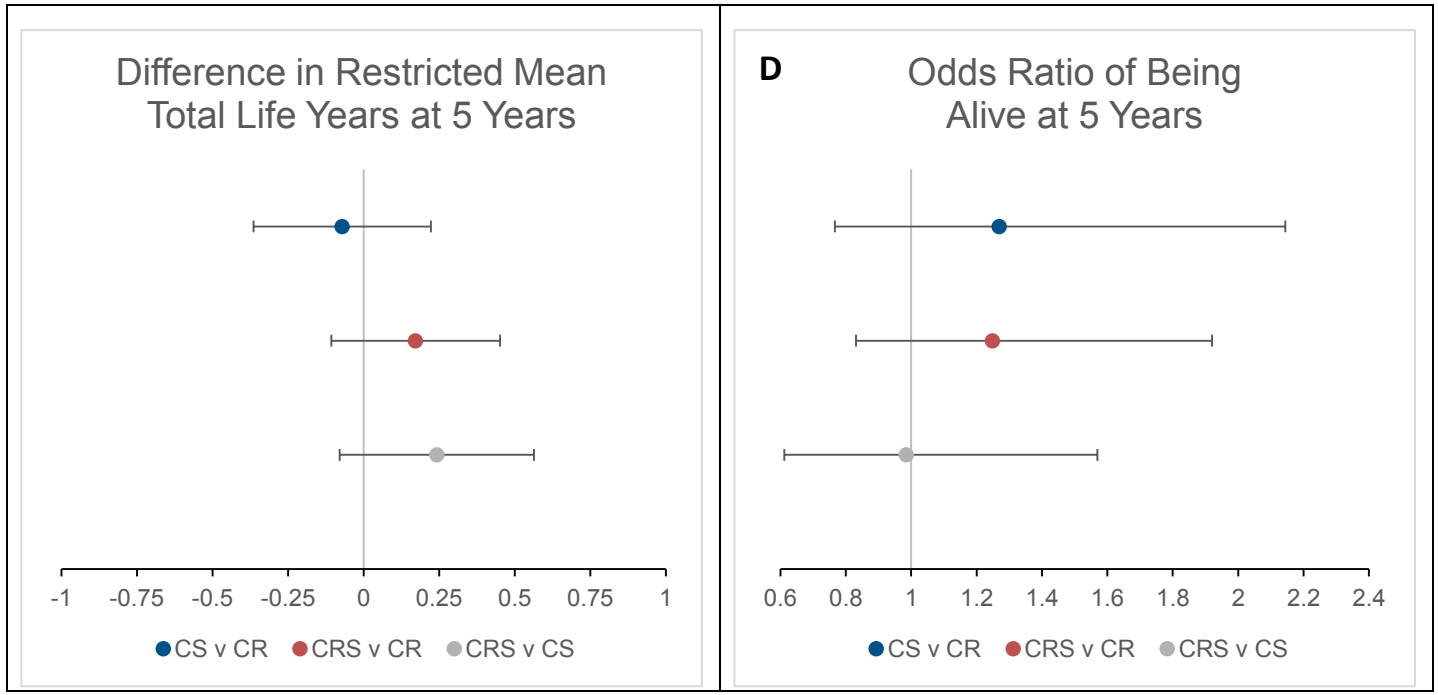


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

Table 14: Treatment differences in restricted mean survival times (RMST) up to 5 years, odds ratios of being alive at 5-years, probabilities of ranking best, ranks, and predicted RMST and probability of being alive at 5-years in the UK population for the three interventions.

		Intervention		
		Chemoradiotherapy[a]	Chemotherapy + Surgery	Chemoradiotherapy + Surgery
Difference in RMST (95% CrI[b])	Progression Free Life Years at 5 Years	Reference Treatment	0.01 (-0.27, 0.3)	0.38 (0.12, 0.63)
	Post Progression Life Years at 5 Years		-0.09 (-0.18, 0.01)	-0.2 (-0.33, -0.07)
	Total Life Years at 5 Years		-0.07 (-0.36, 0.22)	0.17 (-0.11, 0.45)
Odds Ratio (95% CrI)	Being Alive at 5 Years		1.27 (0.77, 2.14)	1.25 (0.83, 1.92)
Probability of Ranking Best	Progression Free Life Years at 5 Years	0.2%	1.1%	98.7%
	Post Progression Life Years at 5 Years	95.8%	4.1%	0.1%
	Total Life Years at 5 Years	9.9%	5.4%	84.7%
	Being Alive at 5 Years	6.3%	50.2%	43.6%
Median Rank (95% CrI)	Progression Free Life Years at 5 Years	3 (2, 3)	2 (2, 3)	1 (1, 1)

	Post Progression Life Years at 5 Years	1 (1, 2)	2 (1, 3)	3 (2, 3)
	Total Life Years at 5 Years	3 (1, 3)	2 (1, 3)	1 (1, 3)
	Being Alive at 5 Years	3 (1, 3)	1 (1, 3)	2 (1, 3)
Predicted RMST and Probability of Being Alive in UK at 5 Years[c]	Mean Progression Free Life Years	1.5 (1.28, 1.71)	1.51 (1.29, 1.73)	1.87 (1.57, 2.17)
	Mean Post Progression Life Years	0.58 (0.51, 0.65)	0.49 (0.42, 0.56)	0.37 (0.24, 0.51)
	Mean Total Life Years	2.07 (1.85, 2.29)	2 (1.77, 2.23)	2.24 (1.93, 2.56)
	Probability of Being Alive at 5 Years	0.13 (0.08, 0.18)	0.16 (0.11, 0.21)	0.16 (0.1, 0.23)

[a] Relative treatment effects presented for comparisons versus chemoradiotherapy. Point estimates are based on posterior medians.
[b] CrI = Credible Interval
[c] Baseline based on posterior distributions of outcomes for van Meerbeeck 2007.

Sensitivity analyses

As part of an assessment of the sensitivity of the results to the selected follow-up time, we also synthesised data based on a shorter follow-up period of 4-years, which allowed the inclusion of all 6 studies, including Girard 2009. Model fit statistics for the fixed and random effects models based on the 4-year follow-up data are given in Table 15Table 15. Convergence was satisfactory for the both models after a burn-in of 20,000 iterations and results are based on a further 40,000 samples on two chains.

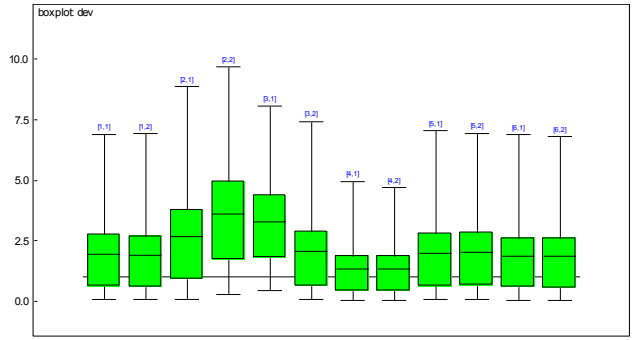
Table 15: Model fit statistics based on 4-year follow-up data

Model			DIC

		Posterior Median Between-Study SD (95% CrI)	Posterior mean residual deviance	
Fixed effect	P(Survival)	---	13.22	-27.429
	AUC		25.84	-20.356
Random effects	P(Survival)	0.24 (0.02, 1.63)	14.29	-25.090
	AUC	PFS: 0.12 (0.01, 0.76) PPS: 0.14 (0.01, 0.59)	23.61	-18.623

Total number of data points for P(survival) is 12 and for AUC is 24.

There were no meaningful differences between the fixed and random effects models in terms of the posterior mean residual deviance and DIC (Table 15). The plots of the posterior deviance values for each study arm in Figure 7 show that the probability of survival up to 4 years in Girard 2009 is not predicted well and this study is a possible outlier. Fitting a random effects model did not help in the prediction of data points in this study (Figure 8). The simpler fixed effect model is therefore preferred.



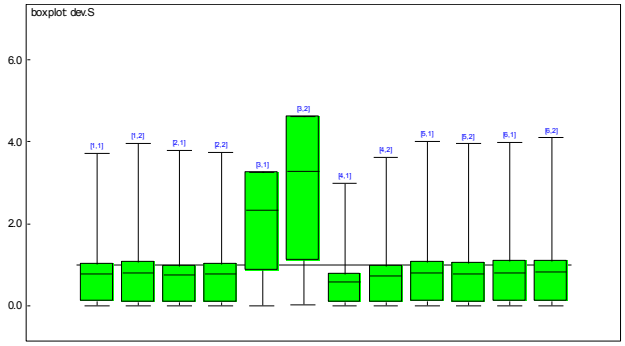


Figure 7: Posterior deviance values for each study arm for the area under the Kaplan Meier curves (left) and probability of survival (right) – fixed effect model.

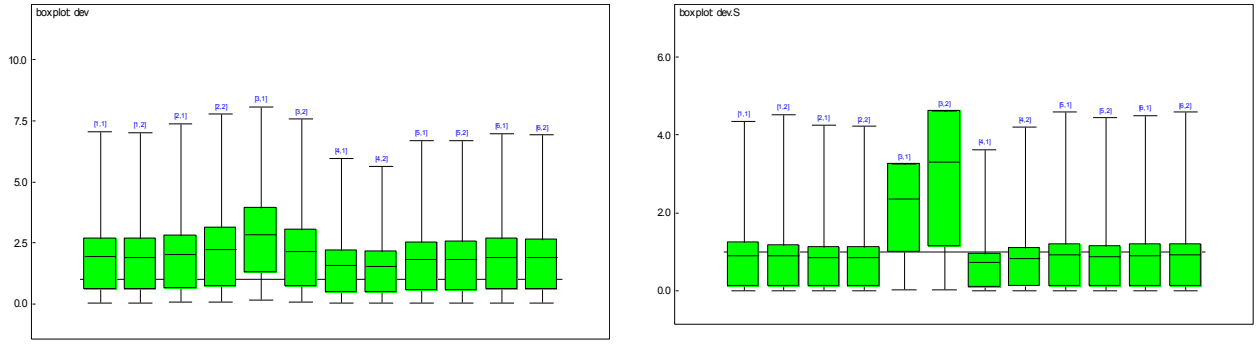


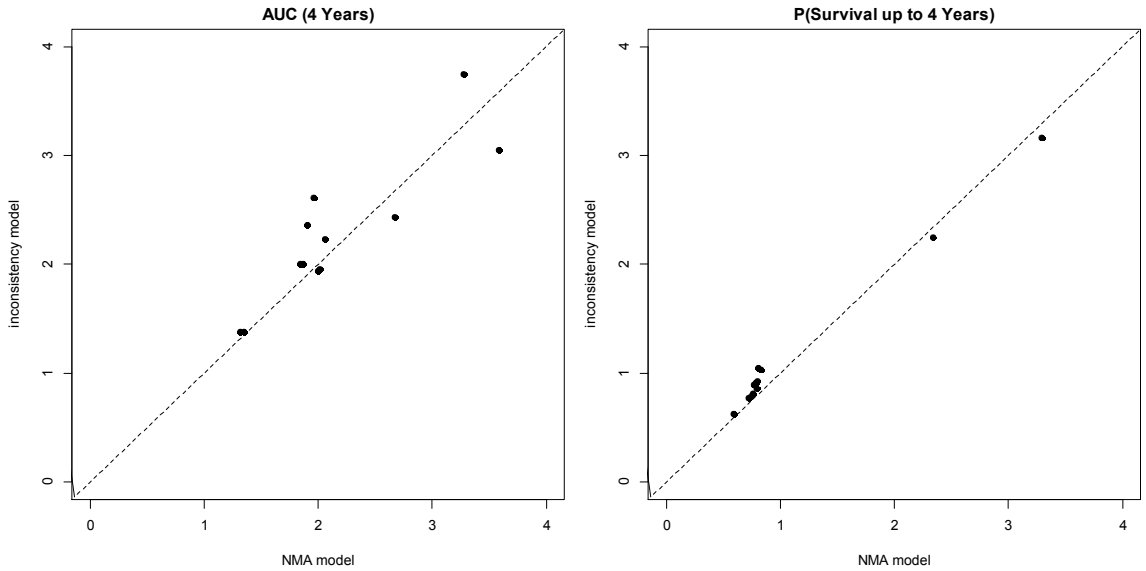
Figure 8: Posterior deviance values for each study arm for the area under the Kaplan Meier curves (left) and probability of survival (right) – random effects model.

No evidence of inconsistency was found through comparison of the consistency and inconsistency random effects models, as little difference was observed between the fit of the models (Table 16). The area below the line of equality in Figure 9 highlights where the inconsistency model better predicted data points, but any improvements were minimal.

Table 16: Model fit statistics for consistency and inconsistency fixed effect models based on 4-year follow-up data

Model		Posterior mean residual deviance	DIC
Fixed effect - consistency	P(Survival)	13.22	-27.429
	AUC	25.84	-20.356
Fixed effect - inconsistency	P(Survival)	14.07	-25.773
	AUC	27.07	-17.115

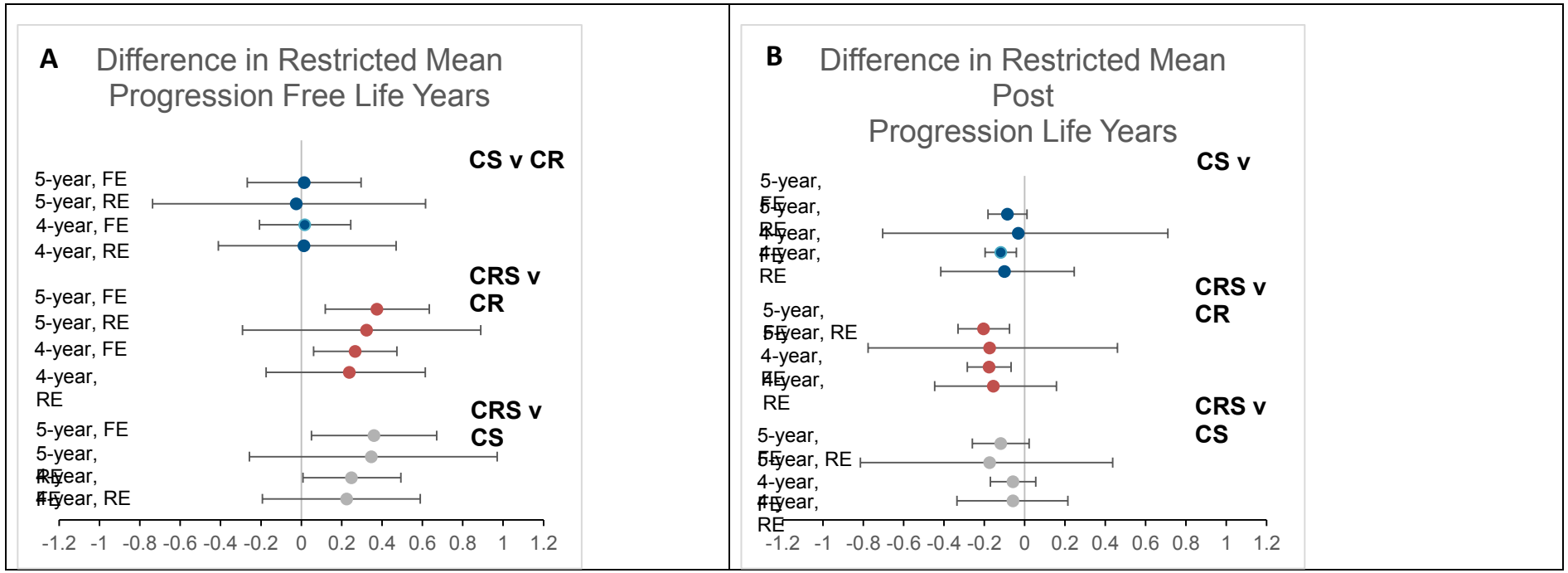
Total number of data points for P(survival) is 12 and for AUC is 24.



645

646 **Figure 9: Deviance contributions from the fixed effect consistency and inconsistency models for area under the Kaplan Meier curves**
647 **(left) and probability of survival (right).**

648 Treatment effects estimated by the fixed and random effects models based on the 4- and 5-year follow up data are presented in


A
Difference in Restricted Mean Progression Free Life Years
CS v CR
5-year, FE
5-year, RE
4-year, FE
4-year, RE
CRS v CR
5-year, FE
5-year, RE
4-year, FE
4-year, RE
CRS v CS
5-year, FE
5-year, RE
4-year, FE
4-year, RE
-1.2 -1 -0.8 -0.6 -0.4 -0.2 0 0.2 0.4 0.6 0.8 1 1.2
B
Difference in Restricted Mean Post Progression Life Years
CS v
CRS v CR
CRS v CS
-1.2 -1 -0.8 -0.6 -0.4 -0.2 0 0.2 0.4 0.6 0.8 1 1.2

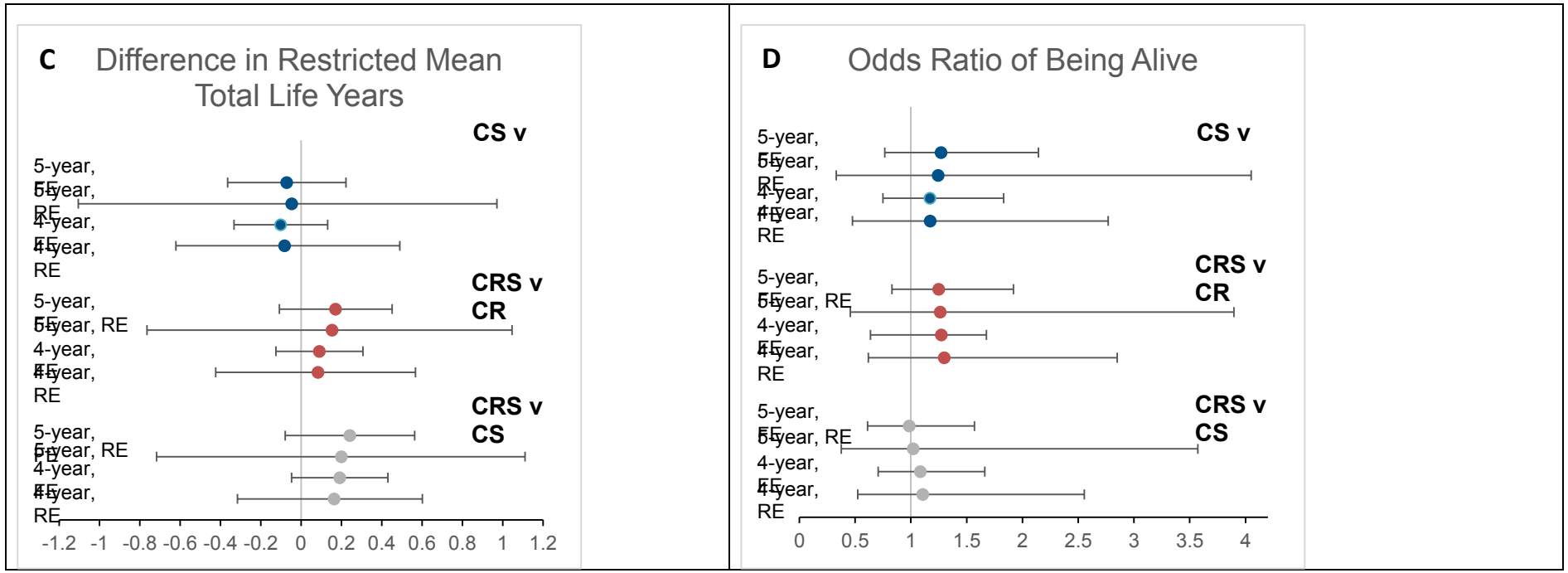
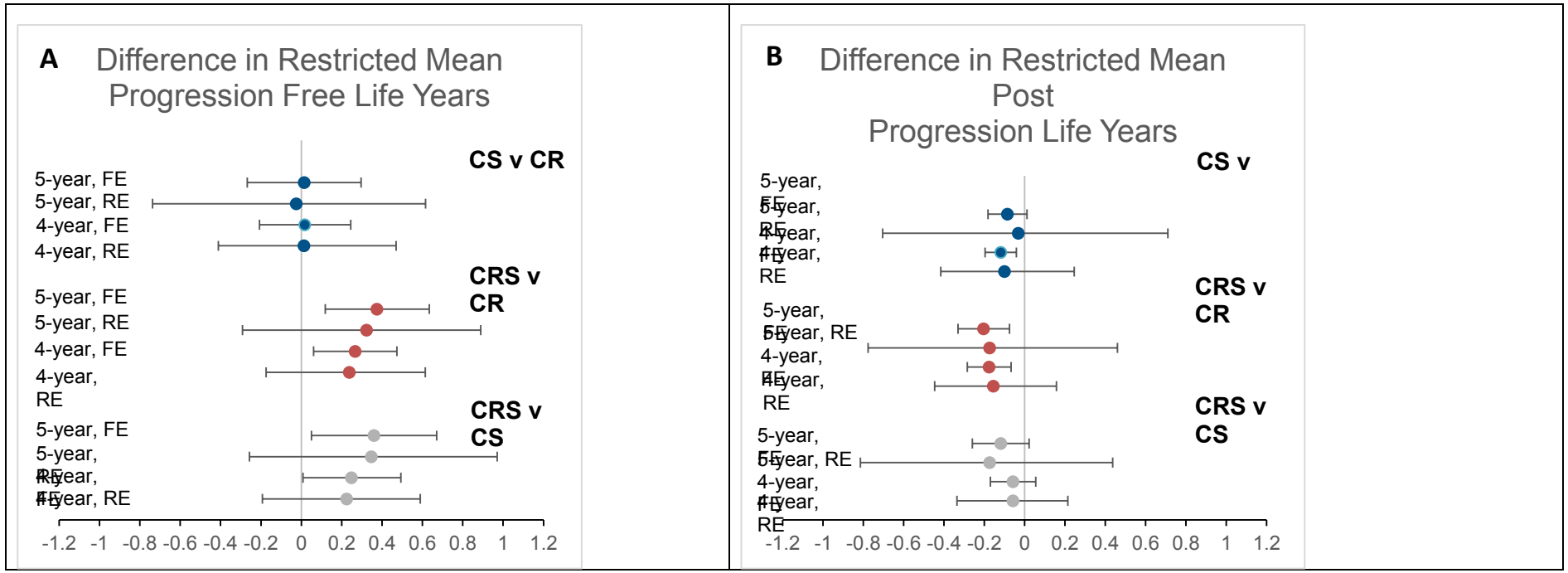


Figure 10. The point estimates of the treatment effects are similar, and the width of the credible intervals reflect that random effects models estimate the treatment effects with more uncertainty, and that there is additional data included in the 4-dataset compared with the 5-year dataset.

Noting that

4. the model fit assessment supports the use of the fixed effect model in both datasets,
5. the assumption that non-progressors by *T*-years do not progress (are "cured") is more reasonable at 5-years than at 4-years,
6. the 5-year dataset excludes the Girard (2009) study, which seems to be an outlier and is based on small numbers

supports the use of the fixed effect model based on the 5-year dataset for the base-case. Results from the random effects model based on the 5-year dataset are presented as a sensitivity analysis.


A
Difference in Restricted Mean Progression Free Life Years
CS v CR
5-year, FE
5-year, RE
4-year, FE
4-year, RE
CRS v CR
5-year, FE
5-year, RE
4-year, FE
4-year, RE
CRS v CS
5-year, FE
5-year, RE
4-year, FE
4-year, RE
-1.2 -1 -0.8 -0.6 -0.4 -0.2 0 0.2 0.4 0.6 0.8 1 1.2
B
Difference in Restricted Mean Post Progression Life Years
CS v
CRS v CR
CRS v CS
-1.2 -1 -0.8 -0.6 -0.4 -0.2 0 0.2 0.4 0.6 0.8 1 1.2

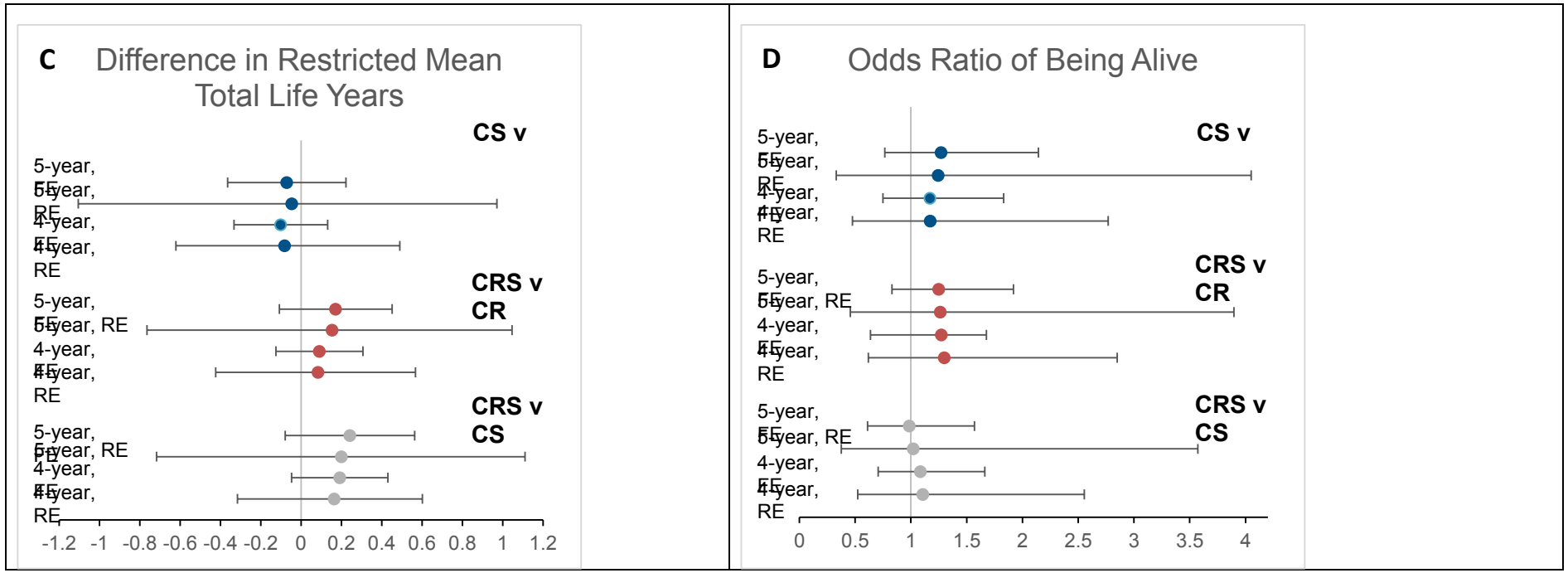


Figure 10: Forest plots of fixed and random effects estimates at 5- and 4-year follow up for (A) differences in restricted mean progression free life years at T-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at T-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at T-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at T-years follow-up relative to

chemoradiotherapy. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

Results: Inputs for Economic Model

Discounted Area Under the Kaplan Meier Curves and Probability of Survival

The fit of the NMA models based on the discounted AUC was also assessed and were in line with the results presented in Section 0 For both the 4-year and 5-year follow-up data, there were no meaningful differences between the fit of the fixed and random effects models (Table 17), and thus the fixed effect model was preferred.

Table 17: Model fit statistics based on 5-year follow-up data, discounted at 3.5% annual rate

Follow-Up Period	Model		Posterior Median Between-Study SD (95% CrI)	Posterior mean residual deviance	DIC
5 years [a]	Fixed effect[c]	P(Survival)	---	9.27	-24.85
		AUC		23.18	-14.69
	Random effects[d]	P(Survival)	0.33 (0.01, 2.34)	9.57	-22.94
		AUC	PFS: 0.17 (0.01, 1.25) PPS: 0.23 (0.03, 1.29)	18.86	-15.24
4 years [b]	Fixed effect[c]	P(Survival)	---	13.35	-27.18
		AUC		24.86	-23.87
	Random effects[e]	P(Survival)	0.22 (0.01, 1.56)	14.31	-25.08
		AUC	PFS: 0.11 (0.00, 0.68) PPS: 0.12 (0.01, 0.54)	23.34	-21.59

[a] Total posterior mean residual deviance compared to total number of data points for P(survival): 10 and AUC: 20
[b] Total posterior mean residual deviance compared to total number of data points for P(survival): 12 and AUC: 24
[c] Burn-in: 20,000 iterations, results based on: 40,000 samples, 2 chains
[d] Burn-in: 50,000 iterations, results based on: 100,000 samples, 2 chains
[e] Burn-in: 30,000 iterations, results based on: 60,000 samples, 2 chains

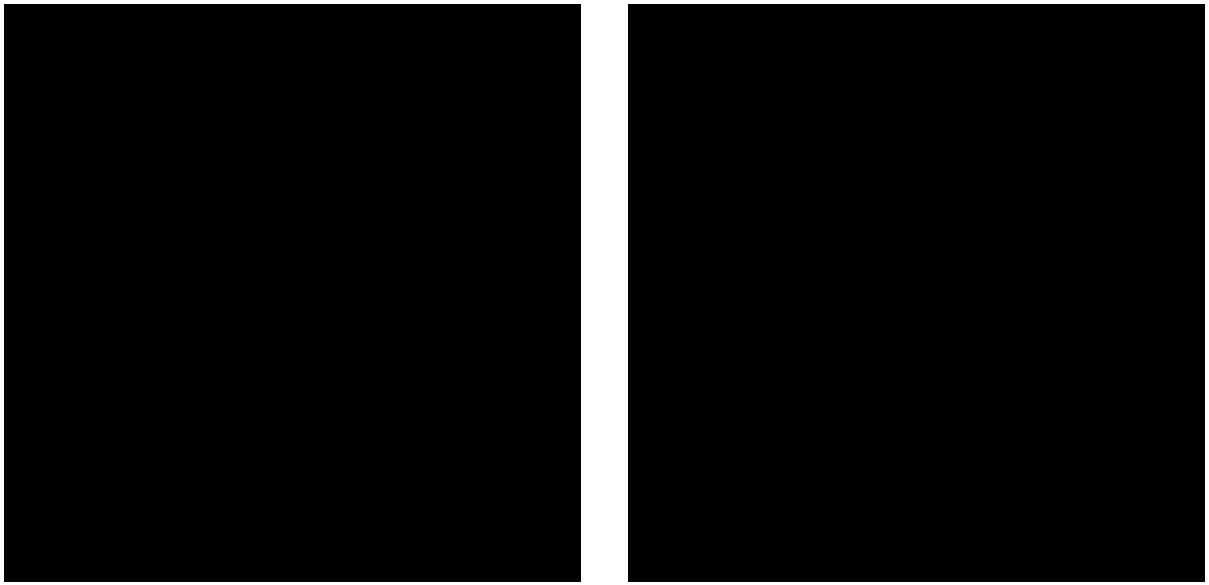
675 Similarly, the fit of the consistency and inconsistency models for both 4- and 5-year follow-up data were compared (Table 18). There is no
676 evidence of inconsistency as no meaningful differences were found in the fit of the models for both datasets. The area below the line of equality in
677 Figure 11 and Figure 12 highlights where the inconsistency model better predicted data points, but any improvements were minimal.
678

679 **Table 18: Model fit statistics for consistency and inconsistency fixed effect models based on 4-year follow-up data, discounted at 3.5%**
680 **annual rate**

Follow-Up Period	Model[c]		Posterior mean residual deviance	DIC
5 years [a]	Fixed effect - consistency	P(Survival)	9.27	-24.85
		AUC	23.18	-14.69
	Fixed effect – inconsistency	P(Survival)	10.17	-22.87
		AUC	23.43	-12.42
4 years [b]	Fixed effect – consistency	P(Survival)	13.35	-27.18
		AUC	24.86	-23.87
	Random effects - inconsistency	P(Survival)	14.15	-25.62
		AUC	26.12	-20.59

681 [a] Total posterior mean residual deviance compared to total number of data points for P(survival): 10 and AUC: 20
682 [b] Total posterior mean residual deviance compared to total number of data points for P(survival): 12 and AUC: 24
683 [c] Burn-in: 20,000 iterations, results based on: 40,000 samples, 2 chains
684

685 **Figure 11: Deviance contributions from the fixed effect consistency and inconsistency models for area under the Kaplan Meier curves**
686 **discounted at 3.5% annual rate (left) and probability of survival (right).**



687 **Figure 12: Deviance contributions from the fixed effect consistency and inconsistency models for area under the Kaplan Meier curves**
688 **discounted at 3.5% annual rate (left) and probability of survival (right).**

689 ***Proportion of Events Occurring each Year***

690 The proportion of events occurring each year pooled across studies is given in Table 19. The estimated proportions are similar across the 5-year
691 and 4-year follow-up datasets.

692

693

694 **Table 19: Pooled proportion of events occurring each year**

Follow-Up Period	Event Type	Year	Median Proportion of Events (95% CrI)
5-year	PFS[a]	1	0.63 (0.59, 0.67)
		2	0.23 (0.19, 0.28)
		3	0.08 (0.03, 0.13)
		4	0.04 (0.00, 0.09)
		5	0.01 (0.00, 0.07)
	OS[b]	1	0.38 (0.34, 0.42)
		2	0.32 (0.27, 0.38)
		3	0.16 (0.10, 0.22)
		4	0.11 (0.04, 0.17)
		5	0.03 (0.00, 0.10)
4-year	PFS[c]	1	0.65 (0.61, 0.69)
		2	0.24 (0.19, 0.30)
		3	0.09 (0.00, 0.14)
		4	0.01 (0.00, 0.08)
	OS[c]	1	0.39 (0.35, 0.43)
		2	0.35 (0.29, 0.41)
		3	0.17 (0.11, 0.23)
		4	0.10 (0.00, 0.15)

695 [a] Burn-in: 500,000 iterations, results based on: 1,000,000 samples, 2 chains
696 [b] Burn-in: 2,000,000 iterations, results based on: 4,000,000 samples, 2 chains
697 [c] Burn-in: 100,000 iterations, results based on: 100,000 samples, 2 chains
698

699 **NMA for Adverse Events**

700 The base case approach used in the economic model for adverse events used pairwise meta-analyses but data then became available that
701 allowed us to fit an NMA for use in sensitivity analyses.

The studies had reported adverse events heterogeneously; in some studies the reporting was comprehensive and in others scant or no details were available. Furthermore, events were classified heterogeneously across studies, being grouped under narrow or broad classes that made event-specific pooling difficult. The committee decided that adverse events should be included in the economic model if possible and we agreed an aggregate approach with them. This involved grouping all adverse events of grade 3+ as homogenously requiring one hospital admission, but having no long term clinical effects or detriment to quality of life. The committee thought it possible that grade 4 adverse events would affect quality of life but these occurred to sparsely to be meaningfully included in the model. Because of the wide disparity between the frequency of adverse events reported among the studies, we selected Pless 2015, Eberhardt 2015, Albain 2009 and van Meerbeeck 2007 for the analysis. These studies were the largest and best conducted studies in the network and had reported event rates that the committee found credible. The data from van Meerbeeck was not reported in the published paper but provided to us upon request by the EORTC, who hold the trial data. We obtained the person years at risk by multiplying the total number of patients in each arm by the mean AUC for total life years at 5 years. The data are in Table 20.

Table 20: Adverse Event NMA Input Data

Treatment Arm 1	Events Arm 1	TatRisk Arm 1	Treatment Arm 2	Events Arm 2	TatRisk Arm 2	Study		Treatments
2	182	285.2	3	141	299.52	Pless 2015		1=CR
3	482	434.3	1	608	409.34	Albain 2009		2=CS
1	137	214.4	3	150	230.04	Eberhardt 2015		3=CRS
1	98	321.75	2	108	298.93	van Meerbeerck 2007		

We assumed that adverse events were treatment related and therefore that it was appropriate to assume a homogenous follow-up time. Since this meant that we did not have to account for variable study endpoints in our pooling of the data, we selected a poisson likelihood, log link NMA model and copied the code directly from NICE TSD2 (citation). The results of the fixed and random effects models are in Table 21. Models were run using 50,000 burn-in iterations and 50,000 iterations to generate the posterior distributions.

Table 21: Adverse Event NMA Results

All Adverse Events	estimate	LCL	UCL	DIC
Fixed effects				74.44
HR of CS vs CR	1.132	0.9382	1.354	
HR of CR vs CRS	1.2425447	1.125112511	1.377221	
HR of CS vs CRS	1.3970383	1.174950065	1.67336	

Random effects				72.627
HR of CR vs CS	1.166	0.3146	4.654	
HR of CR vs CRS	1.176886	0.374531835	3.354579	
HR of CS vs CRS	1.3696754	0.361663653	5.186722	

The DIC for the random effects model was not more than 3-5 points lower than the fixed effects model so we preferred it in the base case. The results show that both CR and CS are associated with more adverse events than CRS.

As discussed in the economic modelling report (Appendix J), the NMA data agreed well with the pairwise estimates of adverse events.

References and Code

References

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12. Spiegelhalter, D.J., et al., *Bayesian measures of model complexity and fit.* Journal of the Royal Statistical Society (B), 2002. **64**(4): p. 583-616.

Code

SEER dataset

Selection criteria:

{Age at Diagnosis.Age recode with <1 year olds} = '30-34 years','35-39 years','40-44 years','45-49 years','50-54 years','55-59 years','60-64 years','65-69 years','70-74 years','75-79 years'

AND ({Site and Morphology.CS Schema v0204+} = 'Lung'

OR {Site and Morphology.CS Schema - AJCC 6th Edition} = 'Lung')

AND ({Stage - AJCC.Derived AJCC Stage Group, 7th ed (2010+)} = 'IIIA'

OR {Stage - AJCC.Derived AJCC Stage Group, 6th ed (2004+)} = 'IIIA'

OR {Stage - AJCC.AJCC stage 3rd edition (1988-2003)} = ' 31'

OR {Stage - AJCC.SEER modified AJCC stage 3rd (1988-2003)} = ' 31')

AND ({Stage - TNM.Derived AJCC N, 7th ed (2010+)} = 'N2','N2a','N2b','N2c'

OR {Stage - TNM.Derived AJCC N, 6th ed (2004+)} = 'N2','N2a','N2b','N2c'

OR {Stage - TNM.N value - based on AJCC 3rd (1988-2003)} = 'N2')

NMA Model for Adverse Events – Fixed Effects

```
# Poisson likelihood, log link
# Fixed effects model for multi-arm trials
model{ # *** PROGRAM STARTS
```

```
764  for(i in 1:ns){ # LOOP THROUGH STUDIES
765   mu[i] ~ dnorm(0,.0001) # vague priors for all trial baselines
766   for (k in 1:na[i]) { # LOOP THROUGH ARMS
767   r[i,k] ~ dpois(theta[i,k]) # Poisson likelihood
768   theta[i,k] <- lambda[i,k]*E[i,k] # event rate * exposure
769   log(lambda[i,k]) <- mu[i] + d[t[i,k]] - d[t[i,1]] # model for linear predictor
770   dev[i,k] <- 2*((theta[i,k]-r[i,k]) + r[i,k]*log(r[i,k]/theta[i,k])) #Deviance contribution
771   }
772   resdev[i] <- sum(dev[i,1:na[i]]) # summed residual deviance contribution for this trial
773   }
774  totresdev <- sum(resdev[]) #Total Residual Deviance
775  d[1]<-0 # treatment effect is zero for reference treatment
776  for (k in 2:nt){ d[k] ~ dnorm(0,.0001) } # vague priors for treatment effects
777
778
779
780
781  sd ~ dunif(0,5) # vague prior for between-trial SD
782  tau <- pow(sd,-2) # between-trial precision = (1/between-trial variance)
783
```

```
# pairwise HRs and LHRs for all possible pair-wise comparisons, if nt>2
for (c in 1:(nt-1)) {
 for (k in (c+1):nt) {
 lhr[c,k] <- (d[k]-d[c])
 log(hr[c,k]) <- lhr[c,k]
 }
 }

} # *** PROGRAM ENDS

list(ns=4, nt=3)

t[,1]  r[,1]  E[,1]  t[,2]  r[,2]  E[,2]  na[]
2      182    285.2  3      141    299.52 2
3      482    434.3  1      608    409.34 2
1      137    214.4  3      150    230.04 2
1      98     321.75 2      108    298.93 2

END
```

```
#chain 1
list(d=c( NA, 0, 0), mu=c(0, 0, 0, 0))
#chain 2
list(d=c( NA, -1, 1), mu=c(-3, -3, -3, -3))
#chain 3
list(d=c( NA, 2, 2),  mu=c(-3, 5, -1, -3))
```

NMA Model for Adverse Events - Random Effects

```
# Poisson likelihood, log link
# Random effects model for multi-arm trials
model{ # *** PROGRAM STARTS
for(i in 1:ns){ # LOOP THROUGH STUDIES
 w[i,1] <- 0 # adjustment for multi-arm trials is zero for control arm
 delta[i,1] <- 0 # treatment effect is zero for control arm
 mu[i] ~ dnorm(0,.0001) # vague priors for all trial baselines
 for (k in 1:na[i]) { # LOOP THROUGH ARMS
 r[i,k] ~ dpois(theta[i,k]) # Poisson likelihood
 theta[i,k] <- lambda[i,k]*E[i,k] # failure rate * exposure
 log(lambda[i,k]) <- mu[i] + delta[i,k] # model for linear predictor
```

```
824  dev[i,k] <- 2*((theta[i,k]-r[i,k]) + r[i,k]*log(r[i,k]/theta[i,k])) #Deviance contribution
825  }
826  resdev[i] <- sum(dev[i,1:na[i]]) # summed residual deviance contribution for this trial
827  for (k in 2:na[i]) { # LOOP THROUGH ARMS
828  delta[i,k] ~ dnorm(md[i,k],taud[i,k]) # trial-specific LOR distributions
829  md[i,k] <- d[t[i,k]] - d[t[i,1]] + sw[i,k] # mean of LOR distributions (with multi-arm trial correction)
830  taud[i,k] <- tau *2*(k-1)/k # precision of LOR distributions (with multi-arm trial correction)
831  w[i,k] <- (delta[i,k] - d[t[i,k]] + d[t[i,1]]) # adjustment for multi-arm RCTs
832  sw[i,k] <- sum(w[i,1:k-1])/(k-1) # cumulative adjustment for multi-arm trials
833  }
834  }
835
836
837  totresdev <- sum(resdev[]) #Total Residual Deviance
838  d[1]<-0 # treatment effect is zero for reference treatment
839  for (k in 2:nt){ d[k] ~ dnorm(0,.0001) } # vague priors for treatment effects
840  sd ~ dunif(0,5) # vague prior for between-trial SD
841  tau <- pow(sd,-2) # between-trial precision = (1/between-trial variance)
842  # pairwise HRs and LHRs for all possible pair-wise comparisons, if nt>2
843  for (c in 1:(nt-1)) {
```

```
844     for (k in (c+1):nt) {
845     lhr[c,k] <- (d[k]-d[c])
846     log(hr[c,k]) <- lhr[c,k]
847     }
848     }
849
850   } # *** PROGRAM ENDS
851
852   list(ns=4, nt=3)
853
854   t[,1]    r[,1]    E[,1]    t[,2]    r[,2]    E[,2]    na[]
855   2        182      285.2    3        141      299.52   2
856   3        482      434.3    1        608      409.34   2
857   1        137      214.4    3        150      230.04   2
858   1        98       321.75   2        108      298.93   2
859
860   END
861
862   #chain 1
863   list(d=c( NA, 0, 0), sd=1, mu=c(0, 0, 0, 0))
```

#chain 2

list(d=c(NA, -1, 1), sd=4, mu=c(-3, -3, -3, -3))

#chain 3

list(d=c(NA, 2, 2), sd=2, mu=c(-3, 5, -1, -3))

Network meta-analysis for PFS and OS

Let $y_{i,k}^{PFS}$ and $y_{i,k}^{OS}$ be the estimated AUC up to *T* years for study *i*, arm *k*, for PFS and OS respectively, with covariance matrix $V_{i,k}$ for the PFS and OS AUC(*T*) outcomes. We assume the AUCs follows a Bivariate Normal likelihood:

$$\begin{pmatrix} y_{i,k}^{PFS} \\ y_{i,k}^{OS} \end{pmatrix} \sim N\left(\begin{pmatrix} \theta_{i,k}^{PFS} \\ \theta_{i,k}^{OS} \end{pmatrix}, V_{i,k} \right)$$

For PFS, the NMA model is:

$$\theta_{i,k}^{PFS} = \mu_i^{PFS} + \delta_{i,k}^{PFS}$$

where μ_i^{PFS} is the baseline AUC for PFS in study i, and $\delta_{i,k}^{PFS}$ the difference in AUC for treatment in arm k relative to the treatment in arm 1 in study i, which may be modelled as either a fixed or random effect:

$$\delta_{i,k}^{PFS} = d_{t_{i,k}}^{PFS} - d_{t_{i,1}}^{PFS} \quad \text{Fixed effect model}$$
$$\delta_{i,k}^{PFS} \sim N\left(d_{t_{i,k}}^{PFS} - d_{t_{i,1}}^{PFS}, \sigma_{PFS}^2\right) \quad \text{Random effects model}$$

where d_k^{PFS} is the difference in AUC for treatment k relative to treatment 1 ($d_1^{PFS} = 0$), and σ_{PFS} is the between-study standard deviation in treatment differences in AUC. For OS, the AUC is defined as the sum of the AUC for PFS and post-progression survival (PPS):

$$\theta_{i,k}^{OS} = \theta_{i,k}^{PFS} + \theta_{i,k}^{PFS}$$

A NMA model is given to PPS, as for PFS:

$$\theta_{i,k}^{PPS} = \mu_i^{PPS} + \delta_{i,k}^{PPS}$$

$$\delta_{i,k}^{PPS} = d_{t_{i,k}}^{PPS} - d_{t_{i,1}}^{PPS} \quad \text{Fixed effect model}$$

$$\delta_{i,k}^{PPS} \sim N\left(d_{t_{i,k}}^{PPS} - d_{t_{i,1}}^{PPS}, \sigma_{PPS}^2\right) \quad \text{Random effects model}$$

Normal(0,10000) prior distributions are given to the trial-specific baselines μ_i^{PFS}, μ_i^{PPS} and for the treatment effects on the AUCs d_k^{PFS}, d_k^{PPS}. In the case of random effects models, the between study standard deviations $\sigma_{PFS}, \sigma_{PPS}$ for the treatment effects on AUC for PFS and PPS were assigned Uniform(0,5) priors.

For an assumed restricted mean PFS time over *T*-years on reference treatment 1 in a UK population, μ_{UK}^{PFS} , we can derive the mean time spent progression free up to *T*-years for treatment *k* in a UK population:

$$meanPFS_k(T) = \mu_{UK}^{PFS} + d_k^{PFS}$$

Similarly, for an assumed mean PPS time over *T*-years on reference treatment 1 in a UK population, μ_{UK}^{PPS} , we can derive the mean time spent in PPS for treatment *k* in a UK population:

$$meanPPS_k(T) = \mu_{UK}^{PPS} + d_k^{PPS}$$

μ_{UK}^{PFS} and μ_{UK}^{PPS} over 4- and 5- years were set to be the posterior distributions of the mean PFS and PPS in the group receiving chemoradiotherapy in the van Meerbeeck 2007 study, since this was the largest study and did not have the limitations of the other studies with chemoradiotherapy arms, Eberhardt (partially indirect population) and Albain (US setting).

Predicted Mean Survival Time

To predict lifetime mean survival time beyond the truncated study periods (*T* = 4 or 5 years), required extrapolation using long-term survival data from an external source. Let C be the area under the Kaplan Meier curve obtained from an appropriate external source of data conditional on having survived *T*-years, which can be interpreted as life-expectancy conditional on surviving the first *T* years.

Assuming that all those who are alive at *T*-years are progression free, and remain progression free thereafter, the mean time spent progression free for treatment *k* in a UK population is:

$$meanPFS_k = meanPFS_k(T) + S_k(T) * C$$

where $S_k(T)$ is the probability of surviving to *T* years.

Under the assumption that those who survive to *T*-years remain progression-free, no further time spent in PPS is obtained after *T*-years so that:

$$meanPPS_k = meanPPS_k(T).$$

Visual inspection of the Kaplan Meier curves for each study suggested this assumption was reasonable.

1.1.1. Probability of Surviving up to T years, $S_k(T)$

The probability of surviving up to *T* years (*T* = 4 or 5 years) for each treatment group was pooled across trials in a separate NMA. Let $y_{i,k}^S = S_{i,k}(T)$ be the estimated survival probability at *T*-years in study *i*, arm *k*, with standard error $se_{i,k}$. Assuming the survival probabilities at *T*-years follow a Normal likelihood:

$$y_{i,k}^S \sim N\left(\pi_{i,k}, se_{i,k}^2\right)$$

The NMA model is put on the logit-scale:

$$logit(\pi_{i,k}) = \mu_i^S + \delta_{i,k}^S$$

$$\delta_{i,k}^S = d_{t_{i,k}}^S - d_{t_{i,1}}^S \quad \text{Fixed effect model}$$

$$\delta_{i,k}^S \sim N\left(d_{t_{i,k}}^S - d_{t_{i,1}}^S, \sigma_S^2\right) \quad \text{Random effects model}$$

where μ_i^S are the study-specific log-odds of survival to T years and d_k^S is the log-odds ratio of survival to T years for treatment k relative to treatment 1.

Trial-specific baseline μ_i^S and treatment effects d_k^S for probability of survival up to 4 or 5 years were assigned Normal(0,10000) prior distributions.

In the case of random effects models, the between study standard deviation σ_S was assigned a Uniform(0,5) prior.

External Survival Data

To estimate mean survival time beyond T years conditional on surviving to T years, we made use of survival data collected from the Surveillance Epidemiology and End Results (SEER) cancer incidence database [8]. A subset of the incidence database was extracted to ensure patients matched those include in the NMA in terms of age at diagnosis (30 – 79 years), cancer site (lung), and stage of cancer (IIIA-N2). Exact selection criteria are given in Section 8. This produced a dataset of 23,602 patients with a maximum observed survival time of 25.7 years. Since the SEER dataset was used to predict survival beyond the truncated study period, we were interested in the SEER data conditional on patients being alive at the end of the truncated study period. After conditioning survival on being alive at 4 and 5 years after diagnosis, data on the remaining 3,703 and 2,865 patients, respectively, were used to calculate the area under the conditional SEER Kaplan Meier curves using the methods described in Section 2.2. Several parametric survival curves were fitted to the SEER data: exponential, Weibull, gamma, log-normal, Gompertz, and log-logistic. The fit of each curve was compared using the Akaike information criterion (AIC) and Bayesian information criterion (BIC). For the SEER data conditional on being alive at 5 years, a Weibull distribution with a shape parameter of 0.88 and scale parameter of 7.37 gave the lowest AIC (Figure 1). For the SEER data conditional on being alive at 4 years, a Weibull distribution with a shape parameter of 0.85 and scale parameter of 6.88 gave the lowest AIC.

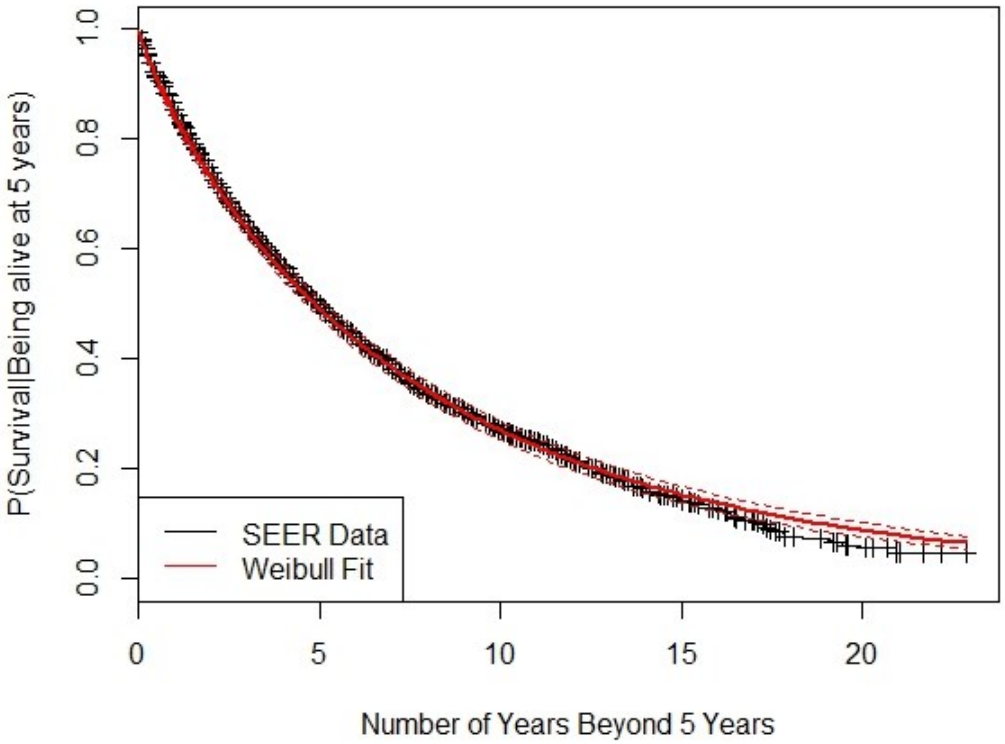


Figure 1: Kaplan Meier Curve for SEER data conditional on being alive at 5 years with fitted Weibull curve superimposed

Additional Requirements for Economic Model

Discounting Area Under the Kaplan Meier Curves

The economic evaluation required the area under the Kaplan Meier curve to be discounted at an annual rate of 3.5% [7]. The discounted area (up to *T* years) for each treatment group within each trial, as well as the SEER dataset, was calculated as

938 $$AUC_{disc_T} = \sum_{i=1}^{n_j} (t_i - t_{i-1}) \hat{S}_{KM}(t_{i-1}) + \sum_{j=2}^{T} \rho^{j-1} \sum_{i=n_{j-1}+1}^{n_j} (t_i - t_{i-1}) \hat{S}_{KM}(t_{i-1})$$

939 where

940 $\rho = \frac{1}{1.035}$, n_j is the index marking the end of year j = 1, …, *T*, and $\hat{S}_{KM}(t_{i-1})$ is the probability of surviving up to time t_{i-1}. As part of a sensitivity

941 analysis, the area under the Kaplan Meier curves were also discounted at an annual rate of 1.5% (i.e., $p = \frac{1}{1.015}$).

942 The standard error of, and correlation between, the discounted area under the Kaplan Meier curves for PFS and OS was calculated using non-
943 parametric bootstrapping, constrained to samples where the OS curve was greater than the PFS curve [6]. The discounted areas under the Kaplan
944 Meier curves for each RCT are provided in Table 10.

945 **Table 10: Discounted area under the curve data required for economic modelling**

Discount Rate			***(Treatment 1=CR, 2=CS and 3=CRS)***	**AUC**	**SE**	**AUC**	**SE**	
3.5%	**5-years**	Albain	1	1.55	0.11	2.33	0.15	0.87
			3	1.95	0.13	2.42	0.15	0.91
		Eberhardt	1	2.41	0.23	3.09	0.21	0.95
			3	2.49	0.22	3.30	0.21	0.88
		Katakami	2	1.60	0.28	2.88	0.30	0.85
			3	2.15	0.35	3.19	0.30	0.88

946

		Pless	2	1.86	0.18	2.90	0.19	0.87
			3	2.13	0.19	2.94	0.18	0.87
		van Meerbeeck	1	1.52	0.12	2.11	0.12	0.95
			2	1.48	0.12	1.96	0.13	0.96
	4-years	Albain	1	1.38	0.09	2.04	0.11	0.83
			3	1.66	0.10	2.07	0.11	0.87
		Eberhardt	1	1.97	0.17	2.58	0.15	0.92
			3	2.08	0.16	2.71	0.16	0.85
		Girard	2	2.13	0.42	2.38	0.30	0.96
			3	1.60	0.32	2.07	0.30	0.96
		Katakami	2	1.43	0.23	2.51	0.22	0.79
			3	1.82	0.26	2.70	0.22	0.83
		Pless	1	1.57	0.14	2.38	0.13	0.87
			2	1.83	0.14	2.46	0.13	0.83
		van Meerbeeck	1	1.36	0.09	1.89	0.09	0.94
			3	1.32	0.10	1.73	0.10	0.96
1.5%	**5-years**	Albain	1	1.53	0.11	2.29	0.15	0.87
			3	1.91	0.13	2.37	0.14	0.91
		Eberhardt	1	2.35	0.22	3.02	0.20	0.95
			3	2.43	0.21	3.22	0.21	0.88
		Katakami	2	1.57	0.28	2.82	0.30	0.85
			3	2.10	0.34	3.12	0.29	0.88
		Pless	2	1.82	0.17	2.83	0.18	0.87
			3	2.09	0.18	2.87	0.18	0.87
		van Meerbeeck	1	1.49	0.11	2.07	0.12	0.95
			2	1.46	0.12	1.93	0.12	0.96

947 Abbreviations: AUC – area under the curve, OS – overall survival, PFS – progression free survival, SE – standard error.

To compute discounted costs of death beyond the truncated study periods (*T* = 4 or 5 years), a parametric survival curve was used to model the conditional SEER data, as described in section 0

Discounting one-off costs

The economic model includes one-off costs for progression events, which also require discounting. The non-parametric approach provides the total number of events by time *T*, but does not give the breakdown of these events into 1-year time periods required for discounting. To obtain the proportion of total events falling in each 1-year period, let $y_{i,k,s}$ be the survival probability at s years with standard error $se_{i,k,s}$, in arm *k* of study *i*. We assume the survival probabilities follow a Normal likelihood:

$$y_{i,k,s} \sim N(\pi_{i,k,s}, se^2_{i,k,s})$$

where $\pi_{i,k,s}$ is the survival probability in study i, arm k, and time s.

Let $\rho_{i,k,s}$ be the proportion of events that have occurred by *T* = 5-years in study *i*, arm *k*, that occur in year *s*. Then the proportion surviving to 4-years, $\pi_{i,k,4}$, is the proportion surviving to 5 years, plus for those experiencing an event by year 5 the proportion of those events that occur in the 5th year:

$$\pi_{i,k,4} = \pi_{i,k,5} + \left(1-\pi_{i,k,5}\right)\rho_{i,k,5}$$

Similarly:

$$\pi_{i,k,3} = \pi_{i,k,5} + \left(1-\pi_{i,k,5}\right)\left(\rho_{i,k,4} + \rho_{i,k,5}\right)$$

$$\pi_{i,k,2} = \pi_{i,k,5} + \left(1-\pi_{i,k,5}\right)\left(\rho_{i,k,3} + \rho_{i,k,4} + \rho_{i,k,5}\right)$$

$$\pi_{i,k,1} = \pi_{i,k,5} + \left(1-\pi_{i,k,5}\right)\left(\rho_{i,k,2} + \rho_{i,k,3} + \rho_{i,k,4} + \rho_{i,k,5}\right)$$

Each $\pi_{i,k,5}$ is given a Beta(1,1) prior, so that the 5-year survival probabilities are unconstrained, and the focus of analysis is the distribution of events over the 1-year periods, $\rho_{i,k,s}$, which are modelled with a Dirichlet distribution to ensure they sum to 1:

965 $$\left(\rho_{i,k,1},\rho_{i,k,2},\rho_{i,k,3},\rho_{i,k,4},\rho_{i,k,5}\right)\sim Dirichlet\left(\alpha_{i,k,1},\alpha_{i,k,2},\alpha_{i,k,3},\alpha_{i,k,4},\alpha_{i,k,5}\right)$$

966 The $\alpha_{i,k,s}$ are modelled on the log-scale. We explored a range of assumptions regarding the effects of time period and treatment, but found the
967 additive time model with no study and no treatment effects to give sufficiently good fit based on the posterior mean residual deviance:

968 $$\log(\alpha_{i,k,s})=\beta_s$$

969 Note this does not mean that study and treatment have no effect on survival probability, but that this is already captured in the estimation of the *T*-
970 year survival probability. This model was run separately for PFS and OS events. Normal(0,100) priors were assigned to β_s. The proportion of
971 events occurring each year for each RCT are provided in Table 11.

972 ***Table 11: Proportion of events occurring each year (Treatment 1=CR, 2=CS and 3=CRS)***

973

			P(event)	SE	P(event)	SE	P(event)	SE	P(event)	SE	P(event)	SE
PFS	Albain	1	0.47	0.04	0.24	0.03	0.19	0.03	0.14	0.03	0.12	0.02
		3	0.53	0.04	0.33	0.03	0.28	0.03	0.25	0.03	0.23	0.03
	Eberhardt	1	0.60	0.05	0.42	0.06	0.36	0.06	0.36	0.06	0.36	0.06
		3	0.69	0.05	0.45	0.06	0.40	0.06	0.34	0.06	0.33	0.06
	Girard	2	0.57	0.13	0.43	0.13	0.43	0.13	0.43	0.13	N/A	N/A
		3	0.53	0.12	0.29	0.11	0.24	0.10	0.24	0.10	N/A	N/A
	Katakami	2	0.38	0.09	0.31	0.09	0.17	0.07	0.14	0.06	0.07	0.05
		3	0.55	0.09	0.34	0.09	0.34	0.09	0.31	0.09	0.21	0.08
	Pless	2	0.49	0.05	0.31	0.04	0.26	0.04	0.24	0.04	0.22	0.04
		3	0.52	0.05	0.40	0.05	0.34	0.05	0.28	0.05	0.22	0.05
	van Meerbeeck	1	0.45	0.04	0.24	0.03	0.16	0.03	0.13	0.03	0.12	0.03
		2	0.40	0.04	0.27	0.03	0.17	0.03	0.14	0.03	0.11	0.03
OS	Albain	1	0.69	0.04	0.45	0.04	0.33	0.04	0.25	0.04	0.19	0.04
		3	0.68	0.04	0.48	0.04	0.37	0.04	0.28	0.04	0.26	0.04

	Eberhardt	1	0.83	0.04	0.63	0.05	0.50	0.06	0.41	0.06	0.41	0.06
		3	0.78	0.05	0.69	0.05	0.60	0.06	0.50	0.06	0.44	0.06
	Girard	2	0.93	0.07	0.60	0.14	0.26	0.15	0.26	0.15	N/A	N/A
		3	0.77	0.10	0.45	0.12	0.32	0.12	0.24	0.11	N/A	N/A
	Katakami	2	0.90	0.06	0.64	0.09	0.40	0.10	0.31	0.09	0.26	0.09
		3	0.86	0.06	0.72	0.08	0.52	0.09	0.38	0.09	0.38	0.09
	Pless	2	0.78	0.04	0.55	0.05	0.47	0.05	0.43	0.05	0.41	0.05
		3	0.76	0.04	0.59	0.05	0.51	0.05	0.43	0.05	0.35	0.05
	van Meerbeeck	1	0.70	0.04	0.41	0.04	0.27	0.04	0.18	0.03	0.14	0.03
		2	0.62	0.04	0.35	0.04	0.25	0.03	0.20	0.03	0.16	0.03

Abbreviations: N/A – not applicable, OS – overall survival, P(event) – probability of event occurring, PFS – progression free survival, SE – standard error.

Model Critique

Assessing model fit

The posterior mean of the residual deviance, which measures the magnitude of the differences between the observed data and the model predictions of the data, was used to assess the goodness of fit of each model [12]. Smaller values are preferred, and in a well-fitting model the posterior mean residual deviance should be close to the number of data points in the network (each study arm contributes 1 data point) [12].

In addition to comparing how well the models fit the data using the posterior mean of the residual deviance, models were compared using the deviance information criterion (DIC). This is equal to the sum of the posterior mean deviance and the effective number of parameters, and thus penalizes model fit with model complexity [12]. Lower values are preferred and differences of at least 5 points were considered meaningful [12].

Assessing heterogeneity and inconsistency

Heterogeneity concerns the differences in treatment effects between trials within each treatment contrast, while consistency concerns the differences between the direct and indirect evidence informing the treatment contrasts [9, 10].

Heterogeneity is assessed by comparing the fit of fixed and random effects NMA models. The fixed effect model assumes that all trials are estimating the same treatment effect, regardless of any differences in the conduct of the trials, populations, or treatments. The random effects

NMA model on the other hand accounts for any differences in treatment effects between trials, that are beyond sampling error, by assuming a distribution of study-specific treatment effects with a pooled mean and between-study standard deviation. The estimated between study standard deviation in treatment effects is also inspected to assess heterogeneity.

Inconsistency was assessed by comparing the fit of the chosen consistency model (fixed or random effects) to an “inconsistency”, or unrelated mean effects, model [9, 10]. The latter is equivalent to having separate, unrelated, meta-analyses for every pairwise contrast, with a common variance parameter assumed in the case of random effects models. Note that inconsistency can only be assessed when there are closed loops of direct evidence on 3 treatments that are informed by at least 3 distinct trials [11].

Network meta-analysis: Results of Clinical Evidence Synthesis

5-year Follow-up

Five studies presented survival data up to 5-years, and a network diagram summarizing the evidence is given in Figure 2

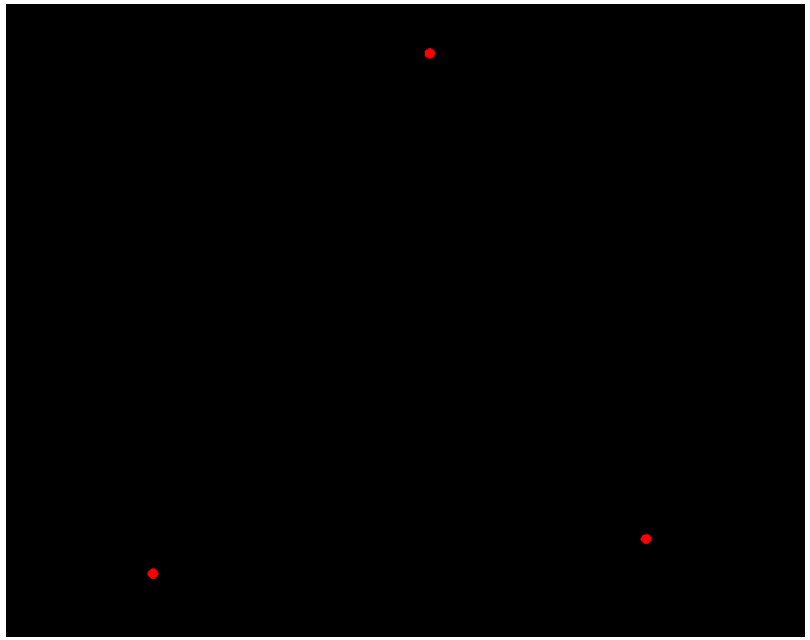


Figure 2: Network diagram of comparisons for which direct evidence on differences in restricted mean survival time up to 5-years is available. Lines are proportional to the number of studies that compare the two connected treatments.

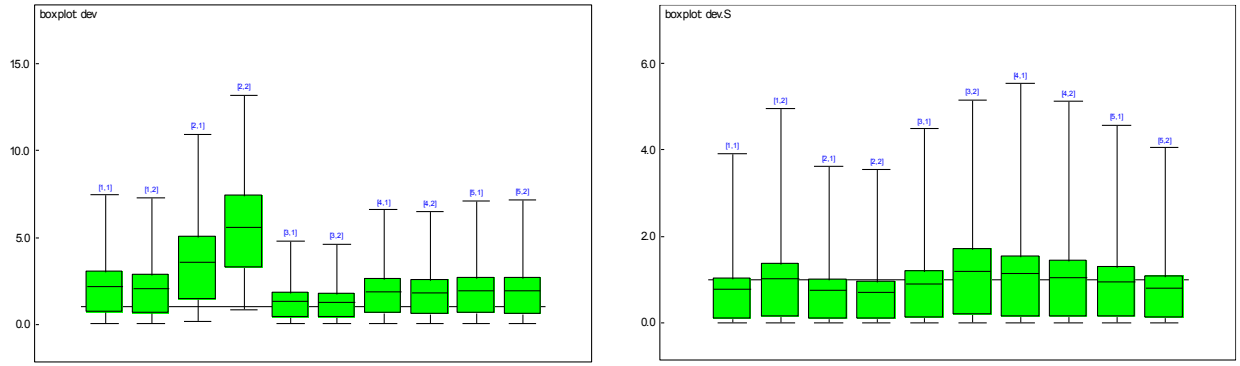
Model fit statistics for the area under the Kaplan Meier curves up to 5-years, as well as the probability of survival are given in Table 12. Convergence was satisfactory for the fixed effect model after a burn-in of 20,000 iterations and results are based on a further 40,000 samples on two chains. For the random effects model, convergence was satisfactory after a burn-in of 30,000 iterations and results are based on a further 60,000 samples on two chains.

Table 12: Model fit statistics based on 5-year follow-up data

Model		**Median Between-Study SD (95% CrI)**	**Posterior mean residual deviance**	**DIC**
Fixed effect	P(Survival)	---	9.267	-24.852
	AUC		23.47	-11.075
Random effects	P(Survival)	0.35 (0.02, 2.41)	9.618	-22.809
	AUC	PFS: 0.18 (0.01, 1.32) PPS: 0.25 (0.03, 1.46)	18.95	-11.781

Total number of data points for P(survival) is 10 and for AUC is 20.

There were no meaningful differences between the fixed and random effects models in terms of the posterior mean residual deviance and DIC for both NMAs (Table 12). The box plots of the posterior deviance values for each study arm in Figure 3 show that the area under the Kaplan Meier curves up to 5 years in Eberhardt 2015 is not predicted well and this study is a possible outlier. Although the prediction of this study improves in the random effects model (Figure 4), this comes at a cost of slight overfit of the model (posterior mean residual deviance = 18.95, compared to 20 datapoints) and additional parameters in the model. In addition, progression events and deaths were rare in the chemoradiotherapy group of this study after 3-years and 4-years, respectively. The simpler fixed effect model was therefore selected in the base-case.



1014 **Figure 3: Posterior deviance values for each study arm for the area under the Kaplan Meier curves (left) and probability of survival**
1015 **(right) – fixed effect model.**

1016

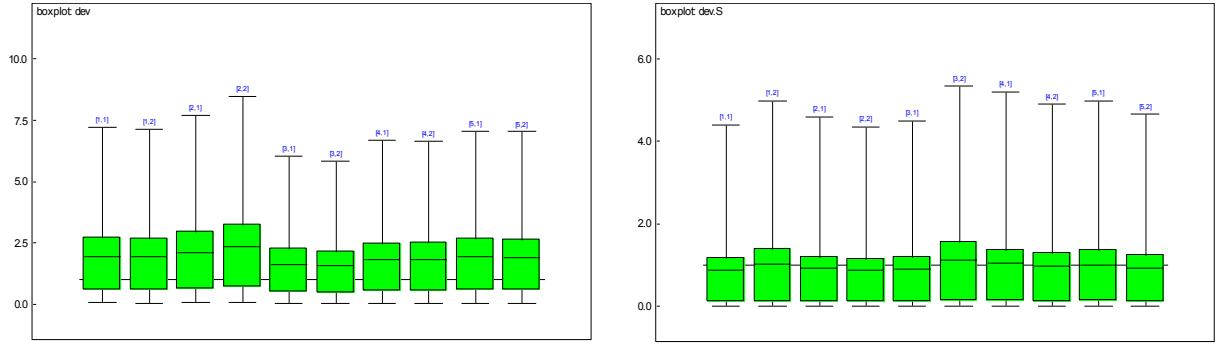


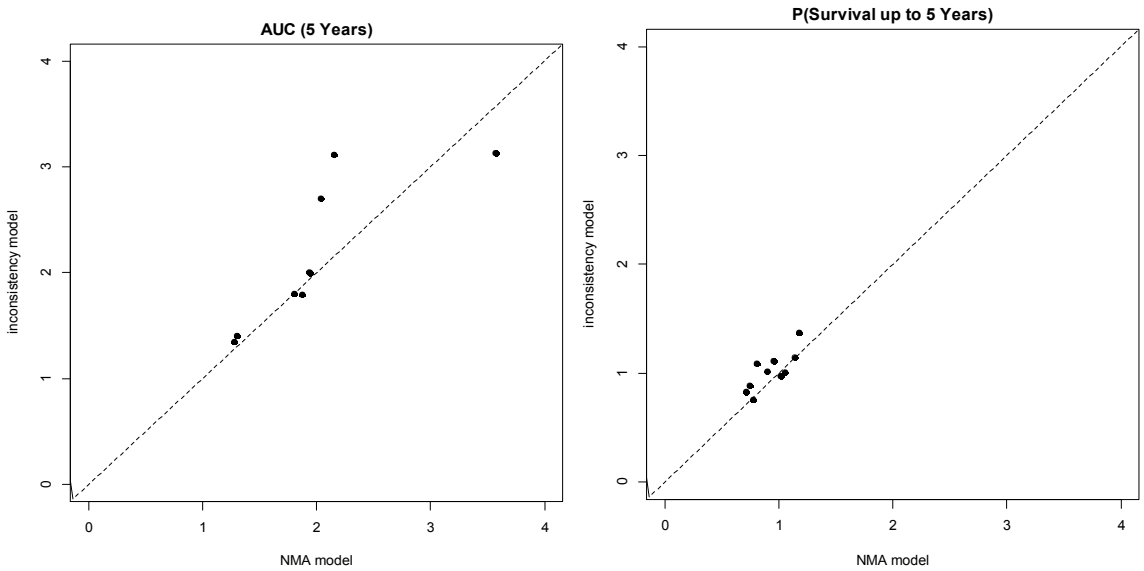
Figure 4: Posterior deviance values for each study arm for the area under the Kaplan Meier curves (left) and probability of survival (right) – random effects model.

No evidence of inconsistency was found, with model fit (posterior mean residual deviance) similar for the consistency and inconsistency (unrelated means) fixed effect models, and a lower DIC for the consistency model (Table 13). The area below the line of equality in Figure 5 highlights where the inconsistency model better predicted data points, and any improvement is minimal.

Table 13: Model fit statistics for consistency and inconsistency fixed effect models based on 5-year follow-up data

Model		Posterior mean residual deviance	DIC
Fixed effect - consistency	P(Survival)	9.267	-24.852
	AUC	23.47	-11.075
Fixed effect - inconsistency	P(Survival)	10.17	-22.867
	AUC	23.65	-8.882

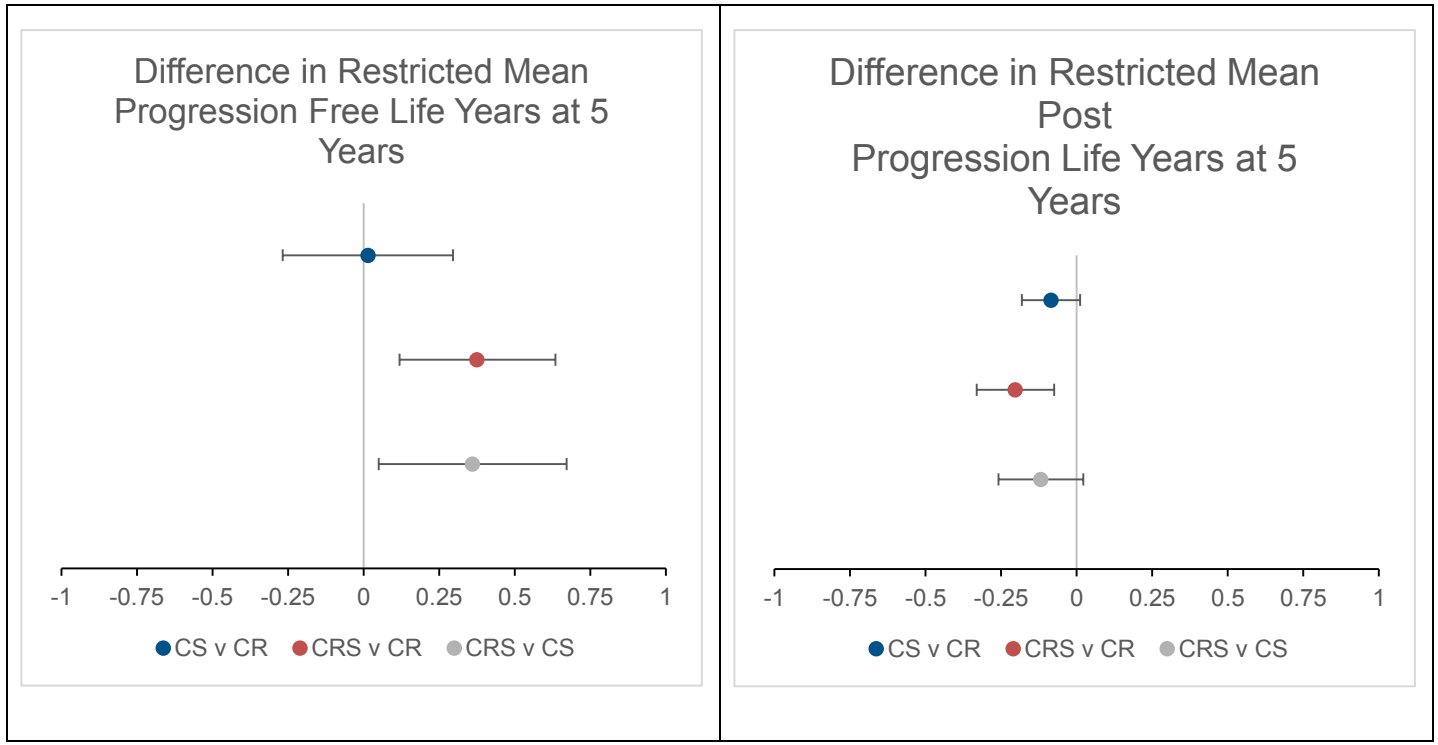
Total number of data points for P(survival) is 10 and for AUC is 20.



1024

1025 **Figure 5: Deviance contributions from the fixed effect consistency and inconsistency models for area under the Kaplan Meier curves**
1026 **(left) and probability of survival (right).**

1027 There is evidence to suggest that chemoradiotherapy + surgery is more effective in increasing progression free life years at 5-year follow-up
1028 compared to chemoradiotherapy alone, while there is no evidence to suggest the effect of chemotherapy + surgery is any different from
1029 chemoradiotherapy (


Difference in Restricted Mean Progression Free Life Years at 5 Years
-1
-0.75
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CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
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CS v CR
CRS v CR
CRS v CS

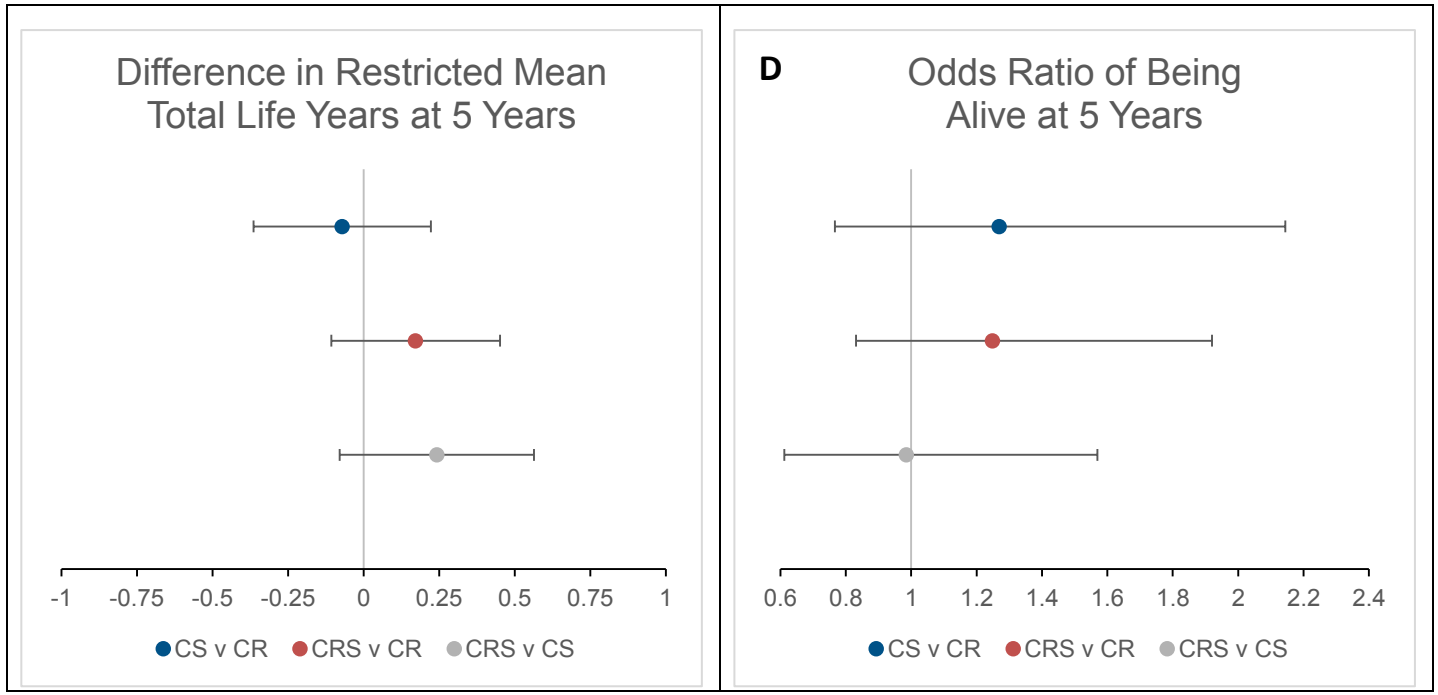
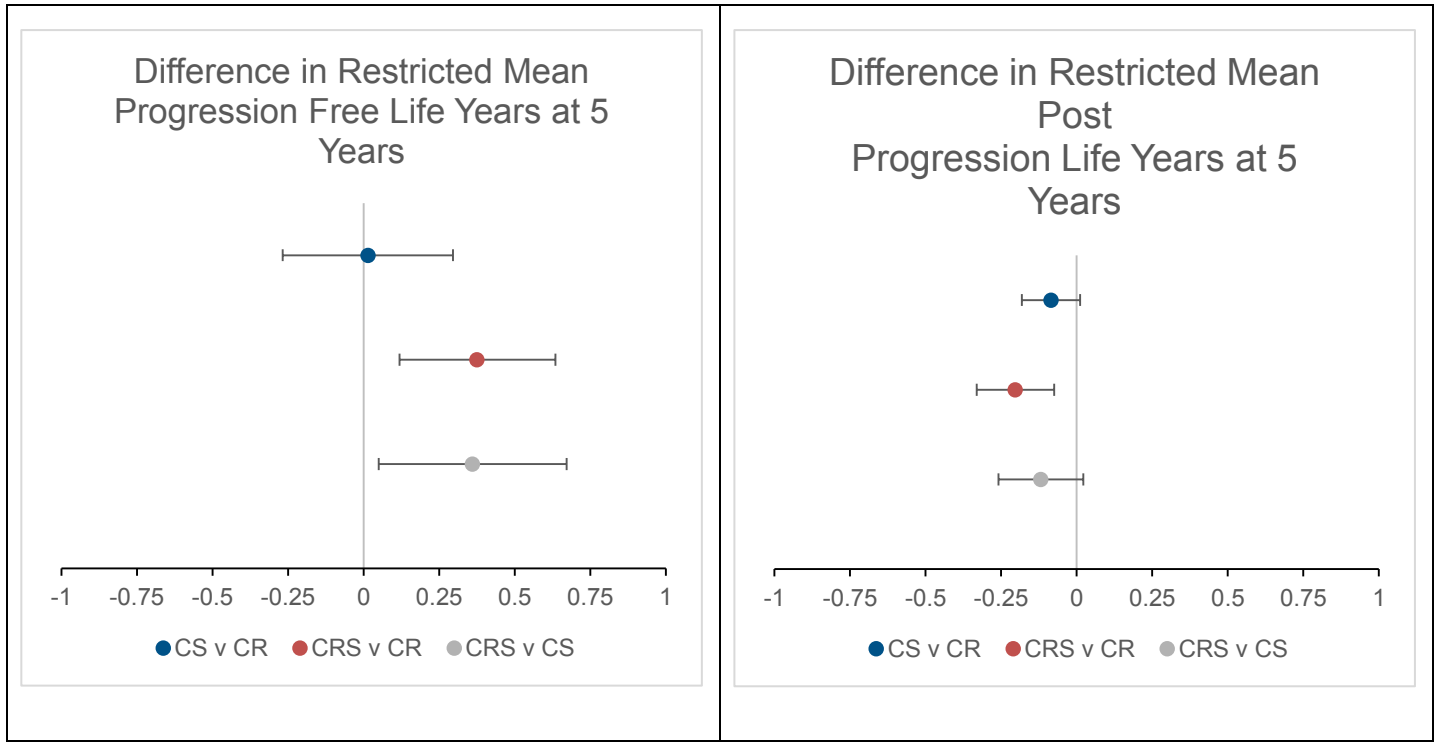


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

A, Table 14). There is also evidence to suggest that chemoradiotherapy + surgery improves progression free life years compared to chemotherapy + surgery (posterior median difference in RMST: 0.36 (95% CrI: 0.05, 0.67)) and it ranked the most effective intervention in increasing progression free life years (Table 14).

In terms of post progression life years at 5-year follow-up, there is evidence suggesting that chemoradiotherapy is more effective than chemoradiotherapy + surgery (


Difference in Restricted Mean Progression Free Life Years at 5 Years
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CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
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CS v CR
CRS v CR
CRS v CS

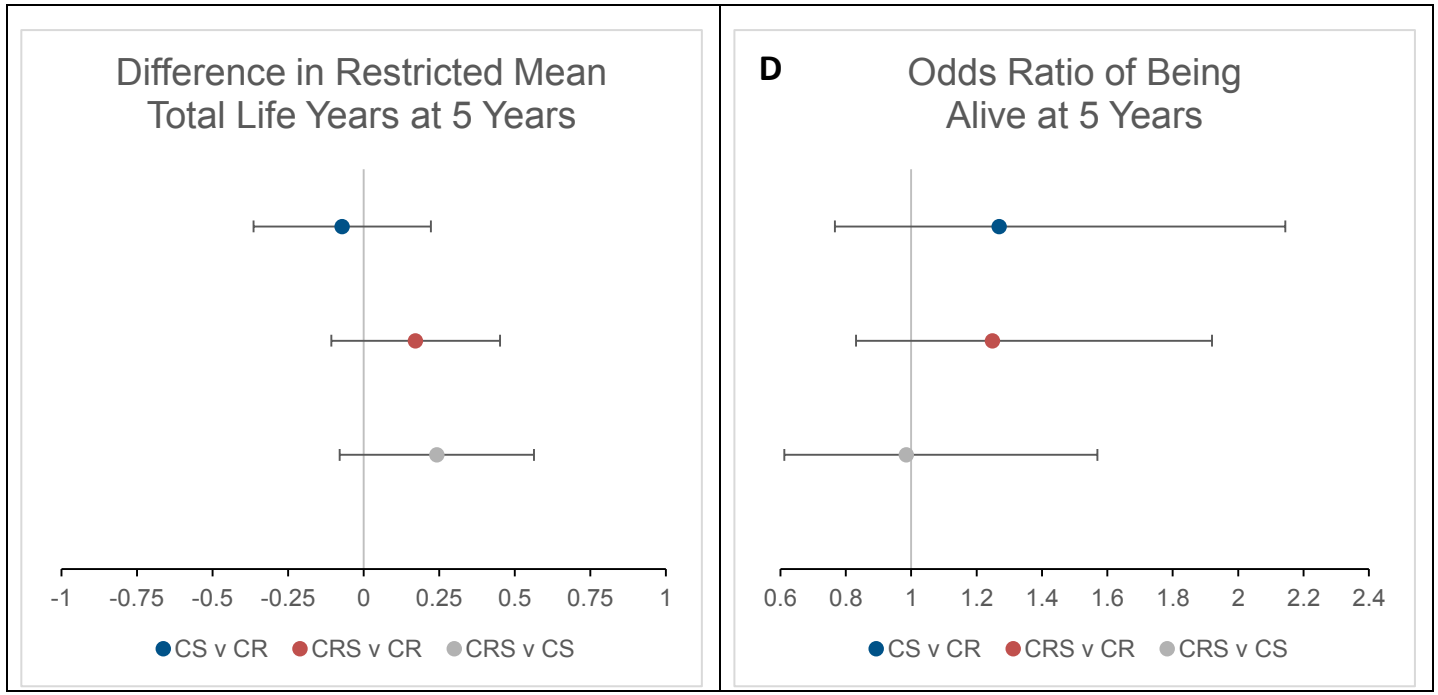
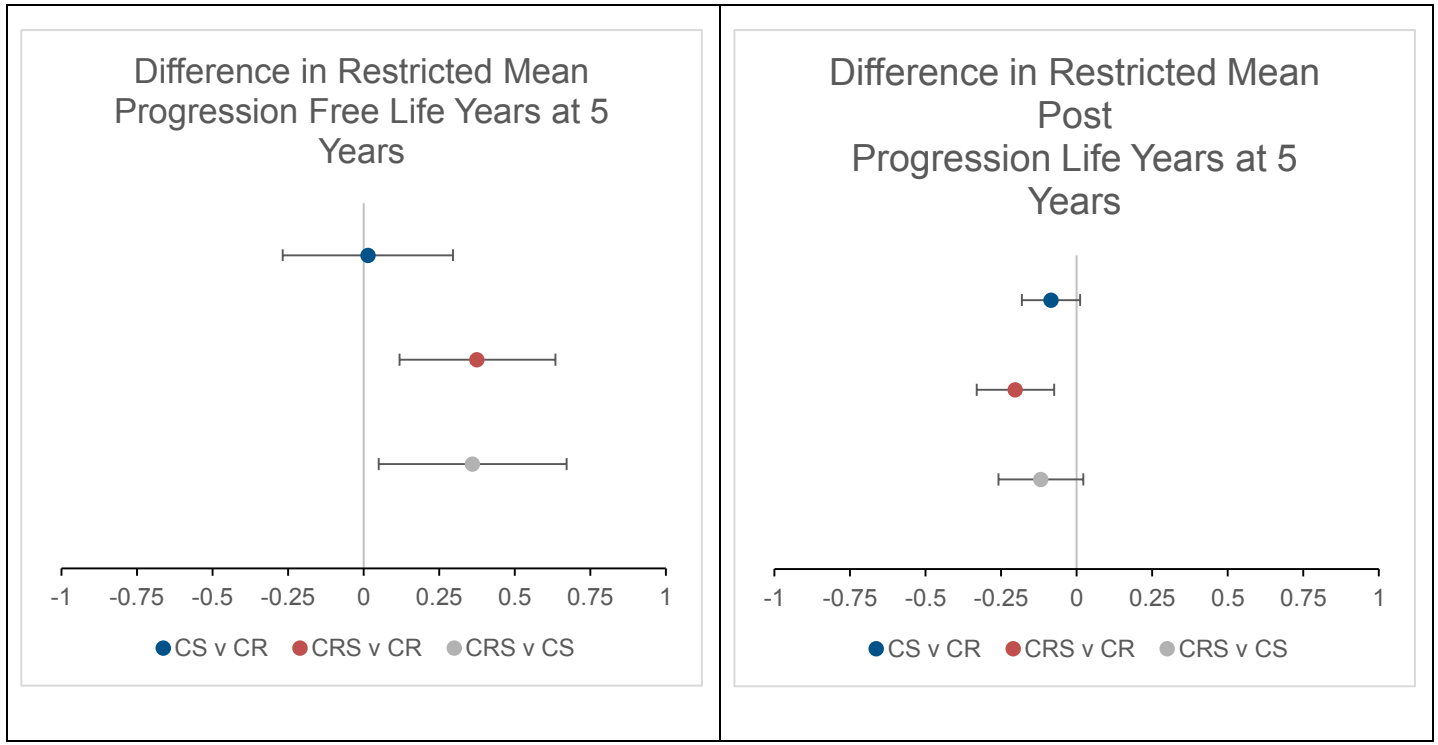


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

B, Table 14). Chemoradiotherapy appears to be more effective than chemotherapy + surgery as well, but this cannot be concluded with high certainty (


Difference in Restricted Mean Progression Free Life Years at 5 Years
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CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
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CS v CR
CRS v CR
CRS v CS

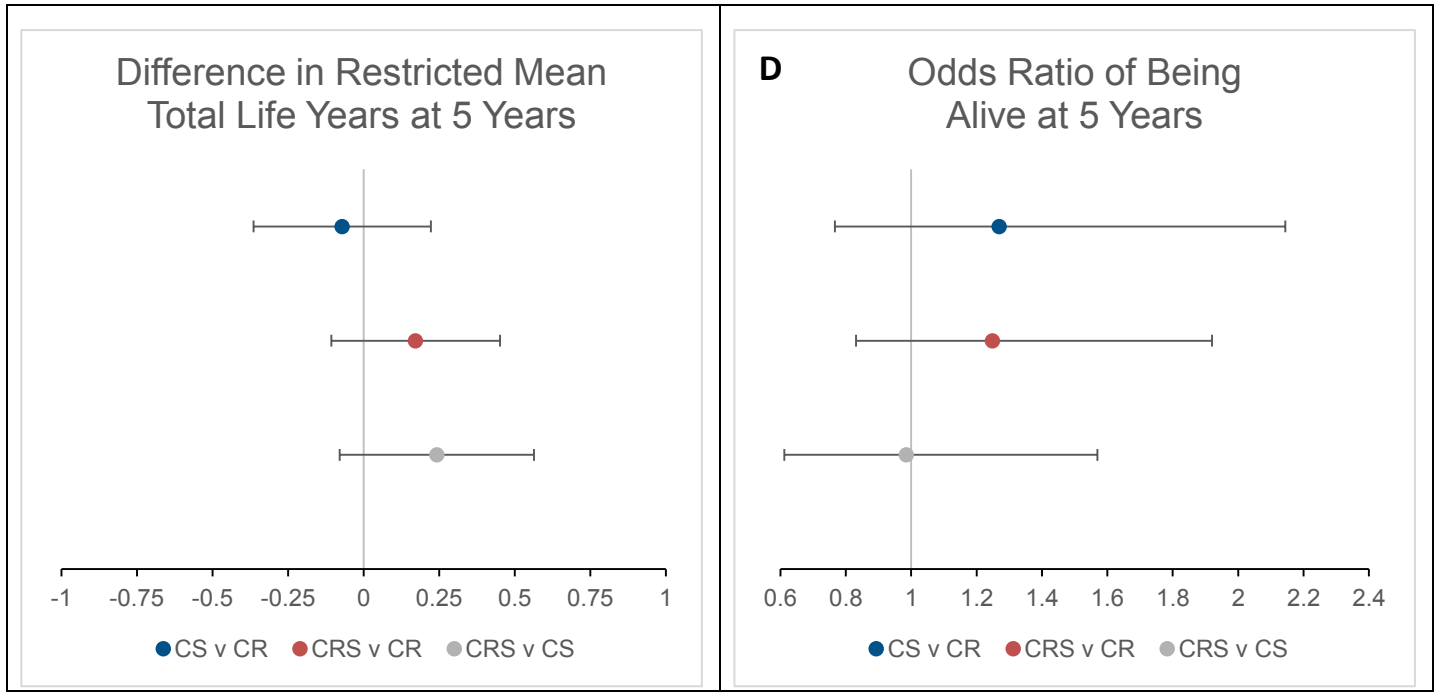
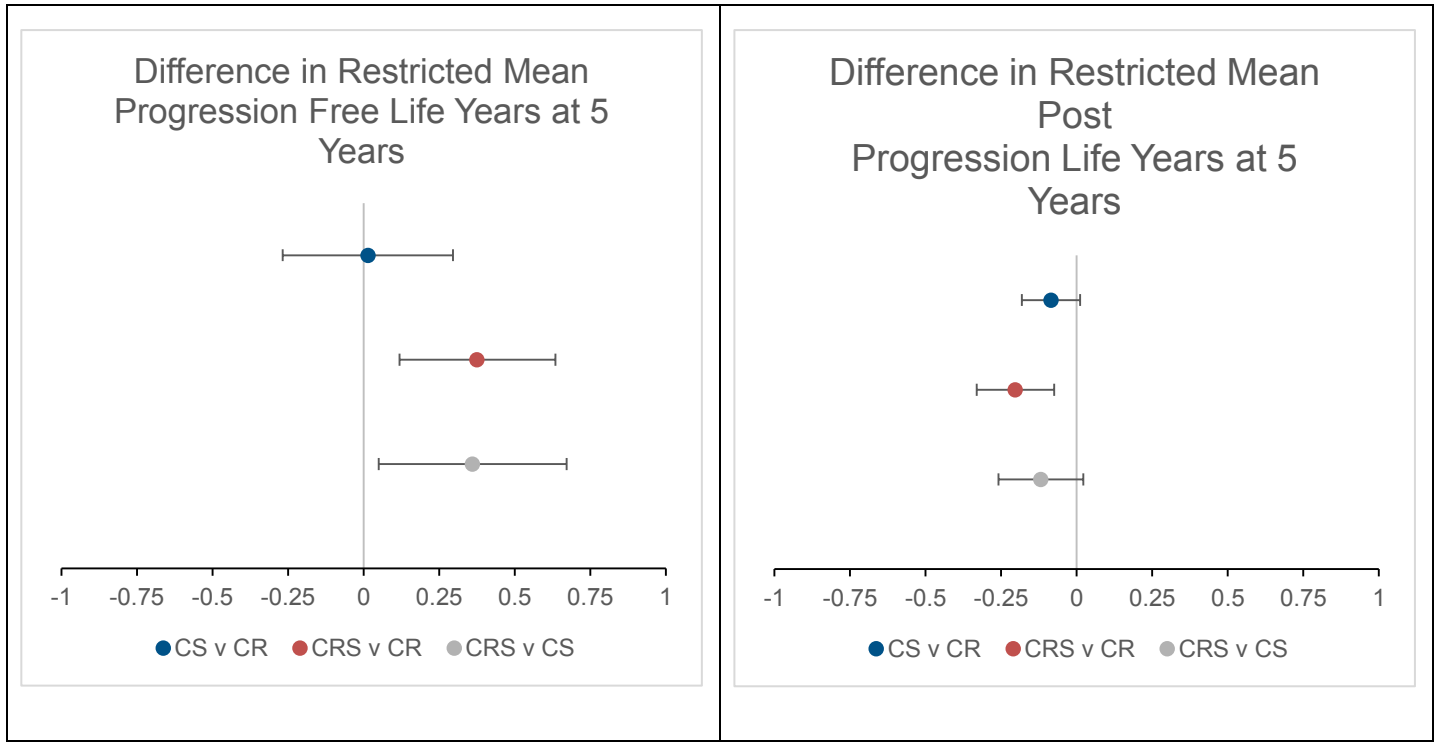


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

B, Table 14). There was not enough evidence to suggest any of the three treatments were different from each other in terms of improving total life years at 5-year follow- up, which is the sum of the progression free and post progression life years (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CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
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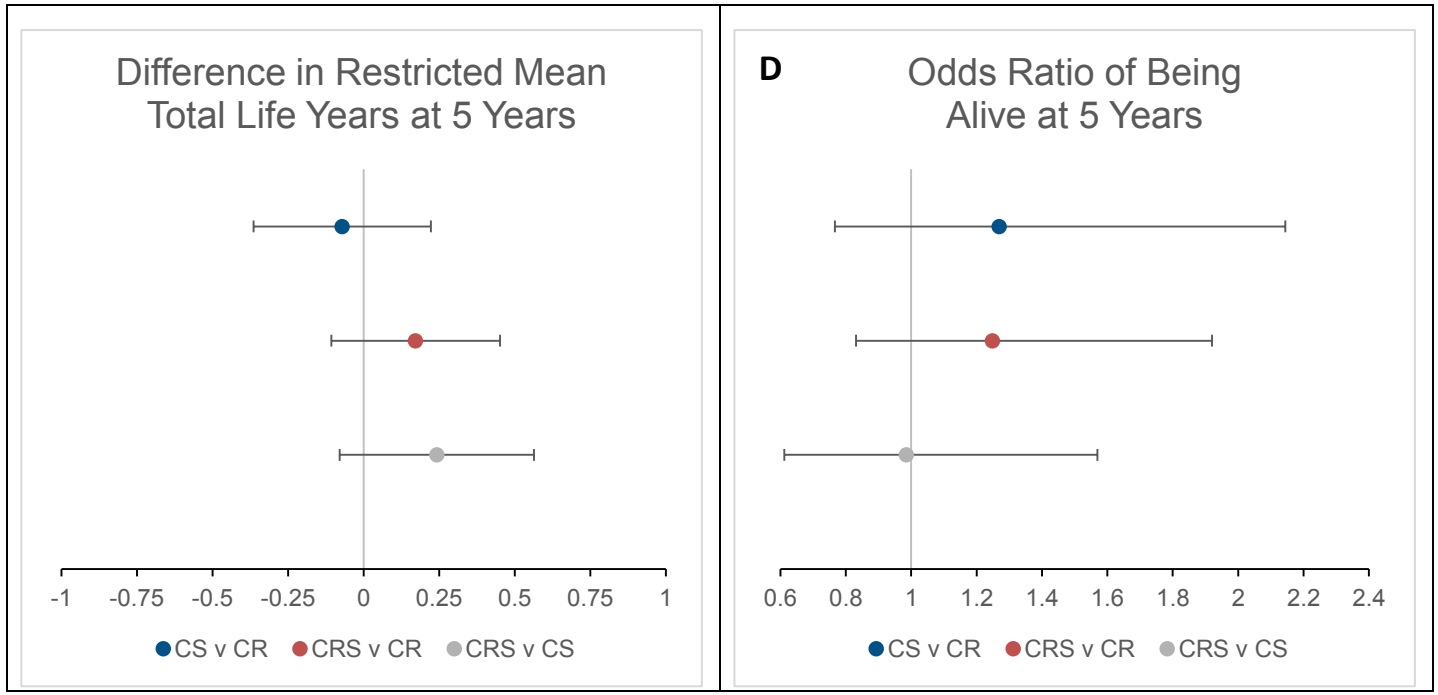
CS v CR
CRS v CR
CRS v CS



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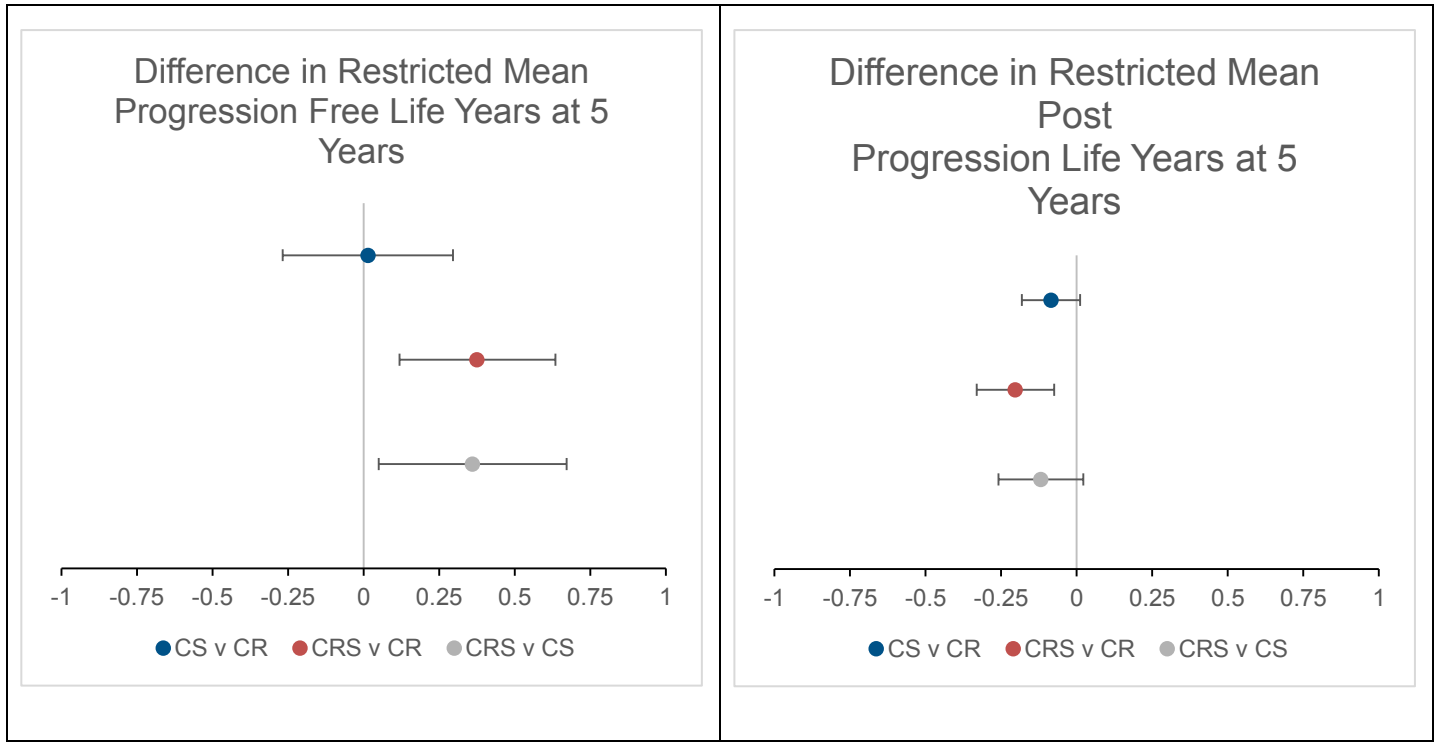
C, Table 14).

Chemotherapy + surgery and chemoradiotherapy + surgery appear to be more likely to improve the odds of being alive at 5-years compared to chemoradiotherapy alone, but there is not enough evidence to infer the direction of effects with certainty (


Difference in Restricted Mean Progression Free Life Years at 5 Years
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CS v CR
CRS v CR
CRS v CS
Difference in Restricted Mean Post Progression Life Years at 5 Years
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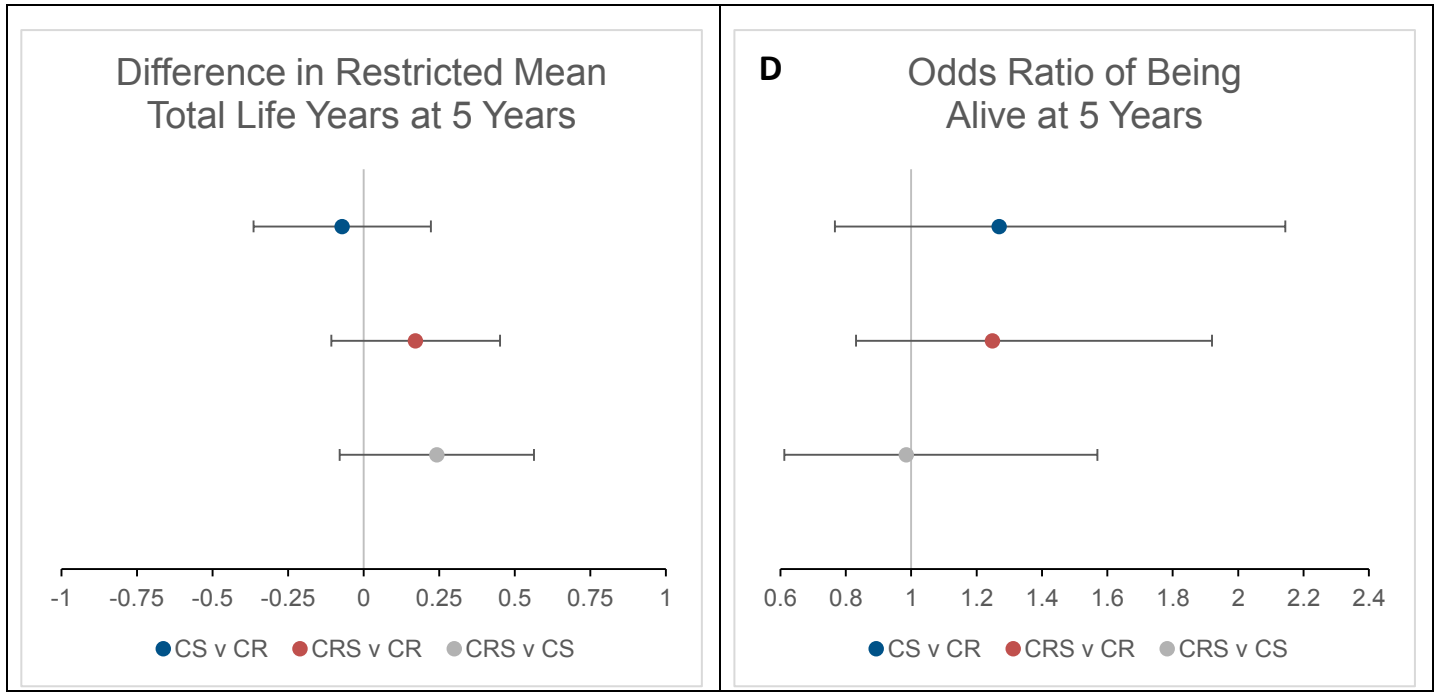
CS v CR
CRS v CR
CRS v CS



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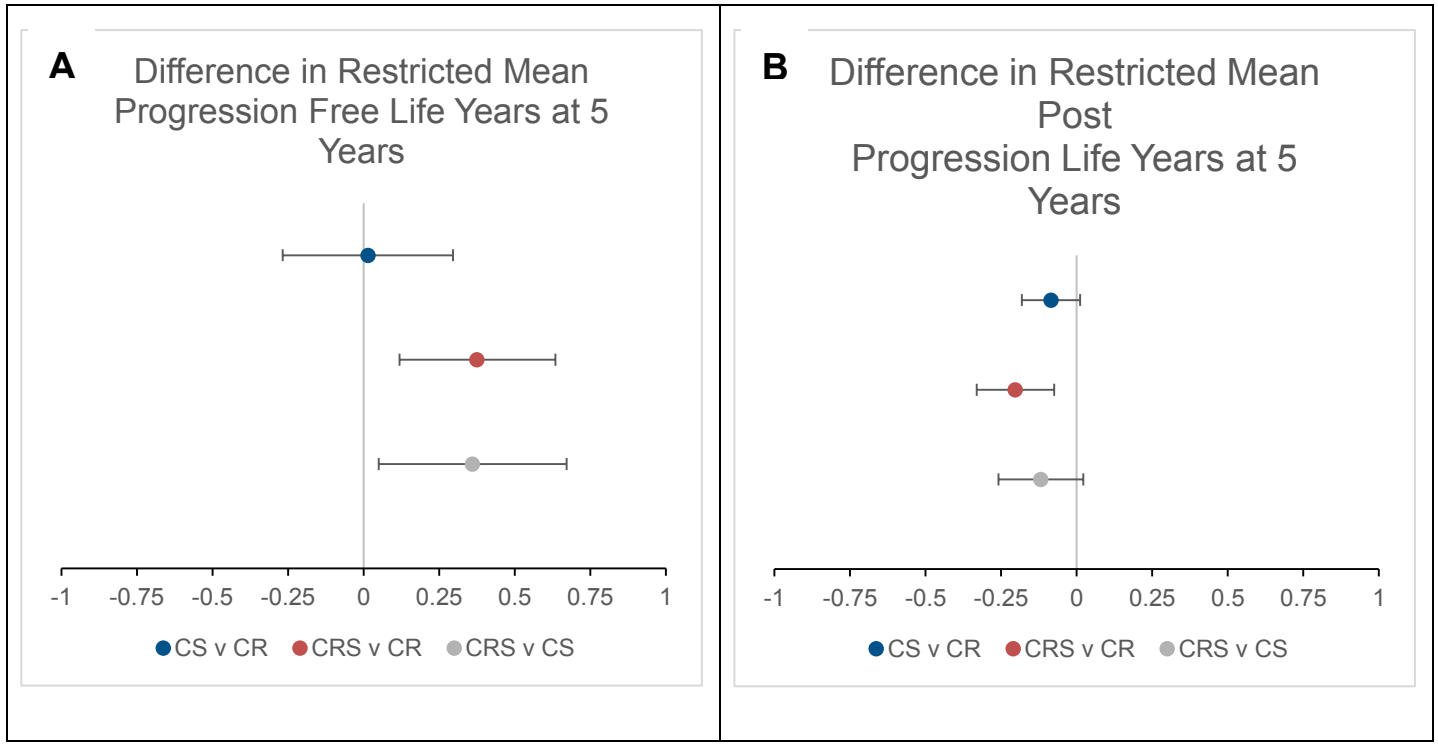
D, Table 14).


A
Difference in Restricted Mean Progression Free Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
0.5
0.75
1
CS v CR
CRS v CR
CRS v CS
B
Difference in Restricted Mean Post Progression Life Years at 5 Years
-1
-0.75
-0.5
-0.25
0
0.25
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1
CS v CR
CRS v CR
CRS v CS

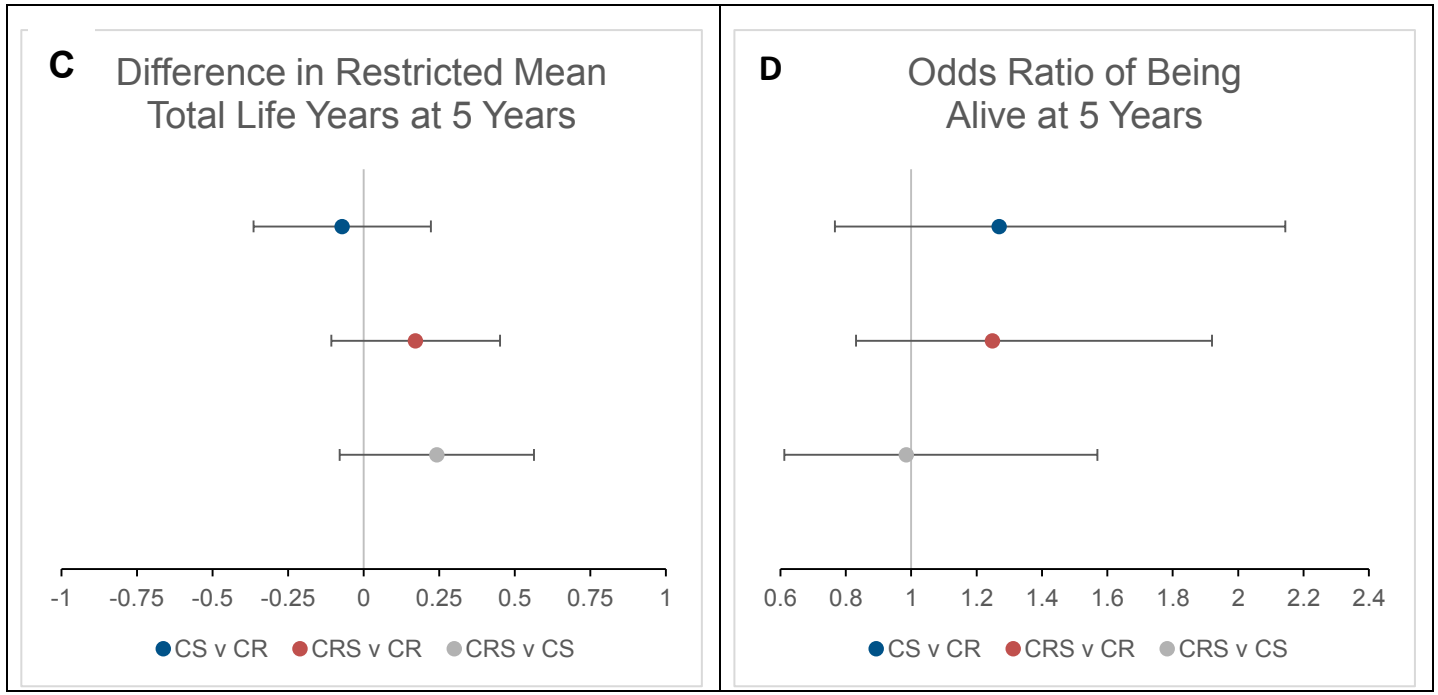


Figure 6: Forest plots of (A) differences in restricted mean progression free life years at 5-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at 5-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at 5-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at 5-years follow-up relative to chemoradiotherapy. Results are presented as the posterior median and 95% credible intervals. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

Table 14: Treatment differences in restricted mean survival times (RMST) up to 5 years, odds ratios of being alive at 5-years, probabilities of ranking best, ranks, and predicted RMST and probability of being alive at 5-years in the UK population for the three interventions.

		Intervention		
		Chemoradiotherapy[a]	Chemotherapy + Surgery	Chemoradiotherapy + Surgery
Difference in RMST (95% CrI[b])	Progression Free Life Years at 5 Years	Reference Treatment	0.01 (-0.27, 0.3)	0.38 (0.12, 0.63)
	Post Progression Life Years at 5 Years		-0.09 (-0.18, 0.01)	-0.2 (-0.33, -0.07)
	Total Life Years at 5 Years		-0.07 (-0.36, 0.22)	0.17 (-0.11, 0.45)
Odds Ratio (95% CrI)	Being Alive at 5 Years		1.27 (0.77, 2.14)	1.25 (0.83, 1.92)
Probability of Ranking Best	Progression Free Life Years at 5 Years	0.2%	1.1%	98.7%
	Post Progression Life Years at 5 Years	95.8%	4.1%	0.1%
	Total Life Years at 5 Years	9.9%	5.4%	84.7%
	Being Alive at 5 Years	6.3%	50.2%	43.6%
Median Rank (95% CrI)	Progression Free Life Years at 5 Years	3 (2, 3)	2 (2, 3)	1 (1, 1)

	Post Progression Life Years at 5 Years	1 (1, 2)	2 (1, 3)	3 (2, 3)
	Total Life Years at 5 Years	3 (1, 3)	2 (1, 3)	1 (1, 3)
	Being Alive at 5 Years	3 (1, 3)	1 (1, 3)	2 (1, 3)
Predicted RMST and Probability of Being Alive in UK at 5 Years[c]	Mean Progression Free Life Years	1.5 (1.28, 1.71)	1.51 (1.29, 1.73)	1.87 (1.57, 2.17)
	Mean Post Progression Life Years	0.58 (0.51, 0.65)	0.49 (0.42, 0.56)	0.37 (0.24, 0.51)
	Mean Total Life Years	2.07 (1.85, 2.29)	2 (1.77, 2.23)	2.24 (1.93, 2.56)
	Probability of Being Alive at 5 Years	0.13 (0.08, 0.18)	0.16 (0.11, 0.21)	0.16 (0.1, 0.23)

[a] Relative treatment effects presented for comparisons versus chemoradiotherapy. Point estimates are based on posterior medians.
[b] CrI = Credible Interval
[c] Baseline based on posterior distributions of outcomes for van Meerbeeck 2007.

Sensitivity analyses

As part of an assessment of the sensitivity of the results to the selected follow-up time, we also synthesised data based on a shorter follow-up period of 4-years, which allowed the inclusion of all 6 studies, including Girard 2009. Model fit statistics for the fixed and random effects models based on the 4-year follow-up data are given in Table 15Table 15. Convergence was satisfactory for the both models after a burn-in of 20,000 iterations and results are based on a further 40,000 samples on two chains.

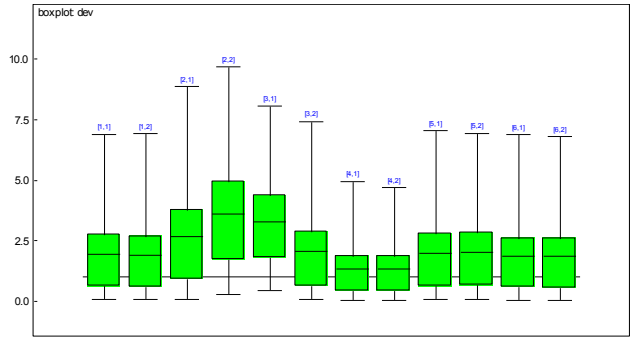
Table 15: Model fit statistics based on 4-year follow-up data

Model			DIC

		Posterior Median Between-Study SD (95% CrI)	Posterior mean residual deviance	
Fixed effect	P(Survival)	---	13.22	-27.429
	AUC		25.84	-20.356
Random effects	P(Survival)	0.24 (0.02, 1.63)	14.29	-25.090
	AUC	PFS: 0.12 (0.01, 0.76) PPS: 0.14 (0.01, 0.59)	23.61	-18.623

Total number of data points for P(survival) is 12 and for AUC is 24.

There were no meaningful differences between the fixed and random effects models in terms of the posterior mean residual deviance and DIC (Table 15). The plots of the posterior deviance values for each study arm in Figure 7 show that the probability of survival up to 4 years in Girard 2009 is not predicted well and this study is a possible outlier. Fitting a random effects model did not help in the prediction of data points in this study (Figure 8). The simpler fixed effect model is therefore preferred.



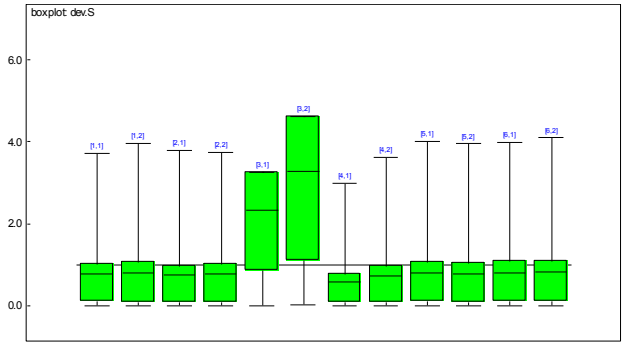


Figure 7: Posterior deviance values for each study arm for the area under the Kaplan Meier curves (left) and probability of survival (right) – fixed effect model.

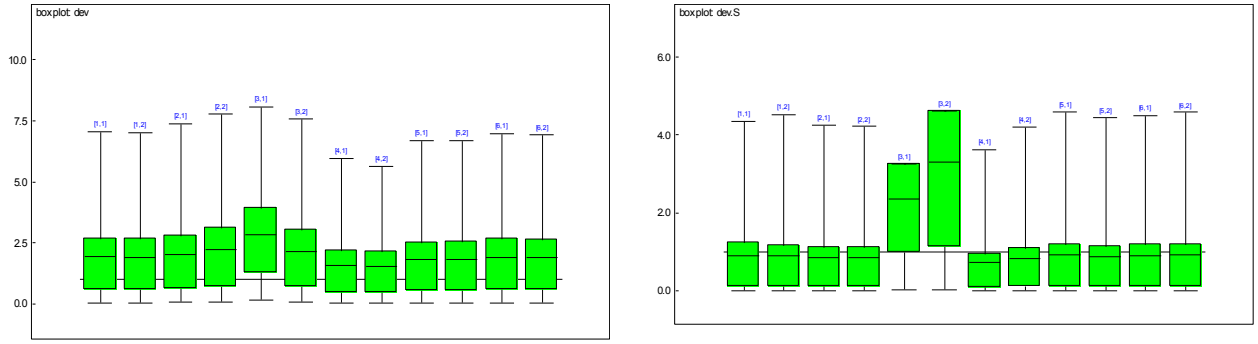


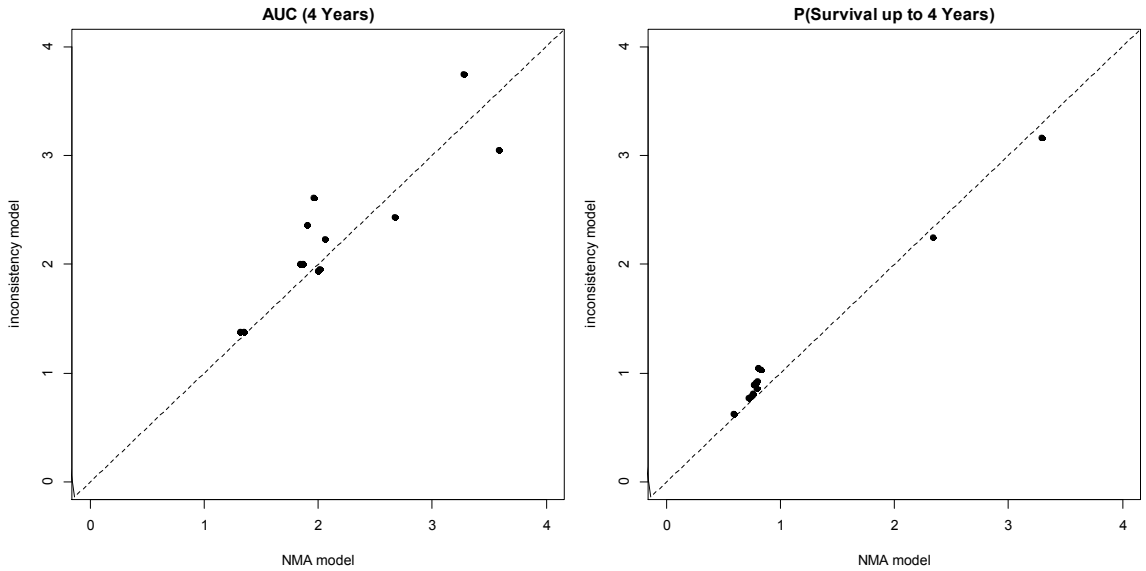
Figure 8: Posterior deviance values for each study arm for the area under the Kaplan Meier curves (left) and probability of survival (right) – random effects model.

No evidence of inconsistency was found through comparison of the consistency and inconsistency random effects models, as little difference was observed between the fit of the models (Table 16). The area below the line of equality in Figure 9 highlights where the inconsistency model better predicted data points, but any improvements were minimal.

Table 16: Model fit statistics for consistency and inconsistency fixed effect models based on 4-year follow-up data

Model		Posterior mean residual deviance	DIC
Fixed effect - consistency	P(Survival)	13.22	-27.429
	AUC	25.84	-20.356
Fixed effect - inconsistency	P(Survival)	14.07	-25.773
	AUC	27.07	-17.115

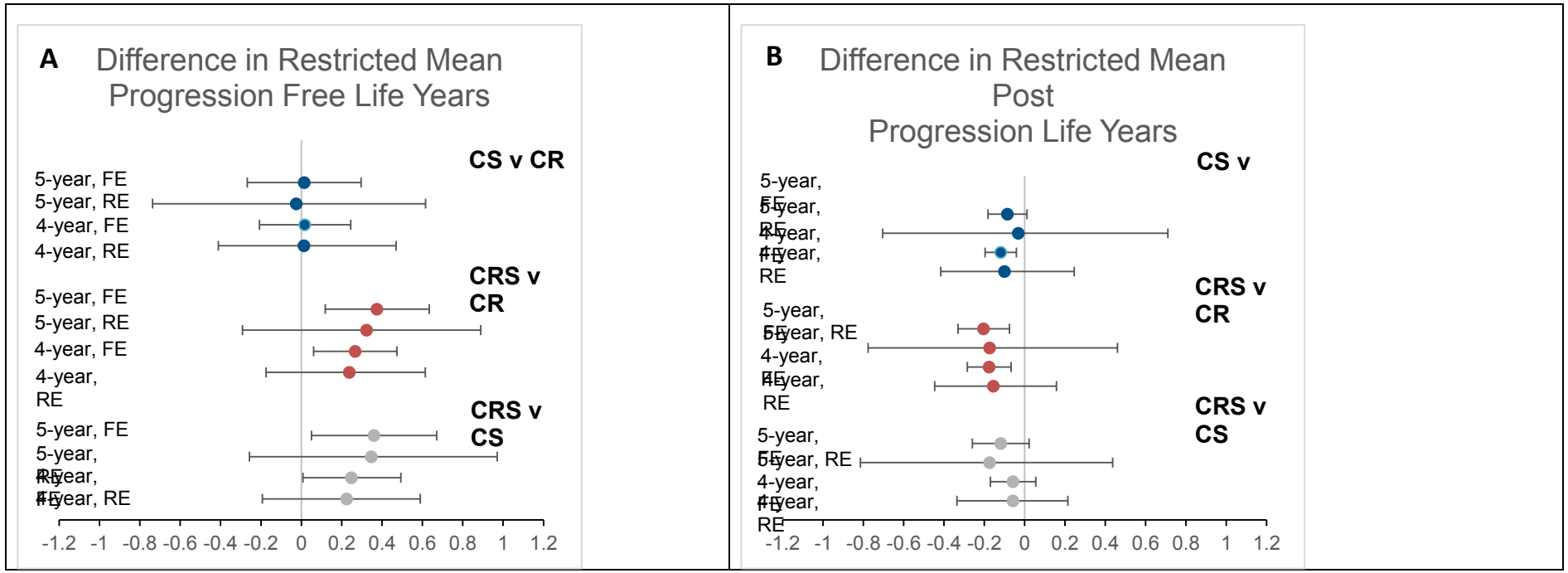
Total number of data points for P(survival) is 12 and for AUC is 24.



1108

1109 **Figure 9: Deviance contributions from the fixed effect consistency and inconsistency models for area under the Kaplan Meier curves**
1110 **(left) and probability of survival (right).**

1111 Treatment effects estimated by the fixed and random effects models based on the 4- and 5-year follow up data are presented in


A
Difference in Restricted Mean Progression Free Life Years
CS v CR
5-year, FE
5-year, RE
4-year, FE
4-year, RE
CRS v CR
5-year, FE
5-year, RE
4-year, FE
4-year, RE
CRS v CS
5-year, FE
5-year, RE
4-year, FE
4-year, RE
-1.2 -1 -0.8 -0.6 -0.4 -0.2 0 0.2 0.4 0.6 0.8 1 1.2
B
Difference in Restricted Mean Post Progression Life Years
CS v
CRS v CR
CRS v CS
-1.2 -1 -0.8 -0.6 -0.4 -0.2 0 0.2 0.4 0.6 0.8 1 1.2

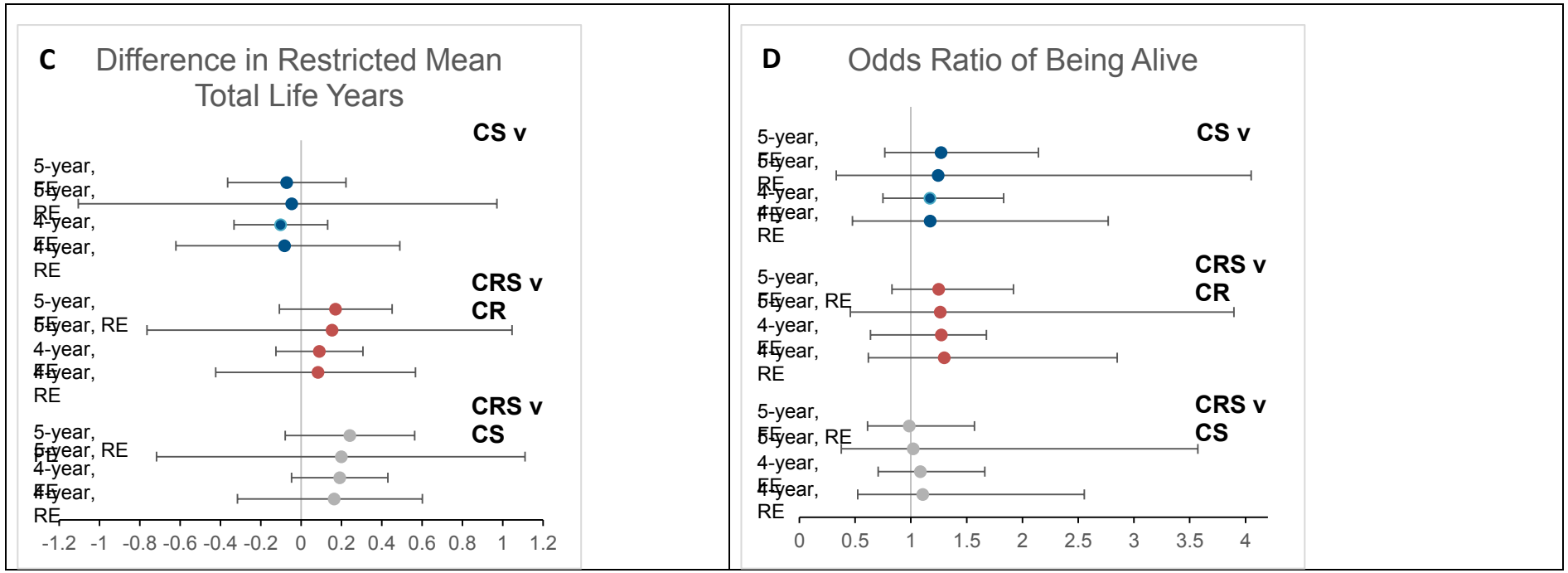
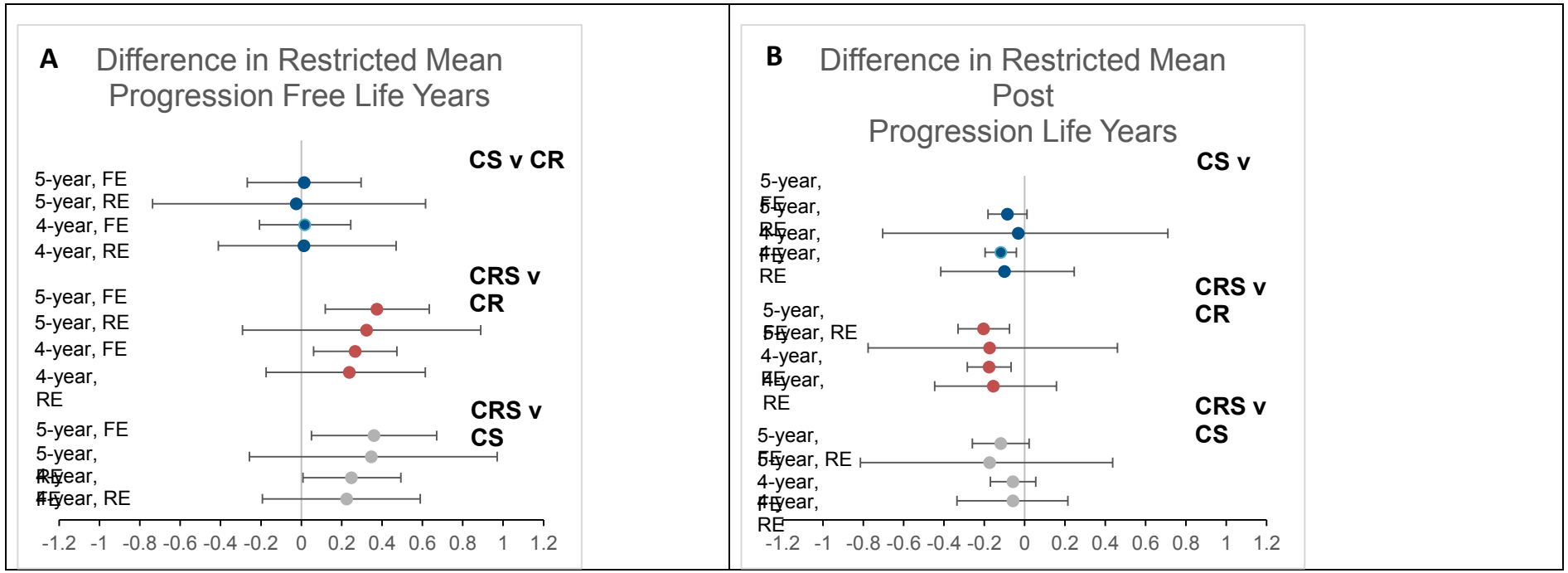


Figure 10. The point estimates of the treatment effects are similar, and the width of the credible intervals reflect that random effects models estimate the treatment effects with more uncertainty, and that there is additional data included in the 4-dataset compared with the 5-year dataset.

Noting that

7. the model fit assessment supports the use of the fixed effect model in both datasets,
8. the assumption that non-progressors by *T*-years do not progress (are "cured") is more reasonable at 5-years than at 4-years,
9. the 5-year dataset excludes the Girard (2009) study, which seems to be an outlier and is based on small numbers

supports the use of the fixed effect model based on the 5-year dataset for the base-case. Results from the random effects model based on the 5-year dataset are presented as a sensitivity analysis.


A
Difference in Restricted Mean Progression Free Life Years
CS v CR
5-year, FE
5-year, RE
4-year, FE
4-year, RE
CRS v CR
5-year, FE
5-year, RE
4-year, FE
4-year, RE
CRS v CS
5-year, FE
5-year, RE
4-year, FE
4-year, RE
-1.2 -1 -0.8 -0.6 -0.4 -0.2 0 0.2 0.4 0.6 0.8 1 1.2
B
Difference in Restricted Mean Post Progression Life Years
CS v
CRS v CR
CRS v CS
-1.2 -1 -0.8 -0.6 -0.4 -0.2 0 0.2 0.4 0.6 0.8 1 1.2

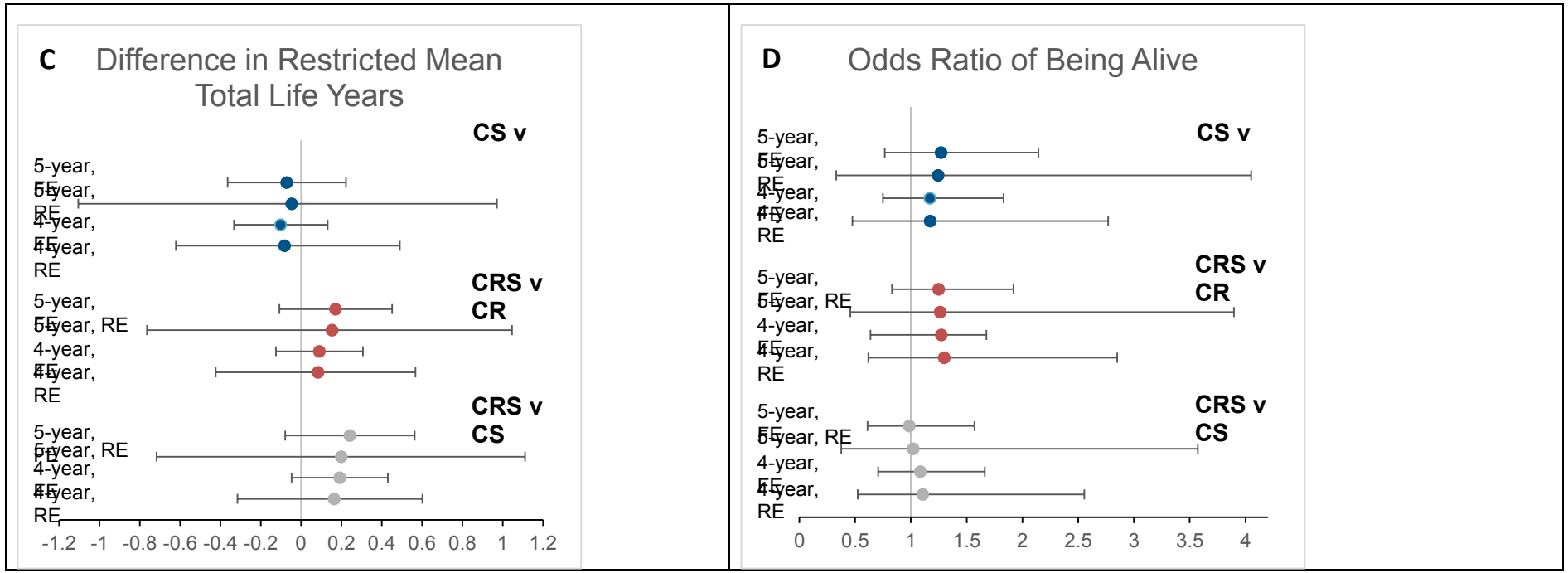


Figure 10: Forest plots of fixed and random effects estimates at 5- and 4-year follow up for (A) differences in restricted mean progression free life years at T-years follow-up relative to chemoradiotherapy, (B) differences in restricted mean post progression life years at T-years follow-up relative to chemoradiotherapy, (C) differences in restricted mean total life years at T-years follow-up relative to chemoradiotherapy, and (D) odds ratios of being alive at T-years follow-up relative to

chemoradiotherapy. Abbreviations: CR – chemoradiotherapy, CS – chemotherapy + surgery, CRS – chemoradiotherapy + surgery.

Results: Inputs for Economic Model

Discounted Area Under the Kaplan Meier Curves and Probability of Survival

The fit of the NMA models based on the discounted AUC was also assessed and were in line with the results presented in Section 0 For both the 4-year and 5-year follow-up data, there were no meaningful differences between the fit of the fixed and random effects models (Table 17), and thus the fixed effect model was preferred.

Table 17: Model fit statistics based on 5-year follow-up data, discounted at 3.5% annual rate

Follow-Up Period	Model		Posterior Median Between-Study SD (95% CrI)	Posterior mean residual deviance	DIC
5 years [a]	Fixed effect[c]	P(Survival)	---	9.27	-24.85
		AUC		23.18	-14.69
	Random effects[d]	P(Survival)	0.33 (0.01, 2.34)	9.57	-22.94
		AUC	PFS: 0.17 (0.01, 1.25) PPS: 0.23 (0.03, 1.29)	18.86	-15.24
4 years [b]	Fixed effect[c]	P(Survival)	---	13.35	-27.18
		AUC		24.86	-23.87
	Random effects[e]	P(Survival)	0.22 (0.01, 1.56)	14.31	-25.08
		AUC	PFS: 0.11 (0.00, 0.68) PPS: 0.12 (0.01, 0.54)	23.34	-21.59

[a] Total posterior mean residual deviance compared to total number of data points for P(survival): 10 and AUC: 20
[b] Total posterior mean residual deviance compared to total number of data points for P(survival): 12 and AUC: 24
[c] Burn-in: 20,000 iterations, results based on: 40,000 samples, 2 chains
[d] Burn-in: 50,000 iterations, results based on: 100,000 samples, 2 chains
[e] Burn-in: 30,000 iterations, results based on: 60,000 samples, 2 chains

Similarly, the fit of the consistency and inconsistency models for both 4- and 5-year follow-up data were compared (Table 18). There is no evidence of inconsistency as no meaningful differences were found in the fit of the models for both datasets. The area below the line of equality in Figure 11 and Figure 12 highlights where the inconsistency model better predicted data points, but any improvements were minimal.

Table 18: Model fit statistics for consistency and inconsistency fixed effect models based on 4-year follow-up data, discounted at 3.5% annual rate

Follow-Up Period	Model[c]		Posterior mean residual deviance	DIC
5 years [a]	Fixed effect - consistency	P(Survival)	9.27	-24.85
		AUC	23.18	-14.69
	Fixed effect – inconsistency	P(Survival)	10.17	-22.87
		AUC	23.43	-12.42
4 years [b]	Fixed effect – consistency	P(Survival)	13.35	-27.18
		AUC	24.86	-23.87
	Random effects - inconsistency	P(Survival)	14.15	-25.62
		AUC	26.12	-20.59

[a] Total posterior mean residual deviance compared to total number of data points for P(survival): 10 and AUC: 20
[b] Total posterior mean residual deviance compared to total number of data points for P(survival): 12 and AUC: 24
[c] Burn-in: 20,000 iterations, results based on: 40,000 samples, 2 chains

Figure 11: Deviance contributions from the fixed effect consistency and inconsistency models for area under the Kaplan Meier curves discounted at 3.5% annual rate (left) and probability of survival (right).

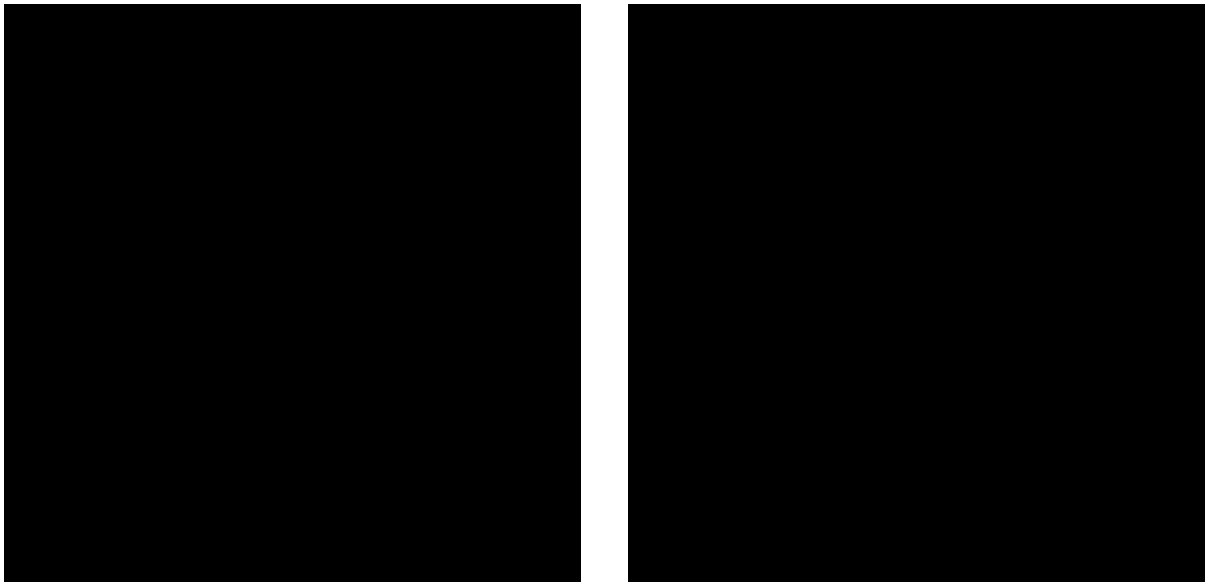


Figure 12: Deviance contributions from the fixed effect consistency and inconsistency models for area under the Kaplan Meier curves discounted at 3.5% annual rate (left) and probability of survival (right).

Proportion of Events Occurring each Year

The proportion of events occurring each year pooled across studies is given in Table 19. The estimated proportions are similar across the 5-year and 4-year follow-up datasets.

Table 19: Pooled proportion of events occurring each year

Follow-Up Period	Event Type	Year	Median Proportion of Events (95% CrI)
5-year	PFS[a]	1	0.63 (0.59, 0.67)
		2	0.23 (0.19, 0.28)
		3	0.08 (0.03, 0.13)
		4	0.04 (0.00, 0.09)
		5	0.01 (0.00, 0.07)
	OS[b]	1	0.38 (0.34, 0.42)
		2	0.32 (0.27, 0.38)
		3	0.16 (0.10, 0.22)
		4	0.11 (0.04, 0.17)
		5	0.03 (0.00, 0.10)
4-year	PFS[c]	1	0.65 (0.61, 0.69)
		2	0.24 (0.19, 0.30)
		3	0.09 (0.00, 0.14)
		4	0.01 (0.00, 0.08)
	OS[c]	1	0.39 (0.35, 0.43)
		2	0.35 (0.29, 0.41)
		3	0.17 (0.11, 0.23)
		4	0.10 (0.00, 0.15)

[a] Burn-in: 500,000 iterations, results based on: 1,000,000 samples, 2 chains
[b] Burn-in: 2,000,000 iterations, results based on: 4,000,000 samples, 2 chains
[c] Burn-in: 100,000 iterations, results based on: 100,000 samples, 2 chains

NMA for Adverse Events

The base case approach used in the economic model for adverse events used pairwise meta-analyses but data then became available that allowed us to fit an NMA for use in sensitivity analyses.

The studies had reported adverse events heterogeneously; in some studies the reporting was comprehensive and in others scant or no details were available. Furthermore, events were classified heterogeneously across studies, being grouped under narrow or broad classes that made event-specific pooling difficult. The committee decided that adverse events should be included in the economic model if possible and we agreed an aggregate approach with them. This involved grouping all adverse events of grade 3+ as homogenously requiring one hospital admission, but having no long term clinical effects or detriment to quality of life. The committee thought it possible that grade 4 adverse events would affect quality of life but these occurred to sparsely to be meaningfully included in the model. Because of the wide disparity between the frequency of adverse events reported among the studies, we selected Pless 2015, Eberhardt 2015, Albain 2009 and van Meerbeeck 2007 for the analysis. These studies were the largest and best conducted studies in the network and had reported event rates that the committee found credible. The data from van Meerbeeck was not reported in the published paper but provided to us upon request by the EORTC, who hold the trial data. We obtained the person years at risk by multiplying the total number of patients in each arm by the mean AUC for total life years at 5 years. The data are in Table 20.

Table 20: Adverse Event NMA Input Data

Treatment Arm 1	Events Arm 1	TatRisk Arm 1	Treatment Arm 2	Events Arm 2	TatRisk Arm 2	Study		Treatments
2	182	285.2	3	141	299.52	Pless 2015		1=CR
3	482	434.3	1	608	409.34	Albain 2009		2=CS
1	137	214.4	3	150	230.04	Eberhardt 2015		3=CRS
1	98	321.75	2	108	298.93	van Meerbeerck 2007		

We assumed that adverse events were treatment related and therefore that it was appropriate to assume a homogenous follow-up time. Since this meant that we did not have to account for variable study endpoints in our pooling of the data, we selected a poisson likelihood, log link NMA model and copied the code directly from NICE TSD2 (citation). The results of the fixed and random effects models are in Table 21. Models were run using 50,000 burn-in iterations and 50,000 iterations to generate the posterior distributions.

Table 21: Adverse Event NMA Results

All Adverse Events	estimate	LCL	UCL	DIC
Fixed effects				74.44
HR of CS vs CR	1.132	0.9382	1.354	
HR of CR vs CRS	1.2425447	1.125112511	1.377221	
HR of CS vs CRS	1.3970383	1.174950065	1.67336	

Random effects				72.627
HR of CR vs CS	1.166	0.3146	4.654	
HR of CR vs CRS	1.176886	0.374531835	3.354579	
HR of CS vs CRS	1.3696754	0.361663653	5.186722	

The DIC for the random effects model was not more than 3-5 points lower than the fixed effects model so we preferred it in the base case. The results show that both CR and CS are associated with more adverse events than CRS.

As discussed in the economic modelling report (Appendix J), the NMA data agreed well with the pairwise estimates of adverse events.

References and Code

References

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10. Dias, S., et al., *Evidence Synthesis for Decision Making 4: Inconsistency in networks of evidence based on randomized controlled trials.* Medical Decision Making, 2013. **33**: p. 641-656.
11. van Valkenhoef, G., et al., *Automated generation of node-splitting models for assessment of inconsistency in network meta-analysis.* Research Synthesis Methods, 2016. **7**: p. 80-93.
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Code

SEER dataset

Selection criteria:

{Age at Diagnosis.Age recode with <1 year olds} = '30-34 years','35-39 years','40-44 years','45-49 years','50-54 years','55-59 years','60-64 years','65-69 years','70-74 years','75-79 years'

AND ({Site and Morphology.CS Schema v0204+} = 'Lung'

OR {Site and Morphology.CS Schema - AJCC 6th Edition} = 'Lung')

AND ({Stage - AJCC.Derived AJCC Stage Group, 7th ed (2010+)} = 'IIIA'

OR {Stage - AJCC.Derived AJCC Stage Group, 6th ed (2004+)} = 'IIIA'

OR {Stage - AJCC.AJCC stage 3rd edition (1988-2003)} = ' 31'

OR {Stage - AJCC.SEER modified AJCC stage 3rd (1988-2003)} = ' 31')

AND ({Stage - TNM.Derived AJCC N, 7th ed (2010+)} = 'N2','N2a','N2b','N2c'

OR {Stage - TNM.Derived AJCC N, 6th ed (2004+)} = 'N2','N2a','N2b','N2c'

OR {Stage - TNM.N value - based on AJCC 3rd (1988-2003)} = 'N2')

NMA Model for Adverse Events – Fixed Effects

```
# Poisson likelihood, log link
# Fixed effects model for multi-arm trials
model{ # *** PROGRAM STARTS
```

```
for(i in 1:ns){ # LOOP THROUGH STUDIES
 mu[i] ~ dnorm(0,.0001) # vague priors for all trial baselines
 for (k in 1:na[i]) { # LOOP THROUGH ARMS
 r[i,k] ~ dpois(theta[i,k]) # Poisson likelihood
 theta[i,k] <- lambda[i,k]*E[i,k] # event rate * exposure
 log(lambda[i,k]) <- mu[i] + d[t[i,k]] - d[t[i,1]] # model for linear predictor
 dev[i,k] <- 2*((theta[i,k]-r[i,k]) + r[i,k]*log(r[i,k]/theta[i,k])) #Deviance contribution
 }
 resdev[i] <- sum(dev[i,1:na[i]]) # summed residual deviance contribution for this trial
 }
totresdev <- sum(resdev[]) #Total Residual Deviance
d[1]<-0 # treatment effect is zero for reference treatment
for (k in 2:nt){ d[k] ~ dnorm(0,.0001) } # vague priors for treatment effects



sd ~ dunif(0,5) # vague prior for between-trial SD
tau <- pow(sd,-2) # between-trial precision = (1/between-trial variance)

```

```
# pairwise HRs and LHRs for all possible pair-wise comparisons, if nt>2
for (c in 1:(nt-1)) {
 for (k in (c+1):nt) {
 lhr[c,k] <- (d[k]-d[c])
 log(hr[c,k]) <- lhr[c,k]
 }
 }

} # *** PROGRAM ENDS

list(ns=4, nt=3)

t[,1]  r[,1]  E[,1]  t[,2]  r[,2]  E[,2]  na[]
2      182    285.2  3      141    299.52 2
3      482    434.3  1      608    409.34 2
1      137    214.4  3      150    230.04 2
1      98     321.75 2      108    298.93 2

END
```

```
#chain 1
list(d=c( NA, 0, 0), mu=c(0, 0, 0, 0))
#chain 2
list(d=c( NA, -1, 1), mu=c(-3, -3, -3, -3))
#chain 3
list(d=c( NA, 2, 2),  mu=c(-3, 5, -1, -3))
```

NMA Model for Adverse Events - Random Effects

```
# Poisson likelihood, log link
# Random effects model for multi-arm trials
model{ # *** PROGRAM STARTS
for(i in 1:ns){ # LOOP THROUGH STUDIES
 w[i,1] <- 0 # adjustment for multi-arm trials is zero for control arm
 delta[i,1] <- 0 # treatment effect is zero for control arm
 mu[i] ~ dnorm(0,.0001) # vague priors for all trial baselines
 for (k in 1:na[i]) { # LOOP THROUGH ARMS
 r[i,k] ~ dpois(theta[i,k]) # Poisson likelihood
 theta[i,k] <- lambda[i,k]*E[i,k] # failure rate * exposure
 log(lambda[i,k]) <- mu[i] + delta[i,k] # model for linear predictor
```

```
1287  dev[i,k] <- 2*((theta[i,k]-r[i,k]) + r[i,k]*log(r[i,k]/theta[i,k])) #Deviance contribution
1288  }
1289  resdev[i] <- sum(dev[i,1:na[i]]) # summed residual deviance contribution for this trial
1290  for (k in 2:na[i]) { # LOOP THROUGH ARMS
1291  delta[i,k] ~ dnorm(md[i,k],taud[i,k]) # trial-specific LOR distributions
1292  md[i,k] <- d[t[i,k]] - d[t[i,1]] + sw[i,k] # mean of LOR distributions (with multi-arm trial correction)
1293  taud[i,k] <- tau *2*(k-1)/k # precision of LOR distributions (with multi-arm trial correction)
1294  w[i,k] <- (delta[i,k] - d[t[i,k]] + d[t[i,1]]) # adjustment for multi-arm RCTs
1295  sw[i,k] <- sum(w[i,1:k-1])/(k-1) # cumulative adjustment for multi-arm trials
1296  }
1297  }
1298
1299
1300  totresdev <- sum(resdev[]) #Total Residual Deviance
1301  d[1]<-0 # treatment effect is zero for reference treatment
1302  for (k in 2:nt){ d[k] ~ dnorm(0,.0001) } # vague priors for treatment effects
1303  sd ~ dunif(0,5) # vague prior for between-trial SD
1304  tau <- pow(sd,-2) # between-trial precision = (1/between-trial variance)
1305  # pairwise HRs and LHRs for all possible pair-wise comparisons, if nt>2
1306  for (c in 1:(nt-1)) {
```

```
 for (k in (c+1):nt) {
 lhr[c,k] <- (d[k]-d[c])
 log(hr[c,k]) <- lhr[c,k]
 }
 }

} # *** PROGRAM ENDS

list(ns=4, nt=3)

t[,1]  r[,1]  E[,1]  t[,2]  r[,2]  E[,2]  na[]
2      182    285.2  3      141    299.52 2
3      482    434.3  1      608    409.34 2
1      137    214.4  3      150    230.04 2
1      98     321.75 2      108    298.93 2

END

#chain 1
list(d=c( NA, 0, 0), sd=1, mu=c(0, 0, 0, 0))
```

#chain 2

list(d=c(NA, -1, 1), sd=4, mu=c(-3, -3, -3, -3))

#chain 3

list(d=c(NA, 2, 2), sd=2, mu=c(-3, 5, -1, -3))

R code to calculate (undiscounted and discounted) area under the Kaplan Meier curves, along with correlation between the areas under PFS and OS curves and standard error based on non-parametric bootstrap sampling.

```
##Load survival package
library("survival")

#############################################################################
## Function to calculate area under a Kaplan Meier curve
## Required Input:
## data - with column names:
## "stime" (survival time for each patient),
## "event" (1 if patient experienced event, 0 if patient censored),
## "treat" (code for treatment patient received)
## rmean - time to restrict curve to
## Outputs: AUC restricted to 'rmean' years and its standard error
#############################################################################
my.AUC<-function(data,rmean){
  fit<-survfit(Surv(stime,event)~1,data=data)
  surv.stats<-summary(fit,print.rmean=TRUE,rmean=rmean)$table[5:6]
  surv.stats
}

#############################################################################
## Function to calculate area and discounted area under a Kaplan Meier curve
## Required Input:
## data - with column names:
## "stime" (survival time for each patient),
## "event" (1 if patient experienced event, 0 if patient censored),
## "treat" (code for treatment patient received)
##       - note should only be 1 treatment in data
## max.time - time to restrict curve to
## dis.fac - discount factor, 1/(1+annual rate)
## Outputs: AUC and discounted AUC restricted to 'rmean' years
```

```
################################################################################
my.disc.AUC<-function(data,max.time=5,disc.fac=1/1.035){
 #Fit Kaplan Meier curve to data
 fit<-survfit(Surv(stime,event)~1,data=data)

 #Calculate AUC in each one-year time interval
 #Check to see if any patient experienced event at the end of a year
 #If so, calculate AUC up to that time point
 #If not, calculate AUC based on time at which an event was last observed before end of year
 time<-0:max.time
 X<-match(fit$time,time)
 X<-X[-which(is.na(X))]
 if(length(X)==0){time=time}else{time=time[-X]}
 sum.fit<-summary(fit)
 #Set up data required to calculate AUC in each one-year time interval
 my.tab<-data.frame(time=sum.fit$time,
                    n.risk=sum.fit$n.risk,
                    n.event=sum.fit$n.event,
                    survival=sum.fit$surv,
                    std.err=sum.fit$std.err,
                    time.diff=rep(NA,length(sum.fit$time)),
                    AUC=rep(NA,length(sum.fit$time)))
 #Add in lines for end of year time point to calculate AUC
 temp.tab<-data.frame(time=time,
                      n.risk=rep(NA,length(time)),
                      n.event=rep(0,length(time)),
                      survival=c(1,rep(NA,length(time)-1)),
                      std.err=rep(NA,length(time)),
                      time.diff=rep(NA,length(time)),
                      AUC=rep(NA,length(time)))
 my.tab<-rbind(my.tab,temp.tab)
 my.tab<-my.tab[order(my.tab$time),]

 #Make sure there are no time points beyond desired cut-off
 test<-length(which(my.tab$time>max.time))>0
 if(test){my.tab<-my.tab[-which(my.tab$time>max.time),]}else{my.tab<-my.tab}

 #Calculate AUC between observed time points
 for(i in 1:(length(time)-1)){
  row.ind<-which(my.tab$time==time[i+1])
  my.tab$survival[row.ind]=my.tab$survival[row.ind-1]
 }
 for(j in 2:length(my.tab[,1])){
  my.tab$time.diff[j]<-my.tab$time[j]-my.tab$time[j-1]
```

```
  my.tab$AUC[j]<-my.tab$survival[j-1]*my.tab$time.diff[j]
 }

 #Which rows contain end of year data
 time.ind<-which(match(my.tab$time,0:max.time)!="NA")

 #Calculate and output the AUC and discounted AUC in each one year time interval
 undisc.AUC<-matrix(nrow=max.time,ncol=2)
 disc.AUC<-matrix(nrow=max.time,ncol=2)
 undisc.AUC[,1]<-1:max.time
 disc.AUC[,1]<-1:max.time
 for(k in 1:max.time){
  undisc.AUC[k,2]<-sum(my.tab$AUC[(time.ind[k]+1):time.ind[k+1]])
  disc.AUC[k,2]<-sum(my.tab$AUC[(time.ind[k]+1):time.ind[k+1]])*(disc.fac^(k-1))
 }
 t(rbind(undisc.AUC,disc.AUC))

}

##############################################################################
## Calculate SE of discounted AUC, correlation between AUC of PFS and OS curves  via bootstrapping
##############################################################################

#Prepare tables to record AUC and Discounted AUC
#AUC at 5 years
AUC.tab.5<-matrix(ncol=24,nrow=5)
colnames(AUC.tab.5)<-c("t1","t2","PFS1.boot","OS1.boot","sePFS1.boot","seOS1.boot","corr1",
                       "PFS2.boot","OS2.boot","sePFS2.boot","seOS2.boot","corr2",
                       "S1","seS1","S2","seS2",
                       "PFS1","OS1","sePFS1","seOS1",
                       "PFS2","OS2","sePFS2","seOS2")

#Discounted AUC at 5 years
disc.AUC.tab.5<-matrix(ncol=20,nrow=5)
colnames(disc.AUC.tab.5)<-c("t1","t2","PFS1.boot","OS1.boot","sePFS1.boot","seOS1.boot","corr1",
                            "PFS2.boot","OS2.boot","sePFS2.boot","seOS2.boot","corr2",
                            "S1","seS1","S2","seS2",
                            "PFS1","OS1","PFS2","OS2")

#Load data for PFS and OS curves

data.pfs <- read.csv("filename.csv", stringsAsFactors=FALSE)
data.os <- read.csv("filename.csv", stringsAsFactors=FALSE)
```

```
#########################################################################
### Bootstrap each curve, for each treatment and outcome separately
#########################################################################

time.horizon<-5    #Cut off time (e.g., 5 years)
B<-5000          #Number of bootstrap samples

#Subset data in first treatment group
treat.num1<-sort(unique(data.pfs$treat))[1]
data.pfs1<-subset(data.pfs,treat==treat.num1)
data.os1<-subset(data.os,treat==treat.num1)

dim(data.pfs1)[1]   #check number of patients
dim(data.os1)[1]    #check number of patients - should equal above

#Create empty matrices to fill in for bootstrapping
boot.auc.pfs1<-matrix(nrow=B,ncol=(2*time.horizon)+2)
colnames(boot.auc.pfs1)<-c(paste(rep("AUC",time.horizon),1:time.horizon,sep="."),
                    paste(rep("dAUC",time.horizon),1:time.horizon,sep="."),
                    "AUC","dAUC")
boot.auc.os1<-matrix(nrow=B,ncol=(2*time.horizon)+2)
colnames(boot.auc.os1)<-c(paste(rep("AUC",time.horizon),1:time.horizon,sep="."),
                   paste(rep("dAUC",time.horizon),1:time.horizon,sep="."),
                   "AUC","dAUC")

#Set the seed
set.seed(1234)

#Bootstrap data, throw out bootstrap samples where OS curve is lower than PFS curve
i<-1
k<-0    #counter for discards
while(i<(B+1)){
  #Calculate maximum number of patients reporting both OS and PFS
  max.samp<-max(dim(data.pfs1)[1],dim(data.os1)[1])
  inds<-sample(1:max.samp,replace=TRUE)
  boot.data.pfs1<-data.pfs1[inds[1:dim(data.pfs1)[1]],]
  boot.data.os1<-data.os1[inds[1:dim(data.os1)[1]],]

  #Fit KM curves to resampled data
  fit.pfs<-survfit(Surv(stime,event)~treat,data=boot.data.pfs1)
  fit.os<-survfit(Surv(stime,event)~treat,data=boot.data.os1)

  #Check to see if P(OS) > P(PFS)
  surv.test<-rep(NA,length(summary(fit.os)$time))
```

```
 for(j in 1:length(summary(fit.os)$time)){
  time.test<-which(summary(fit.os)$time[j]>=summary(fit.pfs)$time)
  surv.test[j]<-summary(fit.os)$surv[j]>=summary(fit.pfs)$surv[max(time.test)]
 }
 surv.test.test<-sum(1*(surv.test=="FALSE"),na.rm=TRUE)

 if(surv.test.test==0){
  boot.auc.pfs1[i,1:(2*time.horizon)]<-my.disc.AUC(boot.data.pfs1,max.time=time.horizon)[2,]
  boot.auc.pfs1[i,((2*time.horizon)+1):((2*time.horizon)+2)]<-c(sum(boot.auc.pfs1[i,1:time.horizon]),sum(boot.auc.pfs1[i,(time.horizon+1):(2*time.horizon)]))
  boot.auc.os1[i,1:(2*time.horizon)]<-my.disc.AUC(boot.data.os1,max.time=time.horizon)[2,]
  boot.auc.os1[i,((2*time.horizon)+1):((2*time.horizon)+2)]<-c(sum(boot.auc.os1[i,1:time.horizon]),sum(boot.auc.os1[i,(time.horizon+1):(2*time.horizon)]))

  i<-i+1
 } else {
  i<-i
  k<-k+1
 }

}

#Number of samples thrown away
k

#Record results, fill in tables
AUC.tab.5[study.num,"t1"]<-treat.num1
disc.AUC.tab.5[study.num,"t1"]<-treat.num1

AUC.tab.5[study.num,"PFS1.boot"]<-mean(boot.auc.pfs1[,((2*time.horizon)+1)])
disc.AUC.tab.5[study.num,"PFS1.boot"]<-mean(boot.auc.pfs1[,((2*time.horizon)+2)])
AUC.tab.5[study.num,"sePFS1.boot"]<-sd(boot.auc.pfs1[,((2*time.horizon)+1)])
disc.AUC.tab.5[study.num,"sePFS1.boot"]<-sd(boot.auc.pfs1[,((2*time.horizon)+2)])

AUC.tab.5[study.num,"OS1.boot"]<-mean(boot.auc.os1[,((2*time.horizon)+1)])
disc.AUC.tab.5[study.num,"OS1.boot"]<-mean(boot.auc.os1[,((2*time.horizon)+2)])
AUC.tab.5[study.num,"seOS1.boot"]<-sd(boot.auc.os1[,((2*time.horizon)+1)])
disc.AUC.tab.5[study.num,"seOS1.boot"]<-sd(boot.auc.os1[,((2*time.horizon)+2)])

AUC.tab.5[study.num,"corr1"]<-cor(boot.auc.pfs1[,((2*time.horizon)+1)],boot.auc.os1[,((2*time.horizon)+1)])
disc.AUC.tab.5[study.num,"corr1"]<-cor(boot.auc.pfs1[,((2*time.horizon)+2)],boot.auc.os1[,((2*time.horizon)+2)])

fit.os1<-survfit(Surv(stime,event)~1,data=data.os1)

AUC.tab.5[study.num,"S1"]<-summary(fit.os1,time=time.horizon)$surv
AUC.tab.5[study.num,"seS1"]<-summary(fit.os1,time=time.horizon)$std.err
```

```
disc.AUC.tab.5[study.num,"S1"]<-summary(fit.os1,time=time.horizon)$surv
disc.AUC.tab.5[study.num,"seS1"]<-summary(fit.os1,time=time.horizon)$std.err

AUC.tab.5[study.num,"PFS1"]<-my.AUC(data.pfs1,rmean=5)[1]
AUC.tab.5[study.num,"sePFS1"]<-my.AUC(data.pfs1,rmean=5)[2]
AUC.tab.5[study.num,"OS1"]<-my.AUC(data.os1,rmean=5)[1]
AUC.tab.5[study.num,"seOS1"]<-my.AUC(data.os1,rmean=5)[2]

disc.AUC.tab.5[study.num,"PFS1"]<-sum(my.disc.AUC(data.pfs1,max.time=5,disc.fac=1/1.035)[2,6:10])
disc.AUC.tab.5[study.num,"OS1"]<-sum(my.disc.AUC(data.os1,max.time=5,disc.fac=1/1.035)[2,6:10])

#Save a copy of results from each bootstrapped sample
write.csv(boot.auc.pfs1,"filename pfs treat 1.csv")
write.csv(boot.auc.os1,"filename os treat 1.csv")

#######################################

#Subset data in first treatment group
treat.num2<-sort(unique(data.pfs$treat))[2]
data.pfs2<-subset(data.pfs,treat==treat.num2)
data.os2<-subset(data.os,treat==treat.num2)

dim(data.pfs2)[1]   #check number of patients
dim(data.os2)[1]    #check number of patients - should equal above

#Create empty matrices to fill in for bootstrapping
boot.auc.pfs2<-matrix(nrow=B,ncol=(2*time.horizon)+2)
colnames(boot.auc.pfs2)<-c(paste(rep("AUC",time.horizon),1:time.horizon,sep="."),
                    paste(rep("dAUC",time.horizon),1:time.horizon,sep="."),
                    "AUC","dAUC")
boot.auc.os2<-matrix(nrow=B,ncol=(2*time.horizon)+2)
colnames(boot.auc.os2)<-c(paste(rep("AUC",time.horizon),1:time.horizon,sep="."),
                   paste(rep("dAUC",time.horizon),1:time.horizon,sep="."),
                   "AUC","dAUC")

#Set the seed
set.seed(1234)

#Bootstrap data, throw out bootstrap samples where OS curve is lower than PFS curve
i<-1
k<-0    #counter for discards
while(i<(B+1)){
  #Calculate maximum number of patients reporting both OS and PFS
  max.samp<-max(dim(data.pfs2)[1],dim(data.os2)[1])
```

```
  inds<-sample(1:max.samp,replace=TRUE)
  boot.data.pfs2<-data.pfs2[inds[1:dim(data.pfs2)[1]],]
  boot.data.os2<-data.os2[inds[1:dim(data.os2)[1]],]

  #Fit KM curves to resampled data
  fit.pfs<-survfit(Surv(stime,event)~treat,data=boot.data.pfs2)
  fit.os<-survfit(Surv(stime,event)~treat,data=boot.data.os2)

  #Check to see if P(OS) > P(PFS)
  surv.test<-rep(NA,length(summary(fit.os)$time))
  for(j in 1:length(summary(fit.os)$time)){
    time.test<-which(summary(fit.os)$time[j]>=summary(fit.pfs)$time)
    surv.test[j]<-summary(fit.os)$surv[j]>=summary(fit.pfs)$surv[max(time.test)]
  }
  surv.test.test<-sum(1*(surv.test=="FALSE"),na.rm=TRUE)

  if(surv.test.test==0){
    boot.auc.pfs2[i,1:(2*time.horizon)]<-my.disc.AUC(boot.data.pfs2,max.time=time.horizon)[2,]
    boot.auc.pfs2[i,((2*time.horizon)+1):((2*time.horizon)+2)]<-c(sum(boot.auc.pfs2[i,1:time.horizon]),sum(boot.auc.pfs2[i,(time.horizon+1):(2*time.horizon)]))
    boot.auc.os2[i,1:(2*time.horizon)]<-my.disc.AUC(boot.data.os2,max.time=time.horizon)[2,]
    boot.auc.os2[i,((2*time.horizon)+1):((2*time.horizon)+2)]<-c(sum(boot.auc.os2[i,1:time.horizon]),sum(boot.auc.os2[i,(time.horizon+1):(2*time.horizon)]))

    i<-i+1
  } else {
    i<-i
    k<-k+1
  }

}

#Number of samples thrown away
k

#Record results, fill in tables
AUC.tab.5[study.num,"t2"]<-treat.num2
disc.AUC.tab.5[study.num,"t2"]<-treat.num2

AUC.tab.5[study.num,"PFS2.boot"]<-mean(boot.auc.pfs2[,((2*time.horizon)+1)])
disc.AUC.tab.5[study.num,"PFS2.boot"]<-mean(boot.auc.pfs2[,((2*time.horizon)+2)])
AUC.tab.5[study.num,"sePFS2.boot"]<-sd(boot.auc.pfs2[,((2*time.horizon)+1)])
disc.AUC.tab.5[study.num,"sePFS2.boot"]<-sd(boot.auc.pfs2[,((2*time.horizon)+2)])

AUC.tab.5[study.num,"OS2.boot"]<-mean(boot.auc.os2[,((2*time.horizon)+1)])
disc.AUC.tab.5[study.num,"OS2.boot"]<-mean(boot.auc.os2[,((2*time.horizon)+2)])
```

```
AUC.tab.5[study.num,"seOS2.boot"]<-sd(boot.auc.os2[,((2*time.horizon)+1)])
disc.AUC.tab.5[study.num,"seOS2.boot"]<-sd(boot.auc.os2[,((2*time.horizon)+2)])


AUC.tab.5[study.num,"corr2"]<-cor(boot.auc.pfs2[,((2*time.horizon)+1)],boot.auc.os2[,((2*time.horizon)+1)])
disc.AUC.tab.5[study.num,"corr2"]<-cor(boot.auc.pfs2[,((2*time.horizon)+2)],boot.auc.os2[,((2*time.horizon)+2)])

fit.pfs2<-survfit(Surv(stime,event)~1,data=data.pfs2)
fit.os2<-survfit(Surv(stime,event)~1,data=data.os2)

AUC.tab.5[study.num,"S2"]<-summary(fit.os2,time=time.horizon)$surv
AUC.tab.5[study.num,"seS2"]<-summary(fit.os2,time=time.horizon)$std.err
disc.AUC.tab.5[study.num,"S2"]<-summary(fit.os2,time=time.horizon)$surv
disc.AUC.tab.5[study.num,"seS2"]<-summary(fit.os2,time=time.horizon)$std.err

AUC.tab.5[study.num,"PFS2"]<-my.AUC(data.pfs2,rmean=5)[1]
AUC.tab.5[study.num,"sePFS2"]<-my.AUC(data.pfs2,rmean=5)[2]
AUC.tab.5[study.num,"OS2"]<-my.AUC(data.os2,rmean=5)[1]
AUC.tab.5[study.num,"seOS2"]<-my.AUC(data.os2,rmean=5)[2]

disc.AUC.tab.5[study.num,"PFS2"]<-sum(my.disc.AUC(data.pfs2,max.time=5,disc.fac=1/1.035)[2,6:10])
disc.AUC.tab.5[study.num,"OS2"]<-sum(my.disc.AUC(data.os2,max.time=5,disc.fac=1/1.035)[2,6:10])


#Save a copy of results from each bootstrapped sample
write.csv(boot.auc.pfs2,"filename pfs treat 2.csv")
write.csv(boot.auc.os2,"filename os treat 2.csv")
```

WinBUGS code for NMA of area under the Kaplan Meier curves and Probability of Surviving up to *5* years – Fixed effect model. Notes: WinBUGS files, including data and initial values are available upon request. Same code may be used for 4-year and discounted AUC data.

```
model{

#Code for 5-year Survival
for (i in 1:ns){
        mu.S[i]~dnorm(0,.0001)
        for (k in 1:na[i]){
                prec.S[i,k]<-pow(se.S[i,k],-2)
                y.S[i,k]~dnorm(pi[i,k],prec.S[i,k])
                dev.S[i,k]<-(y.S[i,k]-pi[i,k])*(y.S[i,k]-pi[i,k])*prec.S[i,k]
                logit(pi[i,k])<-mu.S[i] + delta.S[i,k]
                delta.S[i,k]<- d.S[t[i,k]] - d.S[t[i,1]]
        }
   resdev.S[i] <- sum(dev.S[i,1:na[i]])
}
totresdev.S<-sum(resdev.S[])


#Code for 5-year AUCs (Bivariate for PFS and OS)
for (i in 1:ns){
        mu.PFS[i]~dnorm(0,.0001)
        mu.PPS[i]~dnorm(0,.0001)
        for (k in 1:na[i]){
                #Set precision matrix
                Sigma[i,k,1,1]<-pow(se.PFS[i,k],2)
                Sigma[i,k,2,2]<-pow(se.OS[i,k],2)
                Sigma[i,k,1,2]<-corr[i,k]*se.PFS[i,k]*se.OS[i,k]
                Sigma[i,k,2,1]<-Sigma[i,k,1,2]
                Prec[i,k,1:2,1:2]<-inverse(Sigma[i,k,1:2,1:2])

                y[i,k,1:2]~dmnorm(theta[i,k,1:2],Prec[i,k,1:2,1:2])
                for (j in 1:2){
                        diff[i,k,j]<- y[i,k,j]-theta[i,k,j]
                        z[i,k,j]<- inprod2(Prec[i,k,j,1:2],diff[i,k,1:2])
                }
```

```
            dev[i,k]<-inprod2(diff[i,k,1:2],z[i,k,1:2])

            theta[i,k,1]<- mu.PFS[i] + delta.PFS[i,k]
            theta[i,k,2]<- theta[i,k,1] + phi[i,k]
            phi[i,k]<- mu.PPS[i] + delta.PPS[i,k]

            delta.PFS[i,k]<- d.PFS[t[i,k]] - d.PFS[t[i,1]]
            delta.PPS[i,k]<-  d.PPS[t[i,k]] - d.PPS[t[i,1]]

            }

    resdev[i] <- sum(dev[i,1:na[i]])
    }
totresdev<-sum(resdev[])

#Chemoradiotherapy (treatment code 1) is reference
d.S[1]<-0
d.PFS[1]<-0
d.PPS[1]<-0

for (k in 2:nt){
            d.S[k]~dnorm(0,.0001)
            d.PFS[k]~dnorm(0,.0001)
            d.PPS[k]~dnorm(0,.0001)
            }

#Assumed log odds of survival, mean PPS and PFS time over 5-years on reference treatment 1 in UK
m.S<-mu.S[5]
m.PFS<-mu.PFS[5]
m.PPS<-mu.PPS[5]

#Predicted probability of survival and mean survival times in UK population for each treatment
for (k in 1:nt){
            #Up to 5 years
            logit(S5[k])<- m.S + d.S[k]
            meanPFS5[k]<- m.PFS + d.PFS[k]
            meanPPS5[k]<- m.PPS + d.PPS[k]
            meanOS5[k]<-meanPFS5[k]+meanPPS5[k]
```

```
            #Long-term
            meanPFS[k]<- meanPFS5[k] + S5[k]*C
            meanPPS[k]<- meanPPS5[k]
            meanOS[k]<-meanPFS[k]+meanPPS[k]
            }

#Overall Survival at 5 Years, OR of Survival, Overall Survival relative to CR
for (k in 1:nt){
            d.OS5[k]<-d.PFS[k]+d.PPS[k]
            OR.S[k]<-exp(d.S[k])
            d.OS[k]<-(meanPFS[k]-meanPFS[1])+(meanPPS[k]-meanPPS[1])
            }

#Rank treatments
for (k in 1:nt)  {
            # PFS
            rk.PFS[k]  <- nt+1-rank(d.PFS[],k)
            best.PFS[k]  <- equals(rk.PFS[k],1)    # Largest is best (i.e. rank 1)
            # PPS
            rk.PPS[k]  <- nt+1-rank(d.PPS[],k)
            best.PPS[k]  <- equals(rk.PPS[k],1)    # Largest is best (i.e. rank 1)
            # OS at 5 years
            rk.OS5[k]  <- nt+1-rank(d.OS5[],k)
            best.OS5[k]  <- equals(rk.OS5[k],1)    # Largest is best (i.e. rank 1)
            # OR of Survival
            rk.OR.S[k]  <- nt+1-rank(OR.S[],k)
            best.OR.S[k]  <- equals(rk.OR.S[k],1)    # Largest is best (i.e. rank 1)
            # OS
            rk.OS[k]  <- nt+1-rank(d.OS[],k)
            best.OS[k]  <- equals(rk.OS[k],1)    # Largest is best (i.e. rank 1)
            }

}
```

WinBUGS code for NMA of area under the Kaplan Meier curves and Probability of Surviving up to *5* years – Random effects model.

<u>Notes</u>: WinBUGS files, including data and initial values are available upon request. Same code may be used for 4-year and discounted AUC

data.

```
model{

#Code for 5-year Survival
for (i in 1:ns){
        delta.S[i,1]<-0
        mu.S[i]~dnorm(0,.0001)
        for (k in 1:na[i]){
                prec.S[i,k]<-pow(se.S[i,k],-2)
                y.S[i,k]~dnorm(pi[i,k],prec.S[i,k])
                dev.S[i,k]<-(y.S[i,k]-pi[i,k])*(y.S[i,k]-pi[i,k])*prec.S[i,k]
                logit(pi[i,k])<-mu.S[i] + delta.S[i,k]
                }
        resdev.S[i] <- sum(dev.S[i,1:na[i]])

        md.S[i,2] <- d.S[t[i,2]] - d.S[t[i,1]]
        delta.S[i,2] ~ dnorm(md.S[i,2],tau.S)

        }
totresdev.S<-sum(resdev.S[])


#Code for 5-year AUCs (Bivariate for PFS and OS)
for (i in 1:ns){
        delta.PFS[i,1]<-0
        delta.PPS[i,1]<-0
        mu.PFS[i]~dnorm(0,.0001)
        mu.PPS[i]~dnorm(0,.0001)
        for (k in 1:na[i]){
                #Set precision matrix
                Sigma[i,k,1,1]<-pow(se.PFS[i,k],2)
                Sigma[i,k,2,2]<-pow(se.OS[i,k],2)
                Sigma[i,k,1,2]<-corr[i,k]*se.PFS[i,k]*se.OS[i,k]
                Sigma[i,k,2,1]<-Sigma[i,k,1,2]
                Prec[i,k,1:2,1:2]<-inverse(Sigma[i,k,1:2,1:2])

                y[i,k,1:2]~dmnorm(theta[i,k,1:2],Prec[i,k,1:2,1:2])
```

```
                for (j in 1:2){
                        diff[i,k,j]<- y[i,k,j]-theta[i,k,j]

                        z[i,k,j]<- inprod2(Prec[i,k,j,1:2],diff[i,k,1:2])
                }
        dev[i,k]<-inprod2(diff[i,k,1:2],z[i,k,1:2])

        theta[i,k,1]<- mu.PFS[i] + delta.PFS[i,k]
        theta[i,k,2]<- theta[i,k,1] + phi[i,k]
        phi[i,k]<- mu.PPS[i] + delta.PPS[i,k]

        }

        md.PFS[i,2] <- d.PFS[t[i,2]] - d.PFS[t[i,1]]
        md.PPS[i,2] <- d.PPS[t[i,2]] - d.PPS[t[i,1]]
        delta.PFS[i,2] ~ dnorm(md.PFS[i,2], tau.PFS)
        delta.PPS[i,2] ~ dnorm(md.PPS[i,2], tau.PPS)

   resdev[i] <- sum(dev[i,1:na[i]])
   }
totresdev<-sum(resdev[])

#Chemoradiotherapy (treatment code 1) is reference
d.S[1]<-0
d.PFS[1]<-0
d.PPS[1]<-0

#Priors on between-study SDs
sd.S ~ dunif(0,5)
sd.PFS ~ dunif(0,5)
sd.PPS ~ dunif(0,5)
tau.S <- pow(sd.S, -2)
tau.PFS <- pow(sd.PFS, -2)
tau.PPS <- pow(sd.PPS, -2)

for (k in 2:nt){
        d.S[k]~dnorm(0,.0001)
        d.PFS[k]~dnorm(0,.0001)
```

```
        d.PPS[k]~dnorm(0,.0001)
        }

#Assumed log odds of survival, mean PPS and PFS time over 5-years on reference treatment 1 in UK
m.S<-mu.S[5]
m.PFS<-mu.PFS[5]
m.PPS<-mu.PPS[5]

#Predicted probability of survival and mean survival times in UK population for each treatment
for (k in 1:nt){
        #Up to 5 years
        logit(S5[k])<- m.S + d.S[k]
        meanPFS5[k]<- m.PFS + d.PFS[k]
        meanPPS5[k]<- m.PPS + d.PPS[k]
        meanOS5[k]<-meanPFS5[k]+meanPPS5[k]

        #Long-term
        meanPFS[k]<- meanPFS5[k] + S5[k]*C
        meanPPS[k]<- meanPPS5[k]
        meanOS[k]<-meanPFS[k]+meanPPS[k]
        }

#Overall Survival at 5 Years, OR of Survival, Overall Survival relative to CR
for (k in 1:nt){
        d.OS5[k]<-d.PFS[k]+d.PPS[k]
        OR.S[k]<-exp(d.S[k])
        d.OS[k]<-(meanPFS[k]-meanPFS[1])+(meanPPS[k]-meanPPS[1])
        }

# Rank treatments
for (k in 1:nt)  {
        # PFS
        rk.PFS[k]  <- nt+1-rank(d.PFS[],k)
        best.PFS[k]  <- equals(rk.PFS[k],1)    # Largest is best (i.e. rank 1)
        # PPS
        rk.PPS[k]  <- nt+1-rank(d.PPS[],k)
        best.PPS[k]  <- equals(rk.PPS[k],1)    # Largest is best (i.e. rank 1)
        # OS at 5 years
```

```
            rk.OS5[k]  <- nt+1-rank(d.OS5[],k)
            best.OS5[k]  <- equals(rk.OS5[k],1)    # Largest is best (i.e. rank 1)
            # OR of Survival
            rk.OR.S[k]  <- nt+1-rank(OR.S[],k)
            best.OR.S[k]  <- equals(rk.OR.S[k],1)    # Largest is best (i.e. rank 1)
            # OS
            rk.OS[k]  <- nt+1-rank(d.OS[],k)
            best.OS[k]  <- equals(rk.OS[k],1)    # Largest is best (i.e. rank 1)
            # QALY
            }

}
```

WinBUGS code to estimate proportion of events occurring each year up to *5* years. Notes: WinBUGS files, including data and initial values are available upon request.

```
model{

        for (i in 1:ns){
                for (k in 1:na[i]){
                        for (s in 1:5){
                #Likelihood for Survival at times s=1,2,3,4,5
                                prec.S[i,k,s]<-pow(se.S[i,k,s],-2)
                                y.S[i,k,s]~dnorm(pi[i,k,s],prec.S[i,k,s])
                                dev.S[i,k,s]<-(y.S[i,k,s]-pi[i,k,s])*(y.S[i,k,s]-pi[i,k,s])*prec.S[i,k,s]
                        }

#Model for Survival probs, pi, as a function of the proportion of events in each 1-year time period, rho, by treatment
                pi[i,k,5]~dbeta(1,1)
                pi[i,k,4]<- pi[i,k,5] + rho[5]*(1-pi[i,k,5])
                pi[i,k,3]<- pi[i,k,5] + sum(rho[4:5])*(1-pi[i,k,5])
                pi[i,k,2]<- pi[i,k,5] + sum(rho[3:5])*(1-pi[i,k,5])
                pi[i,k,1]<- pi[i,k,5] + sum(rho[2:5])*(1-pi[i,k,5])
                }
   resdev.S[i] <- sum(dev.S[i,1:na[i], 1:5])
        }
                totresdev<- sum(resdev.S[])

#Dirichlet prior (using Gamma formulation)
                for (s in 1:5){
                        x[s]~dgamma(alpha0[s],1)
                        rho[s]<- alpha[s]/sum(alpha[1:5])
                alpha0[s]<- max(alpha[s],0.1)
                log(alpha[s])<- beta[s]
                beta[s]~dnorm(0,.01)
        }

dum<-t[1,1]
}
```

WinBUGS code to estimate proportion of events occurring each year up to *4* years. Notes: WinBUGS files, including data and initial values are available upon request.

```
model{

        for (i in 1:ns){
                for (k in 1:na[i]){
                        for (s in 1:4){
                #Likelihood for Survival at times s=1,2,3,4
                                prec.S[i,k,s]<-pow(se.S[i,k,s],-2)
                                y.S[i,k,s]~dnorm(pi[i,k,s],prec.S[i,k,s])
                                dev.S[i,k,s]<-(y.S[i,k,s]-pi[i,k,s])*(y.S[i,k,s]-pi[i,k,s])*prec.S[i,k,s]
                        }

#Model for Survival probs, pi, as a function of the proportion of events in each 1-year time period, rho, by treatment
                pi[i,k,4]~dbeta(1,1)
                pi[i,k,3]<- pi[i,k,4] + rho[4]*(1-pi[i,k,4])
                pi[i,k,2]<- pi[i,k,4] + sum(rho[3:4])*(1-pi[i,k,4])
                pi[i,k,1]<- pi[i,k,4] + sum(rho[2:4])*(1-pi[i,k,4])
                }
   resdev.S[i] <- sum(dev.S[i,1:na[i], 1:4])
        }
                totresdev<- sum(resdev.S[])

#Dirichlet prior (using Gamma formulation)
                for (s in 1:4){
                        x[s]~dgamma(alpha0[s],1)
                        rho[s]<- alpha[s]/sum(alpha[1:4])
                alpha0[s]<- max(alpha[s],0.1)
                log(alpha[s])<- beta[s]
                beta[s]~dnorm(0,.01)
        }

dum<-t[1,1]
}
```

Appendix K – Cost-Utility Analysis

Background

Stage IIIA-N2 NSCLC is a common presentation but, despite several RCTs investigating different options, the optimal management strategy remains controversial. This stage of NSCLC is generally considered to be the most advanced stage of the disease in which patients would normally still receive radical rather than systemic treatment. Patients with stage IIIA-N2 disease commonly receive chemoradiotherapy (CR) and chemotherapy and surgery (CS) but may receive tri-modality therapy with chemoradiotherapy and surgery (CRS). These are the three treatment options examined in this analysis.

Typically, the chemotherapy and/or radiotherapy components will happen before surgery to make the tumour more operable although patients may receive an amount of either following surgery. Surgery for N2 disease is a complex operation with a high reference cost. The committee prioritised this area for de novo modelling because they wanted to see an analysis that combined progression-free survival (PFS), post-progression survival (PPS), overall survival (OS), adverse event data and costs into a single analysis. The systematic review conducted for this guideline found no published economic evaluations in this area.

Methods

Model Structure

The model is divided into short and long term components. The short term model, covering five years, is based on clinical trial data from six of the studies included in the review, which were prioritised for further analyses based on the relevance of their populations and interventions (Albain 2009, Girard 2009, Eberhardt 2015. Pless 2015, Katakami 2012 and van Meerbeeck 2007[a]). While four years was the longest common follow up time among all six RCTs, we chose five years as the base case because this only meant excluding Girard 2009, which was the smallest and least relevant study. We felt this was a trade-off worth making to make use of more of the available data, while also making certain modelling assumptions discussed later on more likely to be true. Four year data for all parameters that the time period is relevant to were also sourced and used in sensitivity analysis. Patients surviving the short term model enter the long term model, which takes the form of a Partitioned Survival Analysis[b].

The primary clinical evidence for the short term model came from the network meta-analyses (NMAs) of RCTs identified in the clinical review for this guideline. A full write-up of the NMAs can be found in Appendix I but a brief discussion is included here.

[a] Please see the section on 'Clinical Studies – Included' above for full references
[b] NICE DSU TSD 19: Partitioned survival analysis for decision modelling in health care: a critical review (2007)

1993 It is very common for health economic models in lung cancer to divide patients into pre and post-progression health states, assuming some
1994 homogeneity of resource use and utility within those states and that transition between the two indicates something significant in terms of
1995 treatment. Overall survival at study endpoint is another key measure that is often reported in NSCLC RCTs. In order to obtain the average amount
1996 of time a patient undergoing any of the three interventions would spend in the progression free and progressed health states we digitised all the
1997 survival curves in the trials the committee prioritised for inclusion in the NMAs via the use of the Guyot et al algorithm[c]. This algorithm makes use
1998 of digitised survival curves (in this case we used Enguage[d] for this purpose) and the numbers at risk data that are commonly reported underneath
1999 Kaplan-Meier plots in RCTs to generate synthetic individual patient data. The algorithm creates a survival time and a censorship or event variable
2000 for each "patient" in the trial, which is amenable to the usual survival analysis techniques. Survival analysis on the synthetic data has been found to
2001 accurately reproduce the same analysis conducted on the real individual patient data from the trials in a large number of examples[c].

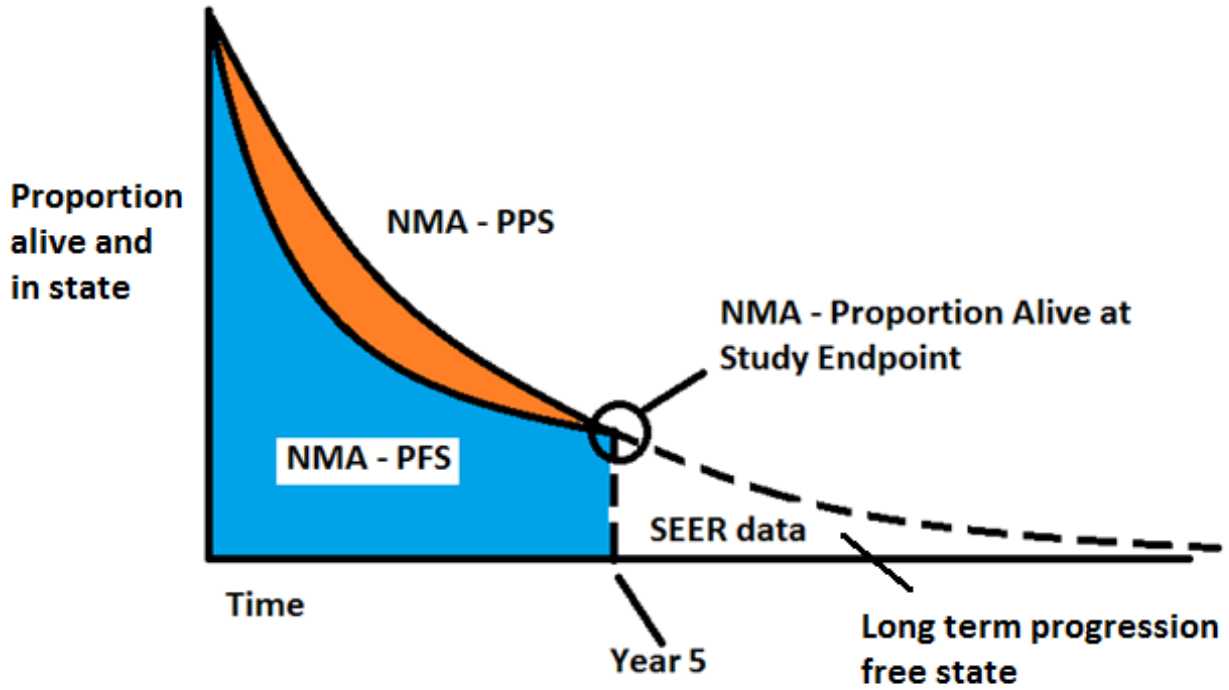
2002 Once the individual patient data had been obtained it was possible to calculate the area under the curve (AUC), which is equivalent to the mean
2003 time in state (restricted by the trial endpoint) and its standard error for both PFS and OS. Since PFS and OS are correlated, a correlation
2004 coefficient between the two was calculated and used in a bivariate NMA model that produced results for both PFS and OS for each of the three
2005 interventions. Since mean PPS would be equal to OS minus PFS for each iteration of the NMA, this statistic was also calculated via simple
2006 subtraction. Since the OS and PFS were obtained over five years of trial data, the AUC statistics were adjusted for discounting. A separate NMA
2007 model also calculated the probability of survival at study endpoint.

2008 All NMAs were conducted separately on two study endpoints; four and five years post treatment. The four year data were available for all six RCTs
2009 but five year data were available for all except the smallest and least relevant RCT so the committee preferred the five year analysis in the base
2010 case, with the four year data being used in sensitivity analysis. In either case, the committee instructed us to assume that all, or at least the vast
2011 majority, of the ~15% of patients who had survived to five years post treatment were in remission and would continue into the long-term model
2012 progression free until death. This assumption may be reasonable, given that the PFS and OS Kaplan-Meier curves reported in the trials showed a
2013 strong tendency toward convergence at five years.

2014 For the long term component of the model, a patient registry containing survival data conditional on NSCLC stage IIIA-N2 patients having survived
2015 for five years was obtained. Survival curves were fitted to this data and used in a long term Partitioned Survival Analysis with only two health
2016 states; (alive and) progression free and dead.

2017 The structure of the model is shown in Figure 13.

[c] Guyot et a (2012) Enhanced secondary analysis of survival data: reconstructing the data from published Kaplan-Meier survival curves. BMC Medical Research Methodology
[d] http://digitizer.sourceforge.net/



2018

2019 **Figure 13: Economic Model Structure (time in state up to 5 years is dictated by NMAs)**

2020

2021 **Model Parameters**

2022 **Utility Data**

2023 No direct health related quality of life data for progression free and post progression survival were available for patients with stage IIIA-N2 NSCLC.
2024 However, a targeted search was undertaken and a large number of potentially relevant data sources were identified that related to people with

Stage III NSCLC undergoing surgery. Of these, the three studies the committee thought the most relevant are displayed in Table 22. A random effects model was chosen to pool these data because the I-squared statistic equalled 80%, indicating high between study heterogeneity.

No relevant post-progression utility estimates were identified so a generic post-progression adjustment value taken from a study widely used in economic models for advanced NSCLC was used (Nafees 2008). The committee agreed that it was likely patients undergoing surgery would experience some reduction in health related quality of life for about three months while they recovered. This was borne out in the evidence from Bendixen 2016[e], a trial that investigated HRQoL in patients having surgery for NSCLC. We used data on EQ-5D measured at various time points in the thoracotomy arm of the trial to calculate the QALY loss from surgery by assuming that any dips below a linear trajectory between the time periods of 0 weeks and 12 weeks were due to surgery. The resulting difference between the areas under the curve for the observed values and the linear trajectory, calculated using simple averaging methods between observed time points, gave a QALY loss due to surgery of -0.012. This value was applied only to people actually undergoing surgery (see the section further down discussing drop-out rates).

Table 22: Utility Parameters

Source	N	Utility/QALYs	SE
Progression Free Survival			
Grutters[f] (2010) "People who had received CRS" (EQ-5D)	19	0.720	0.050
Tramontano[g] (2015) "People receiving CRS" (EQ-5D) Canada	207	0.760	0.013
Yang (2014)[h] "Stage III fit for surgery" (EQ-5D, validated Taiwanese version)	71	0.830	0.020
Random effects meta-analysis (Grutters, Tramontano, Yang)	297	0.779	0.030
Post Progression Adjustment			
Nafees[i] (2008)	100	-0.180	0.022
QALY loss due to surgery (calculated from Bendixen 2016)	~60	-0.012	

[e] Bendixen et al (2016) Postoperative pain and quality of life after lobectomy via video-assisted thoracoscopic surgery or anterolateral thoracotomy for early stage lung cancer: a randomised controlled trial. Lancet Oncology
[f] Grutters et al (2010) Health-related quality of life in patients surviving non-small cell lung cancer. Thorax
[g] Tramontano et al (2015) Catalog and Comparison of Societal Preferences (Utilities) for Lung Cancer Health States: Results from the Cancer Care Outcomes Research and Surveillance (CanCORS) Study. Medical Decision Making
[h] Yang et al (2014) Estimation of loss of quality-adjusted life expectancy (QALE) for patients with operable versus inoperable lung cancer: Adjusting quality-of-life and lead-time bias for utility. Lung Cancer
[i] Nafees et al (2008) Health state utilities for non small cell lung cancer. Health and Quality of Life Outcomes

For the long term portion of the model, in which people were assumed to remain progression-free until death, the progression-free utility value was multiplied by the age specific decrements that would be expected in the general population (Kind et al 1999). More specifically, the age specific value at each cycle was looked up from a table containing general population utility values and divided by the population level age specific utility value at cycle 0 of the long term model. This figure was then multiplied by the progression free survival utility value to give the utility at future cycles including any appropriate decrements for advanced age. Weighted averages were used for men and women assuming 53.4% of people in the model were men (NCLA 2017 data on general lung cancer presentation). To reflect the population in the underpinning trials, the starting age in the model was 60 (and therefore 65 in the long term model).

Table 23: General Population Utility Estimates for Use in Long Term Multiplier

Men	N	Utility	SE	Source
54 < age < 65	196	0.78	0.02	Kind et al 1999[j]
64 < age < 75	228	0.78	0.018543	Kind et al 1999
74 < age	108	0.75	0.026943	Kind et al 1999
Women				Kind et al 1999
54 < age < 65	288	0.81	0.015321	Kind et al 1999
64 < age < 75	260	0.78	0.015504	Kind et al 1999
74 < age	206	0.71	0.018812	Kind et al 1999

Adverse events were assumed to be acute in nature and not contribute meaningfully to QALY losses. Since adverse event rates did not differ greatly between the interventions, this limitation was assessed as minor.

Progression Free and Post Progression Survival Time (Short Term Model)

A single bivariate NMA model produced the estimates for discounted PFS and PPS. A brief discussion of this contained in the Model Structure section above and a full write up of this analysis can be found in Appendix I. The NMA had 50,000 burn-in iterations that were then discarded. 10,000 values that had been thinned by 5 were taken from the next 50,000 iterations and used in the economic model. For each run of the model, discounted PFS and PPS values for all three interventions came from a randomly sampled line of this CODA output. The use of a single line of CODA for all data points was essential to preserve the correlation structure in the posterior distributions.

[j] Kind et al (1999) UK population norms for EQ-5D. University of York

The discounted average pre and post progression survival time were multiplied by the relevant utility values to produce QALYs over 5 years. A surgery specific QALY decrement (see above) was applied to people receiving surgery in the CR and CRS model arms.

Survival at study endpoint

The probability of survival at study endpoint came from the relevant NMA (see Appendix I for a full discussion). The NMA had 50,000 burn-in iterations that were then discarded. 10,000 values that had been thinned by 5 were taken from the next 50,000 iterations and used in the economic model. For each run of the model, probability values for all three interventions came from a randomly sampled line of this CODA output. The use of a single line of CODA for all data points was essential to preserve the correlation structure in the posterior distributions. Patients who survived the short term section of the model proceeded into the long term section.

Table 24: NMA Results - Fixed Effects

Fixed Effects	4 Year Endpoint Data (Undiscounted)			4 Year Endpoint Data (Discounted)			5 Year Endpoint Data (Undiscounted)			5 Year Endpoint Data (Discounted)		
	UCL	Median	LCL	UCL	Median	LCL	UCL	Median	LCL	UCL	Median	LCL
CR - PFS	1.538	1.368	1.196	1.503	1.340	1.176	1.709	1.497	1.283	1.645	1.447	1.245
CS - PFS	1.568	1.383	1.196	1.513	1.339	1.164	1.731	1.506	1.285	1.655	1.446	1.240
CRS - PFS	1.872	1.632	1.396	1.813	1.588	1.366	2.172	1.868	1.567	2.071	1.788	1.509
CR - PPS	0.616	0.552	0.487	0.582	0.523	0.462	0.648	0.577	0.506	0.615	0.550	0.484
CS - PPS	0.483	0.434	0.385	0.467	0.415	0.363	0.560	0.492	0.424	0.534	0.473	0.413
CRS - PPS	0.486	0.376	0.266	0.448	0.344	0.240	0.510	0.373	0.238	0.480	0.355	0.231
CR p Surv	0.234	0.178	0.126	0.232	0.176	0.125	0.179	0.129	0.081	0.179	0.129	0.081
CS p Surv	0.258	0.202	0.145	0.260	0.203	0.148	0.212	0.158	0.107	0.212	0.158	0.107
CRS p Surv	0.299	0.215	0.146	0.299	0.215	0.147	0.230	0.155	0.098	0.230	0.155	0.098

Table 25: NMA Results - Random Effects

Random Effects	4 Year Endpoint Data (Undiscounted)			4 Year Endpoint Data (Discounted)			5 Year Endpoint Data (Undiscounted)			5 Year Endpoint Data (Discounted)		
	UCL	Median	LCL	UCL	Median	LCL	UCL	Median	LCL	UCL	Median	LCL
CR - PFS	1.560	1.382	1.201	1.521	1.350	1.181	1.734	1.514	1.293	1.673	1.461	1.253
CS - PFS	1.874	1.386	1.007	1.771	1.350	0.988	2.142	1.490	0.759	2.032	1.432	0.804

Random Effects	4 Year Endpoint Data (Undiscounted)			4 Year Endpoint Data (Discounted)			5 Year Endpoint Data (Undiscounted)			5 Year Endpoint Data (Discounted)		
CRS - PFS	2.022	1.620	1.220	1.938	1.579	1.203	2.443	1.835	1.191	2.298	1.759	1.180
CR - PPS	0.623	0.557	0.491	0.588	0.527	0.464	0.661	0.587	0.513	0.624	0.557	0.490
CS - PPS	0.821	0.454	0.131	0.748	0.431	0.162	1.261	0.556	0.000	1.169	0.532	0.000
CRS - PPS	0.714	0.399	0.103	0.652	0.365	0.115	1.001	0.408	0.000	0.945	0.390	0.000
CR p Surv	0.234	0.178	0.123	0.234	0.178	0.120	0.188	0.133	0.081	0.189	0.134	0.079
CS p Surv	0.365	0.201	0.097	0.358	0.202	0.102	0.376	0.157	0.049	0.368	0.158	0.040
CRS p Surv	0.380	0.216	0.113	0.371	0.216	0.115	0.372	0.163	0.063	0.425	0.164	0.061

While the relative effects derived from the NMA are insensitive to the choice of baseline values for chemoradiotherapy for PFS, PPS and probability of survival, the absolute values shown in Table 24 and Table 25 are highly sensitive to this choice. We chose to base this data on van Meerbeeck et al 2007 because it the largest study and because it was not characterised by the limitations of the other chemoradiotherapy studies; Eberhardt 2015 (a partially indirect population) and Albain 2009 (a US healthcare setting). The choice of study is expected to make little difference to the model's results as they relate to PFS and PPS but this is not true for the probability of survival. The relative effect for this outcome is an odds ratio, which is then multiplied by the odds of surviving into the long term model on chemotherapy. If the odds of surviving are very large or very small (prob = 0% or 100%) then the resulting absolute difference in probabilities, and therefore differential number of patients in the long term model, will be small. If the odds are close to even (prob = 50%), as in the case of the Eberhardt data then the resulting differential will be large. We used data from Eberhardt as a sensitivity analysis.

Adverse Events

The committee indicate that we should assume adverse events were acute in nature and that they would be unlikely to materially affect patients' health related quality of life for any extended period. The numbers of reported adverse events at grade 4 were extremely low and therefore it was highly uncertain whether they differed meaningfully between interventions. The committee asked us to account for only grade 3+ adverse events in the model as these would be expected to incur a hospital admission and were therefore would potentially influence the net monetary benefit associated with the interventions. Grade 3+ adverse events were treated homogenously in the model (i.e. no difference between grades 3 and 4 and no difference between the clinical nature of events). This approach was taken for several reasons; as mentioned above, grade 4 events were rare, events were reported heterogeneously among trials and the specific nature of events was not expected to affect the net monetary benefit calculations within the model due to lack of meaningful differences in HRQoL loss or costs accrued.

We examined the data and determined that only the larger trials conducted by Pless 2015, Eberhardt 2015 and Albain 2009 had reported adverse events comprehensively enough to give us some confidence in the homogeneity of their data collection and reporting methods. We fitted a baseline incidence rate meta-analysis to the arms containing CRS (as the intervention with the most data and trial arms) where events were the total number of events at grade 3 and above and person years at risk were determined by multiplying the sample size by the total area under the overall survival curve at 5 years (which is equal to restricted mean person years lived for the patients in those trial arms). The test for heterogeneity was significant ($p<0.0001$) so we preferred to use results from a random effects model for the base case analysis.

We then used the same data on events and person years at risk from both arms of the Pless trial to calculate the incidence rate ratio for CS vs CRS. The incidence rate ratio for CR vs CRS was calculated by pooling the data from the Albain and Eberhardt trials in a meta-analysis with random effects again being preferred due to heterogeneity ($p=0.019$).

Late on in development we received additional data from the EORTC on adverse events in the van Meerbeeck trial. This enabled us to fit a network meta-analysis for this outcome using the data from all four large trials. We decided that because the adverse events would be expected to occur within a reasonably short time frame (certainly those that were directly attributable to the interventions) we could assume a homogenous follow up time in our analysis. We therefore used the person years at risk as detailed above and selected a poisson likelihood, log link model for the analysis (the WinBUGS code is available in Appendix I). The NMA calculated hazard ratios, which we applied directly to the baseline incidence rate and overall survival AUC to calculate total events. The deviance information criterion for the random effects model was only 2 points lower so we preferred the fixed effects model in the base case. The credible intervals for the random effects model are very wide so introduce significant uncertainty into the model but have been examined in a sensitivity analysis. Of note, we decided to use a multivariate normal distribution to incorporate these data into the probabilistic sensitivity analysis rather than using the CODA outputs from the NMA so as not to slow down the model. We do not expect this to have affected the results.

The committee examined the resulting data and noted that the total number of events for CS and CR remained roughly the same and that they were both higher than CRS The committee were unsure about the clinical plausibility of this, given that CRS is the more intense intervention but they noted that it could be explained to some extent by the finding that more people in the CS strategy actually go on to have surgery. Ultimately they decided to prefer the pairwise approach over the NMA in the base case as it introduced less uncertainty into the probabilistic sensitivity analysis but in interpreting the results were mindful that few significant differences has been observed in the GRADE tables. A sensitivity analysis where event rates were equal was therefore also specified.

For the 4-year sensitivity analysis we calculated the baseline incident rates using the same number of adverse events and the 4-year person years at risk data. We assumed the pairwise incident rate ratios would remain the same. These data were multiplied by the total person years at risk to give total adverse events at 4 years. These were very similar to using the 5-year data. We did not fit a 4-year NMA because the base case analysis was chosen to be pairwise.

Table 26: Adverse event output data

Adverse Event Data	Mean	SE	Source
Baseline Adverse Event Rate for CRS (RE model)	0.740	0.191	Meta-analysis (Pless, Eberhardt, Albain)
Baseline Adverse Event Rate for CRS (FE model)	0.698	0.027	Meta-analysis (Pless, Eberhardt, Albain)
Baseline Adverse Event Rate for CRS 4 yr (RE model)	0.775	0.197	Meta-analysis (Pless, Eberhardt, Albain)
Baseline Adverse Event Rate for CRS 4 yr (FE model)	0.728	0.028	Meta-analysis (Pless, Eberhardt, Albain)
Incident Rate Ratio (CR vs CRS) - FE model	1.254	0.054	Meta-analysis (Eberhardt, Albain)
Incident Rate Ratio (CR vs CRS) - RE model	1.164	0.155	Meta-analysis (Eberhardt, Albain)
Incident Rate Ratio (CS vs CRS)	1.335	0.112	Pless
HR of CR vs CRS - RE Model	1.18	0.5861	NMA (Pless, Eberhardt, Albain, van Meerbeeck)
HR of CS vs CRS - RE Model	1.38	0.7143	NMA (Pless, Eberhardt, Albain, van Meerbeeck)
HR of CR vs CRS - FE Model	1.24	0.05198	NMA (Pless, Eberhardt, Albain, van Meerbeeck)
HR of CS vs CRS - FE Model	1.4	0.08944	NMA (Pless, Eberhardt, Albain, van Meerbeeck)
Total Events CRS (preferred assumptions)	1.585		Calculated from above
Total Events CS (preferred assumptions)	1.925		Calculated from above
Total Events CR (preferred assumptions)	1.719		Calculated from above
Total Events CR (NMA Derived) (preferred assumptions)	1.743		Calculated from above
Total Events CS (NMA Derived) (preferred assumptions)	1.958		Calculated from above
Cost of an adverse event	£1,590	£398	National Schedule of Reference Cost 2016/17

Costs of Initial Treatment

The committee examined the dosing regimens in the RCTs and noted that the interventions were delivered quite heterogeneously (varied number of cycles of chemotherapy, grays and fractions of radiotherapy and timing of both interventions). They noted that none of the studies were set in the UK and decided on a set of resource uses that they felt were broadly representative of UK practice as well as being similar to the range observed in the trials. This was four cycles of chemotherapy and 55 grays in 20 fractions for radiotherapy in the base case. There are a large

number of possible platinum doublet chemotherapy combinations used in current UK practice, which all cost a similar amount. As costing all these individually and taking a weighted average would not have meaningfully added to the accuracy of the model, we decided to cost a representative treatment. The committee decided that we should use carboplatin and oral vinorelbine for this purpose and supplied us with the typical doses.

Surgery was costed using the NHS reference cost for "Complex Thoracic Procedures, 19 years and over, with CC Score 3-5". The committee felt this was the most representative cost as the procedure was expected to be more complicated than most lobectomy operations, which were costed at "…CC score 0-2". A proportion of operations for N2 stage disease are pneumonectomies which the committee also felt would be covered by this reference cost.

Costs of Interventions		
Radiotherapy Costs		
Hypofractionated Radiotherapy 55 Gy/20#/4 weeks		
Define volume for simple radiation therapy with imaging and dosimetry	1	Resource use from CG121
Deliver a fraction of complex treatment on a megavoltage machine	1	Resource use from CG121
Deliver a fraction of treatment on a megavoltage machine	19	Resource use from CG121
Define volume for simple radiation therapy with imaging and dosimetry cost - SC03Z	£362.59	National Schedule of Reference Cost 2016/17
Deliver a fraction of complex treatment on a megavoltage machine cost - SC23Z	£138.42	National Schedule of Reference Cost 2016/17
Deliver a fraction of treatment on a megavoltage machine cost - SC22Z	£103.37	National Schedule of Reference Cost 2016/17
Total cost of Standard Fractionated Radiotherapy 60–66 Gy/30–33#/6–6.5 weeks	£2,465.07	Calculated
Proportion Receiving 55 in 20	1	Committee Assumption
Total Radiotherapy Cost	**£2,465.07**	**Calculated**
Systemic Anti-Cancer Therapy (platinum doublet chemotherapy)		
Number of cycles	4	Committee Assumption
Outpatient appointment - SB12Z	£173.99	National Schedule of Reference Cost 2016/17
Administration appointment (0.25 of band 4 time, at £28ph)	£7.00	PSSRU 2017 for band 4 hourly cost
Vinorelbine		
Resource use per cycle		
80mg capsule	2	Committee Assumption
20mg capsule	4	Committee Assumption

Costs of Interventions		
Cost per unit of resource		
80mg capsule	£175.50	NHS Indicative Price (BNF Online)
20mg capsule	£43.98	NHS Indicative Price (BNF Online)
Total cost of Vinorelbine (per cycle)	£526.92	Calculated
Carboplatin		
Resource use per cycle		
Dose of Carboplatin required per cycle (mg)	575	Committee Assumption
Dose per vial Carboplatin 150mg/15ml solution for infusion vials (mg)	150	Committee Assumption
Number of Carboplatin 150mg/15ml solution for infusion vials required	3.83	Committee Assumption
Cost per unit of resource		
Price per vial Carboplatin 150mg/15ml solution for infusion vial	£6.35	eMIT National 2016/2017 NCP Code DHE001
Total cost of Carboplatin (per cycle)	£24.34	Calculated
Dexamethasone 8mg bd, reducing over 4 weeks, top dose 1 week and taper down	£74.34	Drug Tarriff 2018
Total cost of SACT (per cycle)	£750.84	Calculated
Total cost of SACT (all cycles)	**£3,003.36**	**Calculated**
Surgery - Complex Thoracic Procedures, 19 years and over, with CC Score 3-5	**£ 7,562.42**	**National Schedule of Reference Cost 2016/17**

Progressions (costs and events)

Since progression-free survival represents both patients who have not either progressed to a more advanced stage of disease or died, obtaining the number of progressions that are in fact deaths is necessary. These data were only available in the Pless 2015 and in a personal communication from the EORTC, who hold the data for van Meerbeeck 2007. The data from both studies was pooled in a fixed effects meta-analysis (heterogeneity p=0.18) to obtain the proportion of progressions that were deaths for the CS intervention, the relative risk was obtained from the van Meerbeeck data and applied to the pooled CS estimate to calculate the proportion for CR and the relative risk was obtained from the Pless data and applied to the pooled CS estimate to calculate the proportion for CRS. These data were highly uncertain and it was not clear that they had clinical plausibility (there was no obvious reason why the proportion of progressions that were deaths would be higher in CS patients than in CRS patients, for example), so were set equal to one another in sensitivity analysis.

Upon progressions that were not deaths, patients were assumed to be treated with another round of systemic therapy. We had no data on the specific types of progression and it was not clear that progression type or the indicated treatment would be expected to differ significantly between the interventions so the committee thought this simplifying assumption reasonable. There are a very large number of systemic therapy options available in NSCLC (see RQ 3.3 of this update for a full algorithm) so costing them all and factoring in their differential benefits in this patient population would have been impractical and subject to high uncertainty. These treatment options have typically been the subject of NICE Technology Appraisals and therefore represent cost-effective additions to the care pathway, but additions that the committee was aware were unlikely to add much in terms of net monetary benefit. This is because Technology Appraisal approved drugs in advanced cancer rarely have base case ICERs significantly lower than the upper limit of the ICER range normally considered cost effective by NICE. The committee also noted that much of the evidence in this model came from survival data collected before many of these drugs were widely available. They therefore thought that the net monetary benefit associated with systemic therapy could reasonably be approximated using the costs of platinum doublet chemotherapy. Four cycles of oral vinorelbine with carboplatin was again chosen for this purpose and the overall cost of systemic therapy for progression was explored in sensitivity analysis.

Table 27: Progressions that are deaths

Proportion of progressions that are deaths	Mean	SE	Source
RR of CR vs CS	0.516	0.285	van Meerbeeck 2007 (supplimentary data)
RR of CRS vs CS	0.651	0.459	Pless 2015
CS	0.183	0.258	FE MA (Pless + van Meerbeeck)
CR	0.095		Calculated
CRS	0.119		Calculated

The committee noted the convergence of the overall and progression free survival curves and made the assumption that progression-free survival would equal overall survival at the study endpoint of 5 years. They felt that NSCLC would be highly unlikely to recur in the vast majority of patients who were alive and unprogressed at this point. The number of progressions for each intervention during the first 5 years was therefore calculated by multiplying one minus the proportion still alive by one minus the proportion of progressions that were deaths.

The total number of deaths was equal to one minus the probability of survival at study endpoint and a cost of death representing a total package of end-of-life care was applied that was drawn from a study including the costs accrued by cancer patients in their last 90 days of life (Georghiou and

Bardsley 2014[k]). This data source had also been used by NICE's recently published guideline on Early and Locally Advanced Breast Cancer. The cost of existing in the pre and post progression states for 90 days, weighted by the proportion of people who were expected to die directly from each state was then subtracted to give the total death-attributable cost. We assigned the overall value an arbitrary high standard error equal to a quarter of the mean as these data were quite uncertain.

Table 28: Death costs

Death Event Costs	Mean	SE	Source
Hospital Costs	£5,890	-	Georghiou and Bardsley 2014
Local Authority Funded Care	£444	-	Georghiou and Bardsley 2014
District Nursing Care	£588	-	Georghiou and Bardsley 2014
GP Contacts	£365	-	Georghiou and Bardsley 2014
Months death costs apply	3	-	Georghiou and Bardsley 2014
Inflation Factor (average over 4 years)	1.063	-	PSSRU HCHS 2014/15 – 2016/17 * 2
Death Event Total Costs (minus weighted state membership costs)	£4,575	£1,144	Calculated

Discounting

Discounting was implemented at 3.5% throughout the model. While the NMAs already discussed provided discounted values for PFS and PPS and probability of OS, which could be multiplied directly by state membership and utility estimates to produce appropriate discounted values, another solution was needed for event costs. Another two NMAs were therefore conducted (see full discussion in Appendix I) that calculated the proportion of progressions and deaths that occurred in each year. These proportions were multiplied by the total number of deaths and progression events and the appropriate discount factor for each year of the model to give a total weighted discounted average cost for both types of events.

Table 29: Proportion of events occurring in each year

Proportion of events occurring in each year (NMA results)		
Weighting of Progressions (5 Year model)	value	SE
Progs - Year 0	0.632871	0.02003
Progs - Year 1	0.2346	0.02529

[k]Georghiou and Bardsley (2014) Exploring the cost of care at the end of life. Nuffield Trust

Proportion of events occurring in each year (NMA results)		
Progs - Year 2	0.08428	0.02637
Progs - Year 3	0.03868	0.02684
Progs - Year 4	0.009569	0.02145
Weighting of Deaths (5 Year Model)		
Deaths - Year 0	0.3849	0.02891
Deaths - Year 1	0.324	0.03051
Deaths - Year 2	0.1555	0.03051
Deaths - Year 3	0.1103	0.03252
Deaths - Year 4	0.0253	0.03153
Weighting of Progressions (5 Year Model)		
Progs - Year 0	0.6474	0.02094
Progs - Year 1	0.2432	0.02643
Progs - Year 2	0.09203	0.02887
Progs - Year 3	0.01737	0.02494
Weighting of Deaths (4 Year Model)		
Deaths - Year 0	0.3906	0.02107
Deaths - Year 1	0.3471	0.02993
Deaths - Year 2	0.1662	0.03282
Deaths - Year 3	0.0961	0.03303

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2173 **Drop Out Rates**

2174 The overall and progression-free survival curves provided intention-to-treat effectiveness data for each arm of each study. Not all patients in the
2175 surgery arms actually had surgery, however, through either dying, not being fit enough or changing their mind by the end of chemoradiotherapy.
2176 The committee therefore thought that the cost of the strategies including surgery should reflect these data. We were able to obtain the proportion
2177 of people actually undergoing surgery from the CS and CRS arms of all the trials. We pooled the data for proportion of patients undergoing surgery

and used a random effects model due to high statistical heterogeneity. Because the smaller studies were less certain and contributed quite a lot of heterogeneity to this calculation we excluded them and pooled only the large studies in a fixed effects meta-analysis. We repeated this same procedure for CS; both the meta-analyses with and without large trials were fitted using random effects models to account for statistical heterogeneity. In the base case, we used the data containing only large trials because we thought it more reliable but the value obtained using all the trials and a value of 100% were examined in sensitivity analysis.

Table 30: Proportion in surgical arm continuing to surgery

Proportion in surgical arm continuing to surgery	Mean	SE	Source
CRS % still having surgery (all trials)	0.8934	0.0281	RE Meta-analysis (Pless, Albain, Eberhardt, Girard, Katakami)
CRS % still having surgery (large trials only)	0.8349	0.0185	FE Meta-analysis (Pless, Eberhardt, Albain)
CS % still having surgery (all trials)	0.9048	0.04	RE Meta-analysis (van Meerbeeck, Girard, Pless, Katakami)
CS % still having surgery (large trials only)	0.8739	0.0522	RE Meta-analysis (van Meerbeeck, Pless)

Health State Costs

No background healthcare resource use data was available for patients with NSCLC stage IIIA-N2. We examined the literature for inspiration and presented a number of possible resource uses to the committee. The committee debated these data and, incorporating their own clinical experience, settled on the assumptions in Table 31 and Table 32 as being broadly representative of a typical patient in the progression free and progressed states. The total monthly average cost is the sum of the product of % of patients, units and costs for each type of resource.

Table 31: Monthly Progression Free State Costs

Weighted monthly average cost of Progression Free	% patients resource use	Units	Cost	
Hospitalisation	3%	1	£1,590.00	National Schedule of Reference Cost 2016/2016
Cancer Nurse	20%	1	£38.75	National Schedule of Reference Cost 2016/2017
Pallitative Care Nurse	30%	1	£102.41	National Schedule of Reference Cost 2016/2017
Pallitative Care Physician	8%	1	£158.81	National Schedule of Reference Cost 2016/2017
Outpatient	75%	1	£191.11	National Schedule of Reference Cost 2016/2017
GP Visit	10%	1	£38.00	PSSRU 2017 General Practioner
Complete blood count	100%	0.75	£3.06	National Schedule of Reference Cost 2016/2017

Weighted monthly average cost of Progression Free	% patients resource use	Units	Cost	
Palliative radiotherapy	13%	1	£132.40	National Schedule of Reference Cost 2016/2018
CT scan	30%	0.75	£120.07	National Schedule of Reference Cost 2016/2019
X-Ray	100%	0.75	£25.00	FOI Request (23023) Stockport NHS Trust 2014
Biochemistry	100%	0.75	£1.13	National Schedule of Reference Cost 2016/2017
Total Monthly Average Cost			£302.72	Assumed SE = £75.68

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2192 **Table 32: Monthly Progressed State Costs**

Weighted monthly average cost of Progressed	% patients resource use	Units	Cost	Cost Source
Hospitalisation	30%	1	£1,590.00	National Schedule of Reference Cost 2016/2017
Cancer Nurse	10%	1	£38.75	National Schedule of Reference Cost 2016/2017
Pallitative Care Nurse	20%	1	£102.41	National Schedule of Reference Cost 2016/2017
Pallitative Care Physician	80%	1	£158.81	National Schedule of Reference Cost 2016/2017
Outpatient	100%	2	£191.11	National Schedule of Reference Cost 2016/2017
GP Visit	28%	1	£38.00	PSSRU 2017 General Practioner
Stereoids (Dexamethasone 0.5mg tablets)	50%	1	£0.58	Price from May 2018 Drug Tarrif.
NSAIDS (ibuprofen 200mg tablets)	30%	16	£0.03	Price from May 2018 Drug Tarrif.
Morphine (20mg tablets)	75%	60	£0.19	Price from May 2018 Drug Tarrif.
Biphosphonate (5mg risendronate)	8%	21	£0.67	Price from May 2018 Drug Tarrif.
Dietary supplement (350gram can)	40%	28	£2.31	BNF 2018
Complete blood count	100%	20	£3.06	National Schedule of Reference Cost 2016/2017
Palliative radiotherapy	20%	1	£132.40	National Schedule of Reference Cost 2016/2018
Biochemistry	100%	1	£1.13	National Schedule of Reference Cost 2016/2017
CT scan	5%	1	£120.07	National Schedule of Reference Cost 2016/2018
Home oxygen	20%	0.75	£107.84	http://www.emrespiratory.co.uk/downloads/documents/HOSAR-Good-Practice-Guide.pdf

Weighted monthly average cost of Progressed	% patients resource use	Units	Cost	Cost Source
X-Ray	30%	0.7	£25.00	FOI Request (23023) Stockport NHS Trust 2014
Total Monthly Average Cost			£1,173.45	Assumed standard error = £293.36

To calculate total costs for the short term model these costs were multiplied by the average discounted time that patients spent in each state, which was derived from the relevant NMA.

Long Term Model

Patients surviving the short term model entered the long term model, which was a partitioned survival model with two states; dead and alive + progression free. It was assumed that no progressions took place among the surviving patients and they had, to all intents and purposes been cured of their lung cancer. Death events were accrued at a rate equivalent to the difference in the death state membership from cycle to cycle. The long term model was run on a monthly cycle length and a half-cycle correction using the life table method was applied. As discussed earlier, progression-free utility estimates were adjusted to reflect the decline in HRQoL in the general population at older ages. Progression-free costs continued to be applied in the model but at a rate of only 20% to reflect the assumptions that patients would be permanently remitted after 5 years but the committee felt patients would still continue to interact with services to some degree, especially if they had impaired lung function following radical treatment.

In order to obtain appropriate survival curves we interrogated the SEER registry[l], which was chosen because it was the only registry we knew about with the ability to extract the data we needed. The database was queried for survival data for patients who were diagnosed between 1988-2003, aged 35-79, had stage IIIA-N2 lung cancer upon diagnosis and had survived five years after their initial diagnosis. We fit survival curves to the data and selected the two with the lowest AIC statistics for use within the model as the base case and in sensitivity analysis. These were Weibull and exponential curves fitted to data from 2,865 patients. From Figure 14, it can be seen that they fitted the survival data well. The AUC (or mean survival time) for these curves was about seven years. The data were somewhat out of date and we were unable to identify any data that would enable us to differentiate these curves by initial treatment but the committee thought that as they were meant to represent a cured population, these limitations were minor. The same process as this was undertaken to parameterise the 4-year sensitivity analysis, with Weibull and Exponential curves again providing the best fit to the data (N=3,703).

[l] https://seer.cancer.gov/registries/

2214 **Table 33: Long term survival curve parameters**

Proportion still having surgery	Mean	SE	Source
4 Year Weibull Shape	0.8466	0.0144	SEER Data
4 Year Weibull Scale	6.8844	0.1694	SEER Data
5 Year Weibull Shape	0.8846	0.0174	SEER Data
5 Year Weibull Scale	7.3666	0.2016	SEER Data
4 Year Exponential	0.14736	0.00305	SEER Data
5 Year Exponential	0.13808	0.00331	SEER Data

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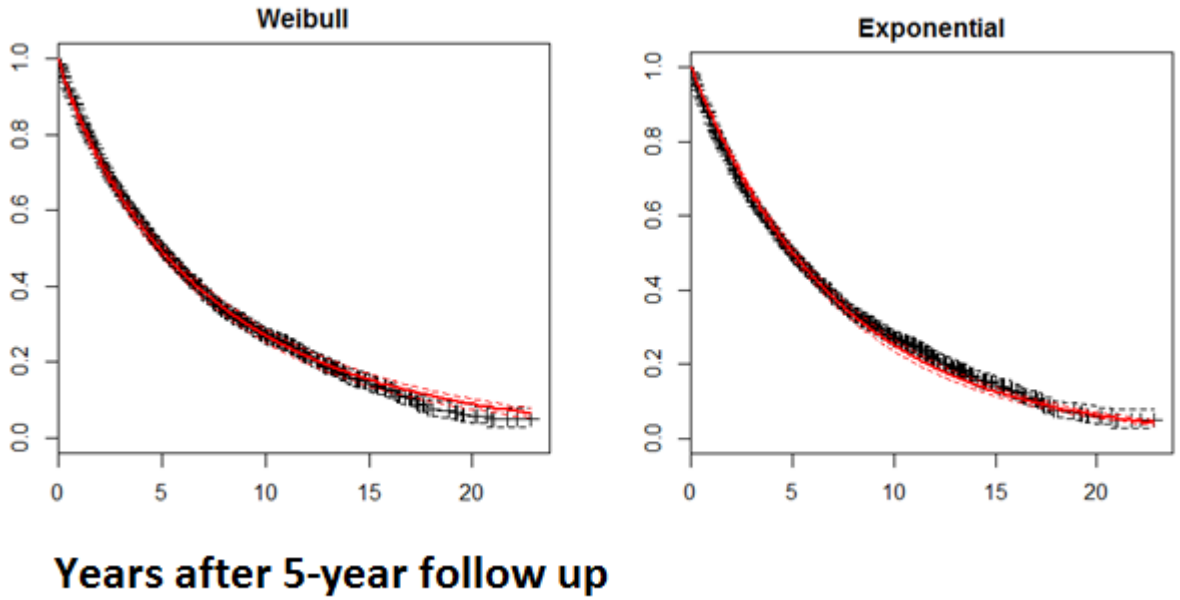


Figure 14: SEER Survival Data and Parametric Models

Sensitivity Analysis

Sensitivity and scenario analyses was conducted by altering key parameters or groups of parameters including changing the short term element of the model to cover four years instead of five, using random effects NMAs instead of fixed effects, changing key cost and utility parameters, setting probability of survival at study endpoints and various other uncertain data equal among interventions, using different survival curves and altering the discount rate.

Probabilistic sensitivity analysis was performed by assigning parameters with appropriate probability distributions that reflected our uncertainty about their mean values. Of note, the NMAs used the relevant CODA. The very bottom end of the posterior distributions for AUC values for PFS and PFS in the random effects models had to be truncated at 0. This was because the NMA input and output data were on the natural scale (i.e. number of years) and so some impossible negative AUC values arose due to the wide credible intervals in the posterior distribution of the random effects models. This was only a small amount of data so was noted as a minor limitation for the PSA in the random effects scenario analysis.

Particularly uncertain costs that were heavily influenced by assumptions (such as the state membership costs and the cost of death) were arbitrarily assigned a high standard error equal to the mean divided by four. As noted in the adverse events section, the hazard ratios derived from NMAs were parameterised using a multivariate normal distribution on the log scale to reduce model size and running time.

Results

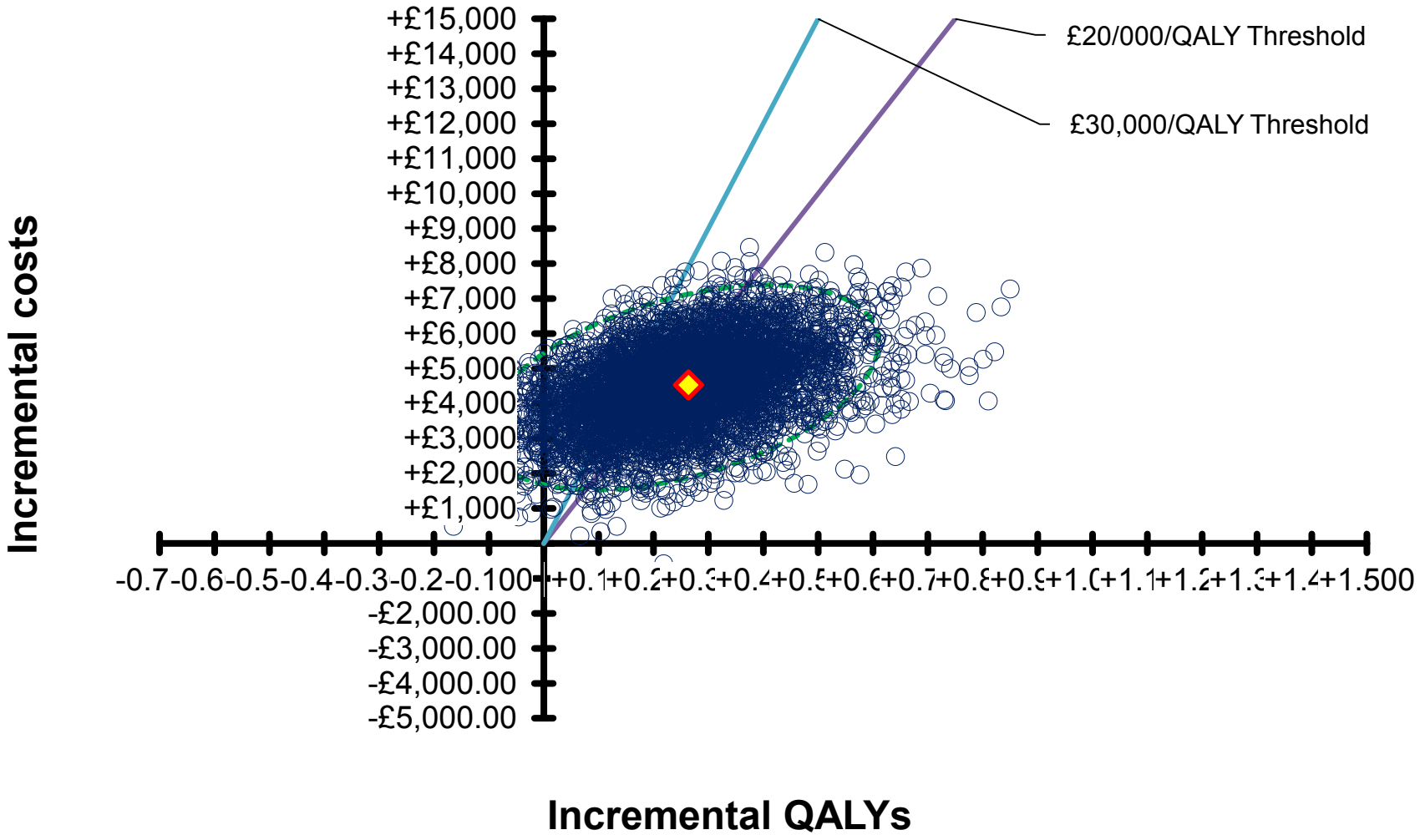
All base case results presented in this section are the mean of 5,000 probabilistic iterations of the model unless otherwise stated. The base case assumptions were; 5 year fixed effects NMA data, random effects pairwise adverse event data.

Table 34: Base Case Results (Fixed Effects NMAs)

Probabilistic										
	Name	Absolute values		Incremental						
				Fully incremental analysis			Compared with:	Chemotherapy and Surgery		
		Costs	QALYs	Costs	QALYs	**ICER**	Costs	QALYs	**ICER**
1	Chemoradiotherapy	£28,359	1.97190				-£3,180	-0.05943	**£53,503**
2	Chemotherapy and Surgery	£31,539	2.03133	£3,180	0.05943	**ext. dom.**	-	-	**ref**
3	Chemoradiotherapy and Surgery	£32,820	2.22299	£4,461	0.25109	**£17,768**	£1,282	0.19166	**£6,687**

Table 35: Base Case Results (Random Effects NMAs)

Deterministic										
	Name	Absolute values		Incremental						
				Fully incremental analysis			Compared with:	Chemotherapy and Surgery		
		Costs	QALYs	Costs	QALYs	**ICER**	Costs	QALYs	**ICER**
1	Chemoradiotherapy	£28,437	2.00003				-£3,818	-0.04389	**£86,996**
2	Chemotherapy and Surgery	£32,255	2.04392	£3,818	0.04389	**ext. dom.**	-	-	**ref**
3	Chemoradiotherapy and Surgery	£33,180	2.23791	£4,744	0.23788	**£19,941**	£926	0.19399	**£4,771**



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2238 **Figure 15: Cost Effectiveness Plane CRS vs CR (base case)**

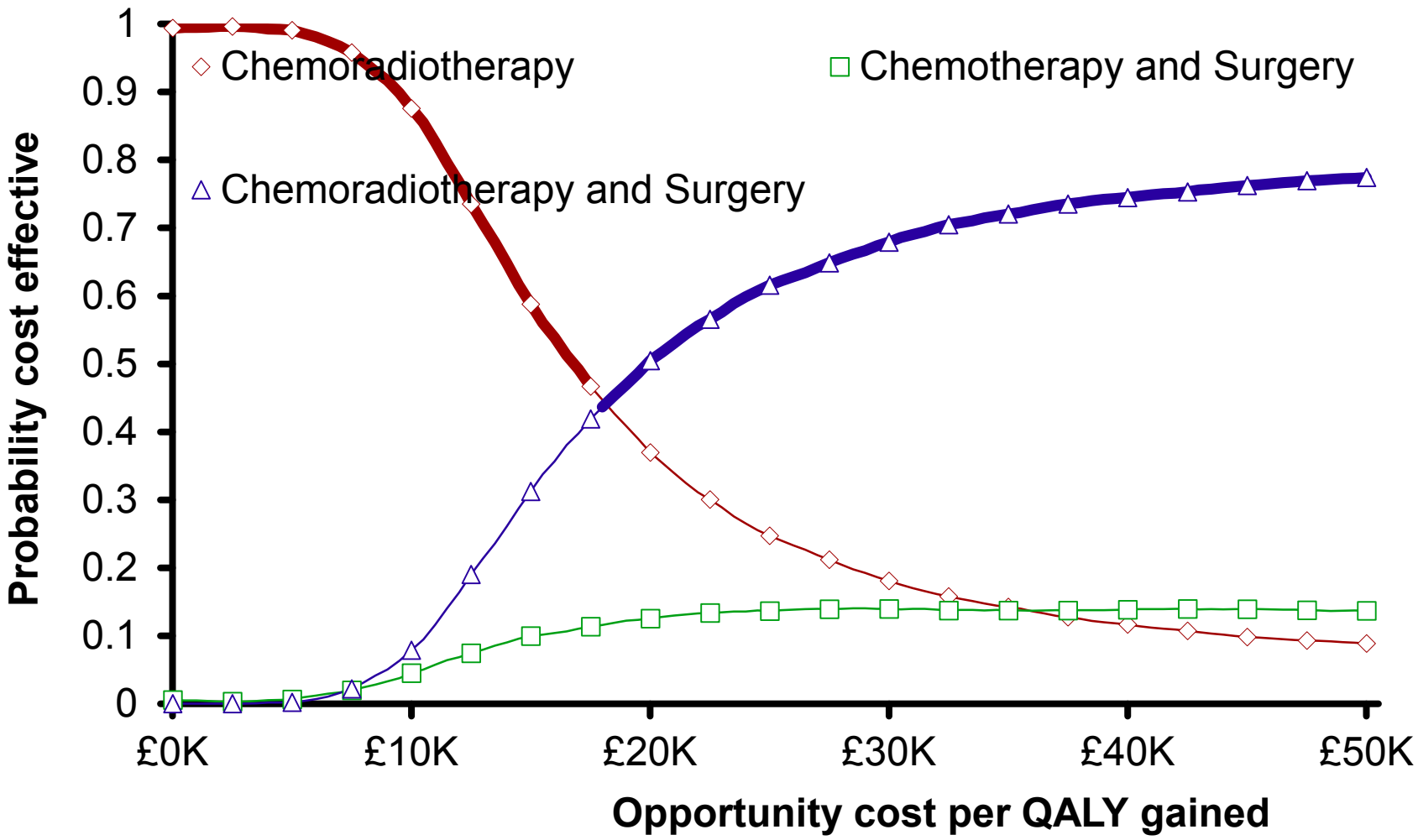

Figure 16: Cost-Effectiveness Acceptability Curve (base case)

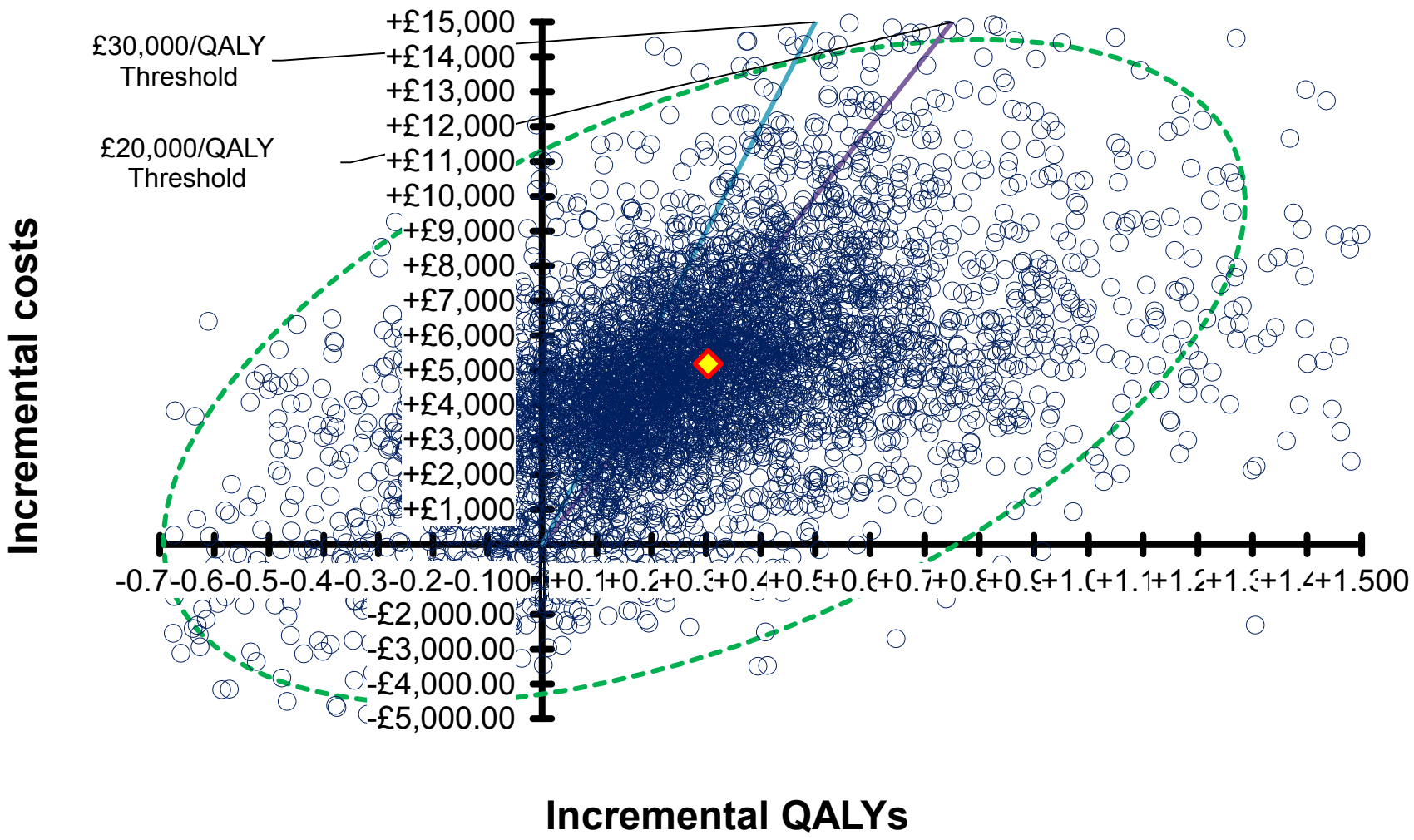

Figure 17: Cost-Effectiveness Plane CRS vs CR (random effects NMAs for PFS, PPS and Prob S)

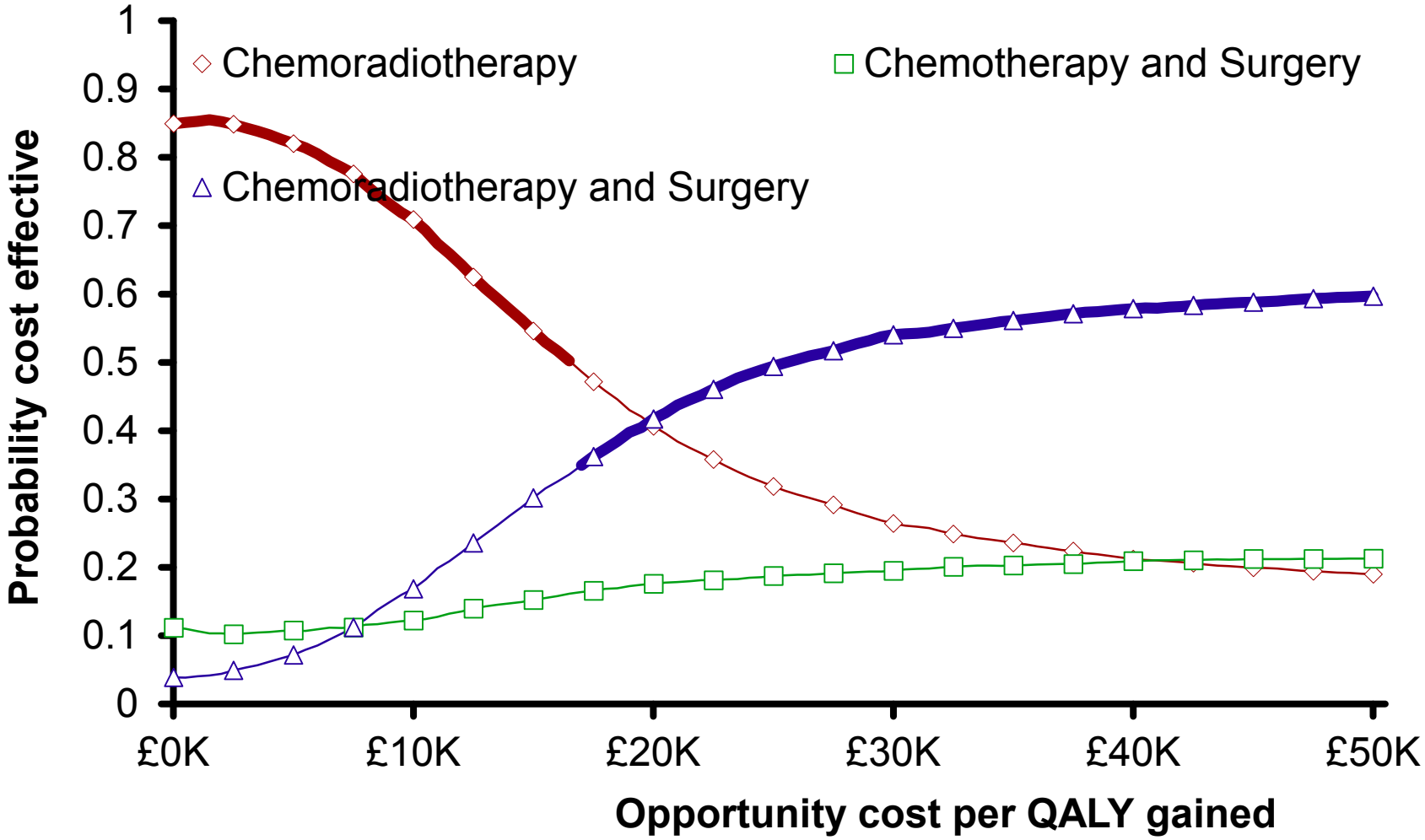

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Figure 18: CEAC (random effects NMAs)

Table 36: Pairwise ICERs from Scenario Analyses (results are deterministic unless otherwise noted)

Scenario	CRS vs CR	CS vs CR	CRS vs CS	Notes
Base Case (5y, FE, disc)	£18,714	£54,673	£7,414	
Base Case PSA	£17,768	£53,503	£6,687	Based on the mean of 5,000 iterations
5Y Random Effects	£19,941	£86,996	£4,771	Random rather than fixed effects NMAs used for first 5 years
No adverse events	£19,597	£49,034	£10,347	Adverse events = 0 for all treatments
Adverse events from NMA	£18,030	£52,988	£7,045	Based on NMA (see appendix J) rather than pairwise data
No treatment disutility	£17,974	£46,327	£7,433	Surgical patients suffer no post-surgery utility decrement
No long term utility decrement	£18,605	£53,219	£7,421	Standard age related utility decrements not applied
Exponential survival curve	£18,946	£58,087	£7,401	Survival in patients alive at 5 years modelled using an Exponential curve
Long term PFS costs = 100%	£20,324	£62,184	£7,170	Costs for patients surviving 5 years progression free = those within the first 5 years
Long term PFS costs = 50%	£19,318	£57,489	£7,414	Costs for patients surviving 5 years progression free half those within the first 5 years
% undergoing surgery MA = all trials	£20,602	£59,089	£8,563	% patients dropping out of surgery after chemotherapy derived from all trials in NMA
% undergoing surgery = 100%	£24,072	£73,059	£9,039	% patients dropping out of surgery after chemotherapy = 0%
Discount rate = 0%	£14,797	£27,145	£7,324	No economic discounting
4y Fixed Effects NMA	£19,696	£152,217	£8,247	NMAs are from 4 year outcomes rather than 5 year. Survival continues from 4 years
Progs that are deaths set equal	£18,973	£58,496	£6,553	% of progressions that are in fact deaths set equal among treatments
PFS Utility = 0.72	£20,077	£58,875	£7,945	Progression free utility set to lowest value from literature review
PFS Utility = 0.83	£17,543	£51,088	£6,957	Progression free utility set to highest value from literature review
Max util, Max decr between PFS and PPS	£18,125	£53,921	£7,140	PFS utility and utility decrement from progression set to highest available values
Min util, Min decr between PFS and PPS	£19,365	£55,513	£7,718	PFS utility and utility decrement from progression set tolowest available values
OR of survival set equal	£32,621	dominated	£6,990	OR of survival = 1 for CS and CRS vs CR
Cost of Surgery = CC 6+	£27,065	£91,230	£6,901	Assume cost of surgery = to most complex in class
Cost of Surgery = CC 0-2	£15,126	£38,968	£7,634	Assume cost of surgery = to least complex in class
Cost of Progressed State Halved	£24,387	£63,985	£11,943	Monthly cost of the post progression state halved
Eberhardt baseline for NMAs	£12,330	£19,423	£5,224	Baseline population CR data from Eberhardt 2015

Discussion

CS produced QALY and life year gains of 0.06 and 0.16 over CR, whereas CRS produced QALY and life year gains of 0.25 and 0.37 over CR. The model results show a high probability that that CRS produces the most life years and QALYs. The probability that CRS generates more QALYs than CR is 97% in the base case analysis and 82% if random effects NMAs are used. There were no plausible and robust sensitivity analyses in which CS would be considered cost-effective compared to CR at £20,000 per QALY gained and the comparison of CRS vs CS uniformly produced ICERs of less than £20,000/QALY. CS produced more QALYs than CR in 64% of model iterations and CRS produced more QALYs than CS in 87%. The model provides evidence that CS is unlikely to be a cost-effective option, being extendedly dominated by the combination of CR and CRS and having a high ICER vs CR, which is subject to high uncertainty. The cost effectiveness acceptability curve always showed CS as having a relatively low probability of being the most cost-effective option, regardless of the value of a QALY.

The model was quite insensitive to a large number of the parameters examined in sensitivity analysis and consistently produced ICERs for CRS vs CR of below £20,000/QALY. One particularly noteworthy source of uncertainty was the sensitivity analysis around the probability of survival at study endpoint, which produced an ICER of slightly over £30,000/QALY for CRS vs CR. The fixed effects NMA for this outcome did not find any significant differences among interventions for this outcome although 86% of the probability mass for the difference in this outcome favoured CRS. In the analysis where the probability of survival at study endpoint is set equal, CRS still produces more QALYs in 92% of model iterations.

The mean ICERs were very similar using random rather than fixed effects NMAs. While these models were not found to be statistically preferable, they might have been more appropriate given some of the heterogeneity in patient populations and interventions in the included studies. The cost-effectiveness plane shows a very wide dispersion of results for the random effects analysis.

CS was always extendedly dominated by the combination of CR and CRS in the scenario analyses. Furthermore, in the majority of these scenario analyses, the ICER for CS vs CR was above £30,000/QALY and was highly sensitive to a number of parameters. This variability in ICERs is due to the small QALY improvement of CS over CR.

Of note, if the Eberhardt data are used as the baseline for PFS, PPS and the probability of survival, the ICERs for the surgical options are much lower. This is because the odds ratio for survival derived from the NMA is applied to a much larger baseline odds of a survival, which produces a greater differential probability of surviving into the long term model. Overall survival in the Eberhardt trial was close to three times that in the van Meerbeeck trial at five years. The choice of trial for the base case analysis is discussed in the methods section but it is likely that the ‘true’ ICERs for the surgical options lie somewhere between the base case and the Eberhardt data i.e. they are likely more cost-effective than our base case results suggest.

Overall, the results of our model suggest that CRS is likely to be a cost-effective improvement over CR but that CS is unlikely to be, albeit with some uncertainty in the underpinning clinical data. This is due largely to the results of the NMAs conducted for this guideline showing that people receiving CRS spend significantly longer progression free and are potentially more likely to be cured of their lung cancer. Differences in adverse events between the different interventions were small and somewhat uncertain and had a fairly significant effect on the results for CS. Adverse

event data did not affect the ICER for CRS vs CR when the rates were set equal. The ICER for CRS vs CR was affected somewhat by the assumption that not all patients would actually continue on to surgery after completing chemoradiotherapy but remained under £30,000 per QALY when this assumption was relaxed. The ICERs were also sensitive to the cost of surgery and the costs of progressed state membership although again remained under £30,000/QALY for CRS vs CR when extreme assumptions were tested.

Strengths and Limitations

Our analysis has a number of important strengths. As far as we are aware is the first cost-effectiveness analysis examining treatment options in people with NSCLC stage IIIA-N2, which is a common presentation that is managed variably across the UK NHS and the world. It is based on novel and high quality methods for synthesising the wealth of data available in the trials conducted to date. In terms of its conclusions for UK practice, the model is insensitive to the vast majority of sensitivity and scenario analyses that were conducted to explore the limitations and uncertainties in the underlying data.

The model also has a number of limitations of varying importance. NSCLC stage IIIA-N2 is a heterogeneous condition and we were unable to find sufficient evidence that enabled us to examine the relative cost-effectiveness of treatment options in different subgroups, for example those indicated for lobectomy versus pneumonectomy, bulky versus non-bulky and multiple versus single-station N2. The model used PFS utility estimates drawn from a potentially clinically and somewhat culturally indirect population, a progression utility adjustment from an indirect population as well as making several strong assumptions about costs and resource use associated with state membership and death events. We were unable to account for advances made in systemic treatment (for example targeted and immunotherapy) although given that these new drugs are usually very expensive, we speculate that surgical options might be more cost-effective because they are associated with a lower probability of disease progression than CR. Most of the data used to drive the model was collected before these drugs were widely available but it is unclear how much survival time, if any, could be attributable to them being used in patients with more advanced disease. Furthermore, people who progress often receive multiple lines of systemic treatment, which was not accounted for at all in our model. Again though, this could make surgical options more cost-effective because more progressions occur in CR and more time is spent in the post-progression state. Adverse events were modelled quite crudely but made little difference to the conclusions. The background resource use of patients surviving into the long term model was uncertain and had a big effect on ICERs. The NMAs driving the model in the base case were fixed effects models with the two statistically significant findings that CRS provided more progression free life years than CR and that CR provided more post-progression life years. While not preferable on grounds of statistical model fit, it might have been more appropriate to use the random effects data, which did not find any statistically significant outcomes (although point estimates remained roughly consistent). The results of the model when driven by the random effects data are more uncertain although the base case ICERs are similar. The model also did not specifically include a strategy of CR followed by immunotherapy as this is currently not a routine option for people with NSCLC stage IIIA-N2 on the UK NHS. The committee were aware of the

existence of relevant data from the PACIFIC[m] trial but the NICE Technology Appraisal on durvalumab, the immunotherapy used in that trial, is not expected to publish until after the publication of this guideline. At that point, there may be another option in this decision space.

[m] Antonia et al (2017) Durvalumab after Chemoradiotherapy in Stage III Non–Small-Cell Lung Cancer. New England Journal of Medicine

2308 Appendix L – Research recommendations

2309

• Question	• What is the effectiveness and cost effectiveness of immunotherapy in people with stage IIIa-N2 NSCLC following multimodality treatment including surgery?
Population	Patients with NSCLC stage IIIA-N2 who have received multimodality treatment (including surgery)
Characteristics of interest	Overall survival Health-related quality of life Adverse events grade 3 or above Safety
Study design	Randomised controlled trial

2310

• Potential criterion	• Explanation
Importance to patients, service users or the population	Immunotherapy has been shown to be effective in a variety of NSCLC indications but there is currently no evidence on whether it is clinically or cost effective for people with stage IIIA-N2 non-small-cell lung cancer following surgery. There is also no evidence on whether it could be used as a replacement or adjunct to current multimodality treatment. The committee made a research recommendation to address this.
Relevance to NICE guidance	Medium priority: a recommendation was made for people with stage III a – N2 who are well enough for multimodality therapy and who can have surgery, to consider chemoradiotherapy with surgery. This updated recommendation could lead to a change in current practice in that more tri-modality therapy might be performed. The role of immunotherapy in current multimodality treatment is worthy of further research to potentially

• Potential criterion	• Explanation
	strengthen this recommendation and provide further treatment options for this presentation where survival is currently poor.
Current evidence base	The updated recommendation is based on statistical and health economic analysis, therefore more RCT studies are required in a UK setting.
Equality	This study could improve equality of access to multimodality treatment for stage IIIa-N2 disease and ensure more people receive this potentially curative treatment.
Feasibility	There is a large enough population of people with this condition and the interventions are available in current clinical practice.

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